

203A Question Bank

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1. (Final 2023): Let $X_n \sim N(0, 1)$ if n is odd, and $X_n \sim N(0, n)$ if n is even. Is $X_n = O_p(1)$?

Solution: No. The definition of $O_p(1)$ is that $\lim_{n \rightarrow \infty} P(|X_n| \leq B) = 1$ for all B . In this case, the limit is not 1 for any B . Note: I'm not sure I like this solution.

2. (Final 2023): Suppose that X_1, X_2, \dots are iid such that their common MGF is

$$E[\exp(tX_i)] = \left(\frac{1}{1-t}\right)^2$$

Let $F_n(x) \equiv P(\frac{1}{n} \sum_{i=1}^n X_i^2 \leq x)$. What is $\lim_{n \rightarrow \infty} F_n(1)$?

Solution: First note that the MGF is that of an exponential distribution with $\lambda = 1$. The mean of an exponential distribution is $\frac{1}{\lambda} = 1$. The variance is $\frac{1}{\lambda^2} = 1$. Since $\text{Var}(X) = E[X^2] - E[X]^2$ That implies $E[X^2] = 2$. Now we use the fact that $\lim_{n \rightarrow \infty} F_n(1) \equiv \lim_{n \rightarrow \infty} P(\frac{1}{n} \sum_{i=1}^n X_i^2 \leq 1) = P(\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X_i^2 \leq 1) = P(E[X^2] \leq 1)$. Since $E[X^2] = 2$, the answer is 0. Note: I'm not sure I like this solution.

3. (Final 2023): Suppose that X has the CDF equal to $\Lambda(t) = \frac{\exp(t)}{1+\exp(t)}$. Let $X_n \equiv \cos(X/n)$. What is $\lim_{n \rightarrow \infty} E[\cos(X_n)]$?

Solution:

4. (Final 2023): Let X_1 denote a random sample of (size 1) from $N(1, \theta)$. We have $H_0 : \theta = 4$ and $H_1 : \theta = 9$. You decided to use the Neyman-Pearson test of size 5%. If you observe $X_1 = 6$, do you reject H_0 or not?

Solution: We do not reject H_0 .

1. identify our test statistic: $Z = (X_1 - \mu)/\sigma = (6 - 1)/4 = 5/4$
2. find the critical value: $z_\alpha = 1.645$
3. compare the test statistic to the critical value: $5/4 \leq 1.645$. So we do not reject H_0 .

5. (Final 2023): Suppose that X_n is a sequence of random variables such that $\sqrt{n}(X_n - 1)$ converges in distribution to $N(0, 1)$. By the multivariate delta method, we can see that

$$\begin{pmatrix} \sqrt{n}(X_n^3 - \mu_1) \\ \sqrt{n}(\ln(X_n) - \mu_2) \end{pmatrix} \xrightarrow{d} N \left[\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \sigma_{1,1} & \sigma_{1,2} \\ \sigma_{2,1} & \sigma_{2,2} \end{pmatrix} \right]$$

for some $\mu_1, \mu_2, \sigma_{1,1}, \sigma_{1,2}, \sigma_{2,1}, \sigma_{2,2}$, and $\sigma_{1,2}$. What are their numerical values?

Solution:

6. (Final 2023): Suppose that a random sample X_1 of size 1 where $X_i \sim N(\mu, 1)$. We want to test the null hypothesis $\mu = 0$ versus $H_1 : \mu > 0$. We will reject the null if $X_1 > c$. Suppose that c was chosen such that the size of the test is 2.5%. For what value of μ is the power of the test 5%?

Solution:

7. (Final 2023): Let X_1 denote a random sample (of size 1) from a Poisson distribution with mean equal to θ . We would like to test $H_0 : \theta = 5$ against $H_1 : \theta \neq 5$.
- (a) Suppose that $X_1 = 25$. What is the value of the LR statistic? Hint 1: The PDF of the Poisson Distribution with mean equal to θ is given by $f(x; \theta) = \frac{\theta^x e^{-\theta}}{x!}$ Hint 2: $\ln 5 = 1.6094$

Solution:

- (b) Do you reject or accept the null at the 5% significance level?

Solution:

8. (Final 2023): Suppose that X_1, \dots, X_{10} are iid $N(0, 1)$. Let $\bar{X} = \frac{1}{10} \sum_{i=1}^{10} X_i$. What is $E[\bar{X}^2 (\sum_{i=1}^{10} (X_i - \bar{X})^2)]$?

Solution:

9. (Final 2018): Let X have the uniform distribution over the $(0, 1)$ interval

- (a) Calculate $E[\log(X)]$

Solution: Use integration by parts: $E[\log(X)] = \int_0^1 \log(x) dx = x \log(x) - x|_0^1 = 0 - 1 = -1$.

- (b) Calculate $\log E[X]$

Solution: $E[X] = \int_0^1 x dx = 1/2$. So $\log E[X] = \log(1/2) = -\log(2)$.

- (c) Which is bigger?

Solution: $\log E[X] = -\log(2) > -1 = E[\log(X)]$. So $\log E[X]$ is bigger.

- (d) Now let g denote some positive valued strictly increasing function defined on $(0, 1)$. Let $Y = g(X)$. Between $E[\log(Y)]$ and $\log(E[Y])$, which one is bigger? Or does the answer depend on g ? Note that there are only three choices.

Solution: Jensen's inequality says that $\phi(E[X]) \leq E[\phi(X)]$ for a convex function ϕ . Since \log is concave, we have that $\log(E[Y]) \geq E[\log(Y)]$. So $\log(E[Y])$ is bigger.

Note: Concavity is the tricky part here. You can convert \log into a convex function by multiplying the \log by -1 , which makes it convex. Then you can apply Jensen's inequality.

10. (Final 2018): Suppose that the support of Y is $\{0, 1\}$, and $Pr[Y = 1|X] = \Phi(X'\beta)$ where Φ is the standard normal cdf. We know that $Pr[Y = 1|(X_1, X_2) = (2, 1)] = .5$ and $Pr[Y = 1|(X_1, X_2) = (2, 2)] = .975$. What is the value of $\beta = (\beta_1, \beta_2)'$?

$\Phi(x)$	Value
$\Phi(0)$	0.5
$\Phi(0.253)$	0.6
$\Phi(0.534)$	0.7
$\Phi(0.842)$	0.8
$\Phi(1.282)$	0.9
$\Phi(1.645)$	0.95
$\Phi(1.960)$	0.975
$\Phi(2.576)$	0.995

The Phi scores you may want to use are:

Solution: $\beta_1 = -.98$ and $\beta_2 = 1.96$

We have $Pr[Y = 1|(2, 1)] = \Phi(2\beta_1 + \beta_2) = .5$ and $Pr[Y = 1|(2, 2)] = \Phi(2\beta_1 + 2\beta_2) = .975$. We can take the inverse of Φ using the table above. So we have $2\beta_1 + \beta_2 = 0$ and $2\beta_1 + 2\beta_2 = 1.960$. By subtraction, we get $\beta_2 = 1.96$. We then solve for β_1 using the first equation to get $\beta_1 = -.98$.

11. (Final 2018): Let X_1, \dots, X_n be a random sample of size n from $N(3, 2)$, and let \bar{X} denote the sample average. what is the asymptotic distribution of $\sqrt{n}(\bar{X}^2 - 9)$? Your answer should be numerical.

Solution: We use the Delta method: $\sqrt{n}[g(\bar{X}) - g(a)] \xrightarrow{d} N(0, g'(a)^2 \sigma^2)$ where $g(x) = x^2$ and $a = 3$. So $g'(a) = 6$ and $\sigma^2 = 2$. So the answer is $N(0, 72)$.

12. (Final 2018): Let X_n denote a sequence of random variables such that the PDF f_n of X_n is given by

$$f_n(x) = \frac{1}{2n} 1(|X| \leq n)$$

- (a) Let $0 < B < \infty$ be given what is $\lim_{n \rightarrow \infty} P(|X_n| \leq B)$?

Solution: $\lim_{n \rightarrow \infty} P(|X_n| \leq B) = \lim_{n \rightarrow \infty} \int_{-B}^B \frac{1}{2n} dx = \lim_{n \rightarrow \infty} \frac{B}{n} = 0$
 Note: I'm not sure I like this solution.

(b) Is $X_n = O_p(1)$?

Solution: X_n is not bounded in probability. The definition of bounded in probability is $\lim_{n \rightarrow \infty} P(|X_n| \leq B) = 1$, which is not the case here.

13. (Final 2018, 2019) Suppose that X and ϵ are independent $N(0, 1)$ variables. Let $Y = X + \epsilon$ What is the correlation between X^2 and Y

Solution: The correlation is zero. I'll write the algebra down later.

14. (Final 2018, 2019): True or False

(a) If $X_n = a + o_p(1)$, then $E[X_n] = a + o(1)$

Solution:

(b) If $X_n = a + o_p(1)$, then $X_n^2 = a^2 + o_p(1)$

Solution:

(c) If $X_n \xrightarrow{d} N(0, 1)$, then $E[X_n] = o(1)$

Solution:

(d) If $X_n \xrightarrow{d} N(0, 1)$, then $E[X_n^2] = 1 + o(1)$

Solution:

(e) If $X_n \xrightarrow{d} N(0, 1)$, $\chi_n^2 \xrightarrow{d} \chi^2(1)$

Solution:

15. (Final 2019): Suppose that $Z \sim N(0, 1)$, and let $X_n = Z^2 1(|Z| \geq \frac{1}{n})$. What is $\lim_{n \rightarrow \infty} E[X_n]$?

Solution: Not finished. I think there's some law that let's us move the limit inside the expectation.

$$\lim_{n \rightarrow \infty} E[X_n] = E[\lim_{n \rightarrow \infty} X_n] = E[\lim_{n \rightarrow \infty} Z^2 1(|Z| \geq \frac{1}{n})] = E[Z^2] = \chi^2(1) = 1$$

16. (Final 2019): Let X_1, \dots, X_n be a random sample of size n from $N(\mu, 1)$, and let \bar{X} denote the sample average.

(a) Compute $\sup_{\mu < 0} \lim_{n \rightarrow \infty} P(\sqrt{n}\bar{X} > 1.645)$

Solution:

(b) $\lim_{n \rightarrow \infty} \sup_{\mu < 0} P(\sqrt{n}(\bar{X} - \mu) > 1.645)$

Solution:

17. (Final 2019) Let X_n denote a sequence of random variables such that the PDF f_n of X_n is given by

$$f_n(x) = \frac{1}{4}1(n \leq |x| \leq n+1) + \frac{1}{4}1(|x| \leq 1)$$

- (a) Let $1 < B < \infty$ be given what is $\lim_{n \rightarrow \infty} P(|X_n| \leq B)$?

Solution:

- (b) Is $X_n = O_p(1)$?

Solution:

18. (Final 2019, 2021) Consider X_1, \dots, X_n iid $N(\mu, \sigma^2)$. We assume that $\sigma^2 = 1$. We have $H_0 : \mu = 0$ vs $H_1 : \mu > 0$. Suppose that C is a critical region such that (1) the test of the form "Reject H_0 if $(X_1, \dots, X_n) \in C$ " is the uniformly most powerful test; and (2) the probability of rejecting H_0 when $\mu = 0$ is 5%.

- (a) Provide a mathematical characterization of rejecting H_0 of such a test as a function of n and $\mu > 0$.

Solution:

- (b) What is the power of the test when $\mu = .1645$ and $n = 100$?

Solution:

19. (Final 2021) Suppose that X_1, X_2, \dots are iid such that $X_i = 0$ with probability $1/2$, and $X_i = 2$ with probability $1/2$. Let $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$, and $F_n(x) \equiv P(\frac{1}{n} \sum_{i=1}^n X_i \leq x)$. What is $\lim_{n \rightarrow \infty} F(1.1)$?

Solution:

20. (Final 2021) Suppose that X has the CDF equal to

$$Pr[X \leq x] = \frac{\exp(x)}{1 + \exp(x)} \equiv \Lambda(x)$$

Let $\Phi(\cdot)$ denote the CDF of $N(0, 1)$, and let Φ^{-1} denote its inverse. $Y \equiv \Phi^{-1}(\Lambda(X))$. What is $E[Y^4]$?

Solution:

21. (Final 2021) Consider X_1, \dots, X_n iid $N(0, 1)$. Let $Y_n = 1(\bar{X}_n \geq 1)$.

- (a) What is $E[Y_n]$?

Solution:

- (b) What is $Var(Y_n)$?

Solution:

- (c) What is the probability limit of Y_n ?

Solution:

22. (Final 2021) Suppose that X_1, \dots, X_4 are iid and their common distribution is uniform $(0, \theta)$, i.e. their common PDF $f(x)$ is equal to $1/(0 < x < \theta)$. We have $H_0 : \theta = 1$ vs $H_1 : \theta = 2$. Suppose that you decided to reject H_0 if $\max(X_1, \dots, X_4) > 1$

- (a) What is the size of the test?

Solution:

- (b) What is the power of the test? Hint: $P(\max(X_1, \dots, X_4) \leq 1) = P[X_1 \leq 1, X_2 \leq 1, X_3 \leq 1, X_4 \leq 1]$

Solution:

23. (Final 2021) Let X_1, \dots, X_n be a random sample of size n from $N(0, 1)$. For any positive integer k , let $m_k = E[X_i^k]$ and $\hat{m}_k = \frac{1}{n} \sum_{i=1}^n X_i^k$. Assuming that $E[|X_i^k|] < \infty$ for all k , derive the asymptotic distribution $\sqrt{n}(\hat{m}_k^{1/2} - m_k^{1/2})$. The asymptotic distribution is normal with mean zero, so your job is to derive the numerical value of the asymptotic variance. Hint: the MGD of $N(0, 1)$ is $\exp(t^2/2)$.

Solution:

24. (Final 2021) Let X_1, \dots, X_5 denote a random sample from $N(\theta, 1)$. We would like to test $H_0 : \theta = 5$ against $H_1 : \theta \neq 5$. If $X_1 = 2, X_2 = 3, X_3 = 4, X_4 = 5, X_5 = 6$, what is the LR statistic?

Solution:

25. (Final 2022) Let $X_n \sim b(n, \frac{q}{n})$ for some $q > 0$. Is $X_n = Op(1)$

Solution:

26. (Final 2022) Let Y_n denote the maximum of a random sample of size n from a uniform $(0, 1)$ distribution. What is $\lim_{n \rightarrow \infty} Pr[Y_n \leq 0.9]$?

Solution:

27. (Final 2022) Let $\Lambda(t) \equiv \frac{\exp(t)}{1 + \exp(t)}$ denote the CDF of the logistic distribution (with location and scale parameters equal to 0 and 1, although these particular details are irrelevant for this question). Let X_n denote a sequence of random variables such that the CDF F_n of X_n is given by $F_n(x) = \Lambda(nx)$. What is $\lim_{n \rightarrow \infty} E[\cos(X_n)]$?

Solution:

28. (Final 2022) Let X_1 denote a random sample (of size 1) from $\mathcal{N}(0, \sigma^2)$. We have null hypothesis $H_0 : \mu = 1$ and alternative hypothesis $H_1 : \mu = 9$.

Solution:

29. (Final 2022) Suppose that X_n is a sequence of random variables such that $\sqrt{n}(X_n - 1)$ converges in distribution to $N(0, 1)$. By the multivariate delta method, we can see that

$$\begin{pmatrix} \sqrt{n}(X_n^2 - \mu_1) \\ \sqrt{n}(\ln(X_n) - \mu_2) \end{pmatrix} \xrightarrow{d} N \left[\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \sigma_{1,1} & \sigma_{1,2} \\ \sigma_{2,1} & \sigma_{2,2} \end{pmatrix} \right]$$

for some $\mu_1, \mu_2, \sigma_{1,1}, \sigma_{1,2}, \sigma_{2,1}, \sigma_{2,2}$, and $\sigma_{1,2}$. What are their numerical values?

Solution:

30. (Final 2022) Let X denote a random sample (of size 1) from a distribution with the PDF equal to $\lambda \exp(-\lambda x) 1(x > 0)$. We have $H_0 : \lambda = 1$ and $H_1 : \lambda < 1$. Suppose that C is a critical region such that (1) the test of the form "Reject H_0 if $X_1 \in C$ " is the uniformly most powerful test; and (2) the probability of rejecting H_0 when $\lambda = 1$ is α . What is the power of your test when $\lambda = 1/2$ and $\alpha = 5\%$?

Solution:

31. (Final 2022) Let X denote a random sample (of size 1) from a distribution with the PDF equal to $\lambda \exp(-\lambda x) 1(x > 0)$. We would like to test $H_0 : \lambda = 1$ against $H_1 : \lambda \neq 1$. Suppose that $X = e$
- (a) What is the value of the LR statistic?

Solution:

- (b) Do you reject or accept the null at the 5% significance level? You may use the approximation $e = 2.7183$

Solution:

32. (Final 2022) Let (X, Y) be a two dimensional random vector with the joint PDF $f_{X,Y}(x, y) = 4e^{-2y} 1(y > x > 0)$. Let $(\alpha, \beta) \equiv \operatorname{argmin}_{a,b} E[(Y - (a + bX))^2]$ What is (α, β) ? Hint: a small number of students may find it useful to know that $\int_0^\infty x^m \lambda e^{-\lambda x} dx = \frac{m!}{\lambda^{m+1}}$
33. (Final 2015) Suppose that X_n is a sequence of random variables such that $\sqrt{n}(X_n - 3) \xrightarrow{d} N(0, 1)$
- (a) What is the asymptotic distribution of $\sqrt{n}(X_n^2 - 9)$?

Solution:

- (b) What is the asymptotic distribution of $(\sqrt{n}(X_n^2 - 3))^2$?

Solution:

34. (Final 2015) Let $F(y|x)$ denote the conditional CDF of Y given X i.e. $F(y|x) = Pr[Y \leq y|X = x]$. Suppose that $F(y|x)$ is continuous and strictly increasing in y for all x in the support of X . Let $V = F(Y|X)$. (It is not $F(Y|x)$). Derive the conditional distribution of V given x . Prove that V and X are independent.

Solution: