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SDS 383 D - Statistical Modeling 2
Homework

Exercise 1.1

$$X_1, \dots, X_n \sim \text{Bernoulli}(p), \quad h(p) \sim \text{Beta}(\alpha, \beta)$$

$$\begin{aligned} p|X_1, \dots, X_N &\propto f(x_1, \dots, x_n|p) \cdot h(p) \\ &\propto p^{\sum x_i} (1-p)^{n-\sum x_i} \cdot p^{\alpha-1} (1-p)^{\beta-1} \\ &\propto p^{\sum x_i + \alpha - 1} (1-p)^{\beta + n - \sum x_i - 1} \end{aligned}$$

$$p|X_1, \dots, X_N \sim \text{Beta}(\sum x_i + \alpha, \beta + n - \sum x_i)$$

Exercise 1.2

$$\begin{aligned} X_1, \dots, X_N &\stackrel{iid}{\sim} \text{Cat}(p), \quad p = (p_1, \dots, p_K), \quad f(x|p) = \prod_{i=1}^n \prod_{k=1}^K p_k^{I(x_i=k)} \\ p &\sim \text{Dirichlet}(\alpha_1, \dots, \alpha_K) \end{aligned}$$

$$\pi(p) = \frac{\Gamma(\sum_k \alpha_k)}{\prod_K \Gamma(\alpha_k)} \prod_{k=1}^K p^{\alpha_k - 1} I(\sum_{k=1}^{K-1} p_k \leq 1, p_k \geq 0)$$

$$\begin{aligned} p|X_{1:n} &\propto f(x_{1:n}|p) \cdot \pi(p) \\ &\propto \prod_{i=1}^n \prod_{k=1}^K p_k^{I(x_i=k)} \cdot \prod_{k=1}^K p_k^{\alpha_k - 1} \\ &\propto \prod_{k=1}^K p_k^{n_k + \alpha_k - 1}, \quad \text{where } n_k = \sum_{i=1}^N I(x_i = k) \\ p|X_{1:n} &\sim \text{Dir}(n_1 + \alpha_1, \dots, n_K + \alpha_K) \end{aligned}$$

Exercise 1.3

$X \sim \text{Gam}(\alpha, b)$ and $Y \sim \text{Gam}(\beta, b)$ and X is independent of Y .

1.3.a

Let $W = X+Y$ and $Z = X/(X+Y)$. Show that W and Z are independent with $W \sim \text{Gam}(\alpha + \beta, b)$ and $Z \sim \text{Beta}(\alpha, \beta)$.

This is a bivariate random variable transformation:

Change of variables equation: $f_{w,z} = f_{x,y}(h_1(w, z), h_2(w, z))|J|$

$$g_1(X, Y) = W = X + Y$$

$$g_2(X, Y) = Z = X/(X + Y)$$

$$h_1(Z, W) = X = ZW$$

$$h_2(Z, W) = Y = W(1 - Z)$$

$$|J| = \text{abs}\left(\begin{vmatrix} \frac{\partial X}{\partial W} & \frac{\partial X}{\partial Z} \\ \frac{\partial Y}{\partial W} & \frac{\partial Y}{\partial Z} \end{vmatrix}\right) = \begin{vmatrix} Z & W \\ 1 - Z & -W \end{vmatrix} = \text{abs}(-w) = w$$

$$f_{w,z} = f_x(zw)f_y(w - zw)w$$

$$f_{w,z} = \frac{b^\alpha (zw)^{\alpha-1} e^{-bzw} \cdot b^\beta (w(1-z))^{\beta-1} e^{-bw(1-z)} w}{\Gamma(\alpha)\Gamma(\beta)} \cdot I(zw \geq 0) \cdot I(w - zw \geq 0)$$

$$f_{w,z} = \frac{w^{\alpha+\beta-1} b^{\alpha+\beta} e^{-bw} \cdot z^{\alpha-1} (1-z)^{\beta-1}}{\Gamma(\alpha)\Gamma(\beta)} \cdot I(0 \leq z \leq 1) \cdot I(w \geq 0)$$

$$f_{w,z} = \frac{b^{\alpha+\beta} w^{\alpha+\beta-1} e^{-bw} I(w \geq 0)}{\Gamma(\alpha + \beta)} \cdot \frac{\Gamma(\alpha + \beta) z^{\alpha-1} (1-z)^{\beta-1} I(0 \leq z \leq 1)}{\Gamma(\alpha)\Gamma(\beta)}$$

So W and Z are independent by factorization theorem and $W \sim \text{Gam}(\alpha + \beta, b)$ and $Z \sim \text{Beta}(\alpha, \beta)$

1.3.b

Explain how to use a random gamma sampler to sample from a $\text{Beta}(\alpha, \beta)$ random variable using R:

We can use the R function `(rgamma(n, shape, rate))` to sample two gamma random variables with shape parameters $\alpha = 10, \beta = 2$ and rate parameters equal $b = 5, b = 5$ and then use the transformation from the previous step to get a sample from a $\text{Beta}(\alpha = 10, \beta = 2)$ distribution.

```
x <- rgamma(1, 10, 5)
y <- rgamma(1, 2, 5)
z = x / (x + y)
```

Exercise 1.4

Suppose $X_1 \dots X_n \stackrel{iid}{\sim} \text{Normal}(\theta, \sigma_0^2)$ where σ_0^2 is known. Suppose $\theta \sim \text{Normal}(m, v)$, find posterior of $\theta|X_{1:N}$

$$\begin{aligned}
\theta|X_{1:N} &\propto \exp\left\{\frac{-1}{2\sigma_0^2} \sum_{i=1}^N (x_i - \theta)^2\right\} \cdot \exp\left\{\frac{-1}{2v^2} (\theta - m)^2\right\} \\
&\propto \exp\left\{\frac{-1}{2\sigma_0^2} \left(\sum_{i=1}^N (x_i^2 + \theta^2 - 2\theta x_i)\right) + \frac{-1}{2v^2} (\theta^2 + m^2 - 2\theta m)\right\} \\
&\propto \exp\left\{\frac{-\theta^2}{2} \cdot \left(\frac{1}{v^2} + \frac{n}{\sigma_0^2}\right) + \theta \left(\frac{m}{v^2} + \frac{\sum x_i}{\sigma_0^2}\right) - \left(\frac{m^2}{2v^2} + \frac{\sum x_i^2}{2\sigma_0^2}\right)\right\} \\
&\propto \exp\left\{\frac{-\theta^2}{2} \cdot \left(\frac{1}{v^2} + \frac{n}{\sigma_0^2}\right) + \theta \left(\frac{m}{v^2} + \frac{\sum x_i}{\sigma_0^2}\right)\right\} \\
&\propto \exp\left\{\frac{-1}{2} \left(\theta^2 \left(\frac{1}{v^2} + \frac{n}{\sigma_0^2}\right) - 2\theta \left(\frac{m}{v^2} + \frac{\sum x_i}{\sigma_0^2}\right)\right)\right\}
\end{aligned}$$

Completing the square:

$$\begin{aligned}
a &= \frac{1}{v^2} + \frac{n}{\sigma_0^2} \\
b &= \frac{m}{v^2} + \frac{\sum x_i}{\sigma_0^2} \\
a\theta^2 - 2\theta b &= a\left(\theta^2 - \frac{2\theta b}{a}\right) = a\left(\left(\theta - \frac{b}{a}\right)^2 - \frac{b^2}{a^2}\right)
\end{aligned}$$

$$\begin{aligned}
\theta|X_{1:N} &\propto \exp\left\{\frac{-1}{2} a \left(\theta - \frac{b}{a}\right)^2\right\} \\
&\propto \exp\left\{\frac{-1}{2} \frac{\left(\theta - \frac{b}{a}\right)^2}{\frac{1}{a}}\right\} \\
\theta|X_{1:N} &\sim \text{Normal}\left(m = \frac{b}{a}, v = a^{-1}\right) \\
m &= \frac{b}{a} = \frac{\frac{m}{v^2} + \frac{\sum x_i}{\sigma_0^2}}{\frac{1}{v^2} + \frac{n}{\sigma_0^2}}, \quad v = \left(\frac{1}{v^2} + \frac{n}{\sigma_0^2}\right)^{-1}
\end{aligned}$$

Exercise 1.5

Suppose $X_1, \dots, X_N \stackrel{iid}{\sim} \text{Normal}(\theta, \sigma^2)$ with θ known but σ^2 unknown. Suppose that $\omega = \sigma^{-2}$ has a $\text{Gam}(\alpha, \beta)$ prior. Derive posterior of $\omega|X_1, \dots, X_n$

$$\begin{aligned}
h(\omega|X_{1:n}) &\propto f(x_{1:n}|\omega)h(\omega) \\
&\propto \omega^{n/2} \exp\{-w/2 \sum (x_i - \theta)^2\} \cdot \omega^{\alpha-1} \exp\{-\beta\omega\} \\
&\propto \exp\{-\omega(\beta + \frac{\sum (x_i - \theta)^2}{2})\} \omega^{\alpha+n/2-1} \\
\omega|X_{1:n} &\sim \text{Gamma}(\alpha + \frac{n}{2}, \beta + \frac{\sum (x_i - \theta)^2}{2})
\end{aligned}$$

Exercise 1.6

Show that for squared error loss $R(F, a) = \text{Var}_F(a) + \text{Bias}_F(a)^2$. Show that $R(F, a)$ is minimized at $a \equiv \mu(F)$.

$$R(F, a) = \int (\mu(F) - a)^2 dF$$

$$R(F, a) = \int (\mu(F)^2 - 2\mu(F)a + a^2) dF$$

$$R(F, a) = \mu(F)^2 - 2\mu(F) \int a dF + \int a^2 dF$$

$$\text{Var}_F(a) = E_F(a^2) - E_F(a)^2 = \int a^2 dF - (\int a dF)^2$$

$$\text{Bias}_F(a)^2 = E_F(a - \mu(F))^2 = (\int (a - \mu(F)) dF)^2$$

$$\text{Bias}_F(a)^2 = (\int a dF - \mu(F))^2 = (\int a dF)^2 - 2\mu(F) \int a dF + \mu(F)^2$$

$$\text{Var}_F(a) + \text{Bias}_F(a)^2 = \mu(F)^2 - 2\mu(F) \int a dF + \int a^2 dF = R(F, a)$$

Minimizing $R(F, a)$

$$R(F, a) = \mu(F)^2 - 2\mu(F) \int a dF + \int a^2 dF$$

$$\frac{d}{da}(R(F, a)) = 2 \int a dF - 2\mu(F) = 0$$

$$\mu(F) = \int a dF \quad R(F, a) \text{ is minimized at } a = \mu(F)$$

Exercise 1.7 Suppose that $L(F, a)$ is the squared-error loss $(\mu(F) - a)^2$. Show that the Bayes action is given by the posterior mean.

$$\begin{aligned}\tilde{a} &= \min_a \int L(f, a) \Pi(dF) = \min_a \int (\mu(F) - a)^2 \Pi(dF) \\ \tilde{a} &= \min_a \int (\mu(F)^2 - 2a\mu(F) + a^2) \Pi(dF) \\ \tilde{a} &= \min_a \int \mu(F)^2 \Pi(dF) - 2a \int \mu(F) \Pi(dF) + a^2 \\ \frac{d}{da} &= 2a - 2E[\mu(F)|Z = z] = 0 \\ \tilde{a} &= \tilde{\mu} = E[\mu(F)|Z = z]\end{aligned}$$

Exercise 1.8 Refer to exercises 1.4 and 1.1. Find the Bayes estimators for these problems and compare them to the maximum likelihood estimates.

From exercise 1.1:

$$X_1, \dots, X_n \sim \text{Bernoulli}(p) \quad p|X_1, \dots, X_n \sim \text{Beta}(\alpha + \sum x_i, \beta + n - \sum x_i)$$

$$\tilde{p} = \frac{\alpha + \sum x_i}{\alpha + \beta + n} \quad \hat{p} = \frac{\sum x_i}{n}$$

$$\tilde{p} = \frac{\alpha}{\alpha + \beta + n} + \frac{n\hat{\theta}}{\alpha + \beta + n}$$

$$\tilde{p} = \frac{\alpha/\alpha + \beta}{\frac{\alpha + \beta + n}{\alpha + \beta}} + \frac{n\hat{\theta}}{\alpha + \beta + n}$$

$$\tilde{p} = \frac{\alpha + \beta}{\alpha + \beta + n} \cdot \frac{\alpha}{\alpha + \beta} + \hat{\theta} \cdot \frac{n}{\alpha + \beta + n}$$

From exercise 1.4:

$$\begin{aligned}
X_1 \text{ dots } X_N &\sim \text{Normal}(\theta, \sigma_0^2) \quad \theta \sim \text{Normal}(m, v) \\
\theta | X_1, \dots, X_N &\sim \text{Normal}\left(\frac{1}{\frac{1}{v^2} + \frac{n}{\sigma_0^2}} \cdot \left(\frac{m}{v} + \frac{\sum x_i}{\sigma_0^2}\right), \frac{1}{\frac{1}{v^2} + \frac{n}{\sigma_0^2}}\right) \\
\tilde{\theta} &= \frac{1}{\frac{1}{v^2} + \frac{n}{\sigma_0^2}} \cdot \left(\frac{m}{v} + \frac{\sum x_i}{\sigma_0^2}\right) \\
\hat{\theta} &= \bar{x} \\
\tilde{\theta} &= \frac{m/v^2}{1/v^2 + n/\sigma_0^2} + \frac{n\hat{\theta}/\sigma_0^2}{1/v^2 + n/\sigma_0^2} \\
\tilde{\theta} &= \frac{1/v^2 \cdot m}{1/v^2 + n/\sigma_0^2} + \frac{n/\sigma_0^2 \cdot \hat{\theta}}{1/v^2 + n/\sigma_0^2}
\end{aligned}$$