

Bayesian Modeling and Test Planning for Multi-phase Reliability Assessment

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Abstract

We propose a Bayesian hierarchical model to assess the reliability of a family of vehicles, based on the development of the Joint Light Tactical Vehicle (JLTV). The proposed model effectively combines information across three phases of testing and across common vehicle components. The analysis yields estimates of failure rates for specific failure modes and vehicles as well as an overall estimate of the failure rate for the family of vehicles. We are also able to obtain estimates of how well vehicle modifications between test phases improve failure rates. In addition to using all data to improve on current assessments of reliability and reliability growth, we illustrate how to leverage the information learned from the three phases to determine appropriate specifications for subsequent testing that will demonstrate if the reliability meets a given reliability threshold.

Keywords: Assurance Testing, Combining Information, Defense Acquisition, Reliability, Reliability Growth

1 Introduction

Reliability is a high priority in the testing of defense systems. A common metric is *reliability growth*, which tracks the change, and ideally improvement, in the reliability of a system as it moves through testing phases and undergoes periodic corrective actions. To most effectively use resources plans for future test events should be based on what has been observed in previous testing.

In this paper, we propose models for combining information across multiple test phases with the ultimate objective of planning a future test. When combining information, there is no omnibus solution. Rather, models need to be carefully considered and evaluated to ensure that they accurately reflect the data and the underlying physical processes. Our analysis is motivated by the reliability assessment of the Joint Light Tactical Vehicle (JLTV), which is a family of vehicles designed to replace one-third of the legacy high-mobility multipurpose wheeled vehicle (Humvee) fleet. There are four unique types of vehicles, and within types there are many possible variants. The primary mission of the family of vehicles is to provide ground mobility that is deployable worldwide and capable of operating across the range of military roles including combat, sustainment, police action, peacekeeping, and security patrol, in all weather and terrain conditions.

The JLTV was evaluated using multiple test events as it was developed. Engineering and Manufacturing Development (EMD) included three phases of testing, where fixes occurred only during a set Corrective Action Period (CAP) between each phase. For every vehicle, each failure encountered during testing was recorded and attributed to a specific failure mode. Since fixes were delayed to CAPs, analysis generally shows a distinct jump in the system reliability between test phases.

We propose a model that can effectively combine information across the three phases of testing and across common vehicle components. We use a Bayesian hierarchical model

to assess the reliability of the family of vehicles. The model accounts for the commonality across vehicle types and allows for uncertainty quantification. Inferential objectives from the proposed model include an estimate of the mean miles between failure (MMBF) for each vehicle, failure mode, and phase of test, and an assessment of the effectiveness of the corrections completed in the CAPs. Additionally, because future tests are being planned, we propose methodology to leverage the information learned from the three phases to determine the length of testing needed to demonstrate a given reliability threshold in subsequent testing.

2 Data

Eight vehicle prototypes were used for reliability, availability, and maintainability (RAM) testing of the JLTV. There were four two-seat utility vehicles, two four-seat general purpose vehicles, a two-seat heavy guns carrier, and a two-seat close combat weapons carrier. The close combat weapons carrier, one general purpose, and two utility vehicles were tested in Aberdeen, Maryland, and the heavy weapons carrier, one general purpose, and two utility vehicles were tested in Yuma, Arizona. When comparing the test sites, the Maryland test site is much colder in the winter months but the terrain is not as harsh; in Arizona, it is warmer but with more challenging terrain. Using two test sites allows for the system to be tested under a broader range of environmental conditions. The vehicles were driven roughly 3000 to 5000 miles in Phase 1, between 4500 and 7500 miles in Phase 2, and between 5000 and 6700 miles in Phase 3.¹

There are many similarities between vehicle types; in fact, the most dissimilar vehicles have over 80% common parts. Consequently, some failure modes, such as those for brakes and radios, will be common to all eight vehicles. Other failure modes will be related, but not identical, as with hydraulics, frame, and body. There are 1412 observed failures attributed

¹The exact mileage has been altered to protect proprietary information.

to 26 failure modes recorded across all eight vehicles and all three test phases. In any given phase of test, every vehicle is not guaranteed to have a failure of all 26 failure modes, and each test phase does not necessarily end with a failure. Therefore, we have 624 right censored observations, which account for miles driven after the last observed failures.

3 Methodology

A typical reliability analysis employed by the Department of Defense (DoD) test community would consider each test phase independently and use the exponential distribution to model the miles between failure for the system [1]. For many situations, including our example, this traditional analysis is overly simplistic, relying on incorrect modeling assumptions and ignoring valuable information learned about the individual vehicles and their failure modes through multiple phases of testing.

We propose an alternative approach using a Bayesian hierarchical model. The methodology section is organized as follows. In Section 3.1, we introduce a general structure that models the relationships in the data across all test phases, while accounting for CAPs, and incorporates known similarities between vehicle failure modes. Next, we will consider different distributional assumptions and their implications. Using the JLTV as a case study we will illustrate the process of model selection using three different diagnostic checks. In Section 3.2 we discuss methods of designing next stage test plans. First we outline the traditional approaches, then we propose a Bayesian approach that allows the practitioner to leverage the information from previous test phases, using the models developed from Section 3.1, with the hopes of reducing testing resources required while minimizing producer and consumer risks.

3.1 Modeling Reliability

In this reliability setting, we are analyzing time to failure data measured in miles for a multiple component system with the primary objective of assessing the system failure rate and predicting future failure counts. The most common models used in this kind of failure rate analysis are lifetime distributions, and the most common for failure count analysis is the *Poisson process* family of counting models.

3.1.1 Model Structure

For each phase of testing we have failure mileage data for a given vehicle i and failure mode j denoted as $t_{ij0}, t_{ij1}, \dots, t_{ijn_{ij}}, t_{ijc}$, where $t_{ij0} = 0$ and t_{ijc} is the censored observation at the end the testing phase. In each phase of testing, we assume the vehicle miles between failures, $y_{ijk} = t_{ijk} - t_{ijk-1}$, follow a lifetime distribution, $f(\cdot)$, with an unknown set of parameters that include a rate parameter. For the first phase of testing, the rate parameter for vehicle i and failure mode j will be denoted as λ_{ij} . For the next stage, we use the same λ_{ij} but add a multiplicative correction parameter ρ_j^{P2} that represents the change in this rate due to the CAP. This procedure is then repeated for each subsequent stage. The correction parameter will be discussed more in the next section, but the general model is:

$$\text{Phase 1: } y_{ijk} \mid \lambda_{ij} \sim f(\cdot)$$

$$\text{Phase 2: } y_{ijk} \mid \lambda_{ij} \rho_j^{P2} \sim f(\cdot)$$

$$\text{Phase 3: } y_{ijk} \mid \lambda_{ij} \rho_j^{P2} \rho_j^{P3} \sim f(\cdot)$$

$$i = 1, 2, \dots, v \quad j = 1, 2, \dots, s \quad k = 1, 2, \dots, n_{ij}$$

where v represents the number of vehicles, s is the number of failure modes, and n_{ij} are the number of failures of vehicle i failure mode j . The number of vehicles v and number of failure modes s are considered fixed and the same across all phases.

For the model we assume a prior distribution on the failure rate parameter λ_{ij} . This distribution depends on whether failure mode j is considered to be common across vehicles or related but not identical. If failure mode j is considered common across vehicles, we will assume there is a single failure rate λ_j that is shared by all eight vehicles with a prior distribution that has positive support. In the case study, we will use a non-informative gamma distribution. For failure modes that are considered related across vehicles, but that are not expected to be exactly the same, we place a prior distribution on the collection of λ_{ij} ; in other words, we assume each of the vehicles has a distinct failure rate for failure mode j but they arise from a common distribution. This is the common Bayesian hierarchical modeling structure, where we then place hyperprior distributions on the shared distribution's parameters. In the case study, we use a gamma distribution with non-informative hyperpriors for all related failure modes. Many times in practice it is difficult to know ahead of time if individual components should be modeled as common or related. In the case study, we fit models for each component with both model structures and use diagnostic tools that are discussed in Section 3.1.4 to make our final decision.

3.1.2 Fix Effectiveness

After the first CAP, the second test phase begins with repaired vehicles. To plan for the reliability improvement caused by these repairs, the PM2 reliability growth model [2] is often used. This model explicitly captures testing phases, choices about which failure modes to correct, and the potential of not completely eliminating a failure upon repair. One of the features of PM2 is that many parameters of potential interest, such as the Fix Effectiveness Factor (FEF), which measures how much repairs improve failure rates, are typically fixed, as PM2 is often used as a planning model. A common value for FEF is 0.70. We follow the premise of this type of model, but allow a more flexible and data-driven result.

One common assumption in reliability growth modeling is nondecreasing failure rates:

either the fixes were effective or had no effect, but did not degrade the reliability of the family of vehicles. This should generally be the case, but because we are dealing with complex systems and testing conditions are rarely completely consistent, we will sometimes see decreases in failure rates after repairs are made. Thus we choose a prior distribution for the FEF parameter ρ_j^{P2} that has positive support. If ρ_j^{P2} is less than one, this represents an improvement in reliability. The same approach is applied to ρ_j^{P3} for the correction period between phases two and three. With this structure we can use the posterior distributions of the correction parameters to make inference on the effectiveness of the repairs and the joint posterior distributions to evaluate the overall system reliability within each phase.

3.1.3 Distributional Assumption

The next modeling choice that must be made is selecting an appropriate probability model. This includes selecting a sampling distribution for the data and prior distributions for each parameter on which the sampling distribution depends. Some common distributional choices for lifetime data are exponential, Weibull, lognormal, gamma, inverse Gaussian, and normal failure time models. Because all of these distributions have a rate parameter (sometimes called a scale parameter, which is the reciprocal of the rate parameter) they can be applied in the proposed modeling structure. We focus on the exponential and Weibull distributions, which are commonly used for reliability.

The exponential model has probability density function

$$f(y_{ijk}|\lambda_{ij}) = \lambda_{ij} \exp(-\lambda_{ij}y_{ijk})$$

This is by far the most common parametric distribution used in reliability modeling because of its desirable mathematical properties and straightforward interpretations. One convenient property, because we are assuming independence between failure modes, is that

the overall system failures also follow an exponential distribution with a system failure rate $\lambda_S = \sum_j \lambda_j$. A second, potentially less attractive property, is that the hazard rate is constant over miles. For the case study, this implies that a given component is equally likely to fail in the first mile of testing as it is in the thousandth. This may be reasonable for some components, but in many settings we observe either a decreasing hazard rate, where failures are likely to occur early in testing but less likely as the test goes on, or an increasing hazard rate, where we see the system wear down as the test progresses.

Despite its common use, the assumption of a constant hazard rate over time is rarely justifiable. The issues that arise when the modeling assumptions for the exponential distribution are violated are well-documented [3]. Therefore, we will also consider more flexible distributions when analyzing the data. In our analysis of the case study, we also fit the same hierarchical model structure using the Weibull as our lifetime distribution, $f(\cdot)$. We assume that the failure mileage t_{ijk} each follow a *Weibull*(λ, γ) distribution with scale parameter λ and shape parameter γ , with probability density function,

$$f(t_{ijk}|\lambda_{ij}, \gamma_i) = \lambda_{ij} \gamma_i t_{ijk}^{\gamma_i-1} \exp(-\lambda_{ij} t_{ijk}^{\gamma_i}).$$

The exponential is a special case of the Weibull, when $\gamma = 1$.

When a failure occurs during a testing phase, the mileage is recorded and the system is restored to full operation, often by repairing or replacing a component. One must consider how these changes will affect the system reliability moving forward. In repairable system literature, two common modeling assumptions are the extreme cases “good as new” or “bad as old”. “Good as new” assumes that a repaired component has returned to a brand-new state or, in our case, that the mileage is set to zero. On the other hand, “bad as old” assumes the component has returned to the state it was at before the failure occurred. The Weibull distribution implies we assuming “bad as old” intra-test repairs by modeling the actual

failure mileage t_{ijk} instead of the miles between failures y_{ijk} . This distinction is important in the Weibull case because the failure rate is not constant in time. In the JLTv case study we have assumed “bad as old” intra-test repairs based on input from subject matter experts. There will be applications where repairs should be treated as somewhere between “good as new” and “bad as old”. This requires more complex modeling that is beyond the scope of this paper. An overview of repairable system models can be found in Hamada et al. [4]

We index both ρ and γ only by j and not i . In other words, we are assuming a single shape and between phase adjustment parameter for each failure mode across vehicles. All of these parameters could also be indexed by i and modeled hierarchically or we could use a combination of common and “related”. When failure modes do not occur within the testing period, resulting in a single censored observation, this can lead to convergence issues. In the case study, we found 6 of the 26 failure modes had less than 10 observed (non-censored) failures across all 3 phase and 8 vehicle types. These failure modes were attributed to system components such as the engine and armor. To improve convergence we have binned these rare failure modes into a single “other” failure mode. This is also useful for future testing in the case of observing failures not originally accounted for.

3.1.4 Model Diagnostics

The decision about which model to use can have a substantial impact on reliability assessment. We consider a few goodness-of-fit techniques that can be used to help with this decision for Bayesian reliability models. The first model selection question we consider is the parametric form, exponential versus Weibull. This an example of the classic statistical modeling the trade-off between fit and interpretability. Because the Weibull is a more flexible model, containing the exponential as a special case, it will always fit the data at least as well as the exponential.

The first diagnostic we use to decide between the exponential and Weibull models is to

fit the Weibull model and look at the posterior distributions of the shape parameters to determine whether their support contains 1. From Figure 1, which shows the Weibull model fits for the case study for each failure mode, we see that 1 is not a reasonable value for any of the failure modes, which suggests that the exponential model may not fit this data well.

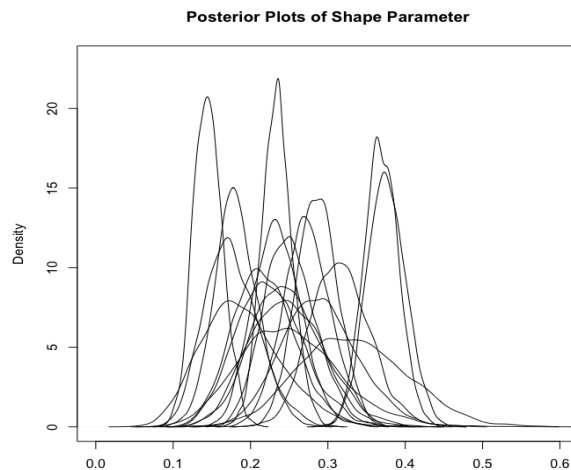


Figure 1: Posterior Plots of Weibull Shape Parameter

We also consider a more formal diagnostic tool, the Deviance information criterion (DIC) [5]. This is a popular method for comparing the goodness of fit of multiple models. The DIC method includes a penalty term that grows with the numbers of unknown parameters in the model. The DIC is unitless, with smaller values indicating a better fit to the data. Table 1 shows the DIC results for the different distribution and structure combinations we considered for the case study. The results show that the hierarchical Weibull model fits the data much better than the exponential or common rate models.

In general, information criteria comparisons do not tell us if the model is truly a good representation of the underlying process – they only compare models. The final model diagnostics method we present is called posterior predictive checking [6]. We start with important features of the data that we would expect to see if the experiment was repeated. These are usually provided by practitioner with knowledge about the random process being

Goodness of Fit

Distribution	Structure	DIC
Exponential	Common Rate Model	21,264.19
	Hierarchical Rate Model	21,244.47
Weibull	Common Rate Model	20,658.25
	Hierarchical Rate Model	17,908.06

Table 1: JLTV DIC Results

studied. In the case study, one feature of interest is the count of miles between failure observations less than 140 miles, for a given vehicle and phase.

We used the final exponential and Weibull models with the lowest DIC to simulate 5,000 new datasets for vehicles tested over the same number of miles for each of the three phases. We plot the resulting distributions for each of the 8 vehicles for all 3 phases. Figure 2 shows two examples of the boxplots produced for one of the vehicle types, one plot from the exponential model (left) and one plot from the Weibull model (right), with the dashed line showing where the value from the observed dataset fell. If the model fits reasonably well, we expect to see the value from the observed data falling within the support of the simulated data. This suggests that the Weibull model fits the case study data better than the exponential data.

3.1.5 Case Study Results

We have introduced several models structures and distributional assumptions that accompany them. Using Markov chain Monte Carlo (MCMC), we fit each of these models to the JLTV case study dataset. (The likelihoods and sampling distributions are available in the appendix.) We use the posterior distributions to estimate reliability-related quantities of interest for each of the eight vehicle types. We are able to examine a number of different summaries related to reliability at both the component and system level.

In this study, one quantity of interest is mean miles to failure (MMTF). This is different

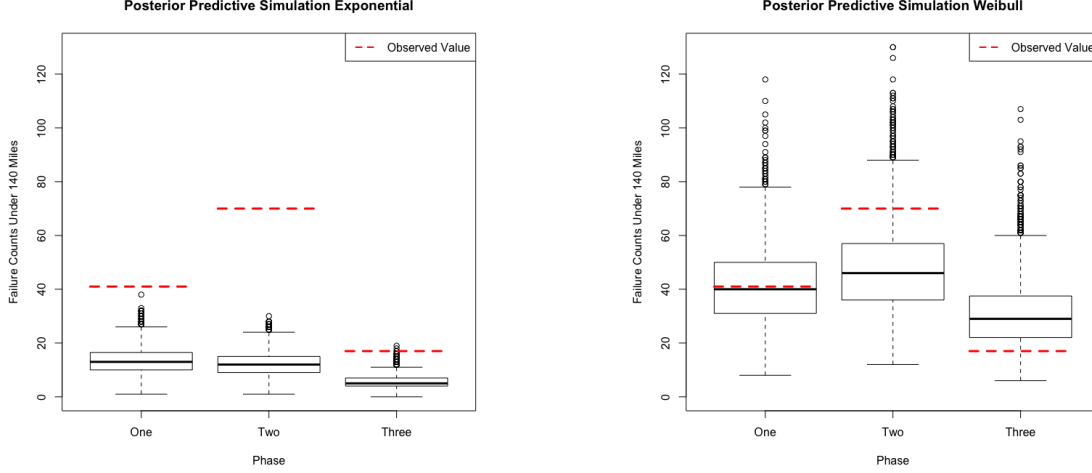
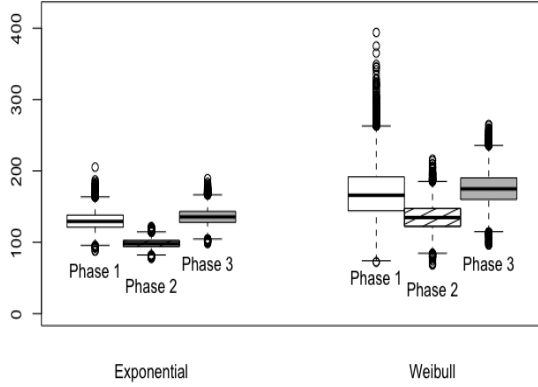


Figure 2: Posterior Predictive Plots
Exponential vs. Weibull

than the commonly used mean miles between failures (MMBF). Due to the failure rate changing over time in the Weibull case MMBF depends on how many miles the vehicle is driven. Instead MMTF can be thought of as the average time until the first failure occurs. Figure 3 shows box plots and summary statistics from the posterior distributions of the expected miles to failure for one of the vehicle types for each of the three phases using the exponential model (left) and the more general Weibull model (right).

Figure 3 shows that the posterior MMTF distributions from the Weibull model consistently have wider uncertainty intervals and larger means and medians than the exponential. We saw this across all eight vehicle types. This is a good illustration of how the choice of model can impact the final inference.

Figure 4 shows the posterior distributions for each of the CAP parameters ρ_j^{P2} and ρ_j^{P3} for the hierarchical Weibull model. These parameters were introduced in Section 3.1.2. They are multiplicative parameters added in between phases to capture the change in failure rate due to improvements by the producer. Values less than one indicate a decrease in failure rate. The distributions look reasonable with most of the mass close to the commonly used



Summary Statistics

Model	Phase	Mean	90% CI Low	90% CI Up
Expo	P1	129.98	110.22	152.23
Expo	P2	98.44	89.00	108.38
Expo	P3	135.89	117.78	155.45
Weibull	P1	169.96	116.95	236.16
Weibull	P2	135.24	104.89	167.48
Weibull	P3	175.51	138.87	214.23

Figure 3: Comparing MMTF for Different Models

FEF value of 0.70. We notice a trend of more improvement in the second CAP between phase two and three compared to the first CAP between phase one and two.

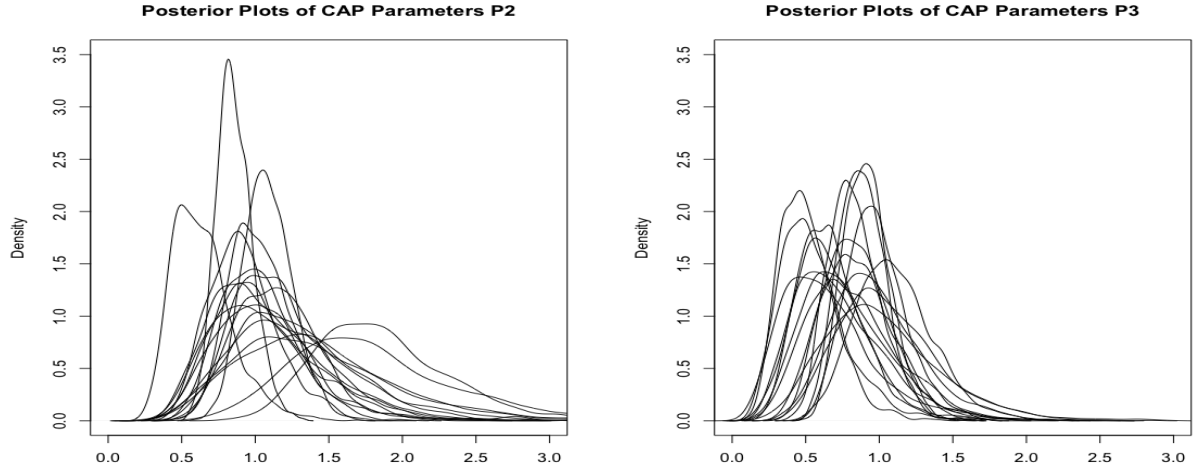


Figure 4: Posterior Distributions for Correction Parameters ρ_j^{P2} and ρ_j^{P3}

3.2 Assurance Test Planning

In addition to fitting the data from the three phases of testing, we also want to use the reliability models to develop plans for future testing. The objective is to demonstrate, at a desired level of confidence, that the system will meet or exceed a specified requirement. The methods to be employed will be similar to the assurance testing as discussed by Hamada et al. [4]. Meeker and Escobar [7] describe the difference between assurance testing and the more traditional *reliability demonstration test*: an assurance test plan incorporates supplementary data and information, where the demonstration test uses only data from the current test.

In most test planning settings we are interested in controlling error rates while minimizing the resources required for testing. In a typical DoD acquisition process, the error rates are referenced in terms of risk. *Consumer's risk* is the probability of purchasing a product that does not meet reliability requirements. *Producer's risk* is the probability that a product with acceptable reliability fails a given test.

In the case study, we determine how many miles on test T , are required and how many system failures to allow c , before the product is considered unacceptable. We define $W(t)$ as a random variable that represents the total number of system failures after t miles. We let π denote a system reliability-related quantity of interest at a given time.

It is common to base both classical and Bayesian test plans on two specified levels of π : π_0 , an *acceptable reliability level* (ARL), and π_1 , a *rejectable reliability level* (RRL), where $\pi_1 \leq \pi_0$. The precise definition of ARL and RRL differ between the classical and Bayesian test criteria and will be described below.

3.2.1 Classical Approach

In the classical (frequentist) test planning setting the risk criteria are as follows. The *producer's risk* is defined as the probability of failing the test under the null hypothesis $\pi = \pi_0$, and the *consumer's risk* is defined as the probability of passing the test under an alternative

hypothesis $\pi = \pi_1$. Suppose that we specify a maximum value, α_c , of the producer's risk and a maximum value, β_c , of the consumer's risk. For test planning, these criteria become

$$\begin{aligned}\text{Producer's Risk} &= P(\text{Test Is Failed} \mid \pi_0) \\ &= P(W(t) > c \mid \pi_0) \leq \alpha_c\end{aligned}$$

and

$$\begin{aligned}\text{Consumer's Risk} &= P(\text{Test Is Passed} \mid \pi_1) \\ &= P(W(t) \leq c \mid \pi_1) \leq \beta_c\end{aligned}$$

To choose a test plan for specified values of $(\alpha, \pi_0, \beta, \pi_1)$, we assume a distributional form that defines the relationship between the number of system failures $W(t)$ and the reliability π . We then find the combinations of time on test and allowed failures that satisfy the risk constraints by simultaneously solving these two equations. In the case where the component failure times are assumed to be exponentially distributed with rate parameter λ_i , the failure counts for each component come from a homogeneous Poisson process $w_i(t) \sim \text{Poisson}(\lambda_i t)$, and the system failure counts from $W(t) \sim \text{Poisson}(\lambda_S t)$. Numerous textbooks provide additional details of this classical test planning approach. [8]

3.2.2 Bayesian Approach

The non-Bayesian methods for determining the testing procedure rely only on distributional assumptions or asymptotic results. With a Bayesian approach, we incorporate the supplementary data from the previous testing phases with the hopes of minimizing the resources needed for additional testing.

We consider fully Bayesian posterior risks, which provides a different interpretation from the corresponding classical risks. While the classical approach provides assurance that satisfactory devices will pass the test and that unsatisfactory devices will fail it, posterior risks provide precisely the assurance that practitioners often desire: if the test is passed, then the

consumer desires a maximum probability β_b that the true reliability is less than the rejectable level, $\pi \leq \pi_1$. On the other hand, if the test is failed, then the producer desires a maximum probability α_b the true reliability is greater than the acceptable level, that $\pi \geq \pi_0$. Notice these posterior risks are fully Bayesian in the sense that they are probability statements about π .

For a test that fails, the *posterior producer's risk* is the probability that $\pi \geq \pi_0$, or $P(\pi \geq \pi_0 | \text{Test Is Failed})$. Notice that this is simply the posterior probability that $\pi \geq \pi_0$ given that we have observed more than c failures. In the exponential case, if we let π be the system failure rate λ_S and π_0 be predetermined acceptable failure rate λ_0 , then using Bayes' Theorem, and assuming a maximum allowable posterior producer's risk α_b , an expression for the posterior producer's risk for the exponential test plan (T, c) is

$$\begin{aligned}
P(\lambda_S \geq \lambda_0 \mid \text{Test Is Failed}) &= P(\lambda_S \geq \lambda_0 \mid W > c) \\
&= \int_{\lambda_0}^{\infty} p(\lambda_S \mid W > c) d\lambda_S \\
&= \int_{\lambda_0}^{\infty} \frac{f(W > c \mid \lambda_S)}{\int_0^{\infty} f(W > c \mid \lambda_S) d\lambda_S} d\lambda_S \\
&= \frac{\int_{\lambda_0}^{\infty} 1 - [\sum_{W=0}^c \frac{(\lambda_S T)^W \exp(-\lambda_S T)}{W!}] p(\lambda_S) d\lambda_S}{\int_0^{\infty} 1 - [\sum_{W=0}^c \frac{(\lambda_S T)^W \exp(-\lambda_S T)}{W!}] p(\lambda_S) d\lambda_S} \leq \alpha_b
\end{aligned}$$

For simplicity, if we set $c = 0$ and then perform Monte Carlo integration using N posterior draws $\lambda_S^{(j)}$

$$\begin{aligned}
P(\lambda_S \geq \lambda_0 \mid W = 0) &= \frac{\int_{\lambda_0}^{\infty} 1 - \exp(-\lambda_S T) p(\lambda_S) d\lambda_S}{\int_0^{\infty} 1 - \exp(-\lambda_S T) p(\lambda_S) d\lambda_S} \\
&\approx \frac{\sum_{j=1}^N 1 - \exp(-\lambda_S^{(j)} T) I(\lambda_S^{(j)} \geq \lambda_0)}{\sum_{j=1}^N 1 - \exp(-\lambda_S^{(j)} T)}
\end{aligned}$$

Similarly, given that the test is passed, the *posterior consumer's risk* is the probability

that $\pi \leq \pi_1$, or $P(\pi \leq \pi_0 | \text{Test Is Passed})$. Let π_1 be a predetermined rejectable failure rate λ_1 . Assuming a maximum allowable posterior consumer's risk β_b , an expression for the posterior producer's risk for the exponential test plan (T, c) is

$$\begin{aligned}
P(\lambda_S \leq \lambda_1 \mid \text{Test Is Passed}) &= P(\lambda_S \leq \lambda_1 \mid W \leq c) \\
&= \int_0^{\lambda_1} p(\lambda_S \mid W > c) d\lambda_S \\
&= \int_0^{\lambda_1} \frac{f(W > c \mid \lambda_S)}{\int_0^\infty f(W > c \mid \lambda_S) d\lambda_S} d\lambda_S \\
&= \frac{\int_0^{\lambda_1} [\sum_{W=0}^c \frac{(\lambda_S T)^W \exp(-\lambda_S T)}{W!}] p(\lambda_S) d\lambda_S}{\int_0^\infty [\sum_{W=0}^c \frac{(\lambda_S T)^W \exp(-\lambda_S T)}{W!}] p(\lambda_S) d\lambda_S} \leq \beta
\end{aligned}$$

For simplicity, if we fix $c = 0$ and then perform Monte Carlo integration using N posterior draws $\lambda_S^{(j)}$

$$\begin{aligned}
P(\lambda_S \leq \lambda_1 \mid W = 0) &= \frac{\int_0^{\lambda_1} \exp(-\lambda_S T) p(\lambda_S) d\lambda_S}{\int_0^\infty \exp(-\lambda_S T) p(\lambda_S) d\lambda_S} \\
&\approx \frac{\sum_{j=1}^N \exp(-\lambda_S^{(j)} T) I(\lambda_S^{(j)} \leq \lambda_1)}{\sum_{j=1}^N \exp(-\lambda_S^{(j)} T)}
\end{aligned}$$

3.2.3 Weibull Case

In the exponential case, we defined our reliability quantity of interest π as the system failure rate λ_S . Assuming an exponential lifetime distribution gave us a convenient distributional form for our system failure counts with $W(t) \sim \text{Poisson}(\lambda_S t)$.

When we assume the miles to failure for each system component follow a Weibull distribution, the system failure counts no longer follow a homogeneous Poisson process. When assuming “bad as old” intra-test repairs, as described in Section 3.1.3, we find that the failure count for a given failure mode follows a nonhomogeneous Poisson process, $w_i(t) \sim \text{Poisson}(\lambda_i t_i^{\gamma_i})$.

The system failure counts for a given mileage t have distribution $W(t) \sim \text{Poisson}(\sum_i (\lambda_i t_i)^{\gamma_i})$. In this case, the acceptable and rejectable parameters π_0 and π_1 become quantities related to $\lambda^* = \sum_i (\lambda_i t_i)^{\gamma_i}$. This leads to integration that is analytically intractable, but Monte Carlo integration allows our evaluation of the posterior producer's risk for the Weibull test plan (T, c) :

$$\begin{aligned} P(\lambda^* \geq \pi_0 \mid W = 0) &= \frac{\int_{\pi_0}^{\infty} 1 - \exp(-\sum_i \gamma_i \lambda_i T^{\gamma_i}) p(\lambda_{1,\dots}) p(\gamma_{1,\dots}) d\lambda_{1,\dots} d\gamma_{1,\dots}}{\int_0^{\infty} 1 - \exp(-\sum_i \gamma_i \lambda_i T^{\gamma_i}) p(\lambda_{1,\dots}) p(\gamma_{1,\dots}) d\lambda_{1,\dots} d\gamma_{1,\dots}} \\ &\approx \frac{\sum_{j=1}^N 1 - \exp(-\sum_i \gamma_i^{(j)} \lambda_i^{(j)} T^{\gamma_i^{(j)}}) I(\lambda^* \geq \pi_0)}{\sum_{j=1}^N 1 - \exp(-\sum_i \gamma_i^{(j)} \lambda_i^{(j)} T^{\gamma_i^{(j)}})} \end{aligned}$$

Again we set $c = 0$ and use N posterior draws for both $\lambda_i^{(j)}$ and $\gamma_i^{(j)}$.

It should be noted that in this test planning process, the practitioner would use the posterior distribution of the rate parameters from the final developmental testing phase $\lambda_i \rho^{P^2} \rho^{P^3}$. It is also common to incorporate a degradation factor to account for an expected reduction in reliability from the developmental phase to the operational phase. In the Bayesian framework, this entails adding another multiplicative parameter to the rate parameter with an appropriate prior. A 10-30 percent reduction from the developmental phase to the operational phase would be a common assumption in many DoD applications.

3.2.4 Case Study Test Plan Results

The first step in applying this method to our case study is defining π_0 , π_1 , α_b and β_b , which are the ARL, RRL, acceptable producer risk and acceptable consumer risk respectively. With input from subject matter experts, we will define the producer's target as the average number of failures in the first 10,000 miles to be less than 155, and the consumer's minimum requirements as the average number of failures in the first 10,000 miles to be greater than 175. The producer's targets and consumer's minimum requirements represent our cut off levels

ARL and RRL respectively. This means if the true average number of failures in the first 10,000 miles is less than 155 the producer would expect his product to fail the test a small proportion of the time, α_b . On the other hand, if the true average number of failures in the first 10,000 miles is greater than 175 the consumer would expect the product to pass the test a small proportion of the time, β_b . We will use $\alpha_b = 0.1$ and $\beta_b = 0.05$, and in practice these values should always be set based on the true risk tolerances of the consumer and producers. This gives us the following probability statements to satisfy when designing the test:

$$\text{Consumer Risk : } P(\sum_i 10,000^{\gamma_i} \lambda_i \geq 175 \mid \text{Test Is Passed}) \leq 0.1$$

$$\text{Producer Risk : } P(\sum_i 10,000^{\gamma_i} \lambda_i \leq 155 \mid \text{Test Is Failed}) \leq 0.05$$

From here we are able to find combinations of failures allowed and test miles required that satisfy both risk requirements.

Test Plans		
Failures Allowed	Assurance Test Plan Miles Required $\alpha_b = 0.1, \beta_b = 0.05$	Classical Test Plan Miles Required $\alpha_c = 0.1, \beta_c = ?$
1	240	225 ($\beta_c = 0.86$)
2	250	305 ($\beta_c = 0.85$)
3	260	385 ($\beta_c = 0.84$)
4	275	460 ($\beta_c = 0.83$)
5	290	530 ($\beta_c = 0.82$)

Table 2: Test Plan Comparison

Table 2 compares the resulting test plans of our proposed method to the classical method. In the classical case we only controlled for the consumer's risk, because with the proposed values of π_0, π_1, α_c and allowed failures there was no test mileage that would satisfy both risk criteria. Notice this leads to extremely high levels of producer risk under the classical plan. In general, we warn against making this comparison of test plans, because each method makes different assumptions and the test plan probability constraints are asking different

questions. The classical plan assumes product is either acceptable or unacceptable and then controls the probability of error in each of those cases. On the other hand, the Bayesian plan puts constraints on the probability of error when the test is passed and failed.

We have provided one example where there was no solution that satisfied all of the constraints using the classical method. This problem does not only come up using the classical method. Both test planning approaches rely heavily on setting reasonable values for π_0 , π_1 , α and β . When these are set to unrealistic values both procedures can result in no plan that satisfies the probability statements. [9]

In this paper we have focused on test plans for single units, but the assurance testing methodology can be extended to find test plans that allow for a changing number of testing units. In this case you would find combinations of failures allowed and test miles required and number of vehicles tested that satisfy the risk requirements. More details and examples of this can be found in Hamada et al. [4]

3.3 Discussion

The JLTV results provide indications that our model is successfully capturing the failure rates across failure modes, inferring logical fix effectiveness factor distributions, and providing reliability estimates that can be leveraged for future test planning. In the traditional analysis, we do not typically look at failure rates of the failure modes, nor would we be able to leverage the commonalities across vehicles. The proposed modeling framework overcomes these limitations and allows for more realistic reliability estimates for vehicles with no observed failures. Additionally, our data-driven approach for reliability growth improves on the current growth models being used in the DoD, which are based solely on fixed quantities set by management.

While our basic approach provides reasonable initial inferences, there are several extensions to incorporate. For example, reliability estimates for vehicles with no observed failures

are sensible; however, we should leverage more information across vehicles. Because our current model assumes independence across failure modes, potential methods that model the correlation structure between failure modes are being investigated. We also need to look more closely at the intra-phase repair assumptions. Applying repairable system methodology in the modeling and reliability assessment stage could be a good approach, but this could add a great deal of complexity to the assurance planning stage.

References

- [1] Director, Defense Test and Evaluation. Test and evaluation of system reliability availability maintainability - A primer; Third Edition 1982.
- [2] P. M. Ellner, J. B. Hall. An approach to reliability growth planning based on failure mode discovery and correction using AMSAA projection methodology. *IEEE Proceedings of the Annual Reliability and Maintainability Symposium* 2006.
- [3] National Research Council. Statistics, Testing, and Defense Acquisition: Background Papers. *National Academies Press* 1999.
- [4] M. Hamada, A. Wilson, C. Reese, H. Martz. Bayesian Reliability. Springer 2008.
- [5] D. Spiegelhalter, N. Best, B. Carlin, A. Van der Linde Bayesian measures of model complexity and fit. *Royal Statistical Society, Series B*, 64: 583-640 2002.
- [6] A. Gelman, M. Xiao-Li, H. Stern. Posterior predictive assessment of model fitness via realized discrepancies (with discussion). *Statist. Sin.* 6, 733-807 1996.
- [7] W. Meeker, L. Escobar. Reliability: the other dimension of quality. *Quality Technology and Quantitative Management* 1, 1-25 2004.

- [8] J. F. Lawless. Statistical Models and Methods for Lifetime Data, 2nd Edition. John Wiley & Sons 2011.
- [9] B. Weaver, M. Hamada, A. Wilson, J. Bakerman. Bayesian Assurance Tests for Degradation Data. *Quality and Reliability Engineering International* 33(8): 2699-2709 2017.

4 Appendix

Exponential Model:

$$L(\lambda_1, \dots, \lambda_8, \rho^{P2}, \rho^{P3}, |Y) = \prod_{i=1}^8 \left[\prod_{j=1}^{n_j} (\lambda_i p^* e^{-\lambda_i p^* y_{ij}}) \prod_{j=1}^R (e^{-\lambda_i p^* y_{ij}}) \left(\frac{\beta^\alpha}{\Gamma(\alpha)} \lambda_i^{\alpha-1} e^{-\beta \lambda_i} \right) \right]$$

$$p^* = \begin{cases} 1 & y_{ij} \in \text{Phase 1} \\ \rho_j^{P2} & y_{ij} \in \text{Phase 2} \\ \rho_j^{P2} \rho_j^{P3} & y_{ij} \in \text{Phase} \end{cases}$$

R is the number of right-censored observations, and n is the number of non-censored observations.

$$\pi(\alpha, \beta, \rho^{P2}, \rho^{P3}) \propto (\alpha^{a_1-1} e^{-\alpha a_2}) (\beta^{b_1-1} e^{-\beta b_2}) (\rho^{P2*c_1-1} e^{-\rho^{P2} c_2}) (\rho^{P3*d_1-1} e^{-\rho^{P3} d_2})$$

Full conditionals for Gibbs sampling:

$$\lambda_i | \cdot \sim \text{Gamma}(n_i + \alpha, \sum_{j=1}^{n_i} y_{ij} p^* + \beta)$$

$$\beta | \cdot \sim \text{Gamma}(8\alpha + b_1, \sum \lambda_i + b_2)$$

$$\rho^{P2} | \cdot \sim \text{Gamma}(N_{1,2} + c_1, \sum y_{ij} \lambda_i + c_2)$$

$$\rho^{P3} | \cdot \sim \text{Gamma}(N_2 + d_1, \sum y_{ij} \lambda_i + d_2)$$

$$\alpha | \cdot \sim \prod_{i=1}^8 \left[\prod_{j=1}^R \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda_i^{\alpha-1} \right] (\alpha^{a_1-1} e^{-\alpha a_2}) (\text{Metropolis-Hastings used for sampling})$$

N_1 and N_2 are the total non-censored observations for phases two and three and phase three, respectively.

Weibull Model (Bad as Old):

$$L(\lambda_1, \dots, \lambda_8, \gamma, \rho^{P2}, \rho^{P3}, |T) = \prod_{i=1}^8 \left[\prod_{j=1}^{n_j} (\lambda_i p^* t_{ij}^{\gamma_i-1} e^{-\lambda_i p^* t_{ij}^{\gamma_i}}) \prod_{j=1}^R (e^{-\lambda_i p^* t_{ij}^{\gamma_i}}) \left(\frac{\beta^\alpha}{\Gamma(\alpha)} \lambda_i^{\alpha-1} e^{-\beta \lambda_i} \right) \right]$$

$$p^* = \begin{cases} 1 & t_{ij} \in \text{Phase 1} \\ \rho^{P2\gamma} & t_{ij} \in \text{Phase 2} \\ \rho^{P2\gamma} \rho^{P3\gamma} & t_{ij} \in \text{Phase 3} \end{cases}$$

R is the number of right-censored observations, and n is the number of non-censored observations.

$$\pi(\alpha, \beta, \rho^{P2}, \rho^{P3}) \propto (\alpha^{a_1-1} e^{-\alpha a_2}) (\beta^{b_1-1} e^{-\beta b_2}) (\rho^{P2*c_1-1} e^{-\rho^{P2} c_2}) (\rho^{P3*d_1-1} e^{-\rho^{P3} d_2}) (\gamma^{g_1-1} e^{-\gamma g_2})$$

Full conditionals for Gibbs sampling:

$$\lambda_i | \cdot \sim \text{Gamma}(n_i + \alpha, \sum_{j=1}^{n_i} t_{ij}^{\gamma_i} p^* + \beta)$$

$$\beta | \cdot \sim \text{Gamma}(8\alpha + b_1, \sum \lambda_i + b_2)$$

$$\rho^{P^2} | \cdot \sim \text{Gamma}(N_{1,2} + c_1, \sum t_{ij}^{\gamma_i} \lambda_i + c_2)$$

$$\rho^{P^3} | \cdot \sim \text{Gamma}(N_2 + d_1, \sum t_{ij}^{\gamma_i} \lambda_i + d_2)$$

$$\alpha | \cdot \sim \prod_{i=1}^8 \left[\prod_{j=1}^R \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda_i^{\alpha-1} \right] (\alpha^{a_1-1} e^{-\alpha a_2}) (\text{Metropolis-Hastings used for sampling})$$

$$\gamma | \cdot \sim \prod_{i=1}^8 \left[\prod_{j=1}^{n_j} (t_{ij}^{\gamma_i-1} e^{-t_{ij}^{\gamma_i}}) \prod_{j=1}^R (e^{-t_{ij}^{\gamma_i}}) \right] (\gamma^{g_1-1} e^{-\gamma g_2}) (\text{Metropolis-Hastings used for sampling})$$

N_1 and N_2 are the total non-censored observations for phases two and three and phase three, respectively.