STAT 542 / CS 598: Homework 1

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Due: Monday, Sep 9 by 11:59 PM Pacific Time

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Question 1 [50 Points] KNN

Write an R function to fit a KNN regression model. Complete the following steps

a. [15 Points] Write a function myknn(xtest, xtrain, ytrain, k) that fits a KNN model that predict a target point or multiple target points xtest. Here xtrain is the training dataset covariate value, ytrain is the training data outcome, and k is the number of nearest neighbors. Use the ℓ_2 norm to evaluate the distance between two points. Please note that you cannot use any additional R package within this function.

```
# euclidean distance function from scratch
euc.dist <- function(x1, x2) {</pre>
  sqrt(sum((x1 - x2) ^ 2))
# myknn function
#-----
myknn <- function(xtest, xtrain, ytrain, k) {</pre>
  ypred <- rep(0, nrow(xtest))</pre>
  for(i in c(1:nrow(xtest))) {
    #calculate the distance from the given row instance of the test data to each
    #row in the training data
    distances <- apply(xtrain, 1, euc.dist, x2=xtest[i,])</pre>
    evalframe <- cbind(xtrain, distances, ytrain)</pre>
    orderedframe <- evalframe[order(evalframe[,6]),c(6,7)]</pre>
    pred = mean(orderedframe[1:k,2])
    ypred[i] <- pred</pre>
  return(ypred)
}
```

b. [10 Points] Generate 1000 observations from a five-dimensional normally distribution:

$$\mathcal{N}(\mu, \Sigma_{5\times 5})$$

where $\mu = (1, 2, 3, 4, 5)^{\mathrm{T}}$ and $\Sigma_{5\times5}$ is an autoregressive covariance matrix, with the (i, j)th entry equal to $0.5^{|i-j|}$. Then, generate outcome values Y based on the linear model

$$Y = X_1 + X_2 + (X_3 - 2.5)^2 + \epsilon$$

where ϵ follows i.i.d. standard normal distribution. Use set.seed(1) right before you generate this entire data. Print the first 3 entries of your data.

```
library(MASS)
set.seed(1)
mu <- c(1,2,3,4,5)
sigma <- matrix(0, nrow=5, ncol=5)
sigma <- outer(1:5, 1:5, function(i, j) 0.5^abs(i-j))

xdata <- mvrnorm(1000, mu=mu, Sigma=sigma)

eps = rnorm(1000)
Y <- matrix(0,nrow=1000, ncol=1)
Y <- xdata[,1] + xdata[,2] + (xdata[,3] - 2.5)^2 + eps

printdata <- cbind(xdata,Y)
colnames(printdata) <- c("X1", "X2", "X3", "X4", "X5", "Y")

printdata[1:3,]</pre>
```

```
## X1 X2 X3 X4 X5 Y
## [1,] 2.0770490 3.555163 2.641969 3.902436 5.108741 4.135994
## [2,] 2.4780195 2.161175 2.188487 2.796376 5.330744 5.365376
## [3,] -0.1413538 2.630428 4.666608 4.493909 5.698190 5.505069
```

c. [10 Points] Use the first 400 observations of your data as the training data and the rest as testing data. Predict the Y values using your KNN function with k=5. Evaluate the prediction accuracy using mean squared error

$$\frac{1}{N}\sum_{i}(y_i-\widehat{y}_i)^2$$

```
# mean square error function
#------
mse = function(actual, predicted) {
  mean((actual - predicted) ^ 2)
}

xtrain <- xdata[1:400,]
xtest <- xdata[401:1000,]
ytrain <- Y[1:400]
ytest <- Y[401:1000]</pre>

ypred <- myknn(xtest, xtrain, ytrain, 5)
```

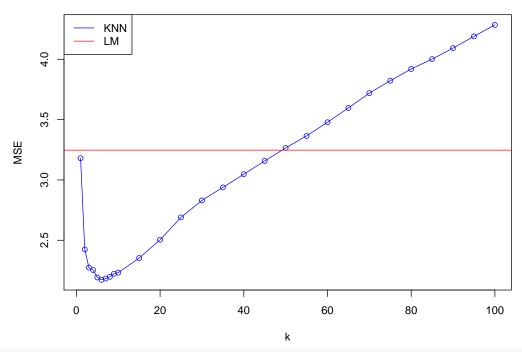
```
MSE_k5 <- mse(ytest,ypred)
print(sprintf("MSE for k=5: %f", MSE_k5))</pre>
```

[1] "MSE for k=5: 2.191387"

d. [15 Points] Compare the prediction error of a linear model with your KNN model. Consider k being 1, 2, 3, ..., 9, 10, 15, 20, ..., 95, 100. Demonstrate all results in a single, easily interpretable figure with proper legends.

```
k \leftarrow c(1:9,seq(10,100,by=5))
n <- length(k)
knn.MSE <- rep(0,n)</pre>
for (i in 1:n) {
  ypred <- myknn(xtest, xtrain, ytrain, k[i])</pre>
  knn.MSE[i] <- mse(ytest, ypred)</pre>
# calculate linear model (LM)
xtrain.df = as.data.frame(xtrain)
xtest.df = as.data.frame(xtest)
colnames(xtest.df) <- c("X1", "X2", "X3", "X4", "X5")</pre>
ytrain.df = as.data.frame(ytrain)
master.df = cbind(xtrain.df, ytrain.df)
colnames(master.df) <- c("X1", "X2", "X3", "X4", "X5", "Y")</pre>
model.lm = lm(Y \sim ., data=master.df)
pred.lm = predict(model.lm, xtest.df)
lm.MSE <- mse(ytest, pred.lm)</pre>
plot(k,knn.MSE, type="o", col="blue", xlab="k", ylab="MSE")
title("KNN Regression", line=3, font.main=1)
legend("topleft", lty = c(1,1), col = c("blue", "red"),
       legend = c("KNN", "LM"))
abline(lm.MSE, 0, col="red")
```

KNN Regression



print(sprintf("lm.MSE: %f", lm.MSE))

[1] "lm.MSE: 3.247147"

Alternatively for this plot, since the LM MSE is a single Y axis point with no valid k to plot for the x, instead of using a straight line across all k for the LM MSE, I could have modified the X-axis to be degree of freedom (N/k) and then plotted the LM MSE as a single point on the graph at degree of freedom = 5, since there are 5 parameters in the linear model.

Question 2 [50 Points] Linear Regression through Optimization

Linear regression is most popular statistical model, and the core technique for solving a linear regression is simply inverting a matrix:

$$\widehat{\boldsymbol{\beta}} = \left(\mathbf{X}^{\mathrm{T}} \mathbf{X} \right)^{-1} \mathbf{X}^{\mathrm{T}} \mathbf{y}$$

However, lets consider alternative approaches to solve linear regression through optimization. We use a gradient descent approach. We know that $\hat{\beta}$ can also be expressed as

$$\widehat{\boldsymbol{\beta}} = \arg\min \ell(\boldsymbol{\beta}) = \arg\min \frac{1}{2n} \sum_{i=1}^{n} (y_i - x_i^{\mathrm{T}} \boldsymbol{\beta})^2.$$

And the gradient can be derived

$$\frac{\partial \ell(\boldsymbol{\beta})}{\partial \boldsymbol{\beta}} = -\frac{1}{n} \sum_{i=1}^{n} (y_i - x_i^{\mathrm{T}} \boldsymbol{\beta}) x_i.$$

To perform the optimization, we will first set an initial beta value, say $\beta = 0$ for all entries, then proceed with the updating

$$oldsymbol{eta}^{
m new} = oldsymbol{eta}^{
m old} - rac{\partial \ell(oldsymbol{eta})}{\partial oldsymbol{eta}} imes \delta,$$

where δ is some small constant, say 0.1. We will keep updating the beta values by setting $\boldsymbol{\beta}^{\text{new}}$ as the old value and calcuting a new one until the difference between $\boldsymbol{\beta}^{\text{new}}$ and $\boldsymbol{\beta}^{\text{old}}$ is less than a prespecified threshold ϵ , e.g., $\epsilon = 10^{-6}$. You should also set a maximum number of iterations to prevent excessively long runing time.

a. [35 Points] Based on this description, write your own R function mylm_g(x, y, delta, epsilon, maxitr) to implement this optimization version of linear regression. The output of this function should be a vector of the estimated beta value.

```
library(mlbench)
mylm_g <- function(x, y, delta, epsilon, maxiter) {

beta_old <- matrix(1,nrow=ncol(x), ncol=1)
beta_new <- matrix(1,nrow=ncol(x), ncol=1)
iter <- 1
beta_diff <- 1

while (beta_diff > epsilon && iter <= maxiter) {

beta_new <- beta_old - (delta / nrow(x)) * (t(x) %*% (x %*% beta_old - y))
beta_diff <- euc.dist(beta_old, beta_new)
beta_old <- beta_new

iter <- iter + 1
}

print(sprintf("Delta: %f Epsilon: %.9f took %d iterations", delta, epsilon,iter))
return(beta_new)
}</pre>
```

b. [15 Points] Test this function on the Boston Housing data from the mlbench package. Documentation is provided here if you need a description of the data. We will remove medv, town and tract from the data and use cmedv as the outcome. We will use a scaled and centered version of the data for estimation. Please also note that in this case, you do not need the intercept term. And you should compare your result to the lm() function on the same data. Experiment on different maxitr values to obtain a good solution. However your function should not run more than a few seconds.

```
library(mlbench)
library(knitr)
data(BostonHousing2)
X = BostonHousing2[, !(colnames(BostonHousing2) %in% c("medv", "town", "tract", "cmedv"))]
X = data.matrix(X)
X = scale(X)
Y = as.vector(scale(BostonHousing2$cmedv))

delta <- 0.1
epsilon <- 10^-6
maxiter <- 5000

beta <- mylm_g(X, Y, delta, epsilon, maxiter)</pre>
```

```
## [1] "Delta: 0.100000 Epsilon: 0.000001000 took 1178 iterations"
##
                 [,1]
## lon
          -0.032319596
## lat
          0.030241889
## crim
          -0.097928372
## zn
          0.118259599
## indus
          0.011349842
## chas
          0.071317218
## nox
          -0.199695261
          0.287239327
## rm
## age
          0.007560419
        -0.321040938
## dis
## rad
         0.290749939
## tax
          -0.236413977
## ptratio -0.206799837
## b
       0.091234909
## lstat -0.417969129
model.lm = lm(Y \sim X)
lm.coeffs = model.lm$coefficients[2:16]
lm.coeffs
##
                                                        Xindus
          Xlon
                      Xlat
                                 Xcrim
                                               Xzn
                                                                     Xchas
## -0.032316441 0.030245087 -0.097935969 0.118273098 0.011390378 0.071312253
          Xnox
                       {\tt Xrm}
                                  Xage
                                              Xdis
                                                          Xrad
Xptratio
                        Xb
                                Xlstat
## -0.206804965 0.091235409 -0.417972819
modeldiff <- beta - lm.coeffs</pre>
modeldiff
##
                  [,1]
## lon
         -3.154922e-06
## lat
         -3.198031e-06
          7.597723e-06
## crim
          -1.349918e-05
## zn
## indus
          -4.053649e-05
## chas
          4.964327e-06
## nox
          8.510655e-06
## rm
          6.515886e-06
## age
         -4.432409e-06
## dis
         -1.596258e-06
## rad
          -1.008163e-04
## tax
          1.121778e-04
## ptratio 5.128383e-06
## b
          -4.999005e-07
## lstat
          3.689583e-06
kable(cbind(beta, as.data.frame(lm.coeffs), modeldiff), digits=9, caption="Result Table for Beta and LM
```

beta	lm.coeffs	modeldiff

Table 1: Result Table for Beta and LM $\,$

	beta	lm.coeffs	modeldiff
lon	-0.032319596	-0.032316441	-0.000003155
lat	0.030241889	0.030245087	-0.000003198
crim	-0.097928372	-0.097935969	0.000007598
zn	0.118259599	0.118273098	-0.000013499
indus	0.011349842	0.011390378	-0.000040536
chas	0.071317218	0.071312253	0.000004964
nox	-0.199695261	-0.199703772	0.000008511
rm	0.287239327	0.287232811	0.000006516
age	0.007560419	0.007564852	-0.000004432
dis	-0.321040938	-0.321039342	-0.000001596
rad	0.290749939	0.290850755	-0.000100816
tax	-0.236413977	-0.236526155	0.000112178
ptratio	-0.206799837	-0.206804965	0.000005128
b	0.091234909	0.091235409	-0.000000500
lstat	-0.417969129	-0.417972819	0.000003690

We set a maxiter of 5000 to limit run time, but with epsilon = 10^-6 , the mylm_g function finishes in 1178 iterations.

The table shows the difference in our Beta values and the values found by running the default lm function provided by R.

Next test with a tighter epsilon of 10⁻8. This converges in only 1909 iterations.

```
epsilon <- 10^-8
beta <- mylm_g(X, Y, delta, epsilon, maxiter)</pre>
```

```
## [1] "Delta: 0.100000 Epsilon: 0.000000010 took 1909 iterations"
modeldiff <- beta - lm.coeffs
kable(cbind(beta, as.data.frame(lm.coeffs), modeldiff), digits=9, caption="Result Table for Beta and LM</pre>
```

Table 2: Result Table for Beta and LM

	beta	lm.coeffs	modeldiff
lon	-0.032316472	-0.032316441	-3.100e-08
lat	0.030245055	0.030245087	-3.200e-08
crim	-0.097935893	-0.097935969	7.600e-08
zn	0.118272962	0.118273098	-1.350e-07
indus	0.011389971	0.011390378	-4.070e-07
chas	0.071312303	0.071312253	5.000e-08
nox	-0.199703687	-0.199703772	8.500e-08
rm	0.287232876	0.287232811	6.500 e-08
age	0.007564807	0.007564852	-4.500e-08
dis	-0.321039359	-0.321039342	-1.700e-08
rad	0.290849743	0.290850755	-1.012e-06
tax	-0.236525028	-0.236526155	1.127e-06
ptratio	-0.206804914	-0.206804965	5.100e-08
b	0.091235404	0.091235409	-5.000e-09

	beta	lm.coeffs	modeldiff
lstat	-0.417972782	-0.417972819	3.700e-08

Bonus Question [5 Points] The Non-scaled Version

When we do not scale and center the data matrix (both X and Y), it could be challenging to obtain a good solution. Try this with your code, and comment on what you observed and explain why. Can you think of a way to calculate the beta parameters on the original scale using the solution from the previous question? To earn a full 5 point bonus, you must provide a rigors mathematical derivation and also validate that by comparing it to the lm() function on the original data.

Answer:

When running the mylm_g function on unscaled data, the beta coefficients quickly diverge and become unstable, resulting in near infinite coeffecients after only running a few dozen iterations.

The problem is we are not calculating an intercept (B-zero) and the data scale varies between parameters by over 4 orders of magnitude, so the gradient coefficient calculation becomes unstable and diverges.

We can use the beta parameters from the scaled solution to calculate the proper beta parameters on unscaled data by "unwinding" the operations performed by scaling and calculating an intercept. The code below calculates the unscaled beta coefficients directly from the scaled beta coefficients and compares them to the coefficients calculated from the default R lm library.

The R code is presented in its entirety below since the bonus question was calculated and tested in a separate R file in my RStudio project:

```
library(mlbench)
euc.dist <- function(x1, x2) {</pre>
  sqrt(sum((x1 - x2) ^ 2))
mylm_g <- function(x, y, delta, epsilon, maxiter) {</pre>
  x \leftarrow cbind(rep(1,nrow(x)), x)
  beta_old <- matrix(1,nrow=ncol(x), ncol=1)</pre>
  beta new <- matrix(1,nrow=ncol(x), ncol=1)
  iter <- 1
  beta_diff <- 1
  while (beta_diff > epsilon && iter <= maxiter) {</pre>
    beta_new <- beta_old - (delta / nrow(x)) * (t(x) %*% (x %*% beta_old - y))
    beta_diff <- euc.dist(beta_old, beta_new)</pre>
    beta_old <- beta_new
    iter <- iter + 1
  }
  return(beta new)
data(BostonHousing2)
X = BostonHousing2[, !(colnames(BostonHousing2) %in% c("medv", "town", "tract", "cmedv"))]
```

```
X_orig = data.matrix(X)
X_scaled = scale(data.matrix(X))
Y_orig = as.vector(BostonHousing2$cmedv)
Y_scaled = as.vector(scale(BostonHousing2$cmedv))
delta <- 0.1
epsilon \leftarrow 10^-9
maxiter <- 10000
beta <- mylm_g(X_scaled, Y_scaled, delta, epsilon, maxiter)
lm.model <- lm(Y_orig ~ X_orig)</pre>
# Convert beta back to original coords by reversing operations and calculating
# beta-zero, the intercept
beta_convert <- beta[c(2:nrow(beta)),] * as.vector(apply(as.matrix(Y_orig),2,sd)) /
                                          as.matrix(apply(X_orig,2,sd))
beta0_convert <- beta[1,1] + mean(Y_orig) - sum(beta[c(2:nrow(beta)),] *
                                (as.matrix(apply(X_orig,2,mean)) *
                                 as.vector(apply(as.matrix(Y_orig),2,sd)) /
                                 as.matrix(apply(X_orig,2,sd))))
convertedBeta <- matrix(c(beta0_convert, beta_convert))</pre>
difference <- abs(lm.model$coefficients - convertedBeta)</pre>
kable(cbind(convertedBeta, as.data.frame(lm.model$coefficients), difference), digits=9, caption="Result
```

Table 3: Result Table for Calculated Beta from scaled values and LM

	convertedBeta	lm.model\$coefficients	difference
(Intercept)	-4.375739e + 02	-4.375739e + 02	9.214e-06
X_origlon	-3.935202e+00	-3.935202e+00	3.820 e-07
X_origlat	4.495441e+00	4.495441e+00	4.780 e-07
X_origcrim	-1.045469e-01	-1.045469e-01	8.000e-09
X _origzn	4.656475 e - 02	4.656476e-02	5.000e-09
$X_{origindus}$	1.524529e-02	1.524535e-02	5.400 e - 08
$X_{origchas}$	2.578020e+00	2.578020e+00	1.800 e-07
$X_{orignox}$	-1.582458e+01	-1.582458e+01	6.690 e-07
X_{origrm}	3.753712e+00	3.753712e+00	8.500 e - 08
X_origage	2.467658e-03	2.467659e-03	1.000e-09
X_origdis	-1.399927e+00	-1.399927e+00	7.000e-09
$X_{origrad}$	3.067144e-01	3.067145 e-01	1.070 e-07
$X_{origtax}$	-1.288632e-02	-1.288633e -02	6.000e-09
$X_{origptratio}$	-8.771212e-01	-8.771212e-01	2.200 e-08
X_{origb}	9.176196e-03	9.176196e-03	0.000e+00
$X_{origlstat}$	-5.374411e-01	-5.374411e-01	5.000e-09

Mathematically, since we scaled both the X and Y variables to calculate the original Beta coefficients, to

recover the Beta for the unscaled data, we simply multiple each X Beta coefficient by the standard deviation of Y by the standard deviation of X, i.e.

 $Beta_convert = Beta_scaled * SD_origY / SD_origX$

To calculate the Beta_zero intercept for the unscaled data is a little more complicated, but effectively we are just reversing the normalization operations on the scaled beta coefficients. We have to add the original intercept from scaled data (which is effectively zero) to the mean of the original Y unscaled data, i.e.:

 $intercept = scaled_intercept (i.e. 0) + y_orig_mean$

and then subtract the summation of the scaled beta values multiplied by x_orig_mean * SD_origY/SD_origX

 $itercept = scaled_intercept \ (i.e. \ 0) \ + \ y_orig_mean \ - \ sum(orig_betax_origmeanSD_origY/SD_origX)$