

A Beta Way: A Tutorial For Using Beta Regression in Psychological Research

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Abstract

Rates, percentages, and proportions are common outcomes in psychology and the social sciences. These outcomes are often analyzed using models that assume normality, but this practice overlooks important features of the data, such as their natural bounds at 0 and 1. As a result, estimates can become distorted. In contrast, treating such outcomes as beta-distributed respects these limits and can yield more accurate estimates. Despite these advantages, the use of beta models in applied research remains limited. Our goal is to provide researchers with practical guidance for adopting beta regression models, illustrated with an example drawn from the psychological literature. We begin by introducing the beta distribution and beta regression, emphasizing key components and assumptions. Next, using data from a learning and memory study, we demonstrate how to fit a beta regression model in R with the Bayesian package `brms` and how to interpret results on the response scale. We also discuss model extensions, including zero-inflated, zero- and one-inflated, and ordered beta models. Basic familiarity with regression modeling and R is assumed. To promote wider adoption of these methods, we provide detailed code and materials at <https://zenodo.org/records/16895241>.

Keywords: beta regression, beta distribution, R tutorial, Psychology, Learning and memory

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Introduction

Many outcomes in psychological research are naturally expressed as proportions or percentages. These include the proportion of correct responses on a test (e.g., Kornell and Bjork (2008)), the proportion of time a participant fixates on a stimulus in an eye-tracking task (e.g., James et al., 2025), or the proportion of respondents endorsing a particular belief (e.g., Costello et al., 2024). A central question for studies involving proportional data is how to analyze these outcomes in a statistically appropriate and interpretable way.

Researchers frequently default to linear models that assume Gaussian (normal) distributions, such as t-tests, ANOVAs, and linear regression. However, these models make strong assumptions: (1) residuals are normally distributed, (2) the outcome is unbounded (from $-\infty$ to ∞), and (3) variance is constant across the range of the data. These assumptions are rarely satisfied in practice (Sladekova & Field, 2024), and they are especially ill-suited for proportional outcomes, which are bounded between 0 and 1 and often exhibit heteroscedasticity—non-constant variance, particularly near the boundaries (Ferrari & Cribari-Neto, 2004; Paolino, 2001; Smithson & Verkuilen, 2006). Violating these assumptions can lead to biased estimates and spurious inferences, especially when many observations cluster near 0 or 1.

In some cases, a generalized linear model (GLM) can relax the assumption of normality. For example, binomial and Bernoulli models—often referred to as logistic regression when used with a logit link—are appropriate when the outcome is binary (0/1) or represents the number of successes out of a fixed number of trials. These models, however, rely on data that arise from discrete trial structures. They tend to perform poorly when the observed proportions are truly continuous or when the data show extra variability (overdispersion), particularly when many values occur near the boundaries of the scale (0 and 1).

The challenges of analyzing proportional data are not new (see Bartlett, 1936). Fortunately, several existing approaches address the limitations of commonly used models. One such approach is beta regression, an extension of the generalized linear model that employs the beta distribution (Ferrari & Cribari-Neto, 2004; Paolino, 2001). Beta regression offers a flexible and robust solution for modeling proportional data directly by accounting for boundary effects and over-dispersion, making it a valuable alternative to traditional binomial models. This approach is particularly well-suited for psychological research because it can handle both the bounded nature of proportional data and the non-constant variance often encountered in these datasets (Sladekova & Field, 2024). In addition, the direct modeling of proportions allows comparability across tasks and scales, and can be particularly valuable when only the proportional data is available, as is often the case with secondary data that lack item-level structure or point values.

While in this paper we will focus on proportional-responses that lie between 0 and 1—it is important to note that our analysis applies to any bounded continuous scale. Any bounded scale can be mapped to lie within 0 and 1 without resulting in a loss of information as the transformation is linear.¹ Consequently, a scale that has natural end points of -1,234 and +8,451—or any other end points on the real number line short of infinity—can be modeled using the approaches we describe in this paper.

A Beta Way Is Possible

With the widespread availability of open-source software such as R (R Core Team, 2024) and its extensive ecosystem of user-developed packages, advanced models like beta regression have become increas-

¹Specifically, for any continuous bounded variable x , we can rescale this variable to lie within 0 and 1 by using the formula $x' = \frac{x - \min(x)}{\max(x) - \min(x)}$ where $0 \leq x' \leq 1$.

40 ingly accessible to applied researchers. Yet, their adoption in psychology remains relatively limited. One
 41 contributing factor may be the lack of domain-specific examples that demonstrate how these models address
 42 common challenges in psychological data. Although recent years have seen a growing interest in beta regres-
 43 sion, and a number of useful tutorials are available (Bendixen & Purzycki, 2023; Coretta & Bürkner, 2025;
 44 Heiss, 2021; e.g., Smithson & Verkuilen, 2006; Vuorre, 2019), these resources often either focus on basic
 45 implementation or briefly mention extensions without detailing how they can be applied to psychologically
 46 relevant research questions.

47 The present tutorial aims to help bridge this gap by providing a comprehensive, applied introduction
 48 to beta regression and several of its extensions. In addition to the standard beta model, we walk through zero-
 49 inflated, zero-one-inflated, and ordered beta regression. These models are particularly useful for researchers
 50 working with proportion outcomes that include boundary values (e.g., exact 0s or 1s) or responses with an
 51 inherent ordinal structure. Our goal is to offer practical guidance that enables psychological researchers to
 52 implement, interpret, and report these models in ways that directly support their empirical questions.

53 Beyond model specification, we place strong emphasis on interpreting results on the response scale—
 54 that is, in terms of probabilities and proportions—rather than relying on often difficult to interpret parameters.
 55 This focus makes the models more accessible and meaningful for psychological applications, where effects
 56 are often easier to communicate when framed on the original scale of the outcome (e.g., changes in recall
 57 accuracy or task performance). Throughout, we provide reproducible code and annotated examples to help
 58 readers implement and interpret these models in their own work.

59 We begin the tutorial with a non-technical overview of the beta distribution and its core pa-
 60 rameters. We then walk through the process of estimating beta regression models using the R package
 61 `brms` (Bürkner, 2017), illustrating each step with applied examples. To guide interpretation, we empha-
 62 size coefficients, predicted probabilities, and marginal effects calculated using the `{marginaleffects}` pack-
 63 age (Arel-Bundock et al., 2024). We also introduce several useful extensions—zero-inflated (ZIB), zero-
 64 one-inflated (ZOIB), and ordered beta regression—that enable researchers to model outcomes that include
 65 boundary values. Finally, all code and materials used in this tutorial are fully reproducible and avail-
 66 able via our GitHub repository: https://github.com/jgeller112/beta_regression_tutorial and on Zenodo
 67 <https://zenodo.org/records/16895241>².

68 Beta Distribution

69 Proportional data pose some challenges for standard modeling approaches: The data are bounded
 70 between 0 and 1 and often exhibit non-constant variance (heteroscedasticity) (Ferrari & Cribari-Neto, 2004;
 71 Paolino, 2001). Common distributions used within the generalized linear model frameworks often fail to
 72 capture these properties adequately, which can necessitate alternative modeling strategies.

73 While we do not have time to delve fully into its derivation, the beta distribution is a preferred
 74 distribution for this type of response because of certain unique properties. The beta distribution is defined as
 75 a distribution of the uncertainty of probabilities, which must lie within 0 and 1. As a consequence, the beta
 76 distribution is the maximum entropy distribution for any bounded continuous random variable, which means

²In this article, we try to limit code where possible; however, the online version has all the code needed to reproduce all analyses herein. Furthermore, to promote transparency and reproducibility, the tutorial was written in R version 4.5.1 (R Core Team (2024)) using Quarto (v.1.5.54), an open-source publishing system that allows for dynamic and static documents. This allows figures, tables, and text to be programmatically included directly in the manuscript, ensuring that all results are seamlessly integrated into the document. In addition, we use the `rix` (Rodrigues & Baumann, 2025) R package which harnesses the power of the `nix` (Dolstra & contributors, 2006) ecosystem to help with computational reproducibility. Not only does this give us a snapshot of the packages used to create the current manuscript, but it also takes a snapshot of system dependencies used at run-time. This way reproducers can easily re-use the exact same environment by installing the `nix` package manager and using the included `default.nix` file to set up the right environment. The README file in the GitHub repository contains detailed information on how to set this up to reproduce the contents of the current manuscript, including a video.

77 that the beta distribution can represent the full range of possibilities of such a scale.³ As a consequence, if we
 78 have a continuous scale with upper and lower bounds—and no other special conditions—the beta distribution
 79 will in principle provide a very good approximation of the uncertainty of the scale.

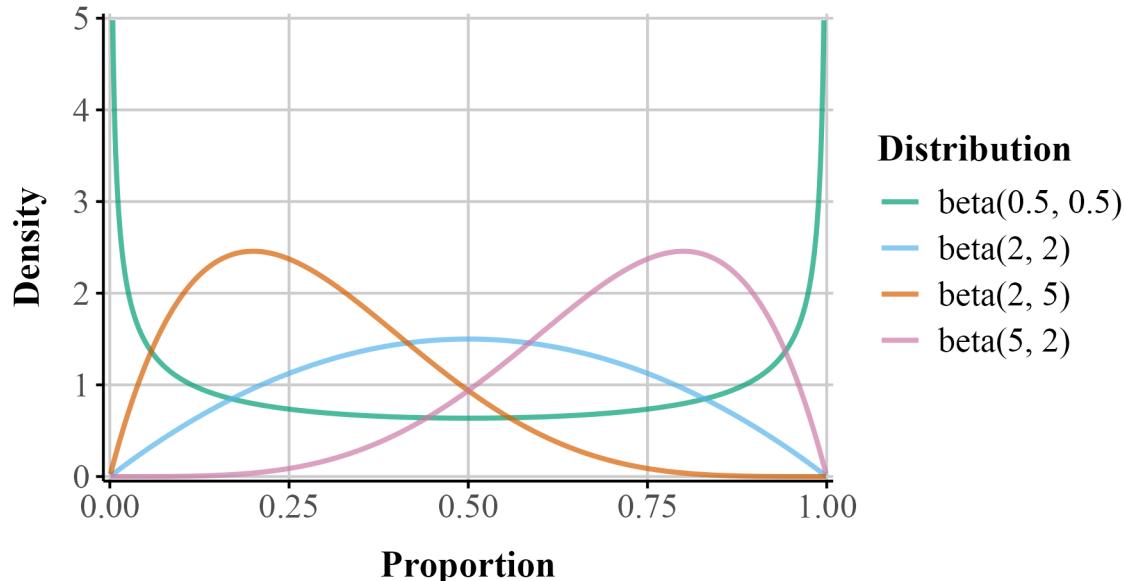
80 Typically, the expected value (or mean) of the response variable is the central estimand scholars want
 81 to estimate. A model should specify how this expected value depends on explanatory variables through two
 82 main components: a linear predictor, which combines the explanatory variables in a linear form ($a + b_1x_1 +$
 83 b_2x_2 , etc.), and a link function, which connects the expected value of the response variable to the linear
 84 predictor (e.g., $E[Y] = g(a + b_1x_1 + b_2x_2)$). In addition, a random component specifies the distribution
 85 of the response variable around its expected value (such as Poisson or binomial distributions, which belong
 86 to the exponential family) (Nelder & Wedderburn, 1972). Together, these components provide a flexible
 87 framework for modeling data with different distributional properties.

88 The beta distribution is continuous and restricted to values between 0 and 1 (exclusive). Its two
 89 parameters—commonly called `shape1` (α) and `shape2` (β)—govern the distribution's location, skewness, and
 90 spread. By adjusting these parameters, the distribution can take many functional forms (e.g., it can be sym-
 91 metric, skewed, U-shaped, or even approximately uniform; see Figure 1).

92 To illustrate, consider a test question worth seven points. Suppose a participant scores five out of
 93 seven. The number of points received (5) can be treated as α , and the number of points missed (2) as β . The
 94 resulting beta distribution would be skewed toward higher values, reflecting a high performance (yellow line
 95 in Figure 1; “`beta(5, 2)`”). Reversing these values would produce a distribution skewed toward lower values,
 96 representing poorer performance (green line in Figure 1; “`beta(2, 5)`”).

Figure 1

beta distributions with different shape1 and shape2 parameters.



97 I Can't Believe It's Not beta

98 While the standard parameterization of the beta distribution uses α and β , a reparameterization to a
 99 mean (μ) and precision (ϕ) is more useful for regression models. The mean represents the expected value
 100 of the distribution, while the dispersion, which is inversely related to variance, reflects how concentrated

³Technically, this maximum entropy condition is satisfied because the $\text{beta}(1,1)$ distribution is uniform over its support.

101 the distribution is around the mean, with higher values indicating a narrower distribution and lower values
 102 indicating a wider one. The connections between the beta distribution's parameters are shown in Equation 1.
 103 Importantly, the variance depends on the average value of the response because uncertainty intervals need to
 104 adjust for how close the value of the response is to the boundary.

$$\begin{array}{ll} \text{Shape 1: } a = \mu\phi & \text{Mean: } \mu = \frac{a}{a+b} \\ \text{Shape 2: } b = (1-\mu)\phi & \text{Precision: } \phi = a+b \\ & \text{Variance: } var = \frac{\mu \cdot (1-\mu)}{1+\phi} \end{array} \quad (1)$$

105 Thus, beta regression allows modeling both the mean and precision of the outcome distribution. To
 106 ensure that μ stays between 0 and 1, we apply a link function, which allows linear modeling of the mean on
 107 an unbounded scale. A common link-function choice is the logit, but other functions such as the probit or
 108 complementary log-log are possible.

109 The logit function, $\text{logit}(\mu) = \log\left(\frac{\mu}{1-\mu}\right)$ links the mean to log-odds which are unbounded, making
 110 linear modeling possible. The logit here no longer carries the same literal *odds* interpretation because there
 111 are no corresponding counts of “successes” and “failures.” Instead, the logit transform here simply maps the
 112 mean of the distribution to the real line. The inverse of the logit, called the logistic function, maps the linear
 113 predictor η back to the original scale of the data $\left(\mu = \frac{1}{1+e^{-\eta}}\right)$. The coefficients describe how predictors shift
 114 the *average proportion* on the logit scale here. Similarly, the strictly positive dispersion parameter is usually
 115 modeled through a log link function, ensuring it remains positive.

116 By accounting for the observations' natural limits and non-constant variance across different val-
 117 ues, the beta distribution is useful in psychology where outcomes like performance rates or response scales
 118 frequently exhibit these features.

119 Bayesian Approach to Beta Regression

120 Beta regression models can be estimated using either frequentist or Bayesian methods. In this paper,
 121 we adopt a Bayesian framework because it facilitates the estimation and interpretation of more complex
 122 models (Gelman et al., 2013; Johnson et al., 2022; McElreath, 2020). Additionally, the use of Bayesian statistics
 123 in psychology has been steadily growing (Pfadt et al., 2025). In principle, most Bayesian models can also be
 124 estimated with frequentist approaches (e.g., maximum likelihood). However, as model complexity increases,
 125 frequentist estimation often requires additional adjustments such as bootstrapping or custom optimization,
 126 whereas Bayesian modeling handles these extensions more naturally through hierarchical structure and prior
 127 specification.

128 A common concern is that Bayesian methods are slower than frequentist ones. While this is true
 129 in general, modern Bayesian computation engines are efficient, and in explanatory modeling our priority
 130 is to specify appropriate estimands rather than to optimize for computation speed. Moreover, we use the
 131 `{brms}` package (Bürkner, 2017), which provides a high-level interface to Stan (Team, 2023) and uses familiar
 132 R formula syntax (similar to `lm()`), making advanced Bayesian modeling accessible to researchers with
 133 standard regression backgrounds. The package also supports parallelization, which substantially reduces
 134 computation time for large datasets.

135 There are several important differences between our Bayesian analysis and the frequentist methods
 136 readers may be more familiar with—most notably, the absence of *t*- and *p*-values. To estimate models, the
 137 `{brms}` package uses Stan's computational algorithms to draw random samples from the posterior distribu-
 138 tion, which represents uncertainty about the model parameters. This posterior is conceptually analogous to a
 139 frequentist sampling distribution. By default, Bayesian models run 4 chains with 2,000 iterations each. The
 140 first 1,000 iterations per chain are warmup and are discarded. The remaining 1,000 iterations per chain are

141 retained as posterior draws, yielding 4,000 total post-warmup draws across all chains. From these draws, we
 142 can compute the posterior mean (analogous to a frequentist point estimate) and the 95% credible interval
 143 (Cr.I.), which is often compared to a confidence interval.

144 In addition, an important part of Bayesian analyses is prior specification. Priors encode our assumptions
 145 about plausible parameter values before observing the data and allow the model to regularize estimates,
 146 especially when data are sparse or parameters are weakly identified. To help bridge the conceptual gap for
 147 users more familiar with frequentist models, we begin with the default priors (flat/non-informative) provided
 148 by `{brms}`. These priors are intentionally non-informative, and in many applications produce results that
 149 closely align with frequentist estimates, while still offering the flexibility and interpretive advantages of a
 150 Bayesian framework. We strongly urge readers to consider prior specification strongly in all their work.

151 To ease readers into Bayesian data analysis we provide a metric known as the *probability of direction*
 152 (*pd*), which reflects the probability that a parameter is strictly positive or negative. When a uniform
 153 prior is used (all values equally likely in the prior), a *pd* of 95%, 97.5%, 99.5%, and 99.95% corresponds
 154 approximately to two-sided p-values of .10, .05, .01, and .001 (i.e., $\text{pd} \approx 1 - p/2$ for symmetric posteriors
 155 with weak/flat priors) (see Figure 2 for an illustrative comparison). For directional hypotheses, the *pd* can be
 156 interpreted as roughly equivalent to one minus the *p*-value (Marsman & Wagenmakers, 2016).

157 For reasons of space, we refer readers unfamiliar to Bayesian data analysis to several existing books
 158 on the topic (Gelman et al., 2013; Kruschke, 2015; McElreath, 2020). In addition, we assume readers are
 159 familiar with R, but those in need of a refresher should find Wickham et al. (2023) useful.

160 Beta Regression Tutorial

161 Example Data

162 Throughout this tutorial, we analyze data from a memory experiment examining whether the fluency
 163 of an instructor's delivery affects recall performance (Wilford et al., 2020, Experiment 1A). Instructor
 164 fluency—marked by expressive gestures, dynamic vocal tone, and confident pacing—has been shown to
 165 influence students' perceptions of learning, often leading learners to rate fluent instructors more favorably
 166 (Carpenter et al., 2013). However, previous research suggests that these impressions do not reliably translate
 167 into improved memory performance (e.g., Carpenter et al., 2013; Toftness et al., 2017; Witherby & Car-
 168 penter, 2022). In contrast, Wilford et al. (2020) found that participants actually recalled more information
 169 after watching a fluent instructor compared to a disfluent one. This surprising finding makes the dataset a
 170 compelling case study for analyzing proportion data, as recall was scored out of 10 possible idea units per
 171 video.

172 In Experiment 1A, participants ($N = 96$) watched two short instructional videos, each delivered
 173 either fluently or disfluently. Fluent videos featured instructors with smooth delivery and natural pacing,
 174 while disfluent videos included hesitations, monotone speech, and awkward pauses. After a distractor task,
 175 participants completed a free recall test, writing down as much content as they could remember from each
 176 video within a three-minute window. Their recall was then scored for the number of idea units correctly
 177 remembered.

Listing 1 Data needed to run examples

```
# get data here from project folder
file <- read_csv(here::here("data", "fluency_data"))
fluency_data <- read.csv(file)
```

178 Our primary outcome variable is the proportion of idea units recalled on the final test, calculated

Figure 2

A Bayesian posterior distribution (assuming a uniform prior) centered at a point estimate chosen so that the probability of direction (pd) equals 0.95, and a frequentist sampling distribution (under the null; centered at 0). In the Bayesian posterior distribution, the blue area represents the pd , and the red area represents the remaining $1 - pd$ of the distribution. In the frequentist sampling distribution, the red tail areas represent the rejection region at $\alpha = 0.10$. In this example, the posterior mean lies exactly at the $1 - \frac{\alpha}{2}$ quantile of the null sampling distribution. For symmetric posteriors with flat priors, the pd is numerically equivalent to the one-sided p -value.

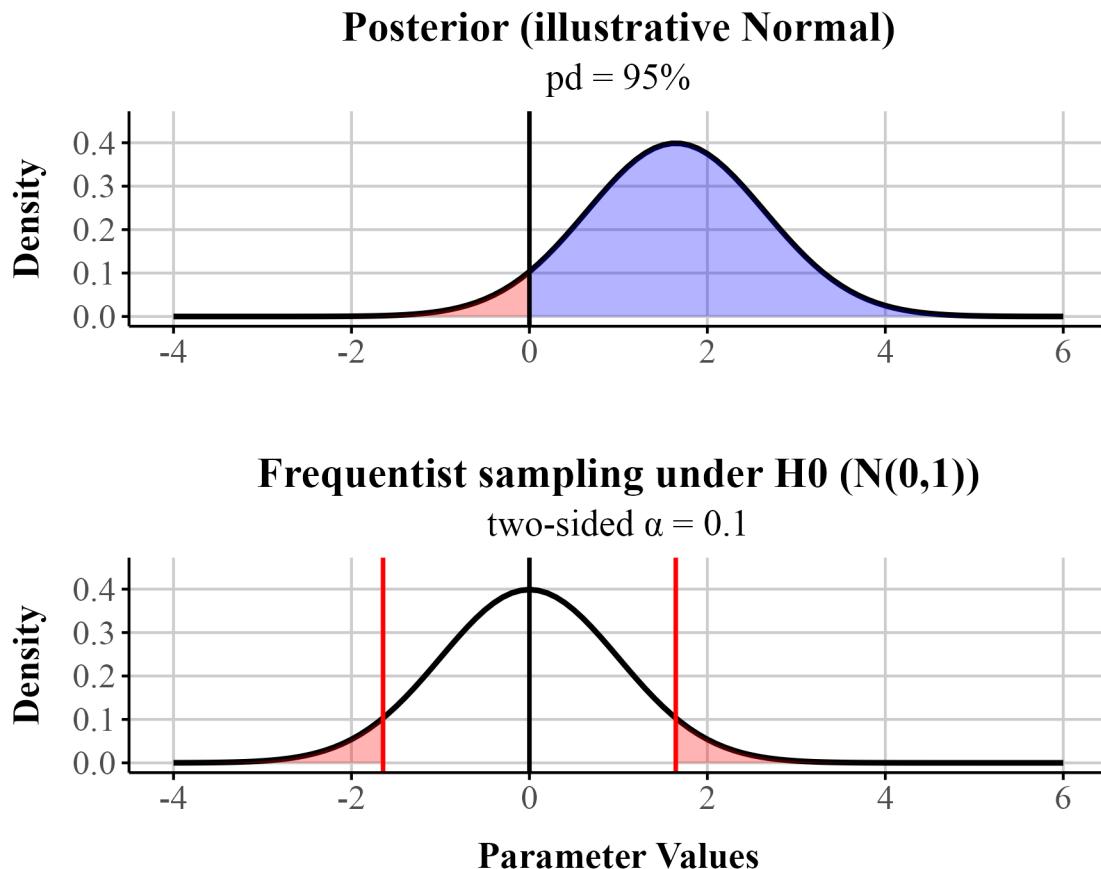


Table 1

Four observations from Wilford et al. (2020). Accuracy refers to the proportion of correctly recalled idea units.

Participant	Fluency	Accuracy
64	Disfluent	0.10
30	Fluent	0.60
12	Fluent	0.10
37	Fluent	0.35

179 by dividing the number of correct units by 10. We show a sample of these data in Table 1. The dataset
 180 can be downloaded from GitHub (Listing 1). Because this is a bounded continuous variable (i.e., it ranges
 181 from 0 to 1), it violates the assumptions of typical linear regression models that treat outcomes as normally
 182 distributed. Despite this, it remains common in psychological research to analyze proportion data using
 183 models that assume normality. In what follows, we reproduce Wilford et al. (2020)'s analysis and then
 184 re-analyze the data using beta regression and highlight how it can improve our inferences.

185 **Reanalysis of Wilford et al. Experiment 1A**

186 In their original analysis of Experiment 1A, Wilford et al. (2020) compared memory performance
 187 between fluent and disfluent instructor conditions using a traditional independent-samples t-test on mean
 188 accuracy for 96 participants. They found that participants who watched the fluent instructor recalled signifi-
 189 cantly more idea units than those who viewed the disfluent version (see Figure 3).

190 We first replicate this analysis in a regression framework using {brms}. We model final test mean
 191 accuracy—the proportion of correctly recalled idea units across the videos—as the dependent variable. Our
 192 predictor is instructor fluency, with two levels: Fluent and Disfluent. We use treatment (dummy) coding,
 193 which is the default in R. This coding scheme sets the first level of a factor (in alphabetical order) as the
 194 reference level. In this case, Disfluent is the reference, and the coefficient for Fluent reflects the contrast
 195 between fluent and disfluent instructor conditions.

196 **Regression Model**

197 We first start by loading the brms (Bürkner, 2017) and cmdstanr (Gabry et al., 2024) packages
 198 (Listing 2). We use the cmdstanr backend for Stan (Team, 2023) because it's faster than the default used to
 199 run models (i.e., rstan),⁴ though all of these models can also be fit with brms defaults.

Listing 2 Load the brms and cmdstanr packages

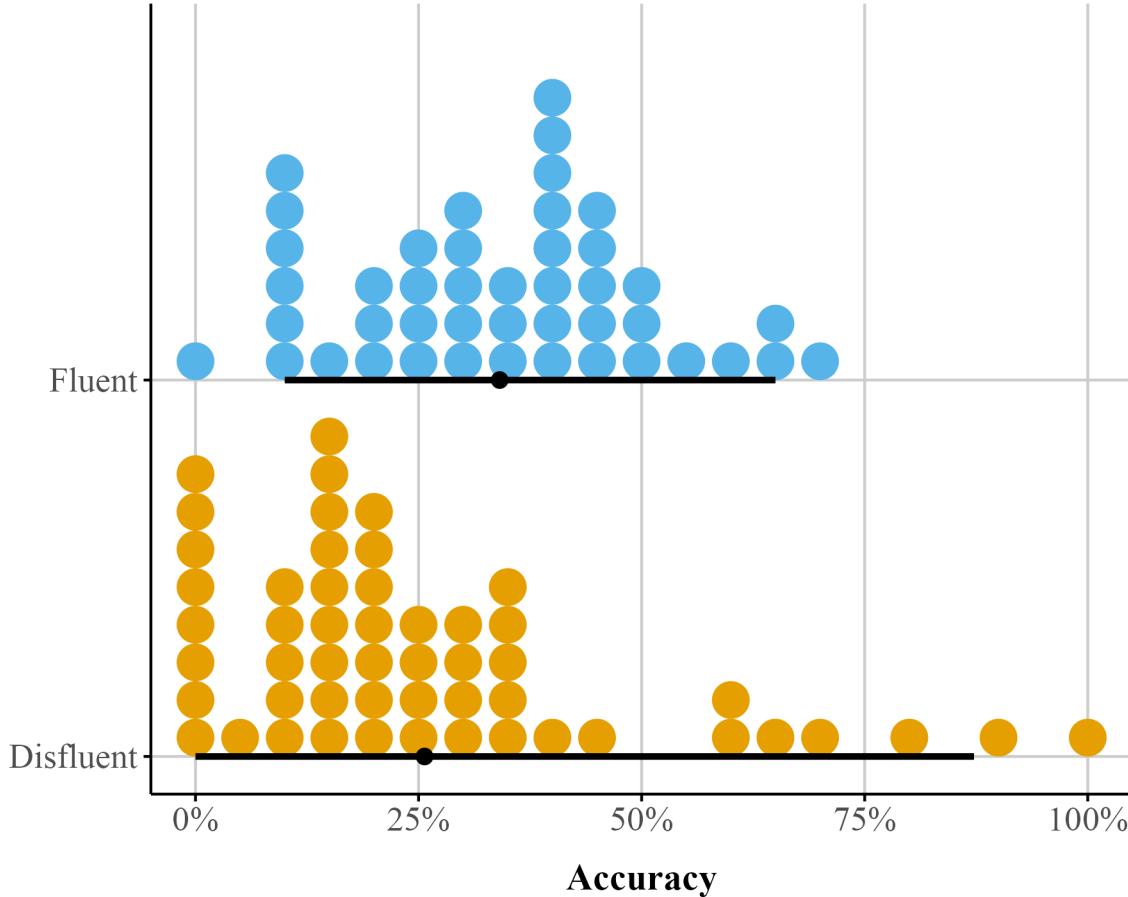
```
library(brms)
library(cmdstanr)
```

200 We fit the model using the brm() function from the brms package (Listing 3). Although not shown
 201 here, we ran the models using four chains (the default), executed in parallel across four cores. When the
 202 model is run in Listing 3, the model summary output will appear in the R console. The output from

⁴In order to use the cmdstanr backend you will need to first install the package (see <https://mc-stan.org/cmdstanr/>) and also run cmdstanr::install_cmdstan() if you have not done so already.

Figure 3

Dot plot depicting accuracy distributions for the Fluent and Disfluent conditions. Each condition shows individual data points and summary statistics (mean and 95% percentile interval) to illustrate variability and central tendency.



203 bayes_reg_model shows each parameter's posterior summary: The posterior distribution's mean and standard deviation (analogous to the frequentist standard error) and its 95% credible interval, which indicate the 204 95% of the most credible parameter values. In brms, the reported Cr.I is an equal-tailed interval, meaning 205 that the probability mass excluded from the interval is split equally between the lower and upper tails. Additionally, the output indicates numerical estimates of the sampling algorithm's performance: Rhat should 206 be close to one, and the ESS (effective sample size) metrics should be as large as possible given the number 207 of iterations specified (default is 4000). Generally, ESS ≥ 1000 is recommended (Bürkner, 2017). For the 208 models we present in this paper, convergence is trivial with standard linear models, though we note that these 209 metrics are still important to pay attention to in case of model misfit.

210 Our main question of interest is: what is the causal effect of instructor fluency on final test performance? In order to answer this question, we will have to look at the output summary produced by Listing 3 211 (also see Table 8 under Bayesian LM). the Intercept refers to the posterior mean accuracy in the disfluent 212 condition, $M = 0.257$, as fluency was dummy-coded. The fluency coefficient (FluencyFluent) reflects the 213 mean posterior difference in recall accuracy between the fluent and disfluent conditions: $b = 0.085$. The 95% 214 Cr.I for this estimate spans from 0.002 to 0.166. These values are shown in the "95% Cr.I" columns of the 215 output. These results closely mirror the findings reported by Wilford et al. (2020) (Experiment 1A).

Listing 3 Fitting a gaussian model with brm().

```

bayes_reg_model <- brm(
  Accuracy ~ Fluency,
  data = fluency_data,
  family = gaussian(),
  file = here::here("manuscript", "models", "model_reg_bayes")
)

219 Family: gaussian
220 Links: mu = identity; sigma = identity
221 Formula: Accuracy ~ Fluency
222 Data: fluency_data (Number of observations: 96)
223
224 Regression Coefficients:
225             Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
226 Intercept      0.26     0.03    0.20    0.32 1.00    3834    2294
227 FluencyFluent  0.08     0.04    0.00    0.17 1.00    4006    2934
228
229 Further Distributional Parameters:
230             Estimate Est.Error 1-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
231 sigma       0.21     0.02    0.18    0.24 1.00    3917    2945

```

232 The output also includes the effective sample size (ESS) and R (R-hat) values, both of which fall
 233 within acceptable ranges, indicating good model convergence. Throughout the tutorial, we focus primarily
 234 on posterior mean estimates and their 95% credible intervals. In addition, we report the pd measure in the
 235 main summary table (Table 8), provided by the {bayestestR} package (Makowski, Ben-Shachar, Chen, et
 236 al., 2019; Makowski, Ben-Shachar, & Lüdecke, 2019). This measure offers an intuitive parallel to *p*-values,
 237 which many readers may find familiar. For example, the fluency effect has a pd of .977, indicating a high
 238 probability that the effect is positive rather than negative (akin to *p* < .05).

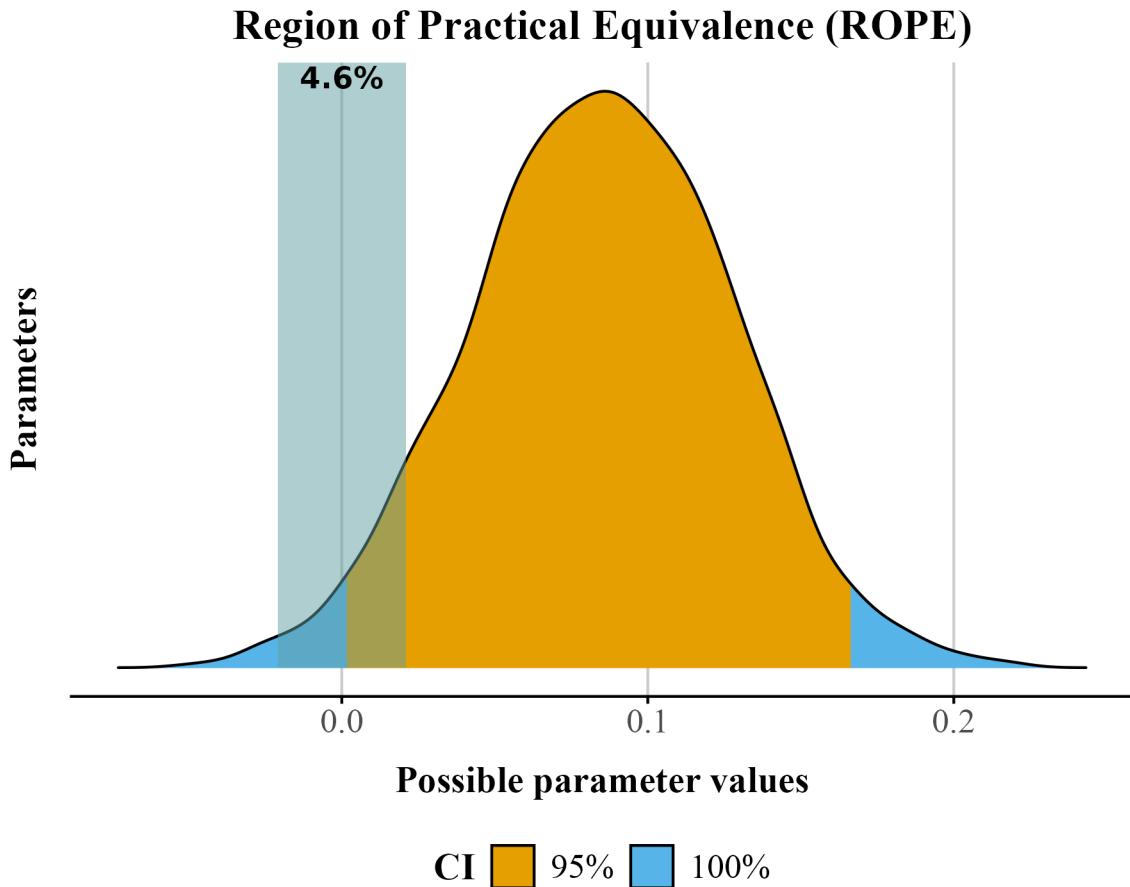
239 Importantly, pd does not indicate whether an effect is meaningfully different from a point value—it
 240 only reflects the proportion of the posterior in one direction. To address questions of practical significance,
 241 we encourage readers to consider the Region of Practical Equivalence (ROPE) with the Cr.Is (Kruschke,
 242 2015). Unlike a hypothesis test of a point null (e.g., exactly zero), the ROPE defines a range of values that
 243 are deemed too small to be of substantive importance. As a rule of thumb (see Kruschke, 2018), if more than
 244 95% of the posterior distribution lies inside the ROPE, the effect can be considered practically equivalent
 245 to that negligible range. If less than 5% lies inside, the effect can be considered meaningfully different.
 246 Intermediate cases are typically labeled undecided.

247 The rope() function in the {bayestestR} package computes the proportion of the posterior within
 248 the ROPE to facilitate this evaluation. By default, from bayesian models fit via brms the package determines
 249 a ROPE based on the data (roughly reflecting “negligible” effects), but these defaults should be used cau-
 250 tiously. The choice of ROPE should ultimately be guided by theoretical considerations, prior research, and
 251 the substantive context of the study. In ?@lst-rope-brms, we show how to compute this using bayestestR.
 252 Running the function with default settings suggests that less than 5% of the posterior distribution lies within
 253 the default ROPE (indicating the effect is larger than .02) (see Figure 4). Going forward we do not include a
 254 discussion of ROPE values, but we encourage readers to adopt it in their own research when appropriate.

255 Wilford et al. (2020) observed that instructor fluency impacts actual learning, using a standard *t*-test

Figure 4

Posterior distribution for the fluency effect showing the ROPE(shaded area) with 95% credible interval (orange) and 100% credible interval (blue). The percentage indicates the proportion of the posterior within the ROPE.



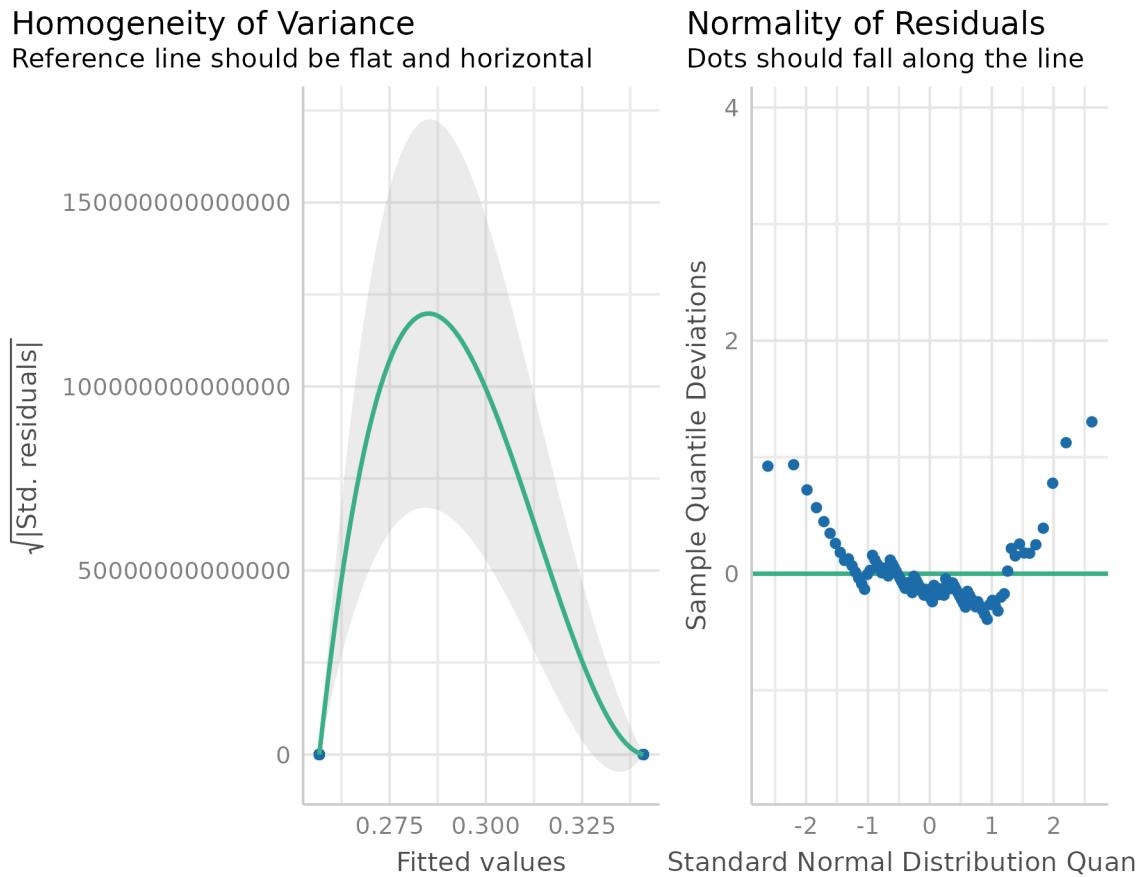
256 on the mean accuracy. But recall this approach assumes normality of residuals and homoscedacity. These
 257 assumptions are unrealistic when the response values approach the scale boundaries (Sladekova & Field,
 258 2024). Does the data we have meet those assumptions? We can use the function `check_model()` from
 259 `{easystats}` (Lüdecke et al., 2022) to check our assumptions easily. The code in Listing 4 automatically
 260 produces Figure 5. We can see some issues with our data. Specifically, there appears to be violations of
 261 constant variance across the values of the scale (homoskedasticity). In plain terms, this type of model mis-
 262 specification means that a standard OLS model can predict non-sensical values outside the bounds of the
 263 scale.

Listing 4 Checking assumptions with the `check_model()` from `easystats` package .

```
check_model(bayes_reg_model, check = c("homogeneity", "qq"))
```

Figure 5

Two assumption checks for our OLS model: Normality (right) and Homoskedasticity (left)



264 We can also examine how well the data fits the model by performing a posterior predictive check
 265 using the `pp_check()` function from `{brms}`. A posterior predictive check involves looking at multiple
 266 draws or repetitions from the posterior distribution and plotting it against the observed data. Ideally, the
 267 predictive draws (the light blue lines) should show reasonable resemblance with the observed data (dark blue
 268 line). In our example (see Figure 12 (A)) the model-predicted density is slightly too peaked and narrow
 269 compared to the data. In addition, some of the draws extend into negative accuracy values.

270 Distributional Regression - Beta Regression

271 It is important to note that there are several justifiable approaches for addressing the distributional
 272 issues observed in the data. For instance, one could analyze median accuracy instead of the mean, use some
 273 type of robust estimator for heterogeneity, or apply non-parametric methods to relax some of the model as-
 274 sumptions. However, in a Bayesian framework, we can address these issues more directly and transparently
 275 by fitting distributional models (Kneib et al., 2023; Kruschke, 2013). A key advantage of Bayesian distribu-
 276 tional modeling approach is that we are not limited to modeling only the mean or median of the outcome;
 277 we can also model parameters such as the variance (or other shape parameters) as functions of predictors.
 278 This allows us to examine how instructor fluency may influence not only average performance, but also the
 279 variability in performance across students. If we wanted to keep our mean accuracy variable and continue to
 280 use a Gaussian model, we could use a distributional approach and model the effect of fluency on σ .

Given the outcome variable is proportional, another solution would be to run a beta regression model. Again, we can create the beta regression model in `brms`. In `brms`, we model each parameter independently. Recall from the introduction that in a beta model we model two parameters— μ and ϕ . Again we do this by using the `bf()` function from `brms` (Listing 5). We fit two formulas, one for μ and one for ϕ and store it in the `model_beta_bayes` object below. In the below `bf()` call, we are modeling Fluency as a function of Accuracy only for the μ parameter. For the ϕ parameter, we are only modeling the intercept value. This is saying dispersion does not change as a function of fluency.

To run our beta regression model, we need to exclude 0s and 1s in our data set. If we try to run a model with our data `data_fluency` we get an error: `Error: Family 'beta' requires response greater than 0`. This is because the beta distribution only supports observations in the 0 to 1 interval *excluding exact 0s and 1s*. We need make sure there are no 0s and 1s in our dataset.

The dataset contains nine 0s and one 1. One approach is to nudge our 0s towards .01 and our 1s to .99, or apply a special formula (Smithson & Verkuilen, 2006) so values fall within the [0, 1] interval. We implore readers not to engage in this practice. Kubinec (2022) showed that this practice can result in serious distortion of the outcome as the sample size grows larger, resulting in ever smaller values that are “nudged”. Because the beta distribution is a non-linear model of the outcome, values that are very close to the boundary, such as 0.00001 or 0.99999, will be highly influential outliers. To run this beta model we will remove the 0s and 1s, and later in this article we will show how to jointly model these scale end points with the rest of the data. The model from Listing 5 uses a transformed `data_fluency` object (called `data_beta`) where 0s and 1s are removed. When we run this code we should not get an error.

Listing 5 Fitting a beta model without 0s and 1s in `brm()`.

```
# set up model formual
model_beta_bayes <- bf(
  Accuracy ~ Fluency, # fit mu model
  phi ~ 1 # fit phi model
)

# transform 0 to 0.1 and 1 to .99
data_beta <- fluency_data |>
  filter(
    Accuracy != 0,
    Accuracy != 1
  )

beta_brms <- brm(
  model_beta_bayes,
  data = data_beta,
  family = Beta(),
  file = here::here("manuscript", "models", "model_beta_bayes_reg_01")
)
```

Model Parameters. In Table 8, under the beta regression column, the coefficient with `b_` represents how fluency of instructor influences the μ parameter estimates (which is the mean of the distribution here). These coefficients are interpreted on the scale of the logit, meaning they represent linear changes on a non-linear space. The intercept term (`b_Intercept`) represents the log odds of the mean on accuracy for the

305 fluent instructor. Log odds that are negative indicate that it is more likely a “success” (like getting the correct
 306 answer) will NOT happen than that it will happen. Similarly, regression coefficients in log odds forms that
 307 are negative indicate that an increase in that predictor leads to a decrease in the predicted probability of a
 308 “success”.

309 The other component we need to pay attention to is the dispersion or precision parameter coefficients
 310 labeled as phi in Table 8. The dispersion (ϕ) parameter tells us how precise our estimate is. Specifically,
 311 ϕ in beta regression tells us about the variability of the response variable around its mean. Specifically, a
 312 higher dispersion parameter indicates a narrower distribution, reflecting less variability. Conversely, a lower
 313 dispersion parameter suggests a wider distribution, reflecting greater variability. The main difference between
 314 a dispersion parameter and the variance is that the dispersion has a different interpretation depending on the
 315 value of the outcome, as we show below. The best way to understand dispersion is to examine visual changes
 316 in the distribution as the dispersion increases or decreases.

317 Understanding the dispersion parameter helps us gauge the precision of our predictions and the con-
 318 sistency of the response variable. In `beta_brms` we only modeled the dispersion of the intercept. When
 319 ϕ is not specified, the intercept is modeled by default (see Table 8). It represents the overall dispersion in
 320 the outcome across all conditions. Instead, we can model different dispersions across levels of the Fluency
 321 factor. To do so, we add Fluency to the phi model in `bf()`. We model the precision (phi) of the Fluency
 322 factor by using a ~ and adding factors of interest to the right of it (Listing 6).

Listing 6 Fitting beta model with dispersion in `brm()`.

```
model_beta_bayes_disp <- bf(
  Accuracy ~ Fluency, # Model of the mean
  phi ~ Fluency # Model of the precision
)

beta_brms_dis <- brm(
  model_beta_bayes_disp,
  data = data_beta,
  family = Beta(),
  file = here::here("manuscript", "models", "model_beta_bayes_dis_run01")
)
```

323 Table 8 displays the model summary with the precision parameter labeled as `phi_Fluency`. Since ϕ
 324 is modeled on the log scale, the coefficients represent changes in $\log-\phi$ rather than ϕ itself. To see the effect
 325 in the original units, we convert the values back by exponentiating. Thus, the effect of the Fluent condition
 326 can be understood by comparing the exponentiated predicted ϕ in the Fluent condition to that in the baseline
 327 condition.

328 The ϕ parameters are estimated on the log scale. The term $\beta_{\text{Intercept}}^{(\phi)}$ represents the log-precision for
 329 the reference (disfluent) condition. The coefficient $\beta_{\text{FluencyFluent}}^{(\phi)}$ represents the change in log-precision when
 330 moving from the disfluent to the fluent condition.

331 To obtain precision on the original scale, we exponentiate the linear predictor:

$$\phi_{\text{disfluent}} = \exp(\beta_{\text{Intercept}}^{(\phi)}), \quad \phi_{\text{fluent}} = \exp(\beta_{\text{Intercept}}^{(\phi)} + \beta_{\text{FluencyFluent}}^{(\phi)}).$$

332 The coefficient $\beta_{\text{FluencyFluent}}^{(\phi)}$ therefore describes a *multiplicative* change in precision. Specifically,

$$\frac{\phi_{\text{fluent}}}{\phi_{\text{disfluent}}} = \exp(\beta_{\text{FluencyFluent}}^{(\phi)}).$$

Because the 95% credible interval for $\beta_{\text{FluencyFluent}}^{(\phi)}$ does not include zero, we infer that there is a credible difference in precision between the fluent and disfluent conditions.

It is important to note that these estimates are not the same as the marginal effects we discussed earlier. Changes in dispersion affect the spread or variability of the response distribution without necessarily altering its mean. This makes dispersion particularly relevant for research questions that focus on features of the distribution beyond the average—such as how concentrated responses are. For instance, high dispersion might indicate that individuals cluster at the extremes (e.g., very high or very low ratings), suggesting clustering in the outcome.

A critical assumption of the linear model is homoscedasticity, which means constant variance of the errors. WIth beta regression model we can include a dispersion parameter for Fluency. Properly accounting for dispersion is crucial because it impacts the precision of our mean estimates and, consequently, the significance of our coefficients. The inclusion of dispersion in the our model increased the uncertainty of the μ coefficient (see Figure 6). This suggests that failing to account for the dispersion of the variables might lead to biased estimates. This highlights the potential utility of an approach like beta regression over a traditional approach as beta regression can explicitly model dispersion and address issues of heteroscedasticity.

While it is advisable to model precision, if there is uncertainty about the best model, a relatively agnostic approach would be to compare models, for example with leave one out (loo) cross validation, to examine if a dispersion parameter should be considered in our model.⁵

351 ***Predicted Probabilities***

Parameter estimates are usually difficult to interpret on their own and can require a lot of mathematical gymnastics to get the estimate you need. We argue that researchers should not spend too much time interpreting raw coefficients from non-linear models. We report them in this tutorial for completeness. Instead researchers should discuss the effects of the predictor on the actual outcome of interest (in this case the 0-1 scale). The logit link allows us to transform back and forth between the scale of a linear model and the nonlinear scale of the outcome, which is bounded by 0 and 1. By using the inverse of the logit, we can easily transform our linear coefficients to obtain average effects on the scale of the proportions or percentages, which is usually what is interesting to applied researchers. In a simple case, we can do this manually, but when there are many factors in your model this can be quite complex.

In our example, we can use the `plogis()` function in base R to convert estimates from the logit scale to the probability scale. The intercept of our model is -0.918, which reflects the logit of the mean accuracy in the disfluent condition. If the estimated difference between the fluent and disfluent conditions is 0.24 on the logit scale, we first add this value to the intercept value (-0.918) to get the logit for the fluent condition: $-0.83 + 0.20 = -0.63$. We then use `plogis()` to convert both logit values to probabilities (Fluent = 35%, Disfluent = 30%).

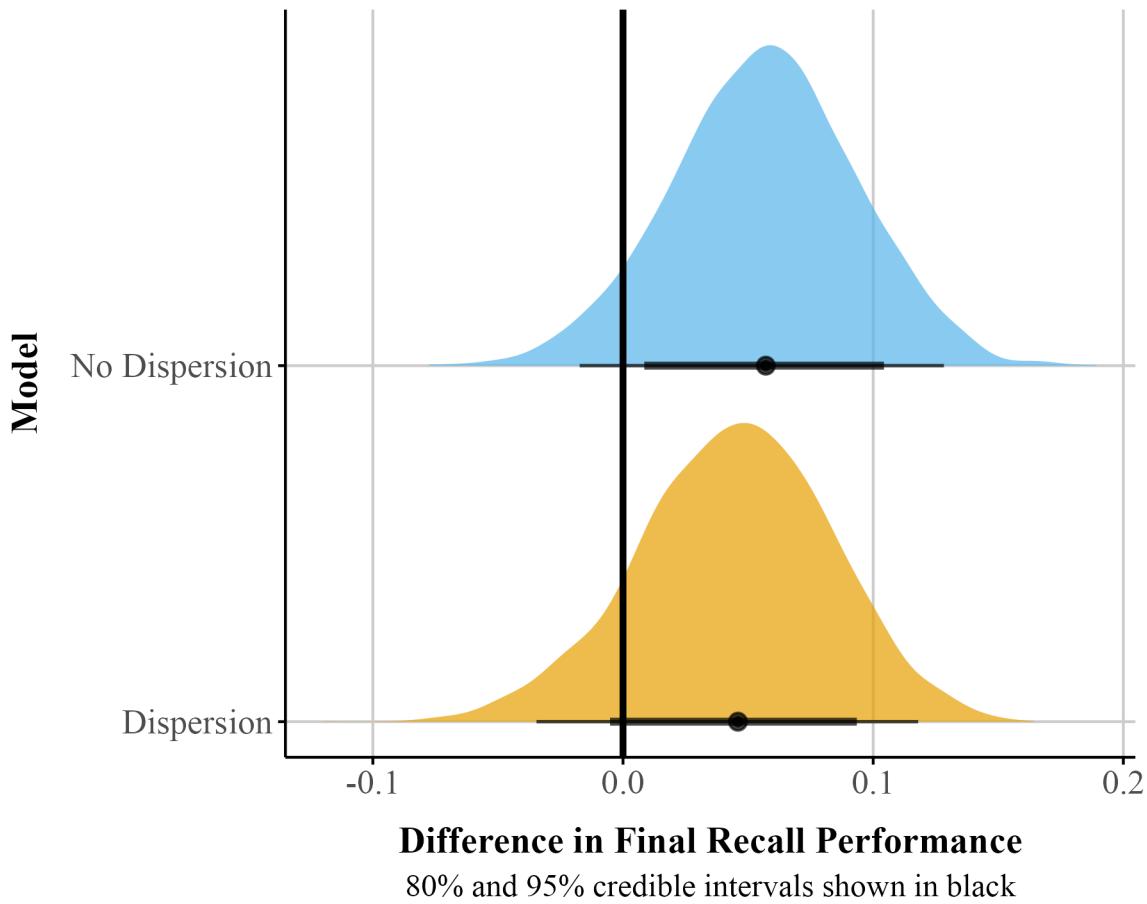
This is pretty easy to do manually, but when your model has many predictors, it can be quite cumbersome. To help us extract predictions from our model and visualize them we will use a package called `{marginaleffects}` (Arel-Bundock et al., 2024) (see Listing 7). To get the proportions for each of our categorical predictors on the μ parameter we can use the function from the package called `predictions()`. These are displayed in Table 2. These probabilities match what we calculated above.

For the Fluency factor, we can interpret Mean as proportions or percentages. That is, participants who watched the fluent instructor scored on average 35% on the final exam compared to 30% for

⁵The model fit statistic LOO-CV can be compared for any set of fitted `brms` models with the function `loo()`.

Figure 6

Comparison of posterior distributions for the risk difference in fluency: Simple model (no dispersion for Fluency) vs. complex model with dispersion

**Table 2**

Predicted probabilities for fluency factor.

Fluency	Mean	95% Cr.I
Disfluent	0.297	[0.25, 0.347]
Fluent	0.354	[0.301, 0.41]

Listing 7 Load the {marginaleffects} package.

```
library(marginaleffects)
options(marginaleffects_posterior_center = mean) # make sure returns mean
```

Listing 8 Predictions from the beta model for each level of Fluency.

```
predictions(
  beta_brms,
  # need to specify the levels of the categorical predictor
  newdata = datagrid(Fluency = c("Disfluent", "Fluent"))
)
```

374 those who watched the disfluent instructor. We can also visualize these from {marginaleffects} using the
375 plot_predictions() function (see Listing 9).

Listing 9 Plot predicted probabilities using plot_predictions() from {marginaleffects}

```
beta_plot <- plot_predictions(beta_brms, by = "Fluency")
```

376 The plot_predictions() function will only display the point estimate with the 95% credible inter-
377 val. However, Bayesian estimation methods generate distributions for each parameter. This approach allows
378 visualizing full uncertainty estimates beyond points and intervals. Using the {marginaleffects} package, we
379 can obtain samples from the posterior distribution with the posterior_draws() function (see Listing 10).
380 We can then plot these results to illustrate the range of plausible values for our estimates at different levels
381 of uncertainty (see Figure 7).

Listing 10 Extracting posterior draws from the beta regression model.

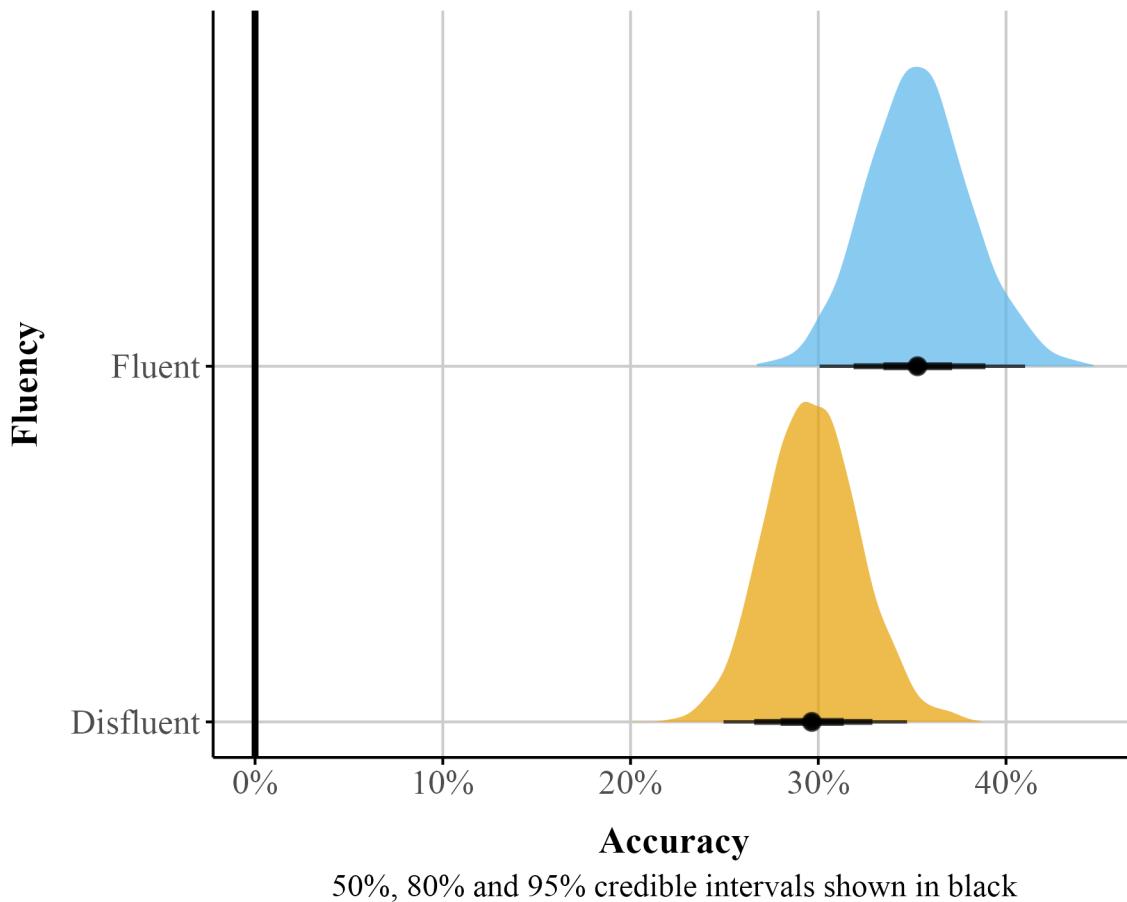
```
# Add a model identifier to each dataset
pred_draws_beta <- avg_predictions(beta_brms, variables = "Fluency") |>
  posterior_draws()
```

382 Marginal Effects

383 Marginal effects offer an interpretable way to quantify how changes in a predictor influence an out-
384 come, while holding other factors constant in a specific manner. **In recent years, there has been a thrust to**
385 **move away from reporting regression coefficients alone, focusing instead on estimates that are easier to**
386 **interpret and communicate—particularly in non-linear models (McCabe et al., 2021; Rohrer & Arel-**
387 **Bundock, 2025).** Technically, marginal effects are computed as partial derivatives for continuous variables
388 or as finite differences for categorical (and sometimes continuous) predictors, depending on the structure
389 of the data and the research question. Substantively, these procedures translate raw regression coefficients
390 into quantities that reflect changes in the bounded outcome—for example, an $x\%$ change in the value of a
391 proportion.

Figure 7

Predicted probability posterior distributions by fluency

**Table 3**

Fluency difference

392 There are various types of marginal effects, and their calculation can vary across software packages.

393 For example, the popular `{emmeans}` package (Lenth, 2025) computes marginal effects by holding all pre-
394 dictors at their means (MEM). In this tutorial, we will use the `{marginaleffects}` package (Arel-Bundock et
395 al., 2024), which focuses on average marginal effects (AMEs) by default. AMEs summarize effects by gen-
396 erating predictions for each row of the original dataset and then averaging these predictions. This approach
397 retains a strong connection to the original data while offering a straightforward summary of the effect of
398 interest.

399 One practical use of AMEs is to estimate the average difference between two groups or conditions

400 which corresponds to the average treatment effect (ATE). Using the `avg_comparisons()` function in the
401 `{marginaleffects}` package (Listing 11), we can compute this quantity directly. By default, the function returns
402 the discrete difference between groups. When we take the difference in proportions between two groups it
403 is called the risk difference. Depending on the audience and modeling goals, the function can also produce
404 alternative effect size metrics, such as odds ratios or risk ratios. This flexibility makes it a powerful approach
405 for summarizing and communicating regression results.

Listing 11 Calculating the difference between probabilities with `avg_comparisons()`

```
# get risk difference by default
beta_avg_comp <- avg_comparisons(beta_brms, comparison = "difference")
```

406 Table 3 presents the estimated difference for the Fluency factor (Mean column). The difference
 407 between the fluent and disfluent conditions is 0.06, indicating that participants who watched a fluent instructor
 408 scored, on average, 6% higher on the final recall test than those who watched a disfluent instructor. However,
 409 the 95% credible interval includes 0 among the most credible values, suggesting we cannot rule out the
 410 possibility of a null or weakly negative effect.

411 We can also use `{marginaleffects}` to get the actual precision difference between the two groups on
 412 ϕ using similar code to above by setting `dpar` to “phi” {Listing 12}.

Listing 12 Calculating ϕ difference with `avg_comparisons()`

```
# get risk difference by default

beta_avg_phi <- avg_comparisons(
  beta_brms_dis,
  dpar = "phi",
  comparison = "difference"
)
```

413 In psychology, it is common to report effect size measures like Cohen’s d (Cohen, 1977). When
 414 working with proportions we can calculate something similar called Cohen’s h . Taking our proportions, we
 415 can use the below equation (Equation 2) to calculate Cohen’s h along with the 95% Cr.I around it. Using this
 416 metric we see the effect size is small (0.107), 95% credible interval [-0.002, 0.361].

$$h = 2 \cdot (\arcsin(\sqrt{p_1}) - \arcsin(\sqrt{p_2})) \quad (2)$$

417 **Posterior Predictive Check**

418 Figure 12 (B) shows the predictive check for our beta model. The model does a pretty good job
 419 at capturing the data (The draws are now between 0-1) and the predicted values from the model follow the
 420 observed data. However, it could be better.

421 **Zero-Inflated beta (ZIB) Regression**

422 A limitation of the beta regression model is that it can only accommodate values strictly between 0
 423 and 1—it cannot handle values exactly equal to 0 or 1. In our dataset, we observed 9 rows where Accuracy
 424 equals zero. To fit a beta regression model, we removed these values, but we have left out potentially valuable
 425 information from our model—especially if the end points of the scale are distinctive in some way. In our case,
 426 these 0s may be structural—that is, they represent real, systematic instances where participants failed to
 427 answer correctly (rather than random noise or measurement error). For example, the fluency of the instructor
 428 might be a key factor in predicting these zero responses. We will discuss two approaches for jointly modeling
 429 these end points with the continuous data. First, we can use a zero-inflated beta (ZIB) model. This model

430 still estimates the mean (μ) and precision (ϕ) of the beta distribution for values between 0 and 1, but it also
 431 includes an additional parameter, α , which captures the probability of observing structural 0s.

432 The zero-inflated beta models a mixture of the data-generating process. The α parameter uses a
 433 logistic regression to model whether the data is 0 or not. Substantively, this could be a useful model when we
 434 think that 0s come from a process that is relatively distinct from the data that is greater than 0. For example,
 435 if we had a dataset with proportion of looks or eye fixations to certain areas on marketing materials, we might
 436 want a separate model for those that do not look at certain areas on the screen because individuals who do
 437 not look might be substantively different than those that look.

438 We can fit a ZIB model using `brms()` and use the `{marginaleffects}` package to make inferences
 439 about our parameters of interest. Before we run a zero-inflated beta model, we will need to transform our
 440 data again and remove the one 1 value in our data—we can keep our 0s. Similar to our beta regression model
 441 we fit in `brms`, we will use the `bf()` function to fit several models. We fit our μ and ϕ parameters as well as
 442 our zero-inflated parameter (α ; here labeled as `zi`). In `brms` we can use the `zero_inflated_beta` family (see
 443 Listing 13).

Listing 13 Fitting zib model with `brm()`

```
# keep 0 but remove 1
data_beta_0 <- fluency_data |>
  filter(Accuracy != 1)

# set up model formula for zero-inflated beta in brm
zib_model <- bf(
  Accuracy ~ Fluency, # The mean of the 0-1 values, or mu
  phi ~ Fluency, # The precision of the 0-1 values, or phi
  zi ~ Fluency, # The zero-or-one-inflated part, or alpha
  family = zero_inflated_beta()
)

# fit zib model with brm
fit_zi <- brm(
  formula = zib_model,
  data = data_beta_0,
  file = here::here("manuscript", "models", "bayes_zib_model0not1.rds")
)
```

444 Posterior Predictive Check

445 The ZIB model does a bit better at capturing the structure of the data than the beta regression model
 446 (see Figure 12). Specifically, the ZIB model more accurately captures the increased density of values near
 447 the lower end of the scale (i.e., near zero), which the standard beta model underestimates. The ZIB model's
 448 predictive distributions also align more closely with the observed data across the entire range, particularly in
 449 the peak and tail regions. This improved fit likely reflects the ZIB model's ability to explicitly model excess
 450 0s (or near-zero values) via its inflation component, allowing it to better account for features in the data that
 451 a standard beta distribution cannot accommodate.

Table 4*Probability fluency difference (μ)*

Contrast	Mean	95% Cr.I	pd
Fluent - Disfluent	0.044	[-0.031, 0.117]	0.872

Table 5*Probability fluency difference (ϕ)*

Contrast	Mean	95% Cr.I	pd
Fluent - Disfluent	2.69	[-0.924, 6.6]	0.926

452 Predicted Probabilities and Marginal Effects

453 Table 8, under the zero-inflated beta regression column, provides a summary of the posterior distribution
 454 for each parameter. As stated before, it is preferable to back-transform our estimates to get probabilities.
 455 To get the predicted probabilities we can again use the `avg_predictions()` and `avg_comparisons()`
 456 functions from `{marginaleffects}` package (Arel-Bundock, 2024) to get predicted probabilities and the prob-
 457 ability difference between the levels of each factor. We can model the parameters separately using the `dpar`
 458 argument setting to: μ , ϕ , α . Here we look at the risk difference for Fluency under each parameter. If one
 459 were interested in the average effect for the entire model, the `dpar` argument could be removed.

460 **Mu.** As shown in Table 4, there is little evidence for an effect of Fluency – the 95% Cr.I includes
 461 zero, suggesting substantial uncertainty about the direction and magnitude of the effect—that is, though most
 462 of the posterior density supports positive effects, nil and weakly negative effects cannot be ruled out.

463 **Dispersion.** As shown in Table 5, the posterior estimates suggest a credible effect of Fluency on
 464 dispersion (ϕ), with disfluent responses showing greater variability. The 95% Cr.I for the fluency contrast
 465 does not include zero, indicating a high probability in differences in precision.

466 Zero-Inflation

467 We can use `{marginaleffects}` to estimate and plot the posterior difference between the fluent and
 468 disfluent conditions (see Figure 8). In Figure 8, the posterior distribution for this contrast lies mostly below
 469 zero, indicating that a fluent instructor is associated with a lower probability of zero responses. The estimated
 470 reduction is approximately 13%. The 95% credible interval does not include zero, which indicates that the
 471 data provide consistent evidence for a reduction in zero responses under fluent instruction.

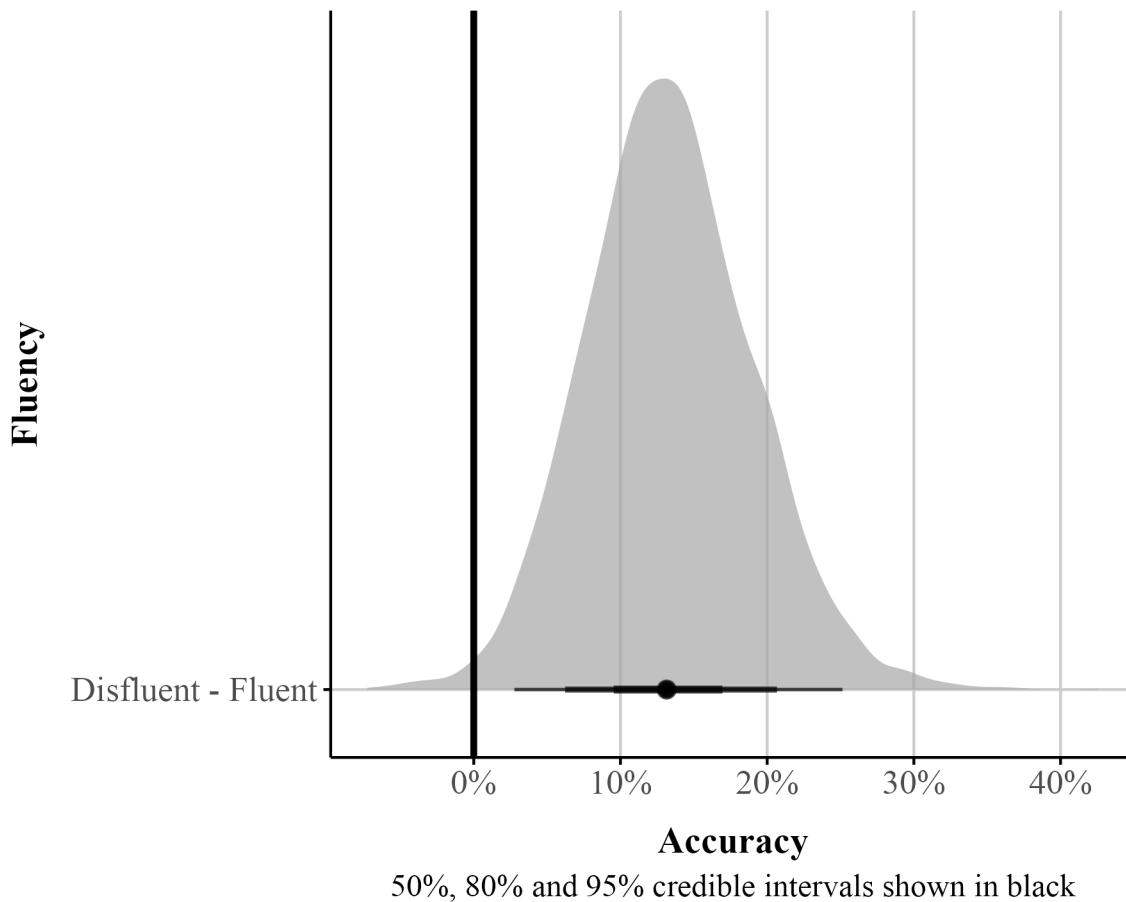
472 Zero-One-Inflated beta (ZOIB)

473 The ZIB model works well if you have 0s in your data, but not 1s.⁶ In our previous examples we
 474 either got rid of both 0s and 1s (beta regression), or removed the 1s (ZIB). Sometimes it is theoretically useful
 475 to model both 0s and 1s as separate processes or to consider these values as essentially similar parts of the
 476 continuous response, as we show later in the ordered beta regression model. For example, this is important
 477 in visual analog scale data where there might be a prevalence of responses at the bounds (Kong & Edwards,

⁶In cases where your data include exact 1s but no 0s, you can fit a one-inflated beta regression model in `brms` by setting the `coi` parameter to 1. This tells the model that all point masses occur at 1, rather than being split between 0 and 1. In other words, `coi = 1` assumes that any inflation in the data is due entirely to values at 1. In our data, we have exactly one value equal to 1^[^6]. While probably not significant to alter our findings, we can model 1s with a special type of model called the zero-one-inflated beta (ZOIB) model (Liu & Kong, 2015) if we believe that both 0s and 1s are distinct outcomes.

Figure 8

Visualization of the predicted difference for zero-inflated part of model



478 2016), in JOL tasks (Wilford et al., 2020), or in a free-list task where individuals provide open responses to
 479 some question or topic which are then recoded to fall between 0-1 (Bendixen & Purzycki, 2023). Here 0s
 480 and 1s are meaningful; 0 means item was not listed and 1 means the item was listed first.

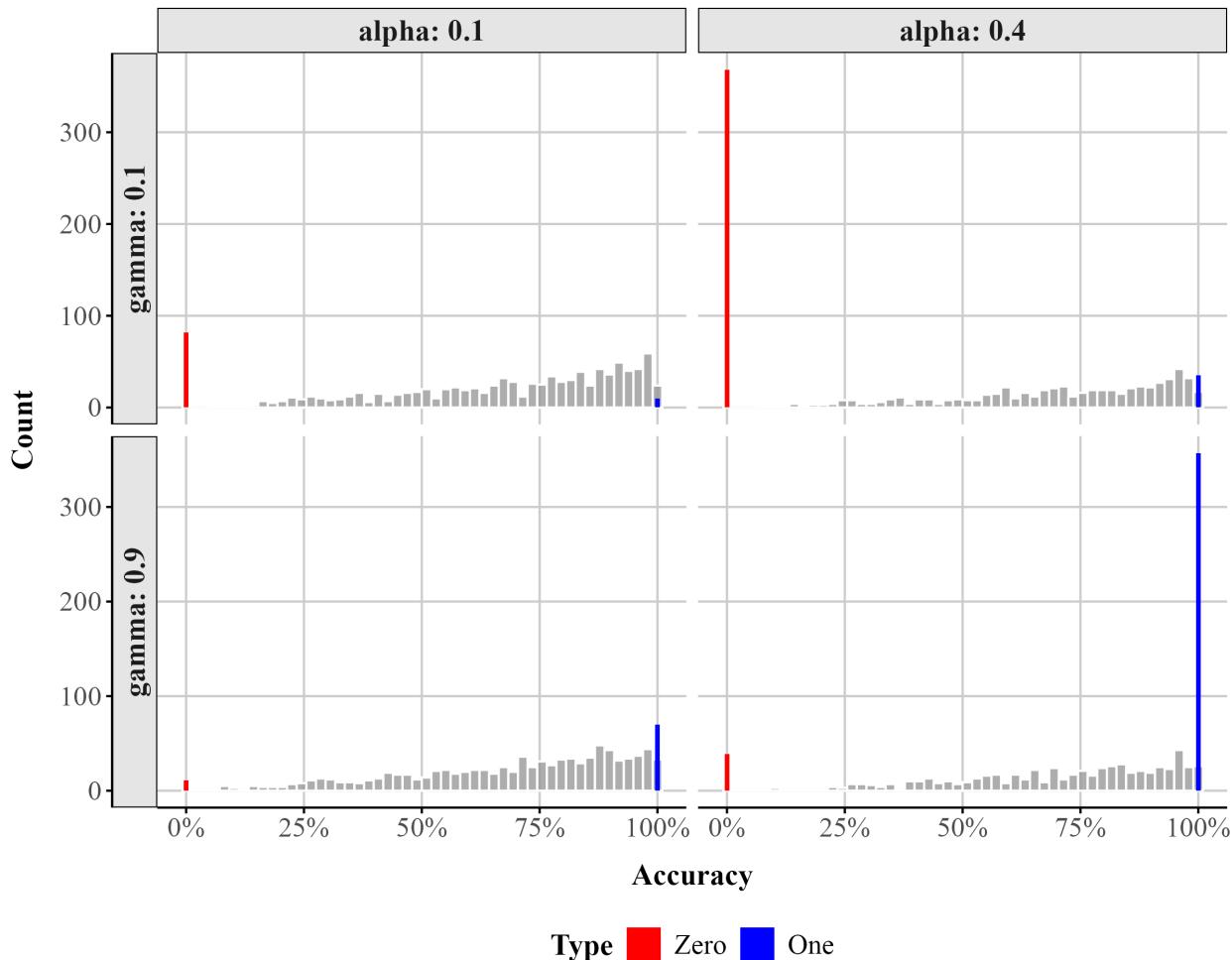
481 Similar to the beta and zero-inflated models discussed above, we can fit a zero-and-one-inflated beta
 482 (ZOIB) model in `{brms}` using the `zero_one_inflated_beta` family. This formulation simultaneously
 483 estimates the mean μ and precision ϕ of the Beta component, as well as two inflation parameters: α , the
 484 probability that an observation is at either boundary (0 or 1), and γ , the conditional probability that, given
 485 an observation falls on a boundary, it takes the value 1 rather than 0. In other words, α determines how often
 486 responses occur exactly at the endpoints, and γ determines the balance between zeros and ones among those
 487 endpoint values. This specification allows the model to capture both the continuous variation in the interior
 488 of the (0, 1) interval and the presence of exact boundary values.

489 To illustrate how α and γ shape the distribution, Figure 9 displays simulated data across a range
 490 of parameter combinations. As α increases, more responses occur at the endpoints. As γ increases, the
 491 proportion of those endpoint responses that are 1 increases relative to 0, producing increasingly pronounced
 492 spikes at 1 as γ approaches 1. Together, these parameters give the ZOIB model the flexibility to represent
 493 datasets with mixtures of continuous values and exact zeros and ones.

494 To fit a ZOIB model we use the `bf()` function. We model each parameter as a function of Fluency.
 495 We then pass the `zoib_model` to our `brm()` function (see Listing 14). The summary of the output is in

Figure 9

Simulated data from a ZOIB model illustrating the effects of the zero-one inflation parameter (α) and the conditional one-inflation parameter (γ).



496 Table 8 (under ZOIB).

497 **Model Parameters**

498 The output for the model is lengthy because we are estimating four distinct components, each with
 499 their own independent responses and sub-models. All the coefficients are on the logit scale, except ϕ , which is
 500 on the log scale. Thankfully drawing inferences for all these different parameters, plotting their distributions,
 501 and estimating their average marginal effects looks exactly the same—all the brms and {marginaleffects}
 502 functions we used work the same.

503 **Predictions and Marginal Effects**

504 With {marginaleffects} we can choose marginalize over all the sub-models, averaged across the 0s,
 505 continuous responses, and 1s in the data, or we can model the parameters separately using the dpar argument
 506 like we did above setting it to: μ , ϕ , α , γ (see below). Using avg_predictions() and not setting dpar we
 507 can get the predicted probabilities across all the sub-models. We can also plot the overall difference between

Listing 14 Fitting a ZOIB model with `brm()`.

```
# fit the zoib model

zoib_model <- bf(
  Accuracy ~ Fluency, # The mean of the 0-1 values, or mu
  phi ~ Fluency, # The precision of the 0-1 values, or phi
  zoi ~ Fluency, # The zero-or-one-inflated part, or alpha
  coi ~ Fluency, # The one-inflated part, conditional on the 1s, or gamma
  family = zero_one_inflated_beta()
)

fit_zoib <- brm(
  formula = zoib_model,
  data = fluency_data,
  file = here::here("manuscript", "models", "bayes_zoib_model")
)
```

508 fluency and disfluency for the whole model with `plot_predictions()`.

509 In addition, we show below how one can extract the predicted probabilities and marginal effects for
510 γ (and a similar process for any other model component, `zoi`, etc.):

Listing 15 Extracting predicted probabilities and marginal effects for conditional-one parameter

```
# get average predictions for coi param
coi_probs <- avg_predictions(fit_zoib, by = c("Fluency"), dpar = "coi")
# get difference between the two conditions
coi_me <- avg_comparisons(fit_zoib, variables = c("Fluency"), dpar = "coi")
```

511 **Ordered Beta Regression**

512 Looking at the output from the ZOIB model (Table 8), we can see how running a model like this
513 can become fairly complex as it is fitting distinct sub-models for each component of the scale. The ability
514 to consider 0s and 1s as distinct processes from continuous values comes at a price in terms of complexity
515 and interpretability. A simplified version of the zero-one-inflated beta (ZOIB) model, known as ordered
516 beta regression (Kubinec, 2022; see also Makowski et al., 2025 for a reparameterized version called the
517 *beta-Gate* model), has been recently proposed. The ordered beta regression model exploits the fact that,
518 for most analyses, the continuous values (between 0-1) and the discrete outcomes (e.g., 0 or 1) are *ordered*.
519 For example, as a covariate x increases or decreases, we should expect the bounded outcome y to increase
520 or decrease monotonically as well from 0 to (0, 1) to 1. The ZOIB model does not impose this restriction;
521 a covariate could increase and the response y could increase in its continuous values while *simultaneously*
522 decreasing at *both* end points.⁷ This complexity is not immediately obvious when fitting the ZOIB, nor is
523 it a potential relationship that many scholars want to consider when examining how covariates influence a
524 bounded scale.

⁷For a more complete description of this issue, we refer the reader to Kubinec (2022).

525 To make the response ordered, the ordered beta regression model estimates a weighted combination
 526 of a standard beta regression model for continuous responses and a logit model for the discrete values of
 527 the response. By doing so, the amount of distinctiveness between the continuous responses and the discrete
 528 end points is a function of the data (and any informative priors) rather than strictly defined as fully distinct
 529 processes as in the ZOIB. For some datasets, the continuous and discrete responses will be fairly distinct,
 530 and in others less so.

531 The weights that average together the two parts of the outcome (i.e., discrete and continuous) are
 532 determined by cutpoints that are estimated in conjunction with the data in a similar manner to what is known
 533 as an ordered logit model. An in-depth explanation of ordinal regression is beyond the scope of this tutorial
 534 (Bürkner & Vuorre, 2019; but see Fullerton & Anderson, 2021). At a basic level, ordinal regression models
 535 are useful for outcome variables that are categorical in nature and have some inherent ordering (e.g., Likert
 536 scale items). To preserve this ordering, ordinal models rely on the cumulative probability distribution.
 537 Within an ordinal regression model it is assumed that there is a continuous but unobserved latent variable
 538 that determines which of k ordinal responses will be selected. For example on a typical Likert scale from
 539 ‘Strongly Disagree’ to ‘Strongly Agree’, you could assume that there is a continuous, unobserved variable
 540 called ‘Agreement’.

541 While we cannot measure Agreement directly, the ordinal response gives us some indication about
 542 where participants are on the continuous Agreement scale. $k - 1$ cutoffs are then estimated to indicate the
 543 point on the continuous Agreement scale at which your Agreement level is high enough to push you into the
 544 next ordinal category (say Agree to Strongly Agree). Coefficients in the model estimate how much differ-
 545 ent predictors change the estimated *continuous* scale (here, Agreement). Since there’s only one underlying
 546 process, there’s only one set of coefficients to work with (proportional odds assumption). In an ordered beta
 547 regression, three ordered categories are modeled: (1) exactly zero, (2) somewhere between zero and one,
 548 and (3) exactly one. In an ordered beta regression, (1) and (2) are modeled with cumulative logits, where
 549 one cutpoint is the the boundary between Exactly 0 and Between 0 and 1 and the other cutpoint is the bound-
 550 ary between *Between 0 and 1* and *Exactly 1*. The continuous values in the middle, 0 to 1 (3), are modeled
 551 as a vanilla beta regression with parameters reflecting the mean response on the logit scale as we have de-
 552 scribed previously. Ultimately, employing cutpoints allows for a smooth transition between the bounds and
 553 the continuous values, permitting both to be considered together rather than modeled separately as the ZOIB
 554 requires.

555 The ordered beta regression model has shown to be more efficient and less biased than some of the
 556 methods discussed (Kubinec, 2022) herein and has seen increasing use across the biomedical and social
 557 sciences (Martin et al., 2024; Nouvian et al., 2023; Shrestha et al., 2024; Smith et al., 2024; Wilkes et al.,
 558 2024) because it produces only a single set of coefficient estimates in a similar manner to a standard beta
 559 regression or OLS.⁸

560 ***Fitting an Ordered Beta Regression***

561 To fit an ordered beta regression in a Bayesian context we use the `{ordbetareg}` (Kubinec, 2023) pack-
 562 age. `{ordbetareg}` is a front-end to the `brms` package that we described earlier; in addition to the functions
 563 available in the package, most `brms` functions and plots, including the diverse array of regression model-
 564 ing options, will work with `{ordbetareg}` models. (We note that the `ordbeta` model is also available as a
 565 maximum-likelihood variant in the R package `{glmmTMB}`.) We first load the `{ordbetareg}` package (see
 566 Listing 16).

567 The `{ordbetareg}` package uses `brms` on the front-end so all the arguments we used previously apply
 568 here. Instead of the `brm()` function we use `ordbetareg()`. To fit a model where dispersion does not vary

⁸Please note that there are other models available that can model this continuous process like the beta-gate model (Makowski et al., 2025) and the censored extended beta regression model (Kosmidis & Zeileis, 2025).

Listing 16 Load {ordbetareg}

```
library(ordbetareg)
```

Table 6*Marginal effect of fluency ordered beta model*

Contrast	Mean	95% Cr.I	pd
Fluent - Disfluent	0.061	[-0.017, 0.137]	0.936

569 as a function of fluency we can use the below code (see Listing 17).

Listing 17 Fitting ordered beta model with ordbetareg()

```
ord_fit_brms <- ordbetareg(
  Accuracy ~ Fluency,
  data = fluency_data,
  file = here::here("manuscript", "models", "bayes_ordbeta_model")
)
```

570 However, if we want dispersion to vary as a function of fluency we can easily do that (see Listing 18).
 571 Note the addition of the `phi_reg` argument in `m.phi`. This argument allows us to include a model that
 572 explicitly models the dispersion parameter. Because we are modeling ϕ as a function of fluency, we set the
 573 the argument to `both`.

Listing 18 Fitting ordered beta model with dispersion using ordbetareg()

```
ord_beta_phi <- bf(Accuracy ~ Fluency, phi ~ Fluency)

m.phi <- ordbetareg(
  ord_beta_phi,
  data = fluency_data,
  phi_reg = 'both',
  file = here::here("manuscript", "models", "bayes_ordbeta_phi_model")
)
```

574 **Marginal Effects.** Table 8 presents the posterior summary. We can use `{marginaleffects}` to calculate differences on the response scale that average over (or marginalize over) all our parameters.

575 In Table 6 the credible interval is close enough to zero relative to its uncertainty that we can conclude
 576 there likely aren't differences between the conditions after taking dispersion and the 0s and 1s in our data
 577 into account.

578 **Cutpoints.** The model cutpoints are not reported by default in the summary output, but we can access them with the R package `posterior` (Bürkner et al., 2025) and the functions `as_draws` and `summary_draws`.

Table 7

Cutzero and cutone parameter summary

Parameter	Mean	95% Cr.I
cutzero	-2.98	[-3.59, -2.42]
cutone	1.85	[1.64, 2.08]

In Table 7, `cutzero` is the first cutpoint (the difference between 0 and continuous values) and `cutone` is the second cutpoint (the difference between the continuous values and 1). These cutpoints are on the logit scale and as such the numbers do not have a simple substantive meaning. In general, as the cutpoints increase in absolute value (away from zero), then the discrete/boundary observations are more distinct from the continuous values. This will happen if there is a clear gap or bunching in the outcome around the bounds. This type of empirical feature of the distribution may be useful to scholars if they want to study differences in how people perceive the ends of the scale versus the middle. It is possible, though beyond the scope of this article, to model the location of the cutpoints with hierarchical (non-linear) covariates in `brms`. In the most recent version of `ordbeta`, you can test the influence of different factors on these boundaries.

591 **Model Fit**

The best way to visualize model fit is to plot the full predictive distribution relative to the original outcome. Because ordered beta regression is a mixed discrete/continuous model, a separate plotting function, `pp_check_ordbetareg`, is included in the `{ordbetareg}` package that accurately handles the unique features of this distribution. The default plot in `brms` will collapse these two features of the outcome together, which will make the fit look worse than it actually is. The `{ordbetareg}` function returns a list with two plots, `discrete` and `continuous`, which can either be printed and plotted or further modified as `{ggplot2}` objects (see Figure 10).

The discrete plot, which is a bar graph, shows that the posterior distribution accurately captures the number of different types of responses (discrete or continuous) in the data. For the continuous plot shown as a density plot with one line per posterior draw, the model does a very good job at capturing the distribution.

Overall, it is clear from the posterior distribution plot that the ordered beta model fits the data well. To fully understand model fit, both of these plots need to be inspected as they are conceptually distinct.

604 **Model Visualization**

`{ordbetareg}` provides a neat visualization function called `plot_heiss()` (Ye & Heiss, 2023) that can represent dispersion in the entire outcome as a function of discrete covariates. This function produces a plot of predicted proportions across the range of our Fluency factor. In Figure 11 we get predicted proportions for Fluency across the bounded scale. Looking at the figure we can see there is much overlap between instructors in the middle portion (μ). However, we do see some small differences at the zero bounds.

610 **Ordered Beta Scale**

In the `{ordbetareg}` function there is a `true_bound` argument. In cases where your data is not bounded between 0-1, you can use this argument to specify the bounds of the argument to fit the ordered beta regression. For example, you data might be bounded between 1 and 7. If so, you can model it as such and `{ordbetareg}` will convert the model predictions back to the true bounds after estimation.

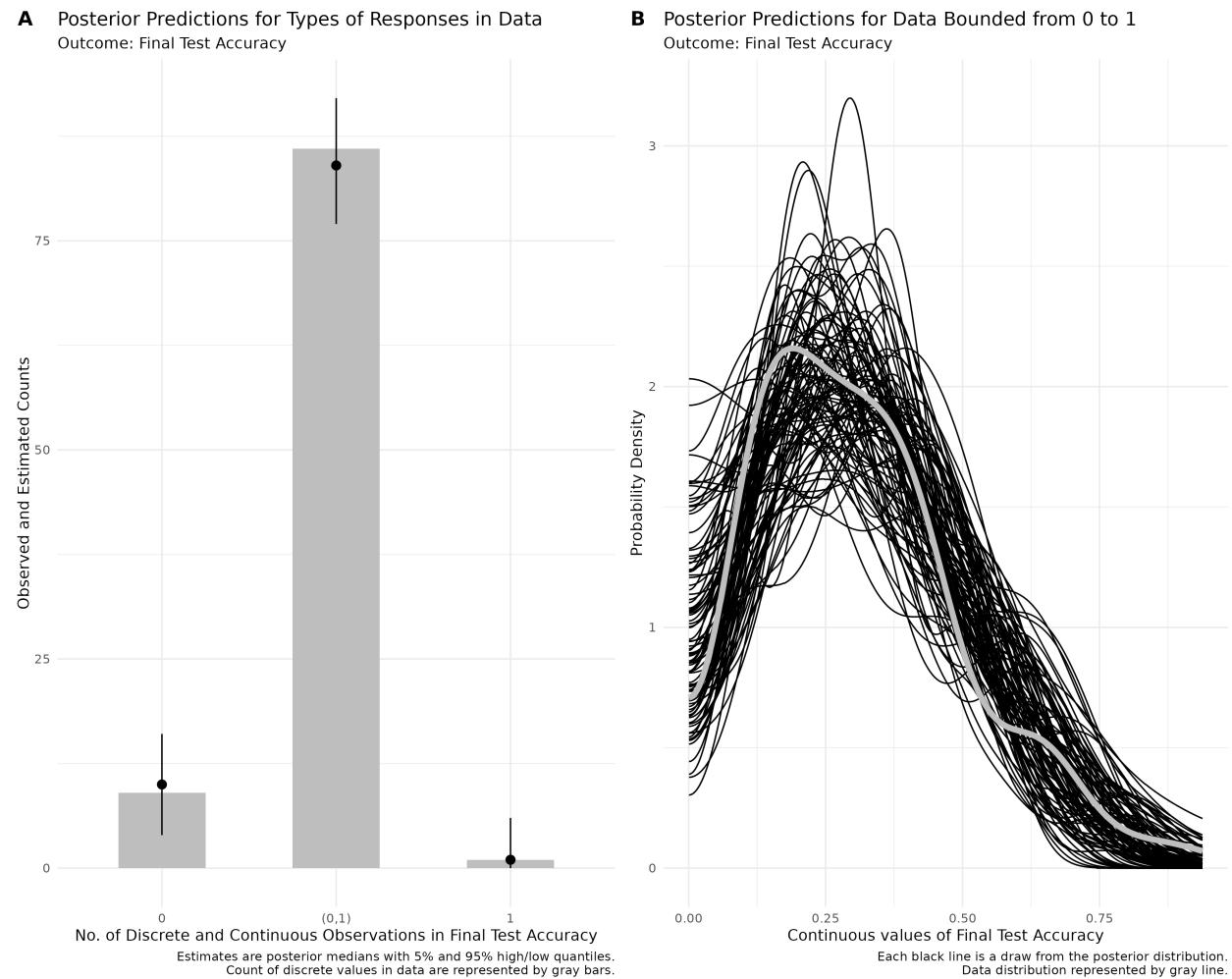
Table 8*Bayesian regression summaries for each model*

Parameter	Stat	Bayesian LM	Beta Regression	ZIB	ZOIB	Ordered Beta
b_Intercept	Mean	0.257	-0.830	-0.832	-0.831	-0.865
	Cr.I	[0.199, 0.315]	[-1.087, -0.55]	[-1.094, -0.552]	[-1.098, -0.559]	[-1.119, -0.596]
	pd	1.000***	1.000***	1.000***	1.000***	1.000***
b_Fluency	Mean	0.085	0.204	0.204	0.203	0.262
	Cr.I	[0.002, 0.166]	[-0.155, 0.539]	[-0.139, 0.545]	[-0.147, 0.541]	[-0.07, 0.598]
	pd	0.977*	0.875	0.872	0.880	0.936
sigma	Mean	0.209	-	-	-	-
	Cr.I	[0.179, 0.245]	-	-	-	-
	pd	1.000***	-	-	-	-
b_phi_Intercept	Mean	-	1.609	1.601	1.604	1.609
	Cr.I	-	[1.193, 2]	[1.187, 1.988]	[1.183, 1.989]	[1.179, 1.993]
	pd	-	1.000***	1.000***	1.000***	1.000***
b_phi_Fluency	Mean	-	0.420	0.425	0.426	0.408
	Cr.I	-	[-0.143, 0.993]	[-0.158, 0.994]	[-0.126, 0.994]	[-0.156, 0.983]
	pd	-	0.931	0.926	0.930	0.918
b_zi_Intercept	Mean	-	-	-1.673	-	-
	Cr.I	-	-	[-2.46, -0.978]	-	-
	pd	-	-	1.000**	-	-
b_zi_Fluency	Mean	-	-	-2.137	-	-
	Cr.I	-	-	[-4.618, -0.34]	-	-
	pd	-	-	0.992**	-	-
b_zoi_Intercept	Mean	-	-	-	-1.549	-
	Cr.I	-	-	-	[-2.339, -0.859]	-
	pd	-	-	-	1.000***	-
b_zoi_Fluency	Mean	-	-	-	-2.201	-
	Cr.I	-	-	-	[-4.449, -0.465]	-
	pd	-	-	-	0.996***	-
b_coi_Intercept	Mean	-	-	-	-2.022	-
	Cr.I	-	-	-	[-4.408, -0.287]	-
	pd	-	-	-	0.991**	-
b_coi_Fluency	Mean	-	-	-	0.245	-
	Cr.I	-	-	-	[-6.946, 5.716]	-
	pd	-	-	-	0.571	-

Note. Link functions: b_mean = logit; b_phi = log; b_zoi (zero-one inflation) = logit; b_coi (conditional one-inflation) = logit. Asterisks reflect approximate two-sided p-values derived from the posterior pd. pd ≥ 0.975 ($p \leq .05$) = *; pd ≥ 0.990 ($p \leq .01$) = **; pd ≥ 0.998 ($p \leq .001$) = ***.

Figure 10

Posterior predictive check for ordered beta regression model. A. Discrete posterior check. B. Continuous posterior check.



615

Discussion

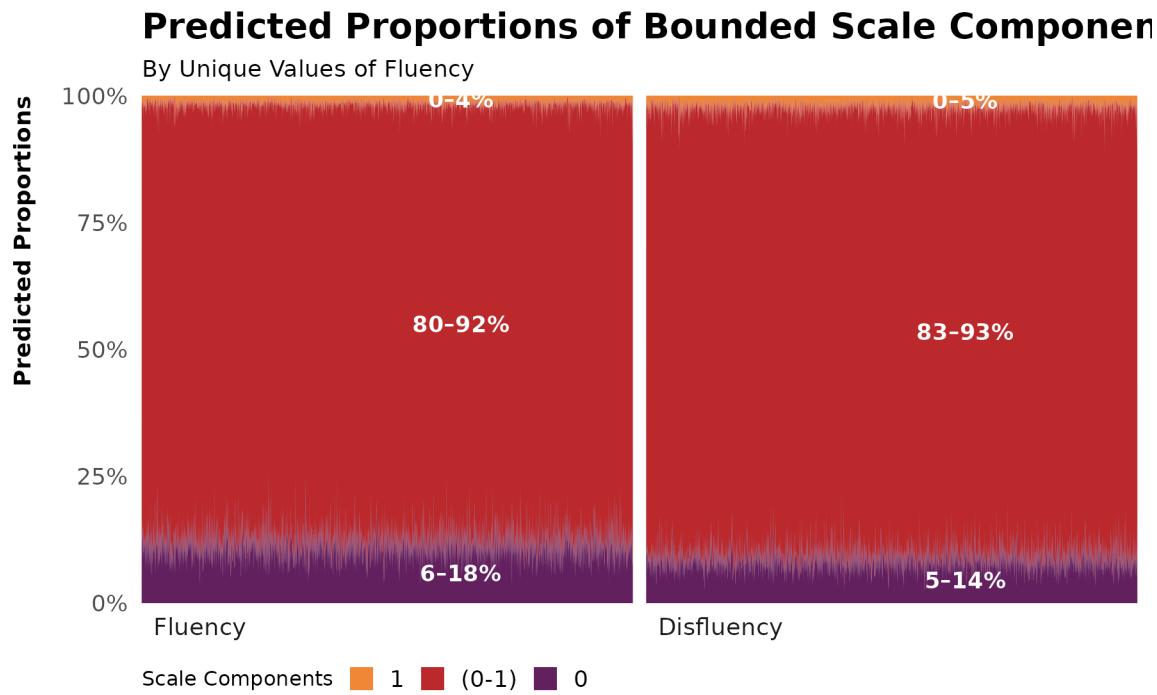
616 The use of beta regression in psychology, and the social sciences in general, is rare. With this tutorial,
 617 we hope to turn the tides. beta regression models are an attractive alternative to models that impose unrealistic
 618 assumptions like normality, linearity, homoscedasticity, and unbounded data. Beyond these models, there
 619 are a diverse array of different models that can be used depending on your outcome of interest.

620 Throughout this tutorial our main aim was to help guide researchers in running analyses with pro-
 621 portional or percentage outcomes using beta regression and some of its alternatives. In the current example,
 622 we used real data from Wilford et al. (2020) and discussed how to fit these models in R, interpret model
 623 parameters, extract predicted probabilities and marginal effects, and visualize the results.

624 Comparing our analysis with that of Wilford et al. (2020), we demonstrated that using a traditional
 625 approache (e.g., t -test) to analyze mean accuracy data can lead to biased inferences. Although we successfully
 626 reproduced one of their key findings, our use of beta regression and its extensions revealed important nuances
 627 in the results. With a traditional beta regression model—which accounts for both the mean and the precision
 628 (dispersion)—we observed similar effects of instructor fluency on performance. However, the standard beta

Figure 11

Heiss plot of predicted probabilities across the scale (0-100)



Plot shows predicted proportions of the components of a bounded scale, i.e. the predicted (expected) probability of the top value of the scale, the intermediate continuous values, and the bottom value of the scale. The predictions are subset for unique values of a grouping factor. The predictions are shown for multiple posterior draws to indicate uncertainty. Labels on components indicate posterior quantiles for the probability of that component for each level of the grouping variable.

629 model does not accommodate boundary values (i.e., 0s and 1s).

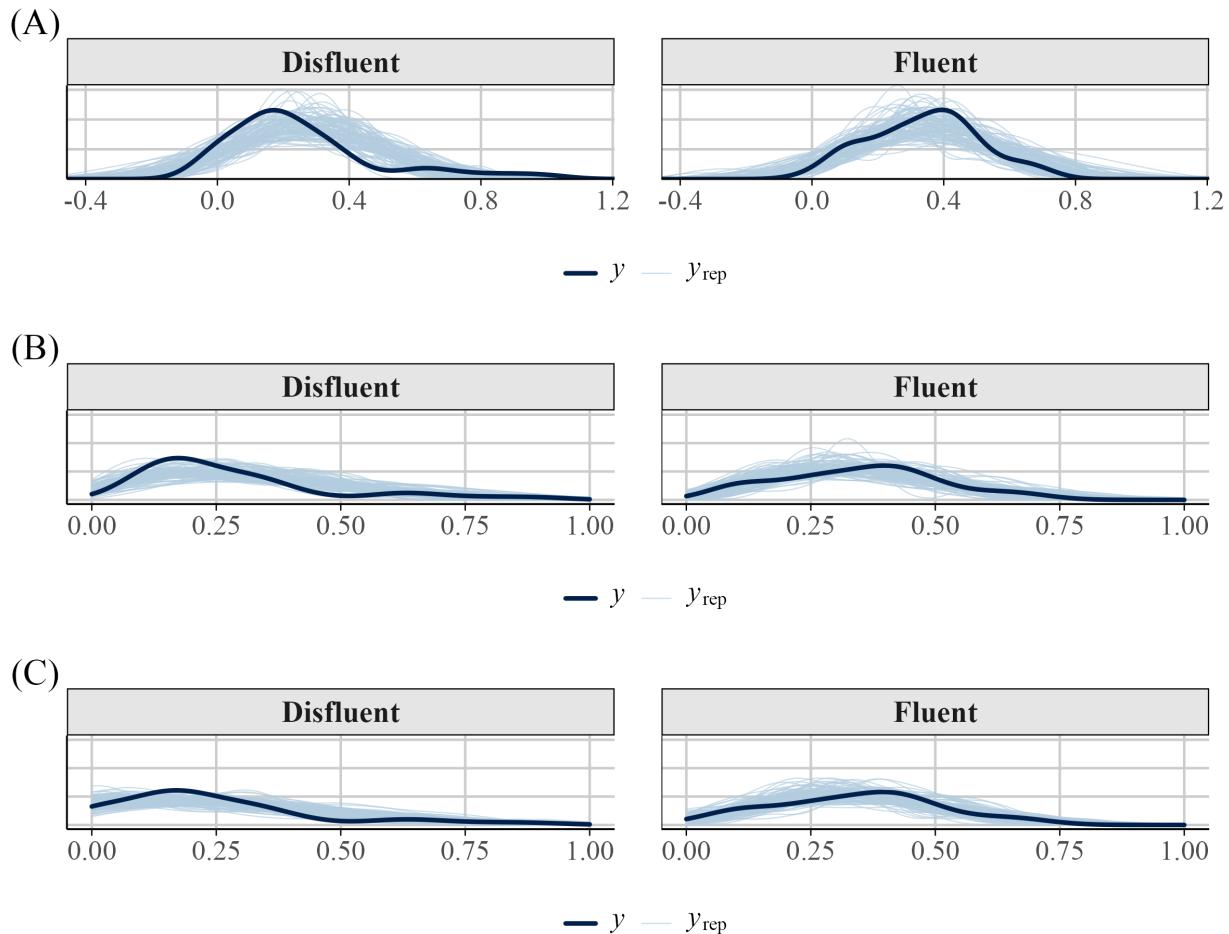
630 When we applied a ZIB model, which explicitly accounts for structural 0s, we found no effect of
 631 fluency on the mean (μ) part of the model. Instead, the effect of fluency emerged in the structural zero
 632 (inflated zero; α) component. This pattern was consistent when using a zero-one-inflated beta (ZOIB) model.
 633 Furthermore, we fit an ordered beta regression model (Kubinec, 2022), which appropriately models the full
 634 range of values, including 0s and 1s. Here, we did not observe a reliable effect of fluency on the mean once
 635 we accounted for dispersion.

636 These analyses emphasize the importance of fitting a model that aligns with the nature of the data.
 637 The simplest and recommended approach when dealing with data that contains 0s and/or 1s is to fit an ordered
 638 beta model, assuming the process is truly continuous. However, if you believe the process is distinct in nature,
 639 a ZIB or ZOIB model might be a better choice. Ultimately, this decision should be guided by theory.

640 For instance, if we believe fluency influences the boundaries (0 and 1), we might want to model
 641 this process separately using a ZIB or ZOIB. With the current dataset, fluency might affect specific aspects
 642 of performance (such as the likelihood of complete failure) rather than general performance levels. This
 643 effect could be due to participant disengagement during the disfluent lecture. If students fail to pay attention
 644 because of features of disfluency, they may miss relevant information, leading to a floor effect at the test.
 645 Following from this, disfluency would be expected to influence the boundary (0) and not the continuous part
 646 of the model. If this is the case, we would want to model this appropriately. However, if we believe fluency

Figure 12

The plots show 100 posterior predicted distributions with the label y_{rep} (light blue), the distribution of accuracy as function of fluency in dark blue for regular regression (A), beta regression (B), and ZIB (C) models



647 effects general performance levels (the continuous part), a model that takes in to account the entire process
 648 accounting for the 0s and 1s might be appropriate.

649 In the discussion section of Wilford et al. (2020), they were unable to offer a tenable explanation for
 650 performance differences based on instructor fluency. A model that accounts for the excess 0s in the dataset
 651 provides one testable explanation: watching a disfluent lecture may lead to lapses in attention, resulting
 652 in poorer performance in that group. These lapses, in turn, contribute to the observed differences in the
 653 fluent condition. This modeling approach opens a promising avenue for future research—one that would have
 654 remained inaccessible otherwise.

655 Not everyone will be eager to implement the techniques discussed herein. In such cases, the key
 656 question becomes: What is the least problematic approach to handling proportional data? One reasonable
 657 option is to fit multiple models tailored to the specific characteristics of your data. For example, if your data
 658 contain 0s, you might fit two models: a traditional linear model excluding the 0s, and a logistic model to
 659 account for the zero versus non-zero distinction. If your data contain both 0s and 1s, you could fit separate
 660 models for the 0s and 1s in addition to the OLS model. There are many defensible strategies to choose from
 661 depending on the context. However, we do not recommend transforming the values of your data (e.g., 0s to

662 .01 and 1s to .99) or ignoring the properties of your data simply to fit traditional statistical models.

663 In this tutorial, we demonstrated how to analyze these models from a Bayesian perspective. While we
664 recognize that not everyone identifies as a Bayesian, implementing these models using a Bayesian framework
665 is relatively straightforward—it requires only a single package, lowering the barrier to entry. For those who
666 prefer frequentist analyses, several R packages are available. For example, the `{betareg}` package (Cribari-
667 Neto & Zeileis, 2010) `{glmmTMB}` (Brooks et al., 2017) and `{gamlss}` (2005) are nice options. To this end,
668 I have included supplemental materials demonstrating how to use frequentist packages to analyze the data
669 presented herein.

670 Conclusion

671 Overall, this tutorial emphasizes the importance of modeling the data you have. Although the ex-
672 ample provided is relatively simple (a one-factor model with two levels), we hope it demonstrates that even
673 with a basic dataset, there is much nuance in interpretation and inference. Properly modeling your data
674 can lead to deeper insights, far beyond what traditional measures might offer. With the tools introduced in
675 this tutorial, researchers now have the means to analyze their data effectively, uncover patterns, make ac-
676 curate predictions, and support their findings with robust statistical evidence. By applying these modeling
677 techniques, researchers can improve the validity and reliability of their studies, ultimately leading to more
678 informed decisions and advancements in their respective fields.

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