

March 16 (1)

Let  $N(t)$  and  $M(t)$  be independent PP's of intensity  $\lambda_1, \lambda_2$ , then  $N+M$  is a PP of intensity  $\lambda_1 + \lambda_2$

Ex: Cars traveling on a highway

$N(t) = \# \text{ cars going east up to } t \rightarrow \text{P.P. (2)}$

$$M(f) = \frac{1}{\text{width}} \rightarrow \text{P.P.}(\mu)$$

Q: What is the distribution of  $N(t)$ , given that  $N(t) + M(t) = C$ ?

A: Internally Binomial  $(C, \frac{d}{2+\mu})$

More precisely:

$$\begin{aligned} P(N(t) = k \mid N(t) + M(t) = C) \\ &= \frac{P(N(t) = k, M(t) = C - N(t))}{P(N(t) + M(t) = C)} \\ &= \frac{P(N(t) = k, M(t) = C - k)}{P(N(t) + M(t) = C)} \\ &= \frac{P(N(t) = k) P(M(t) = C - k)}{P(N(t) + M(t) = C)} \end{aligned}$$

$$= e^{-\lambda t} \cdot (\lambda t)^k / k! \cdot e^{-\mu t} \cdot (\mu t)^{c-k} / (c-k)! \quad (2)$$

$N(t) \sim \text{Poisson}(\lambda t)$

$M(t) \sim \text{Poisson}(\mu t)$

$(N+M)(t) \sim \text{Poisson}((\lambda+\mu)t)$

↑  
superposition  
property

$$= \frac{e^{-\lambda t} e^{-\mu t}}{e^{-(\lambda+\mu)t}} \cdot \frac{t^k t^{c-k}}{t^c} \cdot \frac{\lambda^k \mu^{c-k}}{(\lambda+\mu)^c} \cdot \frac{c!}{k! (c-k)!}$$

↑  
 $\binom{c}{k}$

$$= \binom{c}{k} \left( \frac{\lambda}{\lambda+\mu} \right)^k \left( \frac{\mu}{\lambda+\mu} \right)^{c-k}$$

$$= \binom{c}{k} \left( \frac{\lambda}{\lambda+\mu} \right)^k \left( 1 - \frac{\lambda}{\lambda+\mu} \right)^{c-k}$$

$$= P\left(\text{Bin}\left(c, \frac{\lambda}{\lambda+\mu}\right) = k\right)$$



Remark: More generally if  $N_1, N_2, \dots, N_\ell$  are  $\ell$  independent P.P.'s with rates  $\lambda_1, \dots, \lambda_\ell$  then  $N_1 + \dots + N_\ell$  is also a P.P. of intensity is  $\lambda_1 + \dots + \lambda_\ell$

The following property is a kind of reciprocal of the superposition property (3)

Prop (Thinning): Suppose that each arrival of a P.P.  $N(t)$  (with rate  $\lambda$ ) is classified or labelled of either type 1 or type 2 independently with probability  $p$  and  $1-p$ , respectively

- Let  $N_1(t) = \#$  of type 1 events by time  $t$
- Let  $N_2(t) = \#$  of type 2 events by time  $t$

Then  $N_1$  and  $N_2$  are Poisson processes with rates  $\lambda p$  and  $\lambda(1-p)$ , respectively

Proof: To see this, we check that  $N_1$  and  $N_2$  satisfy the second definition (4 axioms)

1)  $N_1(0) = 0$  ✓

2) Independent increment property.

$N_1(t) - N_1(s)$  is a function of  $N(t) - N(s)$  for  $t > s$ . So the independence of increments for  $N$   $\Rightarrow$  same for  $N_1$  ✓

$$3) P(N_1(t+\Delta t) - N_1(t) = 1) = p\lambda \Delta t + o(\Delta t) \quad ?$$

$$= P(N_1(\Delta t) = 1) \quad (\text{by stationarity})$$

$$= P(N_1(\Delta t) = 1 \mid N(\Delta t) = 1) \cdot P(N(\Delta t) = 1)$$

↑  
we condition on  
 $N(\Delta t)$

$$+ P(N_1(\Delta t) = 1 \mid N(\Delta t) > 1) \cdot P(N(\Delta t) > 1)$$

$$= p \cdot (\lambda \Delta t + o(\Delta t)) + P(N_1(\Delta t) = 1 \mid N(\Delta t) > 1) \cdot o(\Delta t)$$

we used the second definition, knowing that  $N(t)$  is P.P. ( $\lambda$ )

(rule:  $e^{-\lambda \Delta t} (\lambda \Delta t) = \lambda \Delta t + o(\Delta t)$   
because  $e^x = 1 + o(x)$ )

we use  
the second def.  
(4)

$$= p\lambda \Delta t + (p + P(N_1(\Delta t) = 1 \mid N(\Delta t) > 1)) \cdot o(\Delta t)$$

$$= p\lambda \Delta t + o(\Delta t) \quad \checkmark$$

$$4) P(N_1(t+\Delta t) - N_1(t) > 1) \leq P(N(\Delta t) > 1)$$

"  $o(\Delta t)$

$$\Rightarrow P(N_1(t + \Delta t) - N_1(t) > 1) = o(\Delta t) \quad \checkmark \quad (3)$$

Example: A vendor sells fruits, and the sales form a Poisson process of rate 100/day. Suppose each fruit sold is rotten with a probability  $\frac{1}{20}$  independently

Then  $N_1(t)$  = # rotten fruits sold is a Poisson process (thinning property)

with rate  $100 \cdot \frac{1}{20} = 5$  / day

$$\begin{aligned} \text{and Prob (no rotten fruit is sold during 1 day)} \\ &= P(N_1(1) = 0) \quad (N_1(t) \sim \text{Poisson}(5t)) \\ &= P(\text{Poisson}(5,1) = 0) \\ &= e^{-5} \end{aligned}$$

We can also check this by direct computation.

$$\begin{aligned} &P(\text{No rotten fruit sold in a day}) \\ &= \sum_{k=0}^{\infty} P(k \text{ fruits are sold}) \cdot P(\text{No fruit among these } k \text{ is rotten}) \end{aligned}$$

condition on the total # of fruits sold = k

$$= \sum_{k \geq 0} P(k \text{ fruits are sold}) \cdot \left(\frac{19}{20}\right)^k \quad (6)$$

$$= \sum_{k \geq 0} P(\text{Poisson}(100) = k) \cdot \left(\frac{19}{20}\right)^k$$

$$= \sum_{k \geq 0} e^{-100} \frac{(100)^k}{k!} \cdot \left(\frac{19}{20}\right)^k$$

$$= \sum_{k \geq 0} e^{-100} \frac{(100 \cdot 19)^k}{k! \cdot 20^k}$$

$$= \sum_{k \geq 0} e^{-100} \frac{(5 \cdot 19)^k}{k!}$$

$$= e^{-100} \cdot \sum_{k \geq 0} \frac{(5 \cdot 19)^k}{k!}$$

$$= e^{-100} \cdot e^{5 \cdot 19} = e^{-100 + 5 \cdot 19} = e^{-5}$$

More generally (exercise)

$$P(\text{exactly } j \text{ fruits are rotten}) = e^{-5} \cdot \frac{5^j}{j!}$$

( $\rightarrow$  this shows  $N_1(t) \sim \text{Poisson}(\lambda_{pt})$ )

$\uparrow$  100       $\uparrow$   $\frac{1}{20}$