

**NOVEMBER 2004
COURSE 3 SOLUTIONS**

Question #1

Key: D

$$E[Z] = \int_0^\infty b_t v^t {}_t p_x \mu(x+t) dt = \int_0^\infty e^{0.06t} e^{-0.08t} e^{-0.05t} \frac{1}{20} dt$$

$$= \frac{1}{20} \left(\frac{100}{7} \right) \left[-e^{-0.07t} \right]_0^\infty = \frac{5}{7}$$

$$E[Z^2] = \int_0^\infty (b_t v^t)^2 {}_t p_x \mu(x+t) dt = \int_0^\infty e^{0.12t} e^{-0.16t} e^{-0.05t} \frac{1}{20} dt = \frac{1}{20} \int_0^\infty e^{-0.09t} dt$$

$$= \frac{1}{20} \left(\frac{100}{9} \right) \left[e^{-0.09t} \right]_0^\infty = \frac{5}{9}$$

$$Var[Z] = \frac{5}{9} - \left(\frac{5}{7} \right)^2 = 0.04535$$

Question #2

Key: C

Let ns = nonsmoker and s = smoker

$k =$	$q_{x+k}^{(ns)}$	$p_{x+k}^{(ns)}$	$q_{x+k}^{(s)}$	$p_{x+k}^{(s)}$
0	.05	0.95	0.10	0.90
1	.10	0.90	0.20	0.80
2	.15	0.85	0.30	0.70

$$A_{x:\overline{2}|}^{1(ns)} = v q_x^{(ns)} + v^2 p_x^{(ns)} q_{x+1}^{(ns)}$$

$$= \frac{1}{1.02} (0.05) + \frac{1}{1.02^2} (0.95 \times 0.10) = 0.1403$$

$$A_{x:\overline{2}|}^{1(s)} = v q_x^{(s)} + v^2 p_x^{(s)} q_{x+1}^{(s)}$$

$$= \frac{1}{1.02} (0.10) + \frac{1}{(1.02)^2} (0.90 \times 0.20) = 0.2710$$

$$A_{x:\overline{2}|}^1 = \text{weighted average} = (0.75)(0.1403) + (0.25)(0.2710)$$

$$= 0.1730$$

Question #3**Key: A**

$$\bar{P}(\bar{A}_x) = \mu = 0.03$$

$${}^2\bar{A}_x = 0.20 = \frac{\mu}{2\delta + \mu} = \frac{0.03}{2\delta + 0.03}$$

$$\Rightarrow \delta = 0.06$$

$$\text{Var}({}_0L) = \frac{{}^2\bar{A}_x - (\bar{A}_x)^2}{(\delta \bar{a})^2} = \frac{0.20 - \left(\frac{1}{3}\right)^2}{\left(\frac{0.06}{0.09}\right)^2} = 0.20$$

$$\text{where } A = \frac{\mu}{\mu + \delta} = \frac{0.03}{0.09} = \frac{1}{3} \quad \bar{a} = \frac{1}{\mu + \delta} = \frac{1}{0.09}$$

Question #4**Key: B**

$$s(60) = \frac{e^{-(0.1)(60)} + e^{-(0.08)(60)}}{2}$$

$$= 0.005354$$

$$s(61) = \frac{e^{-(0.1)(61)} + e^{-(0.08)(61)}}{2}$$

$$= 0.00492$$

$$q_{60} = 1 - \frac{0.00492}{0.005354} = 0.081$$

Question #5**Key: B**

For Ω , $0.4 = F(\omega) = \left(\frac{\omega}{80}\right)^2$

$$0.6325 = \frac{\omega}{80}$$

$$\omega = 50.6$$

For $T(0)$ using De Moivre, $0.7 = F(t) = \frac{t}{\omega} = \frac{t}{50.6}$
 $t = (0.7)(50.6) = 35.42$

Question #6**Key: C**

$$E[N] = mq = 1.8 \Rightarrow q = \frac{1.8}{3} = 0.6$$

x	$f_N(x)$	$F_N(x)$
0	0.064	0.064
1	0.288	0.352
2	0.432	0.784
3	0.216	1.000

First: $0.432 < 0.7 < 0.784$ so $N = 2$. Use 0.1 and 0.3 for amounts
 Second: $0.064 < 0.1 < 0.352$ so $N = 1$ Use 0.9 for amount
 Third: $0.432 < 0.5 < 0.784$ so $N = 2$ Use 0.5 and 0.7 for amounts

Discrete uniform $\Rightarrow F_X(x) = 0.2x$, $x = 1, 2, 3, 4, 5$

$$0.4 < 0.5 < 0.6 \Rightarrow x_1 = 3$$

$$0.6 < 0.7 < 0.8 \Rightarrow x_2 = 4$$

Aggregate claims = $3+4 = 7$

Question #7**Key: C**

$$E(X \wedge x) = \frac{\theta}{\alpha - 1} \left[1 - \left(\frac{\theta}{x + \theta} \right)^{\alpha - 1} \right] = \frac{2000x}{x + 2000}$$

x	$E(X \wedge x)$
∞	2000
250	222
2250	1059
5100	1437

$$0.75(E(X \wedge 2250) - E(X \wedge 250)) + 0.95(E(X) - E(X \wedge 5100))$$

$$0.75(1059 - 222) + 0.95(2000 - 1437) = 1162.6$$

The 5100 breakpoint was determined by when the insured's share reaches 3600:

$$3600 = 250 + 0.25(2250 - 250) + (5100 - 2250)$$

Question #8**Key: D**

Since each time the probability of a heavy scientist is just half the probability of a success, the distribution is binomial with $q = 0.6 \times 0.5 = 0.3$ and $m = 8$.

$$f(2) = (8 \times 7 / 2) \times (0.3^2) \times (0.7^6) = 0.30$$

Question #9**Key: A**

$$\mu_{xy}(t) = \mu_x(t) + \mu_y(t) = 0.08 + 0.04 = 0.12$$

$$\bar{A}_x = \mu_x(t) / (\mu_x(t) + \delta) = 0.5714$$

$$\bar{A}_y = \mu_y(t) / (\mu_y(t) + \delta) = 0.4$$

$$\bar{A}_{xy} = \mu_{xy}(t) / (\mu_{xy}(t) + \delta) = 0.6667$$

$$\bar{a}_{xy} = 1 / (\mu_{xy}(t) + \delta) = 5.556$$

$$\bar{A}_{xy} = \bar{A}_x + \bar{A}_y - \bar{A}_{xy} = 0.5714 + 0.4 - 0.6667 = 0.3047$$

$$\text{Premium} = 0.304762 / 5.556 = 0.0549$$

Question #10**Key: B**

$$P_{40} = A_{40} / \ddot{a}_{40} = 0.16132 / 14.8166 = 0.0108878$$

$$P_{42} = A_{42} / \ddot{a}_{42} = 0.17636 / 14.5510 = 0.0121201$$

$$a_{45} = \ddot{a}_{45} - 1 = 13.1121$$

$$\begin{aligned} E\left[{}_3L \mid K(42) \geq 3 \right] &= 1000A_{45} - 1000P_{40} - 1000P_{42} a_{45} \\ &= 201.20 - 10.89 - (12.12)(13.1121) \\ &= 31.39 \end{aligned}$$

Many similar formulas would work equally well. One possibility would be

$1000 {}_3V_{42} + (1000P_{42} - 1000P_{40})$, because prospectively after duration 3, this differs from the normal benefit reserve in that in the next year you collect $1000P_{40}$ instead of $1000P_{42}$.

Question #11**Key: E**

For De Moivre's Law:

$$e_x^\circ = \frac{\omega - x}{2}$$

$${}_k|q_x = \frac{1}{\omega - x}$$

$$A_x = \sum_{k=b}^{\omega-x-1} v^{k+1} {}_k|q_x = \frac{1}{\omega - x} \sum_{k=b}^{\omega-x-1} v^{k+1}$$

$$A_x = \frac{a_{\overline{\omega-x}|}}{\omega - x}$$

$$\ddot{a}_x = \frac{1 - A_x}{d}$$

$$e_{50}^\circ = 25 \Rightarrow \omega = 100 \text{ for typical annuitants}$$

$$e_y^\circ = 15 \Rightarrow y = \text{Assumed age} = 70$$

$$A_{70} = \frac{a_{\overline{30}|}}{30} = 0.45883$$

$$\ddot{a}_{70} = 9.5607$$

$$500000 = b \ddot{a}_{20} \Rightarrow b = 52,297$$

Question #12**Key: D**

$$p_x^{(\tau)} = p_x^{(1)} p_x^{(2)} = 0.8(0.7) = 0.56$$

$$q_x^{(1)} = \left[\frac{\ln(p_x^{(1)})}{\ln(p_x^{(\tau)})} \right] q_x^{(\tau)} \text{ since UDD in double decrement table}$$

$$= \left[\frac{\ln(0.8)}{\ln(0.56)} \right] 0.44$$

$$= 0.1693$$

$${}_{0.3}q_{x+0.1}^{(1)} = \frac{0.3q_x^{(1)}}{1 - 0.1q_x^{(\tau)}} = 0.053$$

To elaborate on the last step:

$${}_{0.3}q_{x+0.1}^{(1)} = \frac{\left(\begin{array}{l} \text{Number dying from cause} \\ 1 \text{ between } x + 0.1 \text{ and } x + 0.4 \end{array} \right)}{\text{Number alive at } x + 0.1}$$

Since UDD in double decrement,

$$= \frac{l_x^{(\tau)} (0.3) q_x^{(1)}}{l_x^{(\tau)} (1 - 0.1 q_x^{(\tau)})}$$

Question #13**Key: B**

non absorbing matrix $T = \begin{pmatrix} 0.7 & 0.1 \\ 0.3 & 0.6 \end{pmatrix}$, the submatrix excluding “Terminated”, which is an absorbing state.

$$I - T = \begin{pmatrix} 0.3 & -0.1 \\ -0.3 & 0.4 \end{pmatrix}$$

$$(I - T)^{-1} = \begin{pmatrix} \frac{0.4}{0.09} & \frac{0.1}{0.09} \\ \frac{0.3}{0.09} & \frac{0.3}{0.09} \end{pmatrix} = \begin{pmatrix} 4.\overline{44} & 1.\overline{11} \\ 3.\overline{33} & 3.\overline{33} \end{pmatrix}$$

$$\begin{aligned} \text{Future costs for a healthy} &= 4.\overline{44} \times 500 + 1.\overline{11} \times 3000 \\ &= 5555 \end{aligned}$$

Question #14**Key: D**

$$T = \begin{pmatrix} 0.7 & 0.1 & 0.2 \\ 0.3 & 0.6 & 0.1 \\ 0 & 0 & 1 \end{pmatrix} \quad T^2 = \begin{pmatrix} 0.52 & 0.13 & 0.35 \\ 0.39 & 0.39 & 0.22 \\ 0 & 0 & 1 \end{pmatrix}$$

$$\text{Actuarial present value (A.P.V.) prem} = 800(1 + (0.7 + 0.1) + (0.52 + 0.13)) = 1,960$$

$$\text{A.P.V. claim} = 500(1 + 0.7 + 0.52) + 3000(0 + 0.1 + 0.13) = 1800$$

$$\text{Difference} = 160$$

Question #15**Key: A**

Let N_1, N_2 denote the random variable for # of claims for Type I and II in 2 years

X_1, X_2 denote the claim amount for Type I and II

S_1 = total claim amount for type I in 2 years

S_2 = total claim amount for Type II at time in 2 years

$S = S_1 + S_2$ = total claim amount in 2 years

$\{S_1\} \rightarrow$ compound poisson $\lambda_1 = 2 \times 6 = 12$ $X_1 \sim U(0, 1)$

$\{S_2\} \rightarrow$ compound poisson $\lambda_2 = 2 \times 2 = 4$ $X_2 \sim U(0, 5)$

$$E(N_1) = Var(N_1) = 2 \times 6 = 12$$

$$E(S_1) = E(N_1)E(X_1) = (12)(0.5) = 6$$

$$\begin{aligned} Var(S_1) &= E(N_1)Var(X_1) + Var(N_1)(E(X_1))^2 \\ &= (12)\frac{(1-0)}{12} + (12)(0.5)^2 \\ &= 4 \end{aligned}$$

$$E(N_2) = Var(N_2) = 2 \times 2 = 4$$

With formulas corresponding to those for S_1 ,

$$E(S_2) = 4 \times \frac{5}{2} = 10$$

$$Var(S_2) = 4 \times \frac{(5-0)^2}{12} + 4\left(\frac{5}{2}\right)^2 = 33.\bar{3}$$

$$E(S) = E(S_1) + E(S_2) = 6 + 10 = 16$$

Since S_1 and S_2 are independent,

$$Var(S) = Var(S_1) + Var(S_2) = 4 + 33.\bar{3} = 37.\bar{3}$$

$$\Pr(S > 18) = \Pr\left(\frac{S - 16}{\sqrt{39.\bar{3}}} > \frac{2}{\sqrt{37.\bar{3}}} = 0.327\right)$$

Using normal approximation

$$\begin{aligned} \Pr(S > 18) &= 1 - \Phi(0.327) \\ &= 0.37 \end{aligned}$$

Question #16**Key: D**

Since the rate of depletion is constant there are only 2 ways the reservoir can be empty sometime within the next 10 days.

Way #1:

There is no rainfall within the next 5 days

Way #2

There is one rainfall in the next 5 days

And it is a normal rainfall

And there are no further rainfalls for the next five days

$$\text{Prob (Way \#1)} = \text{Prob}(0 \text{ in } 5 \text{ days}) = \exp(-0.2 \cdot 5) = 0.3679$$

$$\begin{aligned} \text{Prob (Way \#2)} &= \text{Prob}(1 \text{ in } 5 \text{ days}) \times 0.8 \times \text{Prob}(0 \text{ in } 5 \text{ days}) \\ &= 5 \cdot 0.2 \exp(-0.2 \cdot 5) \cdot 0.8 \cdot \exp(-0.2 \cdot 5) \\ &= 1 \exp(-1) \cdot 0.8 \cdot \exp(-1) = 0.1083 \end{aligned}$$

$$\text{Hence Prob empty at some time} = 0.3679 + 0.1083 = 0.476$$

Question #17**Key: C**

Let X be the loss random variable,

So $(X - 5)_+$ is the claim random variable.

$$E(X) = \frac{10}{2.5 - 1} = 6.\bar{6}$$

$$\begin{aligned} E(X \wedge 5) &= \left(\frac{10}{2.5 - 1} \right) \left[1 - \left(\frac{10}{5 + 10} \right)^{2.5 - 1} \right] \\ &= 3.038 \end{aligned}$$

$$\begin{aligned} E(X - 5)_+ &= E(X) - E(X \wedge 5) \\ &= 6.\bar{6} - 3.038 \\ &= 3.629 \end{aligned}$$

$$\begin{aligned} \text{Expected aggregate claims} &= E(N)E(X - 5)_+ \\ &= (5)(3.629) \\ &= 18.15 \end{aligned}$$

Question #18**Key: B**

A Pareto ($\alpha = 2, \theta = 5$) distribution with 20% inflation becomes Pareto with

$$\alpha = 2, \theta = 5 \times 1.2 = 6$$

$$\text{In 2004, } E(X) = \frac{6}{2-1} = 6$$

$$E(X \wedge 10) = \frac{6}{2-1} \left(1 - \left(\frac{6}{10+6} \right)^{2-1} \right) = 3.75$$

$$\begin{aligned} E(X - 10)_+ &= E(X) - E(X \wedge 10) \\ &= 6 - 3.75 = 2.25 \end{aligned}$$

$$\text{LER} = 1 - \frac{E(X - 10)_+}{E(X)} = 1 - \frac{2.25}{6} = 0.625$$

Question #19**Key: A**

Let X = annual claims

$$\begin{aligned} E(X) &= (0.75)(3) + (0.15)(5) + (0.1)(7) \\ &= 3.7 \end{aligned}$$

$$\pi = \text{Premium} = (3.7) \left(\frac{4}{3} \right) = 4.93$$

$$\begin{aligned} \text{Change during year} &= 4.93 - 3 = +1.93 \text{ with } p = 0.75 \\ &= 4.93 - 5 = -0.07 \text{ with } p = 0.15 \\ &= 4.93 - 7 = -2.07 \text{ with } p = 0.10 \end{aligned}$$

Since we start year 1 with surplus of 3, at end of year 1 we have 4.93, 2.93, or 0.93 (with associated probabilities 0.75, 0.15, 0.10).

We cannot drop more than 2.07 in year 2, so ruin occurs only if we are at 0.93 after 1 and have a drop of 2.07.

$$\text{Prob} = (0.1)(0.1) = 0.01$$

Question #20**Key: E**

$$0.96 = e^{-(\mu_1 + \lambda)}$$

$$\mu_1 + \lambda = -\ln(0.96) = 0.04082$$

$$\mu_1 = 0.04082 - \lambda = 0.04082 - 0.01 = 0.03082$$

Similarly

$$\mu_2 = -\ln(0.97) - \lambda = 0.03046 - 0.01 = 0.02046$$

$$\mu_{xy} = \mu_1 + \mu_2 + \lambda = 0.03082 + 0.02046 + 0.01 = 0.06128$$

$${}_5P_{xy} = e^{-(5)(0.06128)} = e^{-0.3064} = 0.736$$

Question #21**Key: C**

$$A_{60} = 0.36913 \quad d = 0.05660$$

$${}^2A_{60} = 0.17741$$

$$\text{and } \sqrt{{}^2A_{60} - A_{60}^2} = 0.202862$$

$$\text{Expected Loss on one policy is } E[L(\pi)] = \left(100,000 + \frac{\pi}{d}\right)A_{60} - \frac{\pi}{d}$$

$$\text{Variance on one policy is } \text{Var}[L(\pi)] = \left(100,000 + \frac{\pi}{d}\right)^2 ({}^2A_{60} - A_{60}^2)$$

On the 10000 lives,

$$E[S] = 10,000E[L(\pi)] \text{ and } \text{Var}[S] = 10,000 \text{Var}[L(\pi)]$$

The π is such that $0 - E[S] / \sqrt{\text{Var}[S]} = 2.326$ since $\Phi(2.326) = 0.99$

$$\frac{10,000 \left(\frac{\pi}{d} - \left(100,000 + \frac{\pi}{d}\right)A_{60} \right)}{100 \left(100,000 + \frac{\pi}{d}\right) \sqrt{{}^2A_{60} - A_{60}^2}} = 2.326$$

$$\frac{100 \left(\frac{\pi}{d} - \left(100,000 + \frac{\pi}{d}\right)(0.36913) \right)}{\left(100,000 + \frac{\pi}{d}\right)(0.202862)} = 2.326$$

$$\frac{0.63087 \frac{\pi}{d} - 36913}{100,000 + \frac{\pi}{d}} = 0.004719$$

$$0.63087 \frac{\pi}{d} - 36913 = 471.9 = 0.004719 \frac{\pi}{d}$$

$$\begin{aligned} \frac{\pi}{d} &= \frac{36913 + 471.9}{0.63087 - 0.004719} \\ &= 59706 \end{aligned}$$

$$\pi = 59706 \times d = 3379$$

Question #22**Key: C**

$$\begin{aligned}
 {}_1V &= ({}_0V + \pi)(1+i) - (1000 + {}_1V - {}_1V) \times q_{75} \\
 &= 1.05\pi - 1000q_{75}
 \end{aligned}$$

Similarly,

$${}_2V = ({}_1V + \pi) \times 1.05 - 1000q_{76}$$

$${}_3V = ({}_2V + \pi) \times 1.05 - 1000q_{77}$$

$$1000 = {}_3V = (1.05^3\pi + 1.05^2 \cdot \pi + 1.05\pi) - 1000 \times q_{75} \times 1.05^2 - 1000 \times 1.05 \times q_{76} - 1000 \times q_{77} \quad *$$

$$\begin{aligned}
 \pi &= \frac{1000 + 1000(1.05^2 q_{75} + 1.05 q_{76} + q_{77})}{(1.05)^3 + (1.05)^2 + 1.05} \\
 &= \frac{1000^x (1 + 1.05^2 \times 0.05169 + 1.05 \times 0.05647 + 0.06168)}{3.310125} \\
 &= \frac{1000 \times 1.17796}{3.310125} = 355.87
 \end{aligned}$$

* This equation is algebraic manipulation of the three equations in three unknowns $({}_1V, {}_2V, \pi)$.

One method – usually effective in problems where benefit = stated amount plus reserve, is to multiply the ${}_1V$ equation by 1.05^2 , the ${}_2V$ equation by 1.05, and add those two to the ${}_3V$ equation: in the result, you can cancel out the ${}_1V$, and ${}_2V$ terms. Or you can substitute the ${}_1V$ equation into the ${}_2V$ equation, giving ${}_2V$ in terms of π , and then substitute that into the ${}_3V$ equation.

Question #23**Key: D**

Actuarial present value (APV) of future benefits =

$$\begin{aligned} &= (0.005 \times 2000 + 0.04 \times 1000) / 1.06 + (1 - 0.005 - 0.04)(0.008 \times 2000 + 0.06 \times 1000) / 1.06^2 \\ &= 47.17 + 64.60 \\ &= 111.77 \end{aligned}$$

$$\begin{aligned} \text{APV of future premiums} &= \left[1 + (1 - 0.005 - 0.04) / 1.06 \right] 50 \\ &= (1.9009)(50) \\ &= 95.05 \end{aligned}$$

$$E \left[{}_1L \mid K(55) \geq 1 \right] = 111.77 - 95.05 = 16.72$$

Question #24**Key: D**

$$\overset{\circ}{e}_0 = \overset{\circ}{e}_{0:\overline{20}|} + {}_{20}p_0 \overset{\circ}{e}_{20}$$

$$\overset{\circ}{e}_0 = E[T] = \frac{50}{3-1} = 25$$

$$\begin{aligned}\overset{\circ}{e}_{0:\overline{20}|} &= E[T \wedge 20] = \frac{50}{3-1} \left(1 - \left(\frac{50}{50+20} \right)^{3-1} \right) \\ &= 12.245\end{aligned}$$

$$\begin{aligned}{}_{20}p_0 &= 1 - F_T(20) = 1 - \left(1 - \left(\frac{50}{50+20} \right)^3 \right) \\ &= 0.3644\end{aligned}$$

$$25 = 12.245 + 0.3644 \overset{\circ}{e}_{20}$$

$$\overset{\circ}{e}_{20} = 35$$

Alternate approach: if losses are Pareto with $\theta = 50$ and $\alpha = 3$, then claim payments per payment with an ordinary deductible of 20 are Pareto with $\theta = 50 + 20$ and $\alpha = 3$.

$$\text{Thus } E(T(20)) = \frac{50+20}{3-1} = 35$$

This alternate approach was shown here for educational reasons: to reinforce the idea that many life contingent models and non-life models can have similar structure. We doubt many candidates would take that approach, especially since it involves specific properties of the Pareto distribution.

Question #25**Key: E**

$Q \geq P$ since in Q you only test at intervals; surplus below 0 might recover before the next test.

In P , ruin occurs if you are ever below 0.

$R \geq P$ since you are less likely to have surplus below 0 in the first N years (finite horizon) than forever.

Add the inequalities

$$Q + R \geq 2P$$

Also (why other choices are wrong)

$S \geq Q$ by reasoning comparable to $R \geq P$. Same testing frequency in S and Q , but Q tests forever.

$S \geq R$ by reasoning comparable to $Q \geq P$. Same horizon in S and R ; R tests more frequently.

$S \geq P$ P tests more frequently, and tests forever.

Question #26**Key: A**

This is a nonhomogeneous Poisson process with intensity function

$$\lambda(t) = 3 + 3t, \quad 0 \leq t \leq 2, \text{ where } t \text{ is time after noon}$$

$$\begin{aligned} \text{Average } \lambda &= \frac{\int_1^2 \lambda(t) dt}{1} = \int_1^2 (3 + 3t) dt \\ &= \left[3t + \frac{3t^2}{2} \right]_1^2 \\ &= 7.5 \end{aligned}$$

$$f(2) = \frac{e^{-7.5} 7.5^2}{2!} = 0.0156$$

Question #27**Key: E**

$X(t) - Y(t)$ is Brownian motion with initial value -2 and $\sigma^2 = 0.5 + 1 = 1.5$

By formula 10.6, the probability that $X(t) - Y(t) \geq 0$ at some time between 0 and 5 is

$$\begin{aligned} &= 2 \times \text{Prob}[X(5) - Y(5) \geq 0] \\ &= 2 \times \left[1 - \Phi\left(\frac{2}{\sqrt{5(1.5)}}\right) \right] = 0.4652 \end{aligned}$$

The 2 in the numerator of $\frac{2}{\sqrt{5(1.5)}}$ comes from $X(0) - Y(0) = -2$; the process needs to move 2 to reach $X(t) - Y(t) \geq 0$.

Question #28**Key: B**

$$\begin{aligned}
{}_2|q_{\overline{80:84}} &= {}_2|q_{80} + {}_2|q_{84} - {}_2|q_{80:84} \\
&= 0.5 \times 0.4 \times (1 - 0.6) + 0.2 \times 0.15 \times (1 - 0.1) \\
&= 0.10136
\end{aligned}$$

Using new p_{82} value of 0.3

$$\begin{aligned}
&0.5 \times 0.4 \times (1 - 0.3) + 0.2 \times 0.15 \times (1 - 0.1) \\
&= 0.16118
\end{aligned}$$

$$\text{Change} = 0.16118 - 0.10136 = 0.06$$

Alternatively,

$$\begin{aligned}
{}_2p_{80} &= 0.5 \times 0.4 = 0.20 \\
{}_3p_{80} &= {}_2p_{80} \times 0.6 = 0.12 \\
{}_2p_{84} &= 0.20 \times 0.15 = 0.03 \\
{}_3p_{84} &= {}_2p_{84} \times 0.10 = 0.003 \\
{}_2p_{\overline{80:84}} &= {}_2p_{80} + {}_2p_{84} - {}_2p_{80} {}_2p_{84} \text{ since independent} \\
&= 0.20 + 0.03 - (0.20)(0.03) = 0.224 \\
{}_3p_{\overline{80:84}} &= {}_3p_{80} + {}_3p_{84} - {}_3p_{80} {}_3p_{84} \\
&= 0.12 + 0.003 - (0.12)(0.003) = 0.12264 \\
{}_2|q_{\overline{80:84}} &= {}_2p_{\overline{80:84}} - {}_3p_{\overline{80:84}} \\
&= 0.224 - 0.12264 = 0.10136
\end{aligned}$$

Revised

$$\begin{aligned}
{}_3p_{80} &= 0.20 \times 0.30 = 0.06 \\
{}_3p_{\overline{80:84}} &= 0.06 + 0.003 - (0.06)(0.003) \\
&= 0.06282 \\
{}_2|q_{\overline{80:84}} &= 0.224 - 0.06282 = 0.16118 \\
\text{change} &= 0.16118 - 0.10136 = 0.06
\end{aligned}$$

Question #29**Key: B**

$$e_x = p_x + p_x e_{x+1} \Rightarrow p_x = \frac{e_x}{1 + e_{x+1}} = \frac{8.83}{9.29} = 0.95048$$

$$\ddot{a}_x = 1 + v p_x + v^2 {}_2p_x + \dots$$

$$\ddot{a}_{\overline{x:2}|} = 1 + v + v^2 {}_2p_x + \dots$$

$$\ddot{a}_{\overline{x:2}|} - \ddot{a}_x = v q_x = 5.6459 - 5.60 = 0.0459$$

$$v(1 - 0.95048) = 0.0459$$

$$v = 0.9269$$

$$i = \frac{1}{v} - 1 = 0.0789$$

Question #30**Key: A**

Let π be the benefit premium

Let ${}_kV$ denote the benefit reserve at the end of year k .

$$\text{For any } n, ({}_nV + \pi)(1+i) = (q_{25+n} \times {}_{n+1}V + p_{25+n} \times {}_{n+1}V) \\ = {}_{n+1}V$$

$$\text{Thus } {}_1V = ({}_0V + \pi)(1+i)$$

$${}_2V = ({}_1V + \pi)(1+i) = (\pi(1+i) + \pi)(1+i) = \pi \ddot{s}_{\overline{2}|}$$

$${}_3V = ({}_2V + \pi)(1+i) = (\pi \ddot{s}_{\overline{2}|} + \pi)(1+i) = \pi \ddot{s}_{\overline{3}|}$$

By induction (proof omitted)

$${}_nV = \pi \ddot{s}_{\overline{n}|}$$

For $n = 35$, ${}_{35}V = \ddot{a}_{60}$ (actuarial present value of future benefits; there are no future premiums)

$$\ddot{a}_{60} = \pi \ddot{s}_{\overline{35}|}$$

$$\pi = \frac{\ddot{a}_{60}}{\ddot{s}_{\overline{35}|}}$$

$$\text{For } n = 20, {}_{20}V = \pi \ddot{s}_{\overline{20}|}$$

$$= \left(\frac{\ddot{a}_{60}}{\ddot{s}_{\overline{35}|}} \right) \ddot{s}_{\overline{20}|}$$

Alternatively, as above

$$({}_nV + \pi)(1+i) = {}_{n+1}V$$

Write those equations, for $n = 0$ to $n = 34$

$$0: ({}_0V + \pi)(1+i) = {}_1V$$

$$1: ({}_1V + \pi)(1+i) = {}_2V$$

$$2: ({}_2V + \pi)(1+i) = {}_3V$$

\vdots

$$34: ({}_{34}V + \pi)(1+i) = {}_{35}V$$

Multiply equation k by $(1+i)^{34-k}$ and sum the results:

$$({}_0V + \pi)(1+i)^{35} + ({}_1V + \pi)(1+i)^{34} + ({}_2V + \pi)(1+i)^{33} + \cdots + ({}_{34}V + \pi)(1+i) = \\ {}_1V(1+i)^{34} + {}_2V(1+i)^{33} + {}_3V(1+i)^{32} + \cdots + {}_{34}V(1+i) + {}_{35}V$$

For $k = 1, 2, \dots, 34$, the ${}_kV(1+i)^{35-k}$ terms in both sides cancel, leaving

$${}_0V(1+i)^{35} + \pi \left[(1+i)^{35} + (1+i)^{34} + \dots + (1+i) \right] = {}_{35}V$$

Since ${}_0V = 0$

$$\begin{aligned} \pi \ddot{s}_{\overline{35}|} &= {}_{35}V \\ &= \ddot{a}_{60} \end{aligned}$$

(see above for remainder of solution)

This technique, for situations where the death benefit is a specified amount (here, 0) plus the benefit reserve is discussed in section 8.3 of Bowers. This specific problem is Example 8.3.1.

Question #31**Key: B**

$$\mu_{xy}^-(t) = \frac{{}_t q_y {}_t p_x \mu(x+t) + {}_t q_x {}_t p_y \mu(y+t)}{{}_t q_x \times {}_t p_y + {}_t p_x \times {}_t q_y + {}_t p_x \times {}_t p_y}$$

For $(x) = (y) = (50)$

$$\mu_{50:50}^-(10.5) = \frac{({}_{10.5} q_{50})({}_{10} p_{50})q_{60} \cdot 2}{({}_{10.5} q_{50})({}_{10.5} p_{50}) \cdot 2 + ({}_{10.5} p_{50})^2} = \frac{(0.09152)(0.91478)(0.01376)(2)}{(0.09152)(0.90848)(2) + (0.90848)^2} = 0.0023$$

where

$${}_{10.5} p_{50} = \frac{\frac{1}{2}(l_{60} + l_{61})}{l_{50}} = \frac{\frac{1}{2}(8,188,074 + 8,075,403)}{8,950,901} = 0.90848$$

$${}_{10.5} q_{50} = 1 - {}_{10.5} p_{50} = 0.09152$$

$${}_{10} p_{50} = \frac{8,188,074}{8,950,901} = 0.91478$$

$${}_{10.5} p_{50} \mu(50+10.5) = ({}_{10} p_{50})q_{60} \quad \text{since UDD}$$

Alternatively, $({}_{10+t}) p_{50} = {}_{10} p_{50} {}_t p_{60}$

$$({}_{10+t}) p_{50:50} = ({}_{10} p_{50})^2 ({}_t p_{60})^2$$

$$\begin{aligned}({}_{10+t}) p_{50:50}^- &= 2 {}_{10} p_{50} {}_t p_{60} - ({}_{10} p_{50})^2 ({}_t p_{60})^2 \\ &= 2 {}_{10} p_{50} (1 - {}_t q_{60}) - ({}_{10} p_{50})^2 (1 - {}_t q_{60})^2 \quad \text{since UDD}\end{aligned}$$

$$\text{Derivative} = -2 {}_{10} p_{50} q_{60} + 2 ({}_{10} p_{50})^2 (1 - {}_t q_{60}) q_{60}$$

Derivative at $10+t=10.5$ is

$$-2(0.91478)(0.01376) + (0.91478)^2 (1 - (0.5)(0.01376))(0.01376) = -0.0023$$

$$\begin{aligned}{}_{10.5} p_{50:50}^- &= 2 {}_{10.5} p_{50} - ({}_{10.5} p_{50})^2 \\ &= 2(0.90848) - (0.90848)^2 \\ &= 0.99162\end{aligned}$$

$$\mu \text{ (for any sort of lifetime)} = \frac{-\frac{dp}{dt}}{p} = \frac{-(-0.0023)}{0.99162} = 0.0023$$

Question #32**Key: E**

$$\begin{aligned}
E(W) &= \frac{1}{4} \int_0^4 \sum_{i=0}^{\infty} 2^i \Pr(N=i|\lambda) d\lambda \quad \left[\frac{1}{4} \text{ is the density of } \lambda \text{ on } [0, 4]. \right] \\
&= \frac{1}{4} \int_0^4 P(2|\lambda) d\lambda \quad [\text{see note}] \\
&= \frac{1}{4} \int_0^4 e^{\lambda(2-1)} d\lambda \quad [\text{using formula from tables for the pgf of the Poisson}] \\
&= \frac{1}{4} e^{\lambda} \Big|_0^4 = \frac{1}{4} (e^4 - 1) \\
&= 13.4
\end{aligned}$$

Note: the probability generating function (pgf) is $P(Z) = \sum_{k=0}^{\infty} p_k Z^k$ so the integrand is $P(2)$, or in this case $P(2|\lambda)$ since λ is not known.

Alternatively,

$$\begin{aligned}
E(W) &= \frac{1}{4} \int_0^4 \sum_{i=0}^{\infty} 2^i \Pr(N=i|\lambda) d\lambda \\
&= \frac{1}{4} \int_0^4 \sum_{i=0}^{\infty} \frac{2^i e^{-\lambda} \lambda^i}{i!} d\lambda \\
&= \frac{1}{4} \int_0^4 \sum_{i=0}^{\infty} \frac{e^{-\lambda} (2\lambda)^i}{i!} d\lambda
\end{aligned}$$

We know $\sum_{i=0}^{\infty} \frac{e^{-2\lambda} (2\lambda)^i}{i!} = 1$ since $\frac{e^{-2\lambda} (2\lambda)^i}{i!}$ is $f(i)$ for a Poisson with mean 2λ

$$\text{so } \sum_{i=0}^{\infty} \frac{e^{-\lambda} (2\lambda)^i}{i!} = \frac{e^{-\lambda}}{e^{-2\lambda}} = e^{\lambda}$$

$$\text{Thus } E(W) = \frac{1}{4} \int_0^4 e^{\lambda} d\lambda$$

$$\begin{aligned}
&= \frac{1}{4} e^{\lambda} \Big|_0^4 = \frac{1}{4} (e^4 - 1) \\
&= 13.4
\end{aligned}$$

Question #33**Key: A or E**

$$E(S) = \lambda E[X] = 2/3(1/4 + 2/4 + 3/2) = 2/3 \times 9/4 = 3/2$$

$$\text{Var}(S) = \lambda E[X^2] = 2/3(1/4 + 4/4 + 9/2) = 23/6$$

So cumulative premium to time 2 is $2\left(3/2 + 1.8\sqrt{23/6}\right) = 10$, where the expression in parentheses is the annual premium

Times between claims are determined by $-1/\lambda \log u$ and are 0.43, 0.77, 1.37, 2.41

So 2 claims before time 2 (second claim is at 1.20; third is at 2.57)

Sizes are 2, 3, 1, 3, where only the first two matter.

So gain to the insurer is $10 - (2+3) = 5$

Note: since the problem did not specify that we wanted the gain or loss from the insurer's viewpoint, we gave credit to answer A; a loss of 5 from the insured's viewpoint.

Question #34**Key: C**

To get number of claims, set up cdf for Poisson:

x	$f(x)$	$F(x)$
0	0.135	0.135
1	0.271	0.406
2	0.271	0.677
3	0.180	0.857

0.80 simulates 3 claims.

$$F(x) = 1 - (500/(x + 500))^2 = u, \text{ so } x = (1 - u)^{-1/2} 500 - 500$$

0.6 simulates 290.57

0.25 simulates 77.35

0.7 simulates 412.87

So total losses equals 780.79

$$\text{Insurer pays } (0.80)(750) + (780.79 - 750) = 631$$

Question #35**Key: D**

$$\mu_x^{(\tau)}(t) = \mu_x^{(1)}(t) + \mu_x^{(2)}(t) = 0.01 + 2.29 = 2.30$$

$$P = P \int_0^2 v^t {}_t p_x^{(\tau)} \mu_x^{(2)}(t) dt + 50,000 \int_0^2 v^t {}_t p_x^{(\tau)} \mu_x^{(1)}(t) dt + 50,000 \int_2^\infty v^t {}_t p_x^{(\tau)} \mu_x^{(\tau)}(t) dt$$

$$P = P \int_0^2 e^{-0.1t} e^{-2.3t} \times 2.29 dt + 50,000 \int_0^2 e^{-0.1t} e^{-2.3t} \times 0.01 dt + 50,000 \int_2^\infty e^{-0.1t} e^{-2.3t} \times 2.3 dt$$

$$P \left[1 - 2.29 \times \frac{1 - e^{-2(2.4)}}{2.4} \right] = 50000 \left[0.01 \times \frac{1 - e^{-2(2.4)}}{2.4} + 2.3 \times \frac{e^{-2(2.4)}}{2.4} \right]$$

$$P = 11,194$$

Question #36**Key: D**

$$\mu^{(accid)} = 0.001$$

$$\mu^{(total)} = 0.01$$

$$\mu^{(other)} = 0.01 - 0.001 = 0.009$$

$$\text{Actuarial present value} = \int_0^\infty 500,000 e^{-0.05t} e^{-0.01t} (0.009) dt$$

$$+ 10 \int_0^\infty 50,000 e^{0.04t} e^{-0.05t} e^{-0.01t} (0.001) dt$$

$$= 500,000 \left[\frac{0.009}{0.06} + \frac{0.001}{0.02} \right] = 100,000$$

Question #37**Key: B**

$$\text{Variance} = v^{30}_{15} p_x {}_{15}q_x \quad \text{Expected value} = v^{15}_{15} p_x$$

$$v^{30}_{15} p_x {}_{15}q_x = 0.065 \quad v^{15}_{15} p_x$$

$$v^{15}_{15} q_x = 0.065 \Rightarrow {}_{15}q_x = 0.3157$$

Since μ is constant

$${}_{15}q_x = (1 - (p_x)^{15})$$

$$(p_x)^{15} = 0.6843$$

$$p_x = 0.975$$

$$q_x = 0.025$$

Question #38**Key: E**

$$(1) \quad {}_{11}V^A = \left({}_{10}V^A + 0 \right) \frac{(1+i)}{p_{x+10}} - \frac{q_{x+10}}{p_{x+10}} \times 1000$$

$$(2) \quad {}_{11}V^B = \left({}_{10}V^B + \pi^B \right) \frac{(1+i)}{p_{x+10}} - \frac{q_{x+10}}{p_{x+10}} \times 1000$$

$$(1) - (2) \quad {}_{11}V^A - {}_{11}V^B = \left({}_{10}V^A - {}_{10}V^B - \pi^B \right) \frac{(1+i)}{p_{x+10}}$$

$$= (101.35 - 8.36) \frac{(1.06)}{1 - 0.004}$$

$$= 98.97$$

Question #39**Key: A**

$$\begin{aligned}\text{Actuarial present value Benefits} &= \frac{(0.8)(0.1)(10,000)}{1.06^2} + \frac{(0.8)(0.9)(0.097)(9,000)}{1.06^3} \\ &= 1,239.75\end{aligned}$$

$$\begin{aligned}1,239.75 &= P \left(1 + \frac{(0.8)}{1.06} + \frac{(0.8)(0.9)}{1.06^2} \right) \\ &= P(2.3955) \\ P &= 517.53 \Rightarrow 518\end{aligned}$$

Question #40**Key: C**

<u>Event</u>	<u>Prob</u>	<u>Present Value</u>
$x = 0$	(0.05)	15
$x = 1$	$(0.95)(0.10) = 0.095$	$15 + 20/1.06 = 33.87$
$x \geq 2$	$(0.95)(0.90) = 0.855$	$15 + 20/1.06 + 25/1.06^2 = 56.12$

$$E[X] = (0.05)(15) + (0.095)(33.87) + (0.855)(56.12) = 51.95$$

$$E[X^2] = (0.05)(15)^2 + (0.095)(33.87)^2 + (0.855)(56.12)^2 = 2813.01$$

$$\text{Var}[X] = E(X^2) - E(X)^2 = 2813.01 - (51.95)^2 = 114.2$$