

# CS 237: Probability in Computing

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## Lecture 12:

- Continuous Distributions
- Exponential
- Uniform Distribution
- Normal Distribution (motivation)

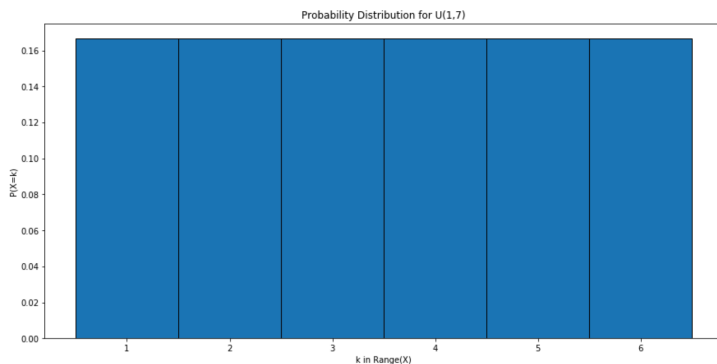
# Review: Cumulative Distribution Functions

The **Cumulative Distribution Function (CDF)** for a random variable  $X$  shows what happens when we keep track of the sum of the probability distribution from left to right over its range:

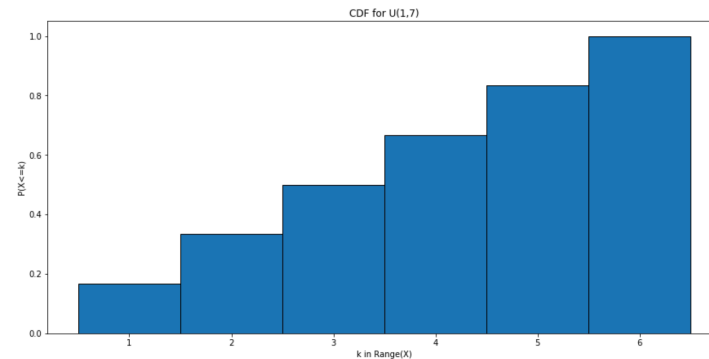
$$F_X(k) = P(X \leq k) = \sum_{a \leq k} f_X(a)$$

Example:  $X$  = “The number of dots showing on a thrown die”

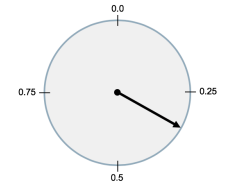
Probability Distribution Function  $f_X$



Cumulative Distribution Function  $F_X$



# Continuous Distributions



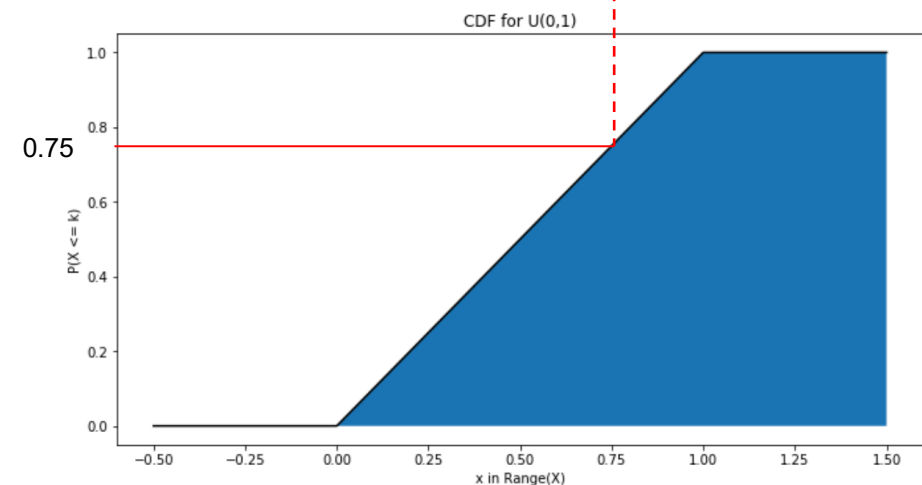
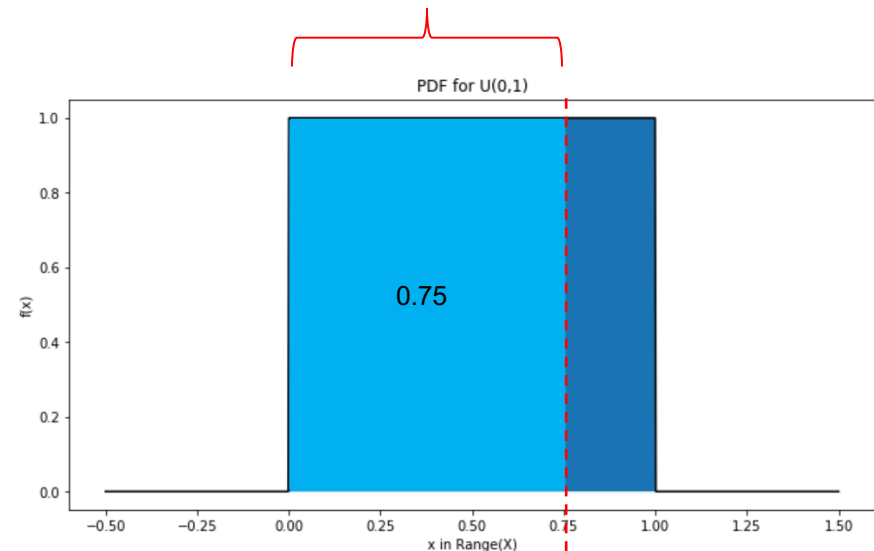
Now recall that the **ONLY** way to deal with continuous probability is to use intervals and to use area (or extent) for the probability. Hence we will calculate probabilities of intervals using the CDF:

$$f(x) = \begin{cases} 1 & \text{if } 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

$$F(a) = \int_0^a 1 \, dx = x \Big|_0^a = a$$

$$F(a) = \begin{cases} 0 & \text{if } a < 0 \\ a & \text{if } 0 \leq a \leq 1 \\ 1 & \text{if } a > 1 \end{cases}$$

$$P(X < 0.75) = F(0.75) = 0.75$$



# Continuous Distributions

## Discrete Random Variables

$$F_X(b) = P(X \leq b) =_{\text{def}} \sum_{x \leq b} f(x)$$

$$P(a \leq X \leq b) =_{\text{def}} \sum_{a \leq x \leq b} f(x)$$

$$E(X) = \sum_{x \in R_X} x \cdot f(x)$$

## Continuous Random Variables

$$F_X(b) = P(X < b) =_{\text{def}} \int_{-\infty}^b f(x) dx$$

$$P(a < X < b) =_{\text{def}} \int_a^b f(x) dx$$

$$E(X) = \int_{-\infty}^{\infty} x \cdot f(x) dx$$

## Same for both Discrete and Continuous Random Variables

$$\text{Var}(X) =_{\text{def}} E[(X - \mu_X)^2]$$

$$\text{Var}(X) = E(X^2) - (\mu_X)^2$$

$$\sigma_X =_{\text{def}} \sqrt{\text{Var}(X)}$$

$$X^* =_{\text{def}} \frac{X - \mu_X}{\sigma_X}$$

# Uniform Distribution

The simplest continuous distribution is similar to the spinner, but with arbitrary endpoints:

If  $X$  = “a random real number uniformly chosen from the interval  $[a..b]$ ”

then  $X$  is a uniform random variable from  $a$  to  $b$ ,  
denoted

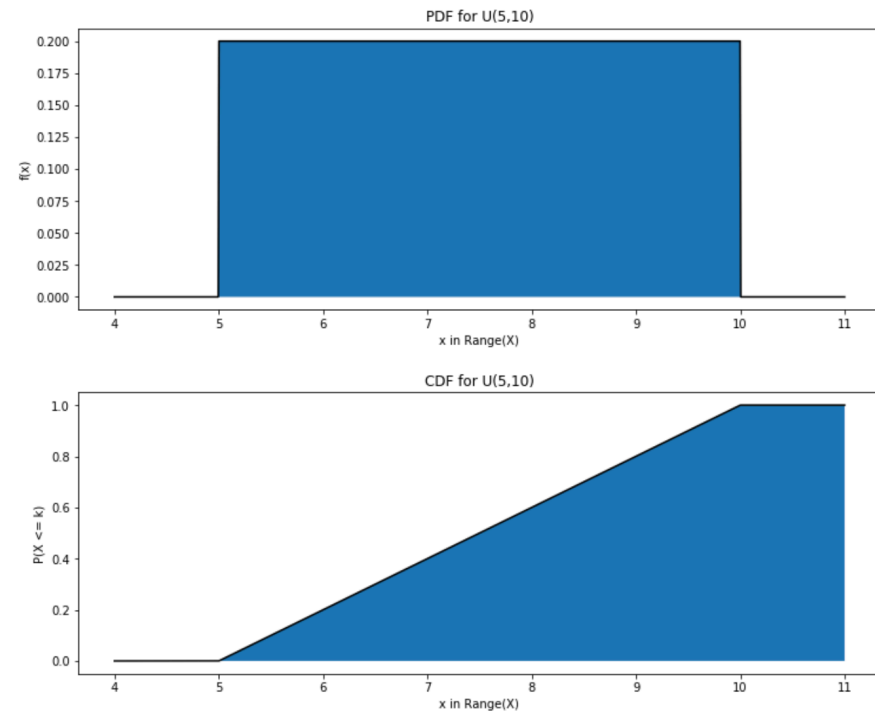
$$X \sim U(5, 10)$$

$$X \sim U(a, b)$$

and where

$$f_X(x) = \begin{cases} \frac{1}{b-a} & \text{if } a \leq x \leq b \\ 0 & \text{otherwise} \end{cases}$$

$$F_X(x) = \begin{cases} 0 & \text{if } x < a \\ \frac{x-a}{b-a} & \text{if } a \leq x \leq b \\ 1 & \text{if } x > b \end{cases}$$

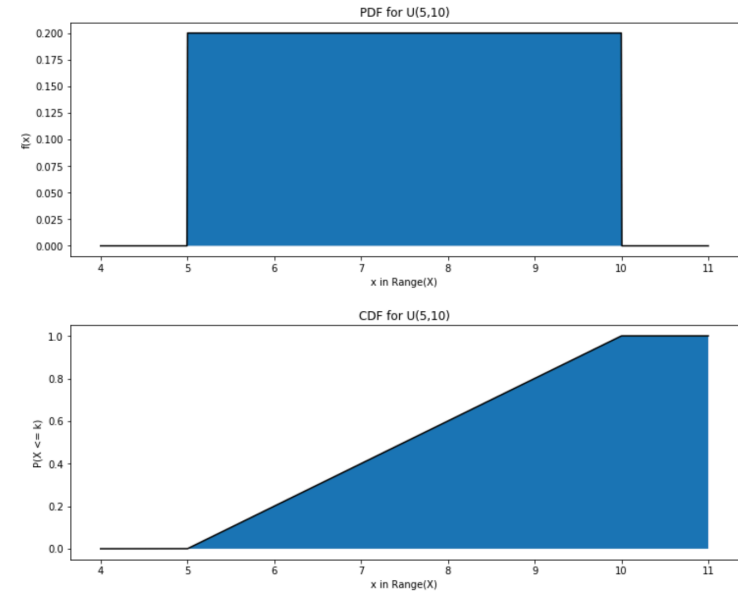


# Uniform Distribution

$$X \sim U(a, b)$$

$$f_X(x) = \begin{cases} \frac{1}{b-a} & \text{if } a \leq x \leq b \\ 0 & \text{otherwise} \end{cases}$$

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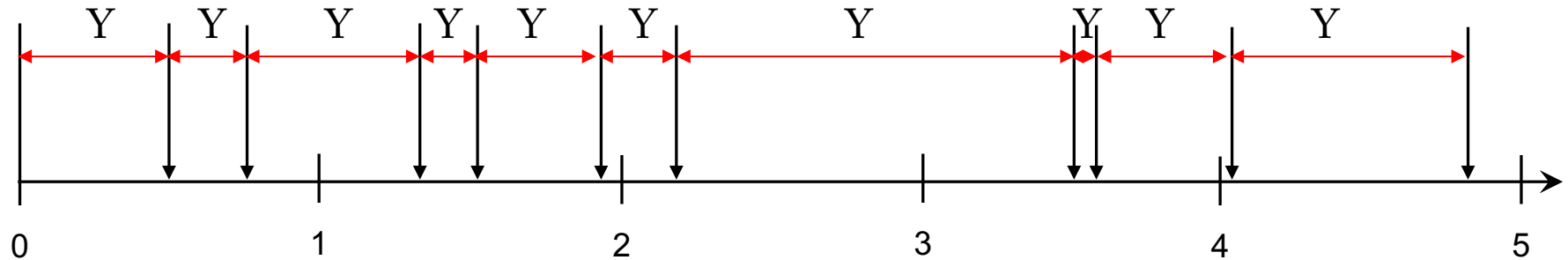


$$E(X) = \int_a^b x \cdot \frac{1}{b-a} dx = \frac{1}{b-a} \cdot \frac{x^2}{2} \Big|_a^b = \frac{b^2 - a^2}{2(b-a)} = \frac{b+a}{2}$$

$$E(X^2) = \int_a^b x^2 \cdot \frac{1}{b-a} dx = \frac{1}{b-a} \cdot \frac{x^3}{3} \Big|_a^b = \frac{b^3 - a^3}{3(b-a)} = \frac{a^2 + ab + b^2}{3}$$

$$Var(X) = E(X^2) - E(X)^2 = \frac{a^2 + ab + b^2}{3} - \frac{a^2 - 2ab + b^2}{4} = \frac{a^2 + 2ab + b^2}{12} = \frac{(b-a)^2}{12}$$

# Review: Inter-Arrival Times of a Poisson Process



What is the distribution of  $Y$ ? Since

$$\lambda = E(N[0..1])$$

and the number of arrivals in an interval is proportional to its length, that is,

$$E(N[0..2]) = 2 * E(N[0..1]), \text{ etc., then } \lambda \cdot t = E(N[0..t])$$

and so the probability that there are exactly  $n$  arrivals by time  $t$  is given by the Poisson:

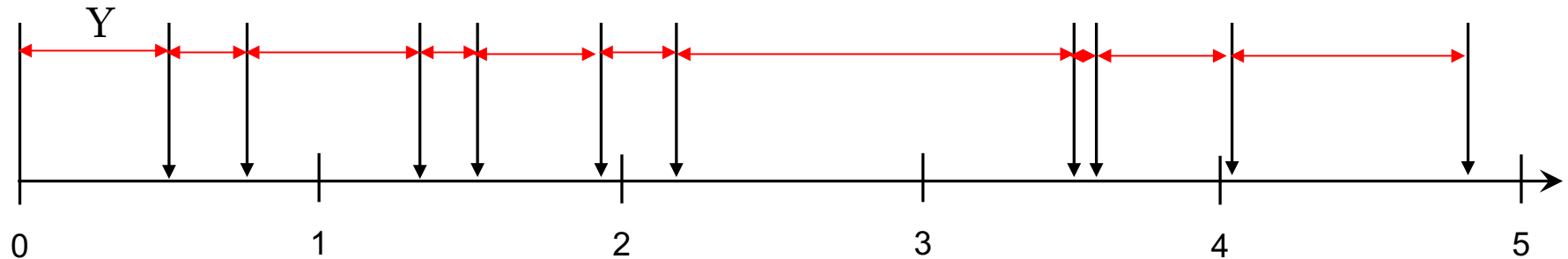
$$P(N[0..t] = n) = \frac{e^{-\lambda t} (\lambda t)^n}{n!}$$

and

$$P(Y > t) = P(N[0..t] = 0) = \frac{e^{-\lambda t} (\lambda t)^0}{0!} = e^{-\lambda t}$$

$$P(Y \leq t) = 1 - e^{-\lambda t}$$

# Distribution of Interarrival Times of a Poisson Process



What is the distribution of Y?

$$P(Y \leq t) = 1 - e^{-\lambda t}$$

Now, this is the formula of a CDF, that is,

$$F(t) = \begin{cases} 1 - e^{-\lambda t} & \text{if } t \geq 0 \\ 0 & \text{if } t < 0 \end{cases}$$

and so if we take the derivative  $0 - (-\lambda)e^{-\lambda t} = \lambda e^{-\lambda t}$

$$f(t) = F'(t) = \begin{cases} \lambda e^{-\lambda t} & \text{if } t \geq 0 \\ 0 & \text{if } t < 0 \end{cases}$$

Recall the derivative of exponential:

$$\frac{d e^{cx}}{dx} = c e^{cx}$$

and the chain rule:

$$h(x) = f(g(x))$$

$$h'(x) = f'(g(x)) \cdot g'(x)$$

we get the PDF:



# Exponential Distribution

This is called the **Exponential Distribution**, and along with the Normal, is one of the most important continuous distributions in probability and statistics.

Formally, then, if the random variable

$Y$  = “the interarrival time between events in a Poisson Process”

we say that  $Y$  is distributed according to the Exponential Distribution with rate parameter  $\lambda$ , denoted

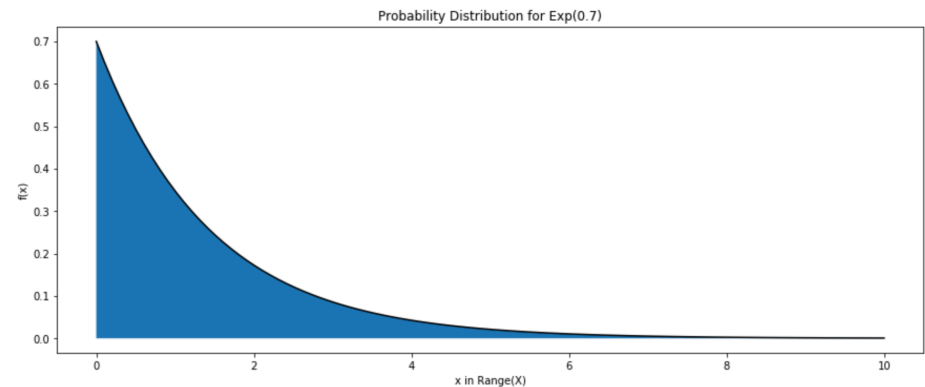
$$Y \sim \text{Exp}(\lambda)$$

if

$$f(t) = \begin{cases} \lambda e^{-\lambda t} & \text{if } t \geq 0 \\ 0 & \text{if } t < 0 \end{cases}$$

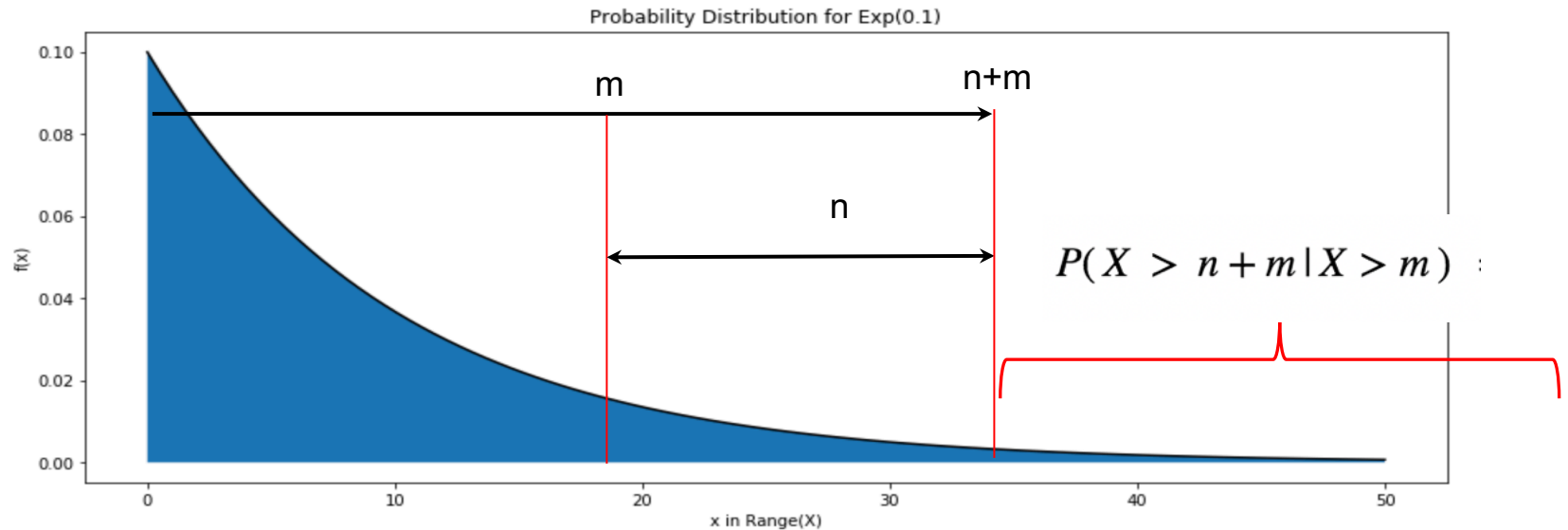
$$F(t) = \begin{cases} 1 - e^{-\lambda t} & \text{if } t \geq 0 \\ 0 & \text{if } t < 0 \end{cases}$$

and where  $E(X) = \frac{1}{\lambda}$  and  $\text{Var}(X) = \frac{1}{\lambda^2}$



$$P(X > t) = e^{-\lambda t}$$
$$P(X \leq t) = 1.0 - e^{-\lambda t}$$

# Exponential Distribution: The Memoryless Property

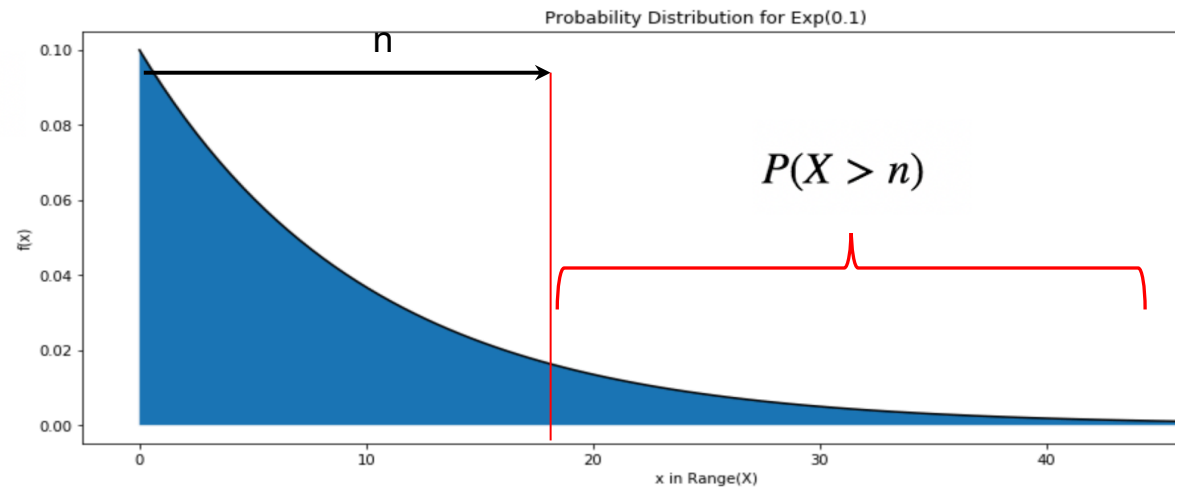


The exponential, like the geometric,  
has the memoryless property,

$$P(X > n+m | X > m) = P(X > n)$$

and the proof is the same!

$$\begin{aligned}
 P(X > n+m | X > m) &= \frac{P(X > n+m \text{ and } X > m)}{P(X > m)} \\
 &= \frac{P(X > n+m)}{P(X > m)} \\
 &= \frac{(1-p)^{(n+m)}}{(1-p)^m} \\
 &= (1-p)^n \\
 &= P(X > n)
 \end{aligned}$$



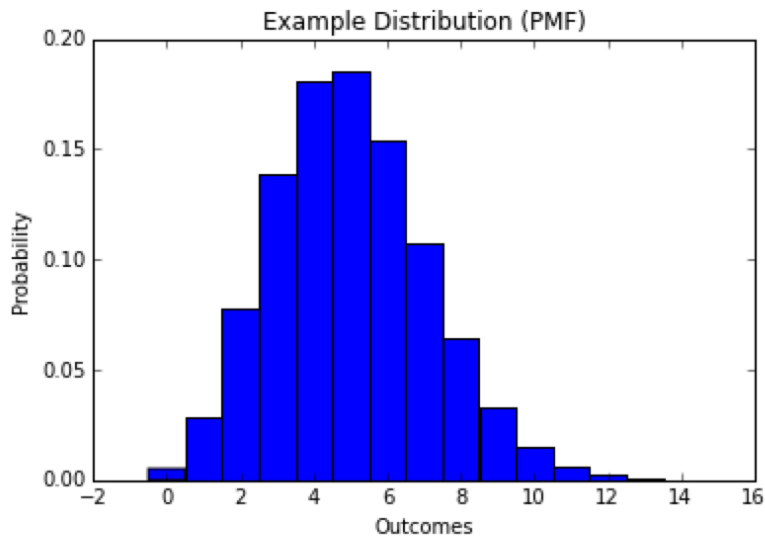
## Optional: The Waiting Time Paradox:

- <https://jakevdp.github.io/blog/2018/09/13/waiting-time-paradox/>

# Normal Distribution

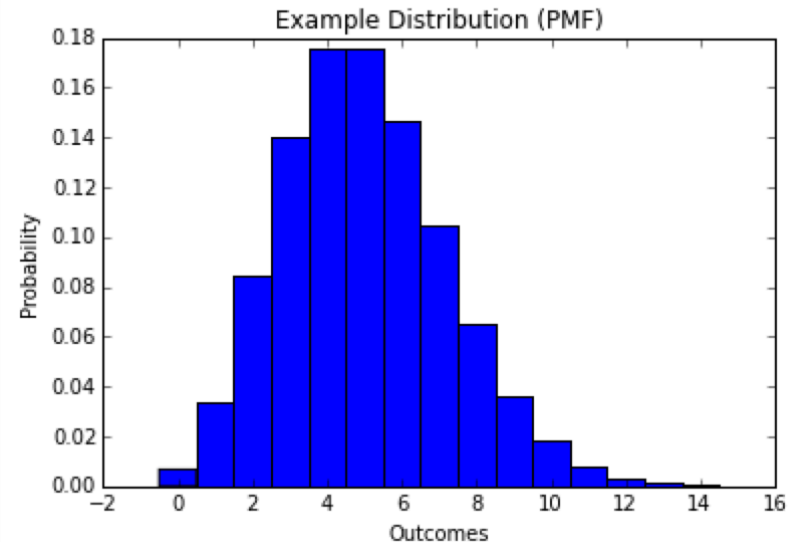
How to approximate the binomial? Under certain conditions (basically when  $p$  is small), the Poisson can be used to approximate the binomial (we will explore this in the next homework).

In [908]: `drawDistribution(B[0][:15],B[1][:15])`



```
E[X]      = 4.9989
Var(X)    = 4.4921
stdev(X)  = 2.1195
Rng(X):   [0, 1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12, 13, 14]
P:        [0.0052, 0.0286, 0.0779, 0.1386, 0.1809, 0.1849, 0.1541, 0.1061, 0.0611, 0.0306, 0.0122, 0.0042, 0.0013, 0.0003, 0.0001]
```

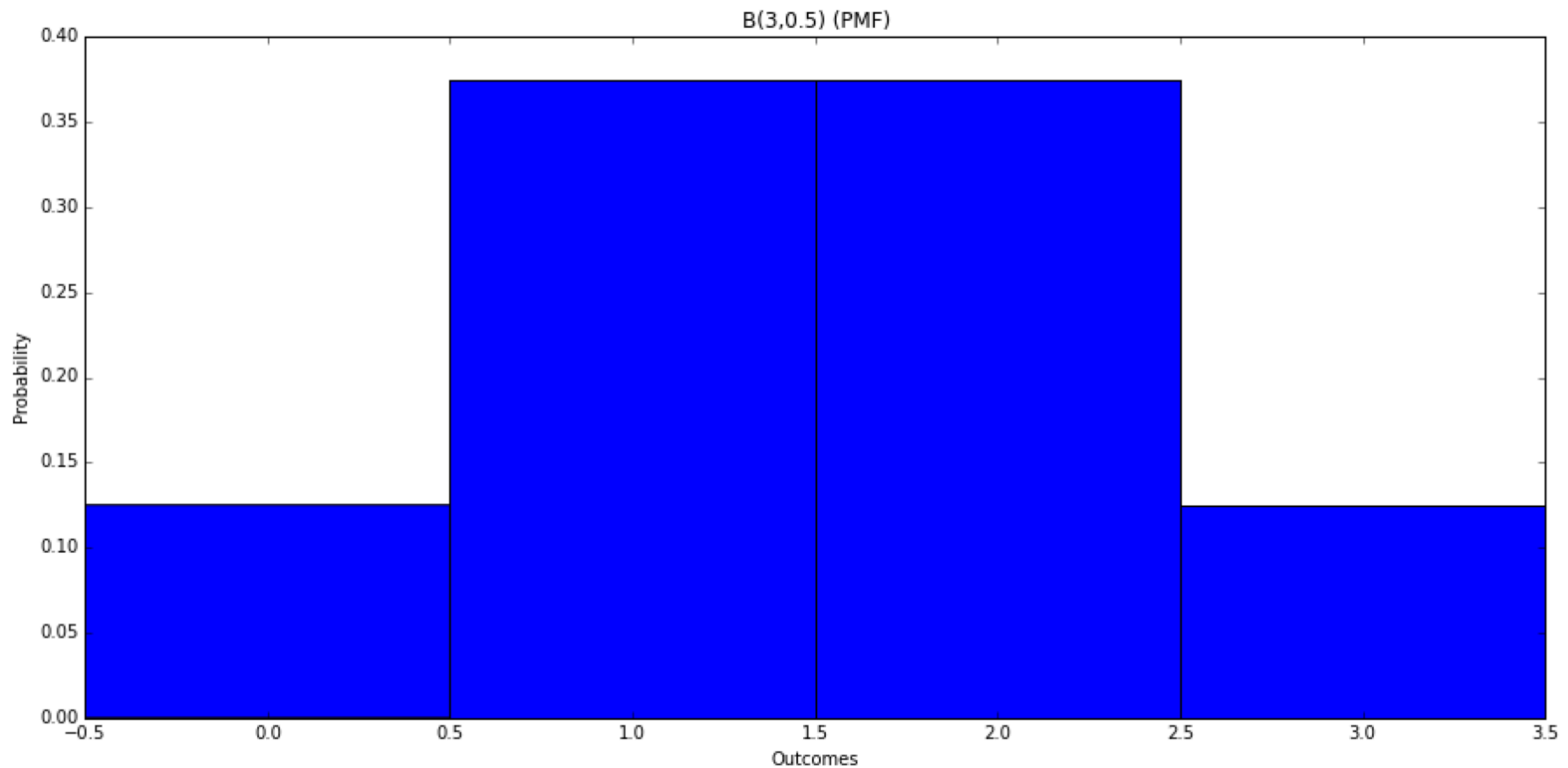
In [909]: `drawDistribution(P[0][:15],P[1][:15])`



```
E[X]      = 4.9965
Var(X)    = 4.9753
stdev(X)  = 2.2305
Rng(X):   [0, 1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12, 13, 14]
P:        [0.0067, 0.0337, 0.0842, 0.1404, 0.1755, 0.1755, 0.1462, 0.1044, 0.0651, 0.0351, 0.0194, 0.0104, 0.0054, 0.0027, 0.0013]
```

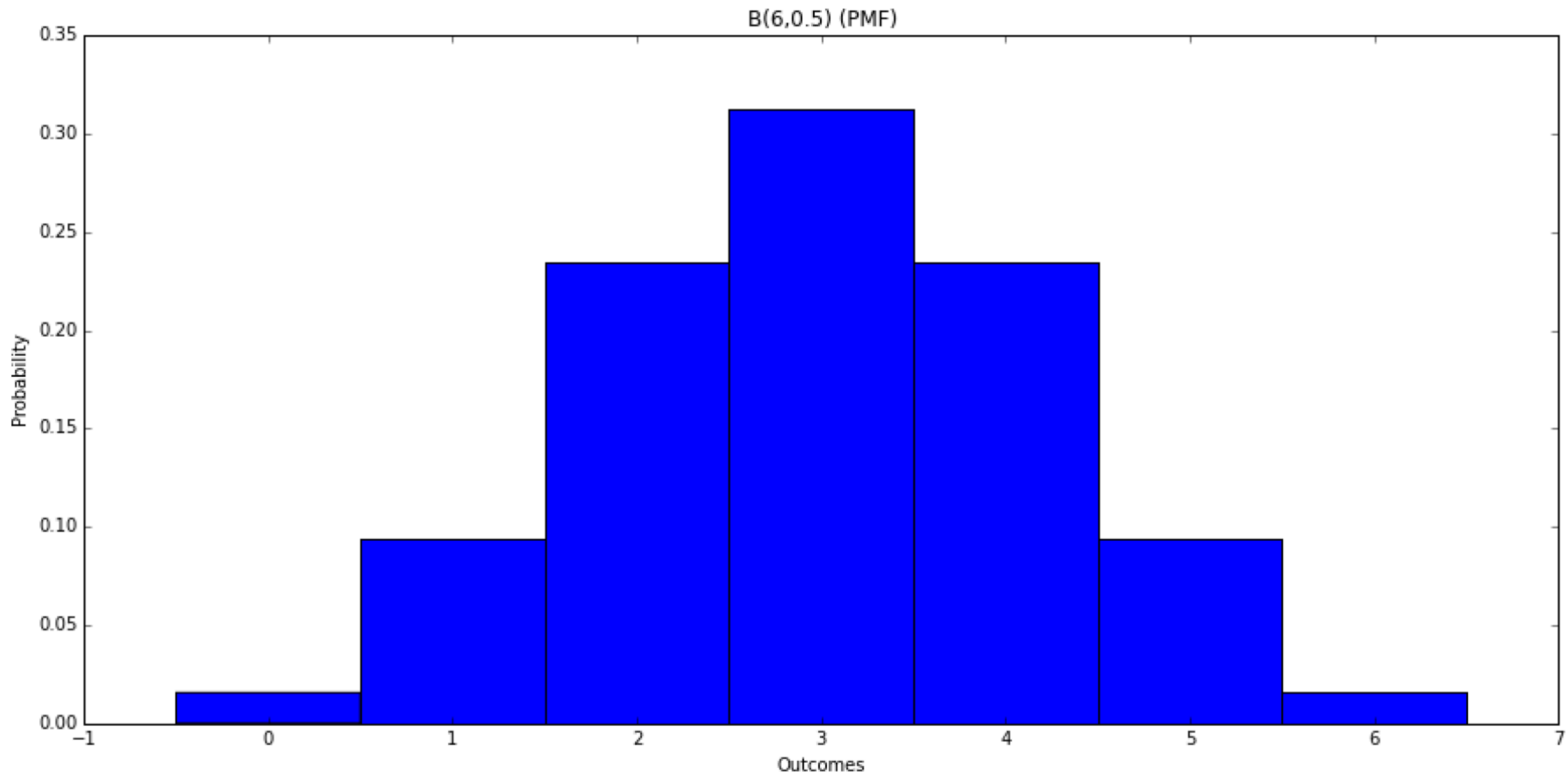
# Normal Distribution

How to approximate the binomial when  $p$  is  $1/2$ ? When we observe the characteristic shape of the Binomial Distribution  $B(N, 0.5)$  as  $N$  approaches Infinity, we see something interesting:



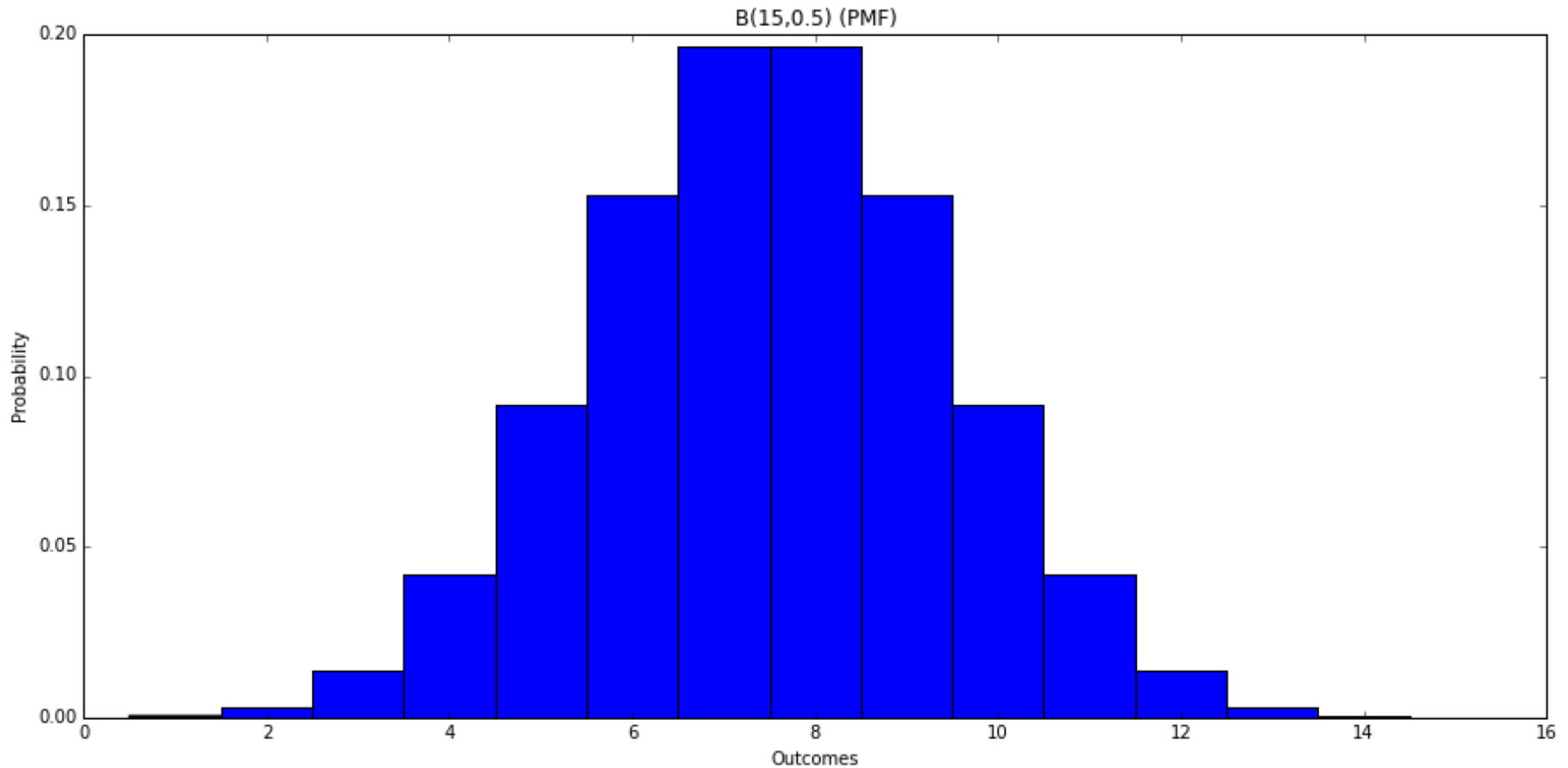
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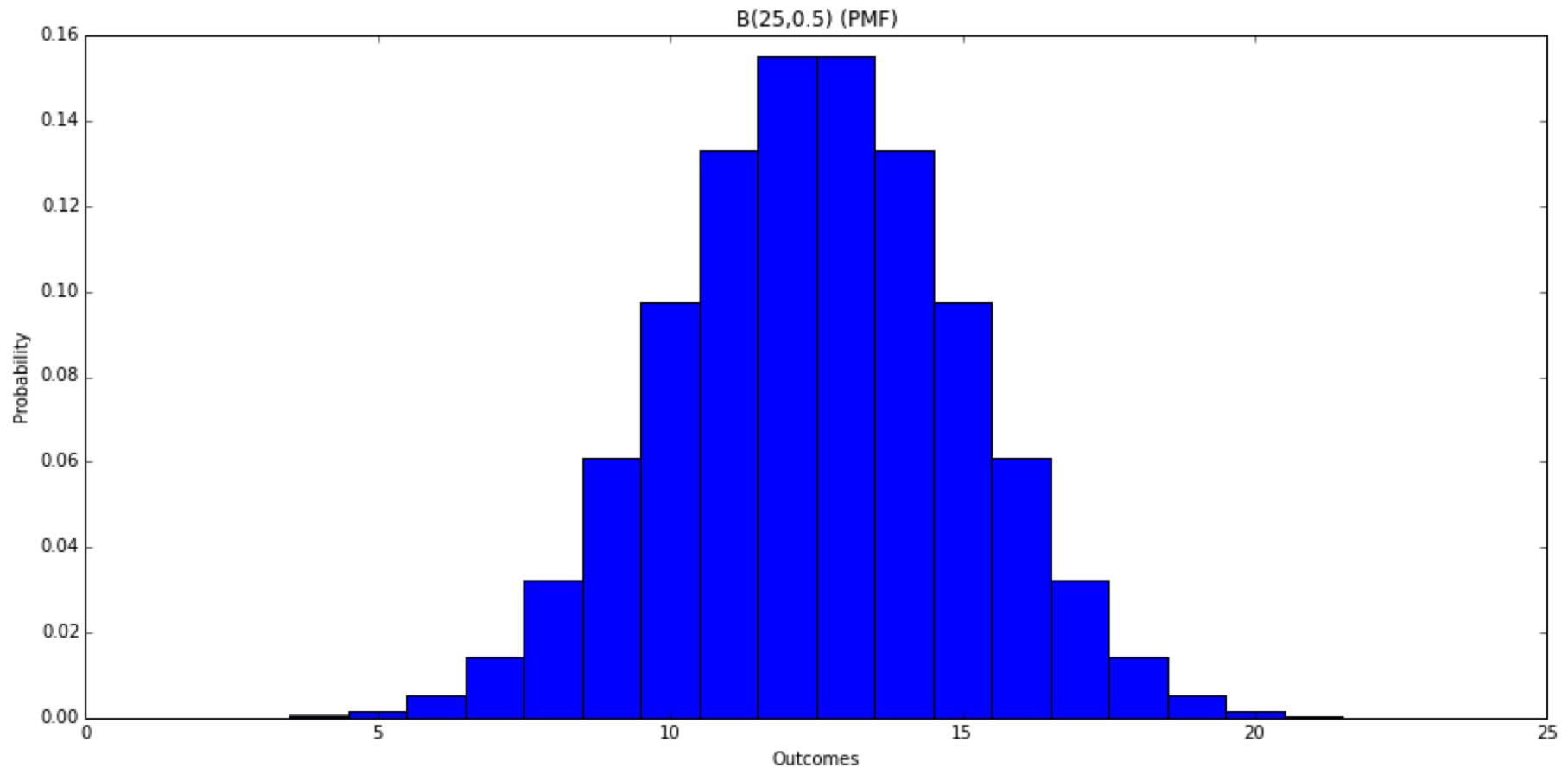
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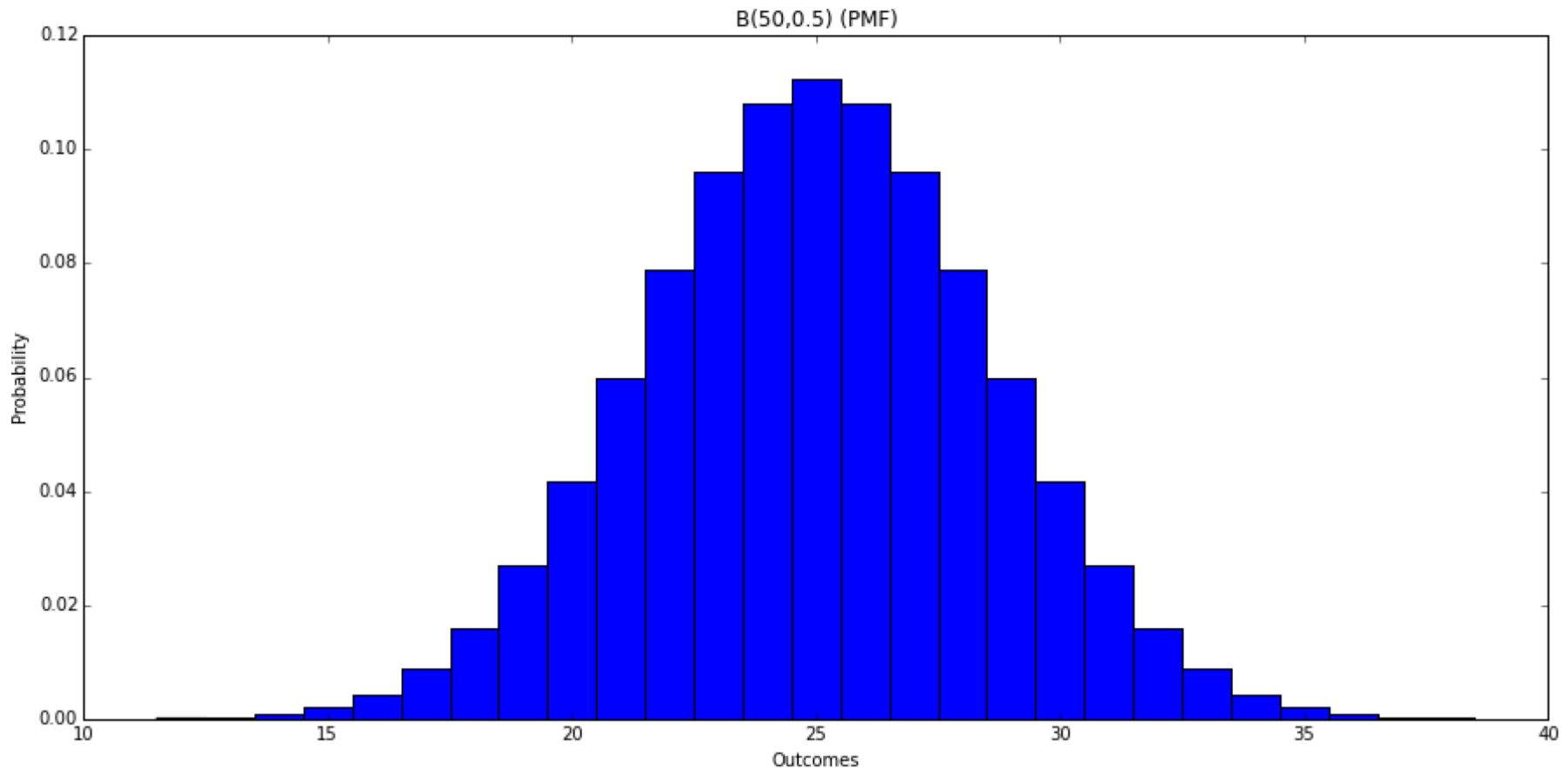
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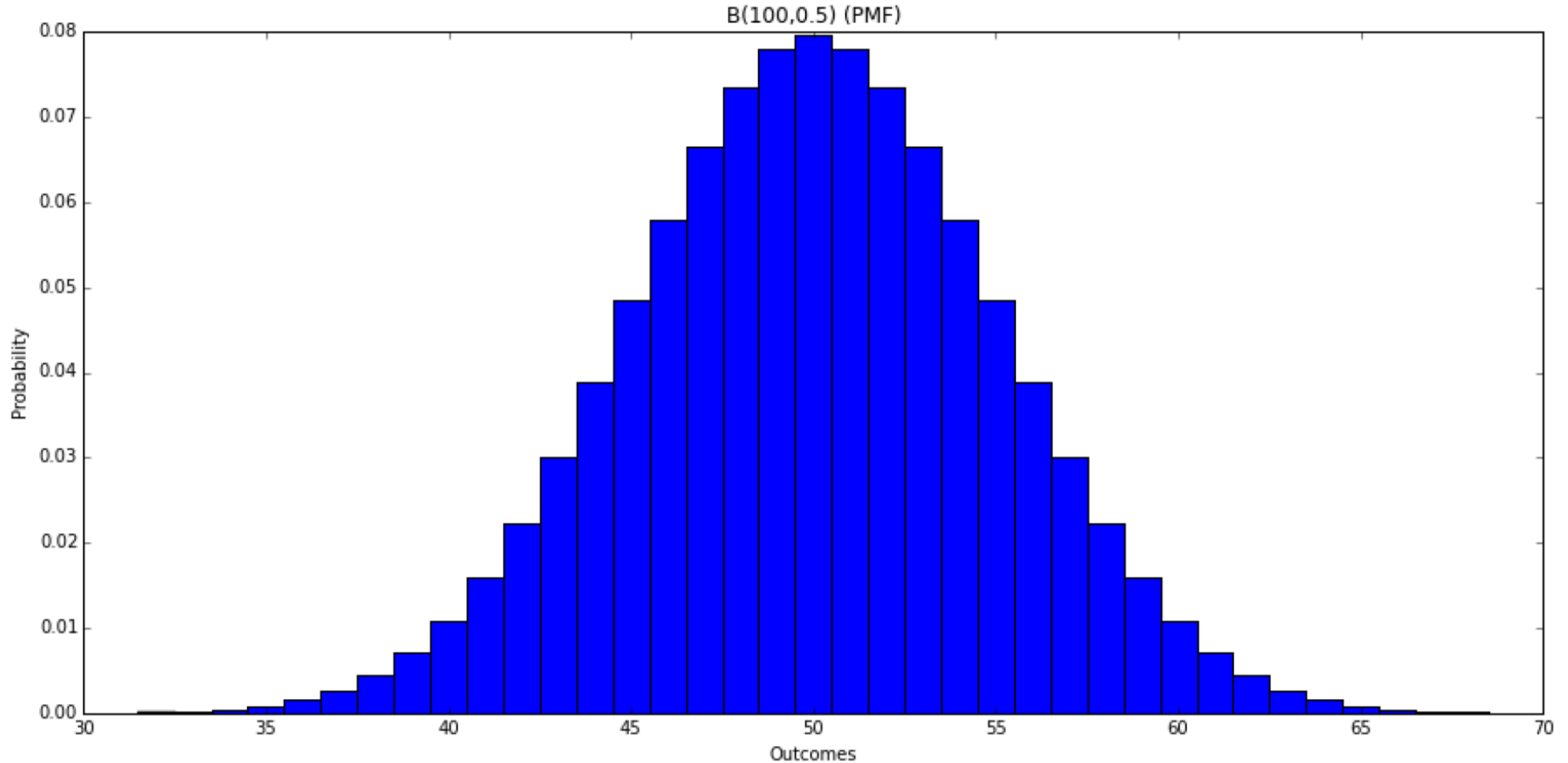
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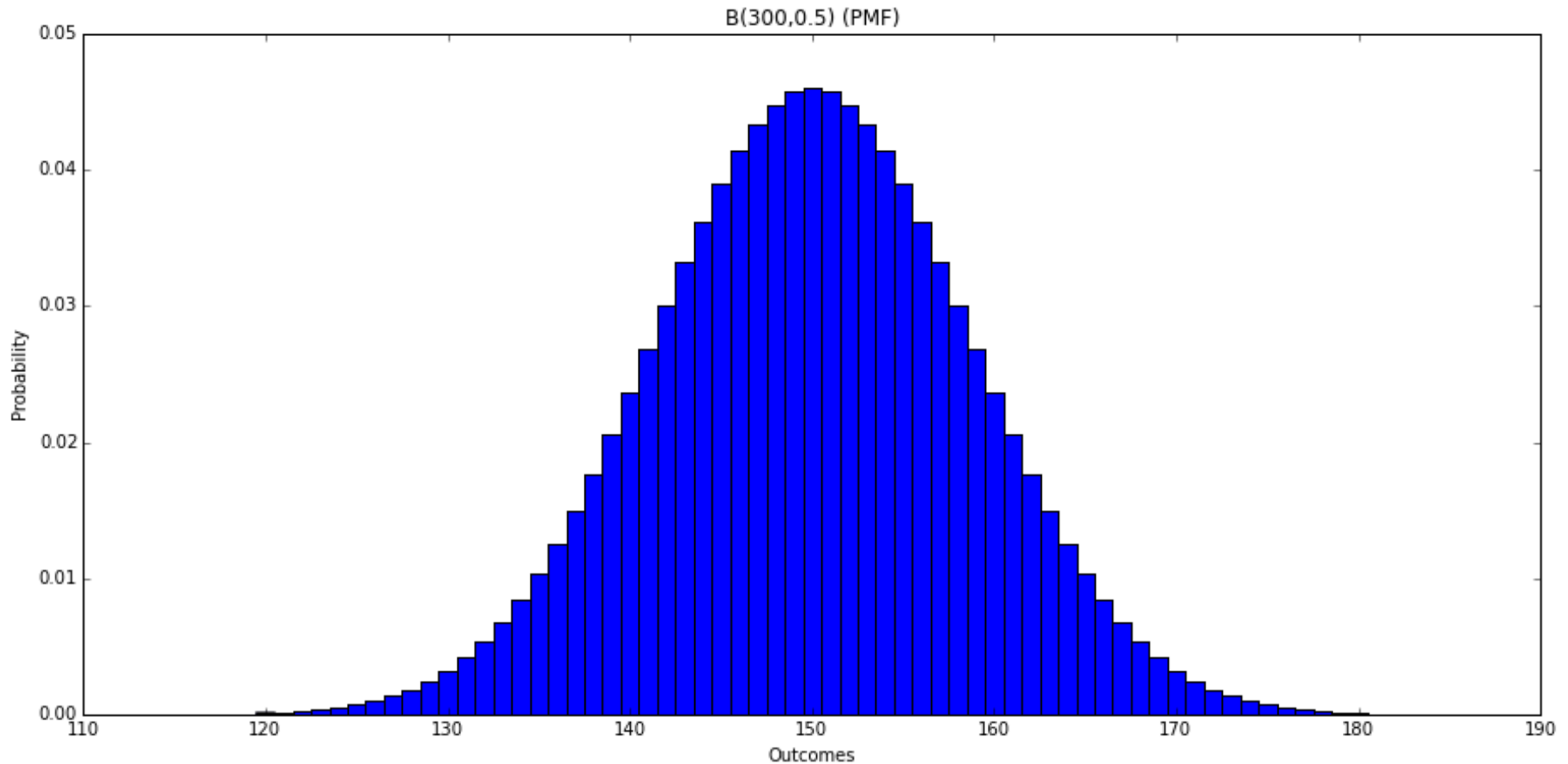
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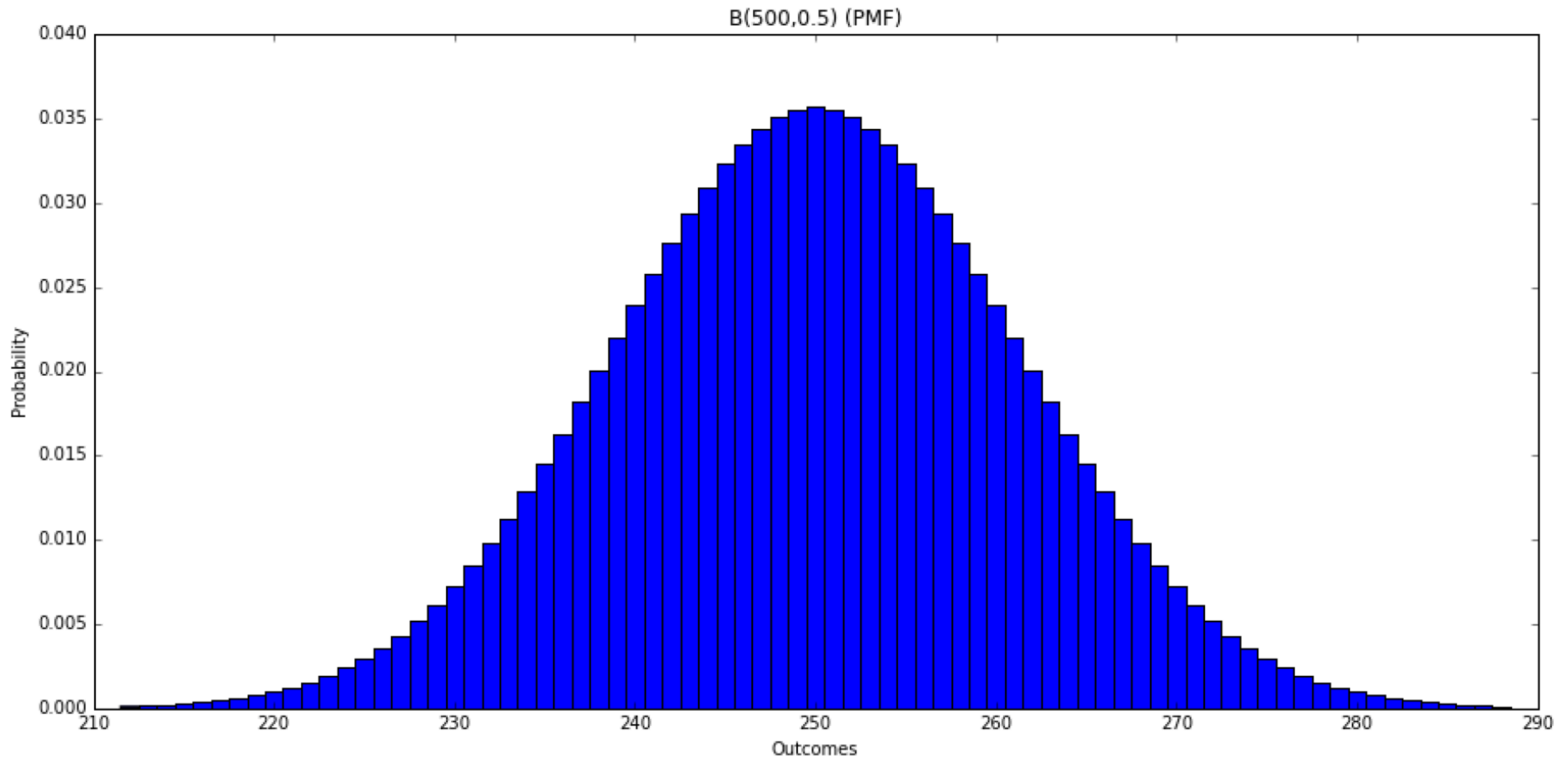
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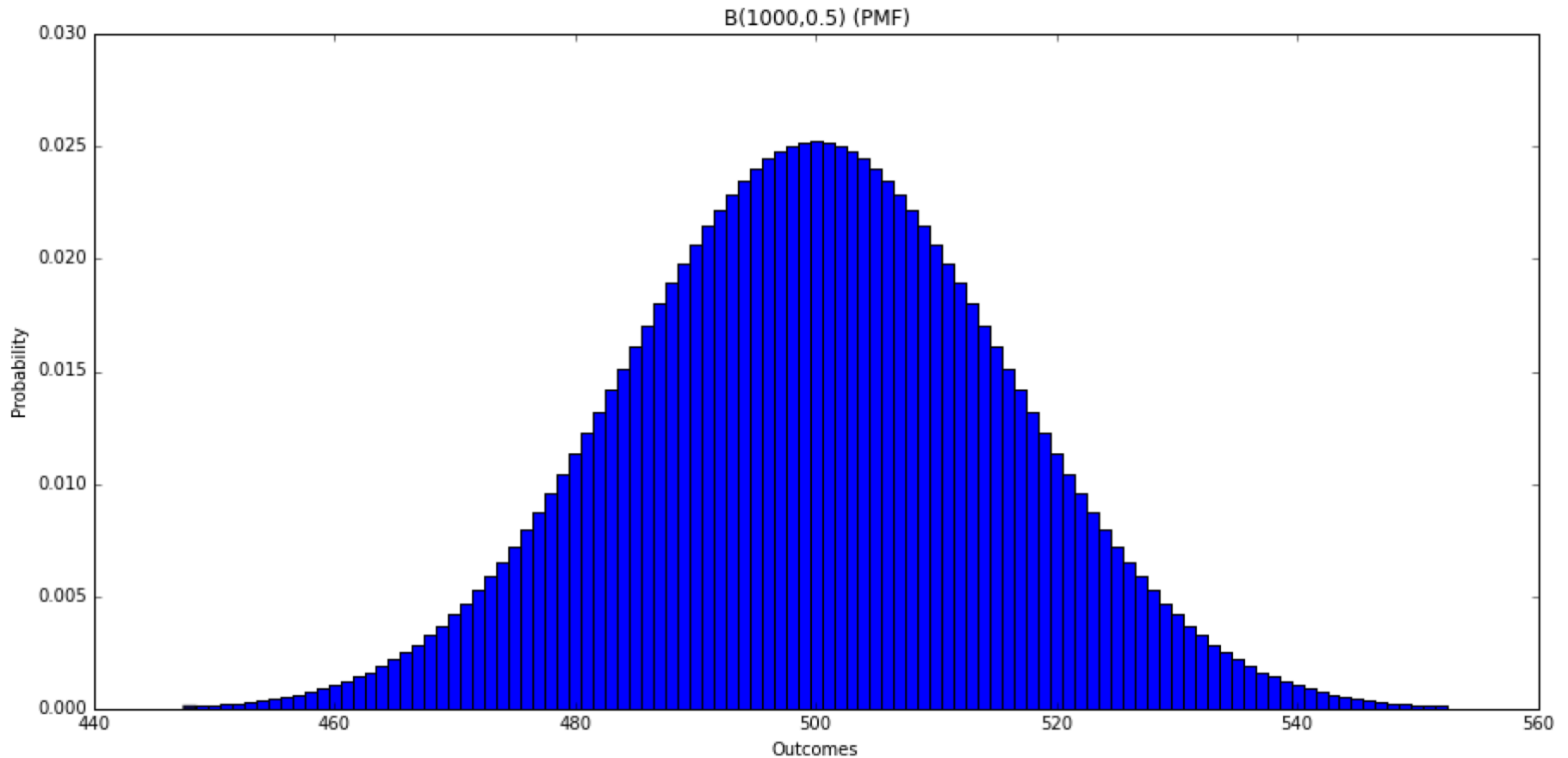
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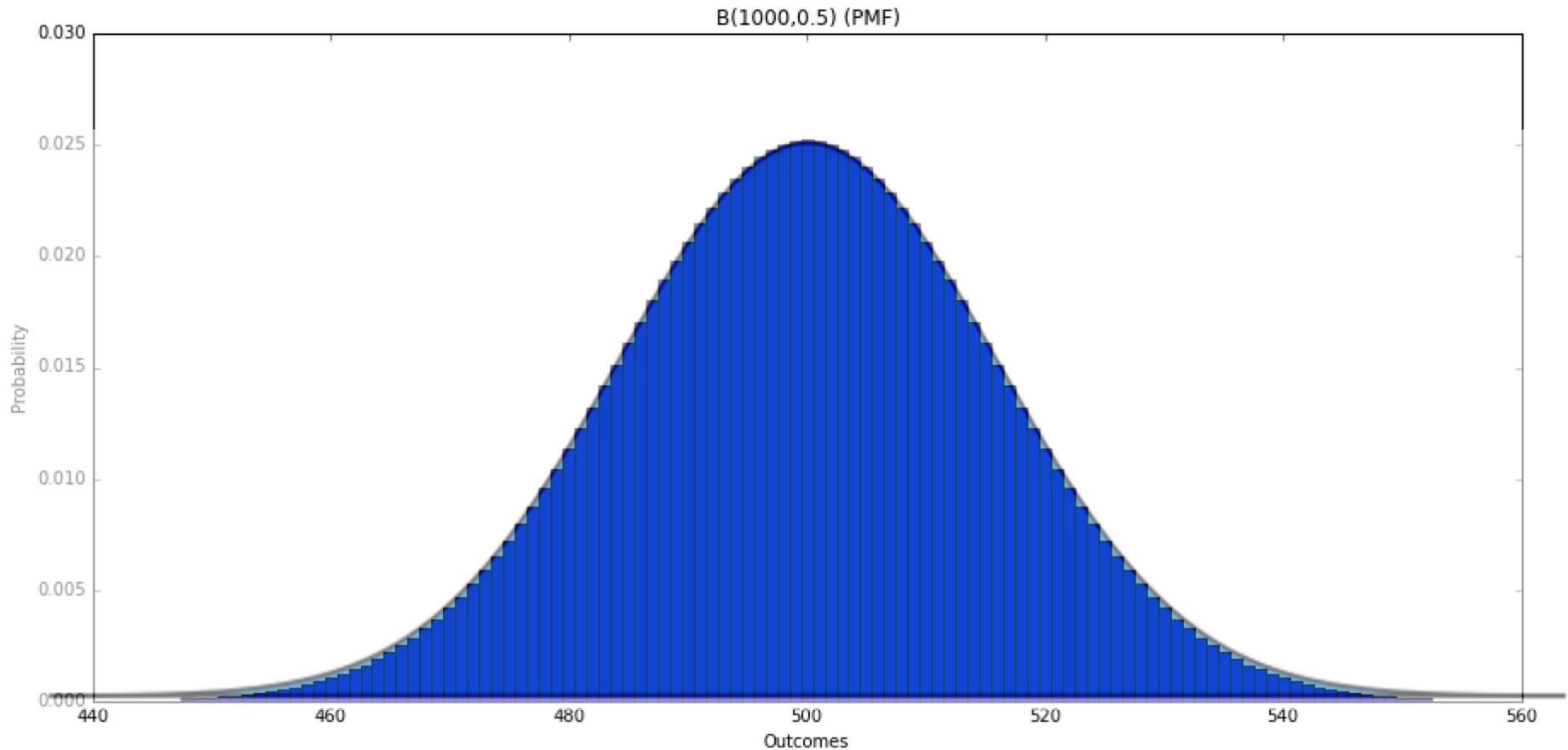
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