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SDS 383 D - Statistical Modeling 2
Homework

Exercise 1.1

$X_1, \dots X_n \sim Bernoulli(p), \quad h(p) \sim Beta(\alpha, \beta)$

$$\begin{aligned} p|X_1, \dots X_N \propto & f(x_1, \dots x_n|p) \cdot h(p) \\ \propto & p^{\sum x_i}(1-p)^{n-\sum x_i} \cdot p^{\alpha-1}(1-p)^{\beta-1} \\ \propto & p^{\sum x_i+\alpha-1}(1-p)^{\beta+n-\sum x_i-1} \end{aligned}$$

$$p|X_1, \dots X_N \sim Beta(\sum x_i + \alpha, \beta + n - \sum x_i)$$

Exercise 1.2

$X_1, \dots X_N \overset{iid}{\sim} Cat(p), \quad p = (p_1, \dots p_K), \quad f(x|p) = \prod_{i=1}^n \prod_{k=1}^K p_k^{I(x_i=k)}$
$p \sim Dirichlet(\alpha_1, \dots \alpha_K)$

$$\pi(p) = \frac{\Gamma(\sum_k \alpha_k)}{\prod_K \Gamma(\alpha_k)} \prod_{k=1}^K p^{\alpha_k - 1} I(\sum_{k=1}^{K-1} p_k \leq 1, p_k \geq 0)$$

$$\begin{aligned} p|X_{1:n} \propto & f(x_{1:n}|p) \cdot \pi(p) \\ \propto & \prod_{i=1}^n \prod_{k=1}^K p_k^{I(x_i=k)} \cdot \prod_{k=1}^K p_k^{\alpha_k-1} \\ \propto & \prod_{k=1}^K p_K^{n_k+\alpha_k-1}, \quad \text{where } n_k = \sum_{i=1}^N I(x_i = k) \\ p|X_{1:n} \sim & Dir(n_1 + \alpha_1, \dots n_k + \alpha_k) \end{aligned}$$

Exercise 1.3

$X \sim Gam(\alpha, b)$ and $Y \sim Gam(\beta, b)$ and X is independent of Y.

1.3.a

Let $W = X+Y$ and $Z = X/(X+Y)$. Show that W and Z are independent with $W \sim Gam(\alpha+\beta, b)$ and $Z \sim Beta(\alpha, \beta)$.

This is a bivariate random variable transformation:

Change of variables equation: $f_{w,z} = f_{x,y}(h_1(w,z), h_2(w,z))|J|$

$g_1(X,Y) = W = X+Y$

$g_2(X,Y) = Z = X/(X+Y)$

$h_1(Z,W) = X = ZW$

$h_2(Z,W) = Y = W(1-Z)$

$|J| = abs(\begin{vmatrix} \frac{\partial X}{\partial W} & \frac{\partial X}{\partial Z} \\ \frac{\partial Y}{\partial W} & \frac{\partial Y}{\partial Z} \end{vmatrix}) = \begin{vmatrix} Z & W \\ 1-Z & -W \end{vmatrix} = abs(-w) = w$

$f_{w,z} = f_x(zw)f_y(w-zw)w$

$f_{w,z} = \dfrac{b^\alpha (zw)^{\alpha-1}e^{-bzw} \cdot b^\beta (w(1-z))^{\beta-1}e^{-bw(1-z)}w}{\Gamma(\alpha)\Gamma(\beta)} \cdot I(zw \geq 0) \cdot I(w - zw \geq 0)$

$f_{w,z} = \dfrac{w^{\alpha+\beta-1}b^{\alpha+\beta}e^{-bw} \cdot z^{\alpha-1}(1-z)^{\beta-1}}{\Gamma(\alpha)\Gamma(\beta)} \cdot I(0 \leq z \leq 1) \cdot I(w \geq 0)$

$f_{w,z} = \dfrac{b^{\alpha+\beta}w^{\alpha+\beta-1}e^{-bw}I(w \geq 0)}{\Gamma(\alpha+\beta)} \cdot \dfrac{\Gamma(\alpha+\beta)z^{\alpha-1}(1-z)^{\beta-1}I(0 \leq z \leq 1)}{\Gamma(\alpha)\Gamma(\beta)}$

So W and Z are independent by factorization theorem and $W \sim Gam(\alpha+\beta, b)$ and $Z \sim Beta(\alpha, \beta)$

1.3.b

Explain how to use a random gamma sampler to sample from a $Beta(\alpha, \beta)$ random variable using R:

We can use the R function (rgamma(n, shape, rate)) to sample two gamma random variables with shape parameters $\alpha = 10, \beta = 2$ and rate parameters equal $b = 5, b = 5$ and then use the transformation from the previous step to get a sample from a $Beta(\alpha = 10, \beta = 2)$ distribution.

```r
x <- rgamma(1, 10, 5)
y <- rgamma(1, 2, 5)
z = x / (x + y)
```

Exercise 1.4

Suppose $X_1 \dots X_n \overset{iid}{\sim} \text{Normal}(\theta, \sigma_0^2)$ where σ_0^2 is known. Suppose $\theta \sim \text{Normal}(m, v)$, find posterior of $\theta | X_{1:N}$

$$\begin{aligned}
\theta|X_{1:N} &\propto \exp\{\frac{-1}{2\sigma_0^2}\sum_{i=1}^{N}(x_i-\theta)^2\} \cdot \ \exp\{\frac{-1}{2v^2}(\theta-m)^2\} \\
&\propto \exp\{\frac{-1}{2\sigma_0^2}\big(\sum_{i=1}^{N}(x_i^2+\theta^2-2\theta_i)\big) + \frac{-1}{2v^2}\big(\theta^2+m^2-2\theta m\big)\} \\
&\propto \exp\{\frac{-\theta^2}{2}\cdot(\frac{1}{v^2}+\frac{n}{\sigma_0^2}) + \theta(\frac{m}{v^2}+\frac{\sum x_i}{\sigma_0^2}) - (\frac{m^2}{2v^2}+\frac{\sum x_i^2}{2\sigma_0^2})\} \\
&\propto \exp\{\frac{-\theta^2}{2}\cdot(\frac{1}{v^2}+\frac{n}{\sigma_0^2}) + \theta(\frac{m}{v^2}+\frac{\sum x_i}{\sigma_0^2})\} \\
&\propto \exp\{\frac{-1}{2}\big(\theta^2(\frac{1}{v^2}+\frac{n}{\sigma_0^2}) - 2\theta(\frac{m}{v^2}+\frac{\sum x_i}{\sigma_0^2})\big)
\end{aligned}$$

Completing the square:

$$\begin{aligned}
a &= \frac{1}{v^2}+\frac{n}{\sigma_0^2} \\
b &= \frac{m}{v^2}+\frac{\sum x_i}{\sigma_0^2} \\
a\theta^2 - 2\theta b &= a(\theta^2 - \frac{2\theta b}{a}) = a((\theta-\frac{b}{a})^2 - \frac{b^2}{a^2})
\end{aligned}$$

$$\begin{aligned}
\theta|X_{1:N} &\propto \exp\{\frac{-1}{2}a(\theta-\frac{b}{a})^2\} \\
&\propto \exp\{\frac{-1}{2}\frac{\left(\theta-\frac{b}{a}\right)^2}{\frac{1}{\sqrt{a}}}\} \\
\theta|X_{1:N} &\sim \text{Normal}(m=\frac{b}{a}, v=a^{-1}) \\
m &= \frac{b}{a} = \frac{\frac{m}{v^2}+\frac{\sum x_i}{\sigma_0^2}}{\frac{1}{v^2}+\frac{n}{\sigma_0^2}}, \quad v = (\frac{1}{v^2}+\frac{n}{\sigma_0^2})^{-1}
\end{aligned}$$