Sum Constructions

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The question we investigate in this note is the following:

Question 0.1. Let $S \subset [0,1]$ be any set of reals. Describe the set of k-sums, or countable sums, namely

$$kS = \{s_1 + s_2 + \dots + s_k : s_1, s_2, \dots, s_k \in S\}$$

$$(0.1)$$

or

$$\operatorname{sum}(S) = \left\{ \sum_{s \in T} s : T \text{ a countable subset of } S \right\}. \tag{0.2}$$

In particular, assuming the sumsets are measurable, what is the measure of kS? Of sum(S)?

(In general, the sumset need not be measurable. If S is countable, then any of the sumsets is also measurable.)

One motivation is the following: suppose you have access to supply of independent samples X of some discrete distribution F, say Poisson(1).

Question 0.2. How many different samples do you need to simulate a Bernoulli(p) event for some $p \in (0,1)$?

Alternatively:

Question 0.3. For which p can you simulate a Bernoulli(p) with just one sample from F? With two? With k?

Without loss of generality, we can assume F takes the form

$$F = \sum_{n \in \mathbb{N}} s_n \delta_n. \tag{0.3}$$

Simulating a Bernoulli(p) event is equivalent to finding an event A, measurable with respect to X, such that $\mathbb{P}(A) = p$. Since F is discrete, this is the same as finding a subset $I \subset \mathbb{N}$ such that

$$\sum_{i \in I} s_i = p. \tag{0.4}$$

There is some literature on the *Bernoulli factory problem*: how do you use an infinite supply of Bernoulli(p)'s to generate a Bernoulli(f(p)), where f is known but p is unknown? There may be some work on other distributions, but I'm not sure.

1 Generating Bernoullis from a single sample

The following will turn out to be a relevant condition for discrete probability measures. Let $S = \{s_n\}_{n\in\mathbb{N}}$ be a countable probability distribution, i.e. $s_n \geq 0$ and

$$\sum_{n} s_n = 1 \tag{1.1}$$

By convention we always write our countable sets S in decreasing order, so that $s_1 \geq s_2 \geq \dots$

Definition 1.1. S has the small tail sum property at level n if

$$s_n > \sum_{m > n} s_m. \tag{1.2}$$

If the above holds for all n, then we simply say that S has the small tail sum property.

Note that if S has the small tail sum property at any level n, then

$$\sum_{m} s_m < \left(\sum_{m \le n} s_m\right) + s_n < \infty,\tag{1.3}$$

so WLOG we may assume that 1.1 holds by scaling.

(Example) The geometric series $s_n = (1 - p)p^n$ has the small tail sum property if and only if p < 1/2. Indeed,

$$(1-p)p^n \ge \sum_{m>n} (1-p)p^m = p^{n+1} \iff 1-p \ge p \iff 1/2 \ge p.$$
 (1.4)

Equality holds for every n exactly when p = 1/2; in that case, $sum(\{1/2, 1/4, 1/8, ...\}) = [0, 1]$, which is equivalent to the fact that every real number in [0, 1] has a binary decomposition.

Let λ denote Lebesgue measure. For countable S and $k \in \mathbb{N}$, kS is countable and thus $\lambda(kS) = 0$, so it is natural to consider sum(S) in the context of Lebesgue measure. We have the following characterization of sum(S) in this case.

Theorem 1.2. Let $S \subset [0,1]$ be a countable probability distribution with the small tail sum property (i.e. satisfying 1.1 and 1.2). We have

$$\lambda(\operatorname{sum}(S)) = \lim_{n \to \infty} 2^{n+1} \left(1 - \sum_{m \le n} s_m \right). \tag{1.5}$$

What makes this case special, and allows this direct computation, is that there are no 'overlaps' between sums over different subsets.

Lemma 1.3. Let $(s_n) \subset [0,1]$ be any countable set with the small tail sum property. If $x \in sum(S)$, then there is a unique subset $I \subset \mathbb{N}$ such that

$$x = \sum_{i \in I} s_i. \tag{1.6}$$

Moreover, I is obtained by applying the greedy algorithm to x: for $i \in \mathbb{N}$,

$$i \in I \iff \left(\sum_{j \in I, j < i} s_j\right) + s_i < x.$$
 (1.7)

Proof. Suppose $I, J \subset \mathbb{N}$ are two distinct subsets, and by reversing I and J if necessary, let

$$k = \min\{l \in \mathbb{N} : l \in I, l \notin J, \text{ and } J \cap [l] \subset I \cap [l]\}. \tag{1.8}$$

By the tail bound 1.2 and the definition of k,

$$\sum_{i \in I} s_i - \sum_{j \in J} s_j \ge s_k - \sum_{j > k} s_j > 0.$$
 (1.9)

So $\sum_{i \in I} s_i \neq \sum_{j \in J} s_j$.

If $x \in \text{sum}(S)$, then the greedy algorithm succeeds. Indeed, choose I such that 1.6 holds, and suppose I does not agree with the greedy algorithm, i.e. for some n,

$$\left(\sum_{i \in I \cap [n-1]} s_i\right) + s_n < x \text{ but } n \notin I.$$
(1.10)

Then by 1.2,

$$\sum_{i \in I} s_i < \left(\sum_{i \in I \cap [n-1]} s_i\right) + \sum_{j > i} s_j \tag{1.11}$$

$$< \left(\sum_{i \in I \cap [n-1]} s_i\right) + s_n \tag{1.12}$$

$$\langle x, \rangle$$
 (1.13)

contradicting 1.6.

An immediate corollary is:

Corollary 1.4. For S satisfying 1.2, the greedy algorithm is a bijection between the power set $2^{\mathbb{N}}$ and sum(S).

The next lemma describes the complement of sum(S) as a countable union of intervals. For $I \subset \mathbb{N}$, use s_I to denote the sum over I:

$$s_I = \sum_{i \in I} s_i. \tag{1.14}$$

Lemma 1.5. Let $S \subset [0,1]$ be a countable probability distribution satisfying 1.2. The set of reals not in the sumset of S can be written as a union of (open) intervals:

$$[0,1] \setminus \text{sum}(S) = \bigcup_{n \ge 0} \bigcup_{I \subset [n-1]} \left(s_I + \sum_{i > n} s_i, s_I + s_n \right) := \bigcup_n \bigcup_{I \subset [n-1]} A_I(n), \tag{1.15}$$

where $[n] = \{0, 1, ..., n\}$, and by convention $[-1] = \emptyset$. Moreover, the intervals appearing in the union are all pairwise disjoint.

Proof. Suppose $y \notin \text{sum}(S)$. By 1.3, the greedy algorithm must fail at some finite stage, i.e. for some n and $I \subset [n-1]$,

$$s_I + \sum_{i > n} s_i < y < s_I + s_n. \tag{1.16}$$

So it suffices to show that the $A_I(n)$ are disjoint. Let $I, J \subset \mathbb{N}$ be distinct finite subsets, and consider the intervals $A_I(n)$ and $A_J(m)$. By reversing I and J if necessary, let

$$k = \min\{l \in \mathbb{N} : l \in I, l \notin J, \text{ and } J \cap [l] \subset I \cap [l]\},\tag{1.17}$$

as in the proof of 1.3. Then

$$s_I \ge s_{I \cap [n]} \ge s_{J \cap [n]} + \sum_{j > n} s_j \ge s_J + s_{m+1},$$
 (1.18)

which implies inf $A_I(n) > \sup A_J(m)$.

Lemma 1.5 immediately leads to a computation for the measure of sum(S).

Proof of 1.2. By lemma 1.5 and 1.1,

$$1 - \lambda(\text{sum}(S)) = \sum_{n>0} \sum_{I \subset [n-1]} \left(s_I + s_n - s_I - \sum_{i>n} s_i \right)$$
 (1.19)

$$=\sum_{n>0} 2^n \left(s_n - \sum_{i>n} s_i \right) \tag{1.20}$$

$$= \sum_{n>0} 2^n \left(s_n - \left(1 - \sum_{i=0}^n s_i \right) \right) \tag{1.21}$$

$$= \lim_{N \to \infty} \left(\sum_{n=0}^{N} 2^n \left(s_0 + s_1 + \dots + 2s_n - 1 \right) \right)$$
 (1.22)

$$= \lim_{N \to \infty} \left(-2^{N+1} + 1 + \sum_{k=0}^{N} s_k (2 \cdot 2^k + 2^{k+1} + \dots + 2^N) \right)$$
 (1.23)

$$= \lim_{N \to \infty} 1 - 2^{N+1} \left(1 - \sum_{k=0}^{N} s_k \right) \tag{1.24}$$

Thus

$$\lambda(\text{sum}(S)) = \lim_{N \to \infty} 2^{N+1} \left(1 - \sum_{n=0}^{N} s_n \right).$$
 (1.25)

For example, when $s_n = (1 - p)p^n$ for p < 1/2, we get

$$\lambda(\text{sum}(S)) = \lim_{N \to \infty} 2^{N+1} \left(1 - (1 - p^{N+1}) \right) = \lim_{N \to \infty} (2p)^{N+1} = 0.$$
 (1.26)

This is somewhat surprising: the measure of the sumset is 1 for p = 1/2, but it jumps down to 0 for p < 1/2 – there is a sharp phase transition. To obtain a 'smoother' transition, we can look at the Hausdorff dimension of sum(S) instead. By the usual ideas to compute the Hausdorff dimension of a Cantor set, and ignoring issues of convergence, we get the general formula

$$\mathcal{H}(\operatorname{sum}(S)) = \frac{\log 2}{-\lim_{n \to \infty} \frac{1}{n} \log \sum_{m > n} s_m}.$$
(1.27)

It seems that this limit should always exist, perhaps as a consequence of:

Lemma 1.6. Any countable discrete probability distribution S with the small tail sum property 1.2 satisfies $s_n \leq 2^{-n}$ for n sufficiently large.

$$Proof.$$
 ????

(**Example**) For the geometric distributions $s_n = (1 - p)p^n$, the Hausdorff dimension is exactly

$$\mathcal{H}(\operatorname{sum}(S)) = \frac{\log 2}{-\log p}.$$
(1.28)

This gives a smooth phase transition for $p \in (0, \frac{1}{2})$.

2 Multiple independent samples

In this section we address the following:

Question 2.1. Given a discrete probability distribution F, how many independent samples of F are needed to generate any Bernoulli(p) for $p \in (0,1)$?

The small tail sum property will be relevant here, too. Given k iid samples X_1, X_2, \ldots, X_k sampled from the distribution F, we can look at events of the form $(X_i = x_i, i \in [k])$, and sum over countably many such events to get an event with probability p. Thus

$$\{p: \exists A \subset \mathbb{N}^k, \mathbb{P}((X_1, \dots, X_k) \in A) = p\} = \text{sum}(S), \tag{2.1}$$

where S is the multiset of all possible atoms $\mathbb{P}(X_i = x_i, i \in [k])$.

Writing $F = \sum_n u_n \delta_n$, the atoms of the product measure are of the form $\prod_{i=1}^k u_{x_i}$ for any $x = (x_i : i \in [k]) \subset \mathbb{N}^k$, and each such atom occurs $\operatorname{perm}(x)$ many times in S, where $\operatorname{perm}(x)$ is the number of distinct permutations of the sequence x (exactly k! occurrances if the entries of x are distinct, but fewer otherwise). Because of these multiplicities in S, S will not have the small tail sum property at most levels, and correspondingly, $\operatorname{sum}(S)$ will be larger:

Proposition 2.2. Suppose S fails the small tail property for all $n \ge n_0$. Then $\lambda(sum(S)) \ge \sum_{m>n_0} s_m > 0$. In particular, if S fails the small tail property for all levels, then sum(S) = [0,1].

Proof. In short: when the small tail property fails, the greedy algorithm succeeds. Note that if the small tail property fails at level m, then $s_m > 0$.

Claim 2.3. For any $x \in [0, \sum_{m \geq n_0} s_m]$, applying the greedy algorithm to x yields a subset $I \subset \{n_0, n_0 + 1, \ldots\}$ with $\sum_{i \in I} s_i = x$.

If not, then the greedy algorithm fails at some stage, i.e. for some set $I \subset \{n_0, n_0 + 1, \dots, \ell\}$ with $\ell \in I$,

$$\sum_{i \in I} s_i + s_{\ell+1} > x > \sum_{i \in I} s_i + \sum_{i > \ell+1} s_i, \tag{2.2}$$

contradicting the fact that the small tail property fails at $\ell + 1$.

2.1 Geometric

Computations for general distributions F are difficult, so we focus on the class of Geometric(p) random variables for $p \in (0, 1/2)$. This case is particularly nice becasuse the atom sizes are constant on hyperplanes normal to $(1, \ldots, 1) \in \mathbb{N}^k$, i.e. with $u_y = qp^y$, we have

$$\prod_{i=1}^k u_{x_i} = q^k p^{\sum_i x_i}$$

And, the total number of such atoms is just the size of the hyperplane, or the number of compositions of $r = \sum_i x_i$ with k parts, namely $\binom{r+k-1}{r}$. Mathematica gives that

Lemma 2.4.
$$p^r \leq \sum_{w>r} {w+k-1 \choose w} p^w$$
 for all $r \geq 0$ if and only if $p \geq 1 - 2^{-k^{-1}}$.

Thus the sequence S of atoms obtained this way fails the small tail sum property at every level if p is large enough. We obtain:

Proposition 2.5. k iid samples of Geometric(p) generate all possible Bernoulli(p') for $p' \in [0,1]$ if and only if $p \ge 1 - 2^{-k^{-1}}$.

2.2 Geometric, k = 2 independent samples

In the special case k=2 we can do some exact computations for the Lebesgue measure of sum(S). Some algebra shows that the set of atoms S satisfies the small tail sum property once for each level q^2p^r with $r=0,1,\ldots,R-1$, where

$$R = \left[\frac{1 - 4p + 2p^2}{p(1 - p)} \right]. \tag{2.3}$$

It follows that for each $r=0,1,\ldots,R-1,$ sum(S) is missing has exactly (r+1)! intervals of size

$$q^{2} \left(p^{r} - \sum_{w>r} (w+1)q^{2}p^{w} \right) \tag{2.4}$$

(There are w + 1 copies of the atom q^2p^w for each w, and thus (w + 1) many choices for each atom size's contribution to the sum). So the Lebesgue measure of the set of Bernoulli's that can be simulated with k = 2 independent Geo(p)'s is

$$1 - \sum_{r \le R} q^2(r+1)! \left(p^r - \sum_{w \ge r} (w+1)p^w \right). \tag{2.5}$$

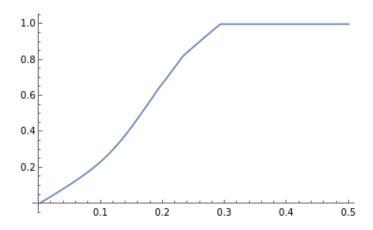


Figure 1: A plot of the Lebesgue measure 2.5 of sum(S) as a function of p, where S is the atoms of the product measure with two independent Geometric(p) random variables. Note that the measure is 1 for $p > 1 - \frac{1}{\sqrt{2}} \approx .29289$ by Proposition 2.5.

3 Further questions

- Is there a countable set $S \subset [0,1]$ such that property 1.2 fails for infinitely many n, and such that $\lambda(\text{sum}(S)) = 0$?
- Is there a discrete distribution F such that k iid random variables distributed according to F are not enough to generate all possible Bernoulli(p)'s? (One possible idea: given p that is not constructible, form a dynamical system by thinking about what happens when $k \to k + 1$. Given atoms S at stage k, for p not to be constructible in the next step, it is sufficient that $(s_{\ell+1})^{-1}(p-\sum_I s_i)$ is not in sum(S), where I is the greedy construction in S which fails at level ℓ . If one can show that this map $p \mapsto g_S(p) = (s_{\ell+1})^{-1}(p-\sum_I s_i)$ which trivially has $g_S(p) \le p$ with strict inequality if $p \ne 0$ always hits 0 in a globally bounded number of iterations, we'd be in business.)
- Prove that if S is uncountable, then $\lambda(\text{sum}(S)) > 0$. Can the measure be arbitrarily close to 0 in this case?
- Is there an uncountable set $S \subset [0,1]$ such that $\lambda(2S) = 0$? (The cantor set C satisfies 2S = [0,2].)
- Given a set S and $p \in \text{sum}(S)$, where $p = \sum_{i \in I} s_i$, give a notion of 'efficiency' of the representation e.g. Komolgorov complexity of the set I. Is there some tradeoff between the worst case or average efficiency and the size of sum(S)? Is the number of ways to write p as a sum of elements of S related to such a tradeoff? What about the entropy of the sequence S?
- Suppose we construct S in a random way: for example, fix a distribution function F on [0,1], sample X_0, X_1, \ldots i.i.d. $\sim F$, set $S_0 = X_0$ and recursively define $S_n = X_n S_{n-1}$; or let (X_n) have the Poisson-Dirichlet distribution. What is the probability that the random sequence $(S_n)_n$ satisfies 1.1? 1.2? What is the distribution of $\lambda(\text{sum}(S))$? (Expectation?)