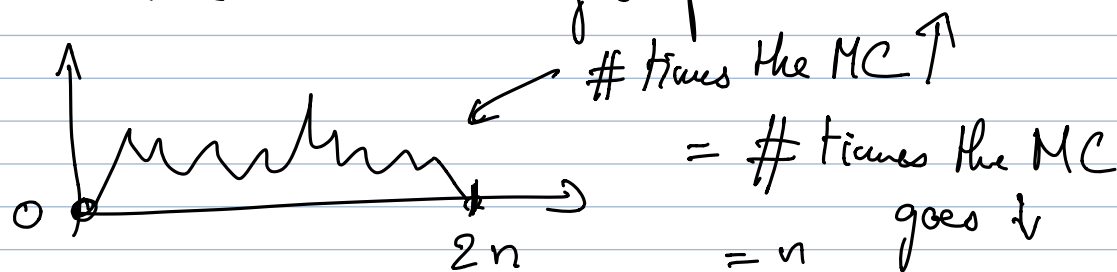


Is the symmetric R.W in \mathbb{Z}^d transient or recurrent?

$d=1$ (Recall: We will use the criterion that $\sum_n P_{ii}^n < \infty$ to show that the M-C is recurrent)

- Starting from 0, one can only come back in an even number of steps



P_{00}^{2n} = Proba. to move n times backward and n times forward.

- What is the probability of observing a specific trajectory of $2n$ steps? $\rightarrow \frac{1}{2^{2n}}$
 (ex: $\uparrow \downarrow \uparrow \uparrow \downarrow \dots$)
 $\underbrace{\hspace{10em}}_{2n \text{ steps}}$

- How many trajectories satisfy the condition above.
 $\rightarrow \binom{2n}{n}$

$$\Rightarrow P_{00}^{2n} = \frac{1}{2^{2n}} \times \binom{2n}{n}$$

$$\Rightarrow \sum_{n=0}^{+\infty} P_{00}^{2n} = \sum_{n=0}^{+\infty} \frac{\binom{2n}{n}}{2^{2n}} = \sum_{n=0}^{+\infty} \frac{2n!}{(n!)^2 2^{2n}} \quad (2)$$

Stirling Formula: $n! \sim \left(\frac{n}{e}\right)^n \sqrt{2\pi n}$ $\xrightarrow{\sim U_n}$

$$\Rightarrow U_n \sim \left(\frac{2n}{e}\right)^{2n} \sqrt{2\pi 2n} \left(\frac{e}{n}\right)^{2n} \frac{1}{2\pi n} \times \frac{1}{2^{2n}}$$

$$U_n \sim \frac{1}{\sqrt{\pi n}} = \frac{1}{(\pi n)^{\frac{1}{2}}} < 1$$

$$\Rightarrow \sum_{n=0}^{+\infty} P_{00}^{2n} = \sum_{n=0}^{+\infty} P_{00}^n = +\infty \text{ so the } \boxed{\text{R-W is recurrent}}$$

Remark: • When the R-W is asymmetric $\left(\frac{1-p}{2}, \frac{p}{2}\right)$ $p \neq \frac{1}{2}$
we can show that the R-W is transient

- For $d=2$, one can use a similar approach (and the walk is recurrent), but the problem gets more and more complex combinatorically as the dimension increases

(For a proof in $d=3$ - using characteristic functions - , see Notes online)

Result: The \uparrow walk is $\left\{ \begin{array}{l} \text{recurrent if } d=1,2 \\ \text{transient if } d \geq 3 \end{array} \right.$
symmetric random

IV - limiting probabilities

(3)

Recall the example of 2 state MC seen in HW 1

$1-p \rightarrow \text{state 0} \xrightarrow{p} \text{state 1} \xrightarrow{1-q} \text{state 0}$
 $\text{state 0} \xrightarrow{q} \text{state 1}$

\rightarrow we saw $\lim_{n \rightarrow \infty} P(X_n = i) = \begin{cases} \frac{q}{p+q} & i=0 \\ \frac{p}{p+q} & i=1 \end{cases}$

This distribution defines some
limiting probabilities of the M.C.

Def: Let $(X_n)_{n \geq 0}$ be a M.C with transition matrix \underline{P} and let μ be a probability distribution on its state space. We say that μ is a stationary distribution of (X_n) if

$$\mu \underline{P} = \mu$$

Prop: If (X_n) converges to a distribution μ as $n \rightarrow \infty$ then μ is stationary

Proof: Let μ_k be the distribution of X_k

$$\begin{array}{ccc} \text{then } \mu_{k+1} & = & \mu_k \underline{P} \\ \downarrow k \rightarrow \infty & & \downarrow k \rightarrow \infty \\ \mu & = & \mu \underline{P} \end{array}$$

□

Remark : The limiting probabilities also account for the proportion of time the MC stays in each state in the long run ④

- Let m_i = mean time to return starting from i . Then the proportion of time in state i must be $\frac{1}{m_i}$ in the long run.

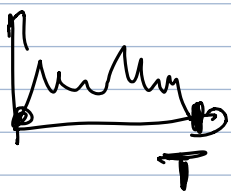
Q : Can we guarantee the existence of a stationary distribution?

Def : • A recurrent state i is **positive recurrent** if $m_i < \infty$ (expected return time to i)
• If i is recurrent but not positive recurrent ($m_i = \infty$) then i is **null-recurrent**

Ex : Symmetric R.W on \mathbb{Z}

let T = first time $X_n = 0$, starting from 0

so $m_0 = \mathbb{E}(T)$



One can show that :

$$\begin{cases} P(T = 2n+1) = 0 \\ P(T = 2n) = \frac{1}{2^{n-1}} \binom{2n}{n} \frac{1}{2^{2n}} \sim \frac{1}{2\sqrt{\pi} n^{3/2}} \end{cases}$$

$$\text{So } m_0 = \sum_n 2n P(T=2n) = +\infty \quad (3)$$

$$\sim \frac{1}{\sqrt{n}}$$

\Rightarrow The chain is null recurrent

Prop (admitted): Positive and null recurrence are class properties.

- If the state space is finite, then all recurrent states are positive recurrent

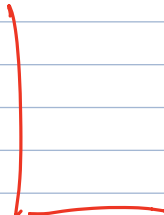
D.f: An aperiodic positive recurrent state is called ergodic. An ergodic MC is a MC whose states are all ergodic

Thm (admitted): For an irreducible ergodic MC $\pi_j = \lim_{n \rightarrow +\infty} P_{ij}^n$ exists for all j , and is independent of i . In addition

(i) $\underline{\pi}$ is the unique solution of
$$\begin{cases} \underline{\pi} = \underline{\pi} P \\ \sum_j \pi_j = 1 \end{cases}$$

(ii) $\pi_j = \frac{1}{m_j}$, where $m_j = \text{mean return time to } j$ (in particular $\pi_j > 0$)

(iii) $\pi_j = \lim_{n \rightarrow \infty} \frac{\# \text{ visits to } j \text{ by time } n}{n}$

 = proportion of time, in the long run, spent in state j (b)

Remark: In practice, if you have an irreducible MC, then you can try to solve

$$(i) \quad \begin{cases} \pi = \pi P \\ \sum_j \pi_j = 1 \end{cases}$$

and if you solve it (or guess a solution that works), then π is the stationary distribution and the chain is positive recurrent.