Stikprøveteori, Brøk

tirsdag 2. uge, juli 2018

Stikprøveteori definitioner Univers $Y_1, Y_2, Y_3, \dots, Y_N$,

$$\overline{Y} = \frac{1}{N} \sum_{j=1}^{N} Y_i$$

$$S^2 = \frac{1}{N-1} \sum_{j=1}^{N} (Y_i - \overline{Y})^2$$

stikprøve $y_1, y_2, y_3, \ldots, y_n$

$$\overline{y}_{si} = \frac{1}{n} \sum_{j=1}^{n} y_i$$

$$V(\overline{y}_{si}) = \frac{(N-n)}{N} \frac{1}{n} S^2$$

tre dele: endelighed, stikprøve, varians i univers

$$\widehat{V(\overline{y}_{si})} = \frac{(N-n)}{N} \frac{1}{n} \widehat{S}^{2}$$

$$\widehat{S}^{2} = s^{2} = \frac{1}{n-1} \sum_{i=1}^{n} (y_{i} - \overline{y}_{si})^{2}$$

Brøk (ratio)

$$R = \frac{Y}{X} = \frac{\frac{1}{N} \sum_{j=1}^{N} Y_{i}}{\frac{1}{N} \sum_{i=1}^{N} X_{i}} = \frac{\overline{Y}}{\overline{X}}$$

stikprøve
$$y_1, y_2, y_3, y_n,$$

stikprøve $x_1, x_2, x_3, x_n,$

$$\widehat{R} = \frac{\overline{y}}{\overline{x}}$$

$$E(\widehat{R}) \approx R$$

$$V(\widehat{R}) \approx \frac{1}{\overline{\chi}^2} \frac{N-n}{N} \frac{1}{n} T_z^2$$

$$T_z^2 = \frac{1}{N-1} \sum_{j=1}^{n} (Y_j - RX_j)^2$$

Brøk

$$T_z^2 = \frac{1}{N-1} \sum_{i=1}^{N} (Y_i - RX_j)^2$$

 $Z_i = Y_i - RX_i$

$$i = 1$$
 $i = 1$ $i = 1$

$$\begin{array}{l} \frac{1}{N-1}\sum\limits_{j=1}^{n}(Y_{j}\text{-}RX_{j})^{2}\text{=}[S_{y}^{2}+R^{2}S_{x}^{2}-2R\rho_{xy}S_{x}S_{y}]\\ \text{estimeres ved}\\ \widehat{S}_{y}^{2}+\widehat{R}^{2}\widehat{S}_{x}^{2}-2\widehat{R}\widehat{\rho}_{xy}\widehat{S}_{x}\widehat{S}_{y} \end{array}$$

$$Z_j = Y_j - RX_j$$
 hvor $R = \frac{\overline{Y}}{\overline{X}}$

$$\overline{Z} = \frac{1}{N} \sum_{j=1}^{N} Z_j = \frac{1}{N} \sum_{j=1}^{N} (Y_j - RX_j) = \frac{1}{N} \sum_{j=1}^{N} Y_j - \frac{1}{N} \sum_{j=1}^{N} RX_j =$$

$$= \overline{Y} - R\overline{X} = \overline{Y} - \frac{\overline{Y}}{\overline{X}}\overline{X} = \overline{Y} - \overline{Y} = 0$$

$$var(Z) = \frac{1}{N-1} \sum_{j=1}^{N} (Z_j - \overline{Z})^2 = \frac{1}{N-1} \sum_{j=1}^{N} (Z_j)^2$$

$$T_z^2 = \frac{1}{N-1} \sum_{i=1}^{N} (Y_i - RX_j)^2$$

Nu haves et univers $Z_1, Z_2, Z_3, \dots, Z_N$ med gennemsnit 0 og varians T_z^2

Brøk

$$T_z^2=rac{1}{N-1}\sum\limits_{j=1}^N(Y_j$$
- $RX_j)^2=rac{1}{N-1}\sum\limits_{j=1}^N(Y_j$ - $\overline{Y}+\overline{Y}-RX_j)^2=$

$$\frac{1}{N-1} \sum_{j=1}^{N} (Y_{j} - \overline{Y})^{2} + \frac{1}{N-1} \sum_{j=1}^{N} (\overline{Y} - RX_{j})^{2} + \frac{1}{N-1} \sum_{j=1}^{N} 2(Y_{j} - \overline{Y})(\overline{Y} - RX_{j})$$

$$\frac{1}{N-1} \sum_{j=1}^{N} (Y_j - \overline{Y})^2 = S_y^2$$

$$rac{1}{N-1}\sum\limits_{j=1}^{N}(\overline{Y}-RX_j)^2=rac{1}{N-1}\sum\limits_{j=1}^{N}[R(\overline{X}-X_j)]^2=$$

$$\frac{1}{N-1}R^2\sum_{j=1}^{N}(-\overline{X}+X_j)^2=R^2S_x^2$$

Brøk

$$\frac{1}{N-1}\sum_{j=1}^{N}2(Y_{j}-\overline{Y})(\overline{Y}-RX_{j})=$$

$$\frac{1}{N-1} \sum_{j=1}^{N} 2R(Y_j - \overline{Y})(\overline{X} - X_j) =$$

$$\frac{1}{N-1} 2R(SAP_{xy}) =$$

$$2Rrac{\mathit{SAP}_{\mathit{xy}}}{\sqrt{\mathit{SAK}_{\mathit{x}}\mathit{SAK}_{\mathit{y}}}}\sqrt{rac{\mathit{SAK}_{\mathit{x}}}{\mathit{N}-1}}rac{\mathit{SAK}_{\mathit{y}}}{\mathit{N}-1}}=$$

 $2R\rho S_x S_y$ samlet fås:

$$var(Z) = T_z^2 = \frac{1}{N-1} \sum_{j=1}^{N} (Y_j - RX_j)^2 = 0$$

$$[S_y^2 + R^2 S_x^2 - 2R \rho_{xy} S_x S_y]$$



BRØK frækt bevis

Fra kap. II fås
$$E(\overline{z}_{si}) = \overline{Z} = 0$$
 og $V(\overline{z}_{si}) = \frac{N-n}{Nn} T_z^2$

$$\widehat{R} - R = \frac{\frac{\frac{1}{n} \sum_{i=1}^{n} y_i}{\frac{1}{n} \sum_{i=1}^{n} x_i} - R = \frac{\frac{\frac{1}{n} \sum_{i=1}^{n} (y_i - Rx_i)}{\overline{x}}}{\overline{x}}$$

$$E(\widehat{R}-R)=E(\frac{\frac{1}{n}\sum\limits_{i=1}^{n}(y_{i}-Rx_{i})}{\overline{x}})\approx\frac{E(\overline{z})}{\overline{X}}=0$$

BRØK frækt bevis

$$\widehat{R} - R = \frac{\frac{\frac{1}{n} \sum_{i=1}^{n} y_i}{\frac{1}{n} \sum_{i=1}^{n} x_i} - R = \frac{\frac{1}{n} \sum_{i=1}^{n} (y_i - Rx_i)}{\overline{x}}$$

$$V(\widehat{R}) = V(\widehat{R} - R) = V(\frac{\frac{1}{n} \sum_{i=1}^{n} (y_i - Rx_i)}{\overline{x}}) \approx \frac{V(\overline{z})}{\overline{X}^2} =$$

$$= \frac{1}{\overline{V}^2} \frac{N - n}{N} \frac{1}{n} T_z^2$$

taylor

$$\begin{split} f(x) &= f(x_0) + f^{'}(x_0)(x - x_0) + \text{ rest} \\ \cdot & E[f(x)] = E[f(x_0)] + f^{'}(x_0)E(x - x_0) + \text{ rest} \approx f(x_0) \\ \cdot & \cdot \\ V[f(x)] &= V[f(x_0)] + [f^{'}(x_0)]^2 V(x - x_0) + \text{ rest} \approx [f^{'}(x_0)]^2 V(x) \end{split}$$

taylor

$$f(x,y) = f(x_0,y_0) + f'_x(x_0,y_0)(x-x_0) + f'_y(x_0,y_0)(y-y_0) + \text{ rest}$$

$$E[f(x,y)] \approx E[f(x_0,y_0)] + f'_x(x_0,y_0)E(x-x_0) + f'_y(x_0,y_0)E(y-y_0)$$

$$= E[f(x_0,y_0)] = f(x_0,y_0)$$

$$V[f(x,y)] \approx V[f(x_0,y_0) + f'_x(x_0,y_0)(x-x_0) + f'_y(x_0,y_0)(y-y_0)] = V[f'_x(x_0,y_0)(x-x_0)] + V[f'_y(x_0,y_0)(y-y_0)] + V[f'_y(x_0,y_0)(x-x_0)] + V[f'_y(x_0,y_0)(y-y_0)]$$

$$f(x,y) = \frac{y}{x}$$
 udvikles i $(x_0, y_0) = (\overline{X}, \overline{Y})$ og $f(\overline{X}, \overline{Y}) = \frac{\overline{Y}}{\overline{X}} = R$

$$f'_{x} = (-x^{-2})y$$

$$f'_{y} = x^{-1}$$

$$f(x,y) = \frac{y_0}{x_0} + (-x_0^{-2})y_0 * (x - x_0) + x_0^{-1} * (y - y_0) = f(x,y) = R + (-Rx_0^{-1}) * (x - x_0) + x_0^{-1} * (y - y_0) = R$$

$$f(x,y) = R + (-Rx_0^{-1}) * (x - x_0) + x_0^{-1} * (y - y_0) =$$

$$E[f(x,y)] = E(R) + (-Rx_0^{-1}) * E((x - x_0)) + x_0^{-1} * E((y - y_0)) =$$

$$= R + (-Rx_0^{-1}) * 0 + x_0^{-1} * 0 = R$$

$$\begin{split} &f(x,y) = R + \left(-Rx_0^{-1}\right) * \left(x - x_0\right) + x_0^{-1} * \left(y - y_0\right) \\ &. \\ &V[f(x,y)] = \\ &. \\ &V(R) + \left(-Rx_0^{-1}\right)^2 * V(\left(x - x_0\right)) + \left(x_0^{-1}\right)^2 * V(\left(y - y_0\right)) + \\ &. \\ &2(-Rx_0^{-1}) * \left(x_0^{-1}\right) * Cov(x,y) = \\ &. \\ &0 + \frac{R^2}{x_0^2} * V(x) + \frac{1}{x_0^2} * V(y) - 2R(\frac{1}{x_0^2}) * Cov(x,y) = \\ &. \\ &. \\ &\frac{1}{x_0^2} * \left[R^2V(x) + V(y) - 2R * Cov(x,y)\right] = \\ &\frac{1}{x_0^2} * \left[R^2V(x) + V(y) - 2R * \rho * S_y * S_x\right] = \\ &S_V = \sqrt{V(y)} \quad S_X = \sqrt{V(x)} \quad \rho = \frac{Cov(x,y)}{S_XS_X} \end{split}$$

$$S_y = \sqrt{V(y)}$$
 $S_x = \sqrt{V(x)}$ $\rho = \frac{Cov(x,y)}{S_y * S_x}$
 $x_0 = \overline{X}$ $y_0 = \overline{Y}$ $R = \frac{\overline{Y}}{\overline{x}}$

$$V(\widehat{R}) \approx (\frac{1}{\overline{X}})^2 * [S_y^2 + R^2 S_x^2 - 2R * \rho * S_y * S_x]$$

Usikkerhed i Lovmodel

Eksempel IV.2.2 side 204

N=2,3 mio. husstande i Danmark

 $X_j = \#$ personer i husstand nr. j

$$X = \sum_{j=1}^{N} X_j = 5,1$$
 mio. personer. Det samlede antal personer i Danmark

 $Y_j = X_j$ hvis husstanden er i Roskilde, ellers er $Y_j = 0$

$$Y = \sum_{j=1}^{N} Y_j$$
 = det samlede antal personer i Roskilde

$$R = \frac{Y}{X} = \frac{antal_personer_i_Roskilde}{antal_personer_i_Danmark} = 0,041950516$$

(opslag i Statistisk Årbog, nu bruges statistikbanken)

Stikprøve på n = 76.784 (husstande)

$$x_1, x_2, x_3, \dots, x_n$$

 $\sum_{i=1}^{76.784} X_i = 171.893$ er det samlede antal personer i de udtrukne husstande

Model 1 (forkert model)

Antag nu at de 171.893 personer er udtrukket simpelt tilfældigt blandt de 5,1 mio. personer i Danmark.

Vi ønsker at estimere andelen af personer der bor i Roskilde.

I stikprøven på de 171.893 personer bor 7.406 personer i Roskilde

$$\hat{R} = \frac{7406}{171896} = 0,04308494$$

$$V(\widehat{R}) = \frac{N-n}{Nn} \frac{N}{N-1} R(1-R)$$

$$\widehat{V(\widehat{R})} = \frac{N-n}{Nn} \frac{n}{n-1} \widehat{R} (1-\widehat{R}) = \frac{N-n}{N} \frac{1}{n-1} \widehat{R} (1-\widehat{R}) =$$

husk at udvælgelsen er 1/30

brug
$$\widehat{R} = \frac{7406}{171896} = 0,04308494$$

$$(1-\frac{n}{N})\frac{1}{n-1}\widehat{R}(1-\widehat{R})=(1-\frac{1}{30})\frac{1}{171893-1}\widehat{R}(1-\widehat{R})=(0,0004815)^2$$

$$H_0: R = 0,041950516$$
 $H_A: R \neq 0,041950516$

$$H_A: R \neq 0,041950516$$

$$U = \frac{\hat{R} - R}{\sqrt{var(\hat{R})}} = \frac{0.04308494 - 0.041950516}{0.0004815} = 2,36$$
 som er udover 1,96 grænsen

model 2 (Brøk estimation) Stikprøve på n= 76.784 (husstande) $y_1, y_2, y_3, \dots, y_n$ (husk at mange af y'erne bliver nul)

$$x_1, x_2, x_3, \dots, x_n$$

$$\widehat{R} = \frac{\frac{\frac{1}{n} \sum_{i=1}^{n} y_i}{\frac{1}{n} \sum_{i=1}^{n} x_i}}{\frac{1}{n} \sum_{i=1}^{n} x_i} = \frac{\overline{y}}{\overline{x}} = \frac{7406}{171896} = 0,04308494$$

vi har at

$$V(\widehat{R}) pprox rac{1}{\overline{\chi}^2} rac{N-n}{N} rac{1}{n} T_z^2$$
 som skal estimeres

vi bruger
$$\overline{x} = \frac{1}{n} \sum_{i=1}^{n} x_i = \frac{171.893}{76.784} = 2,2387$$

$$\frac{N-n}{N} = (1-f) = (1-\frac{1}{30})$$

$$\widehat{T}_{z}^{2} = \frac{1}{n-1} \sum_{i=1}^{n} (y_{i} - \widehat{R}x_{i})^{2} = 0,2905$$

$$V(\widehat{R}) \approx \frac{1}{(2,2387)^2} (1 - \frac{1}{30}) \frac{1}{76.784} * 0,2905 = (0,00085431)^2$$

$$H_0: R = 0,041950516$$

$$H_A: R \neq 0,041950516$$

$$U = \frac{\widehat{R} - R}{\sqrt{var(\widehat{R})}} = \frac{0.04308494 - 0.041950516}{0.00085431} = 1,32 \text{ som er under } 1,96 \text{ grænsen}.$$

ratio estimation

$$R = \frac{Y}{X}$$
 $Y = RX$

$$\hat{Y} = \hat{R}X$$

$$\widehat{Y} = \frac{\overline{y}}{\overline{x}}X$$
 $\frac{1}{N}\widehat{Y} = \frac{\overline{y}}{\overline{x}}\frac{1}{N}X$

$$\overline{y}_R = \frac{\overline{y}}{\overline{x}} \overline{X} = \overline{y}_{si} \frac{\overline{X}}{\overline{x}} = \overline{y}_{si} \frac{\text{gennemsnit_i_univers}}{\text{gns_i_stikprøve}}$$

$$V(\overline{y}_R) = V(\frac{\overline{y}}{\overline{x}}\overline{X}) = V(\widehat{R})(\overline{X})^2 \approx \frac{N-n}{N}\frac{1}{n}T_z^2$$

Regression overspringes

$$\overline{y}_{reg} = \overline{y}_{si} + \beta(\overline{X} - \overline{x})$$

vælg β som hældningskoefficient i universet

$$dvs \beta = \frac{\rho_{xy}S_y}{S_x}$$

$$V(\overline{y}_{\mathit{reg}}) = rac{\mathit{N}-\mathit{n}}{\mathit{N}} rac{1}{\mathit{n}} S_{\mathit{y}}^2 (1-
ho_{\mathit{xy}}^2) = V(\overline{y}_{\mathit{si}}) (1-
ho_{\mathit{xy}}^2)$$

så altid
$$V(\overline{y}_{\mathit{reg}}) <= V(\overline{y}_{\mathit{R}})$$

$$V(\overline{y}_{reg}) <= V(\overline{y}_{si})$$

når
$$ho_{{\scriptscriptstyle xy}}^2$$
 stor så $V(\widehat{R}) <= V(\overline{y}_{{\scriptscriptstyle si}})$

eks. IV.3.2 præmieindtægter s. 220

Univers $Y_1, Y_2, Y_3, \dots, Y_{94}$, indtægt i 1990 $X_1, X_2, X_3, \dots, X_{94}$, indtægt i 1988

Univers oplysninger

· · · · · · · · · · · · · · · · · · ·				
	1988	1990		
gns	$\overline{X} = 58.437$	$\overline{Y} = 67.368$		
S	$S_X = 226.535$	$S_Y = 252.730$		
R	$R = \frac{67368}{58437} = 1,1528$			
korr.	$ ho_{XY} = 0,99556$			
hældning	$\beta = \frac{0,99556*252730}{226535} = 1,1107$			

eks. IV.3.2 præmieindtægter s. 220

Stikprøve n=5 oplysninger

	1988	1990
Asperup	85	78
Hagl	708	810
Lunde	976	1171
Rødvig	1142	1227
Ulfborg	138	125
gns	$\overline{x} = 610$	$\overline{y} = 682$
S	$s_X = 481$	$s_Y = 554$
R	$\widehat{R} = \frac{682}{610} = 1,1180$	
korr.	$\widehat{\rho}_{xy} = 0,99577$	
hældning	$\widehat{\beta} = \frac{0.99577*554}{481} = 1,1469$	

eks. IV.3.2 præmieindtægter s. 220

 $\begin{array}{c} \text{Univers } Y_1, Y_2, Y_3, Y_{94}, \text{indtægt i } 1990 \\ X_1, \ X_2, \ X_3, X_{94}, \ \text{indtægt i } 1988 \end{array}$

	estimator	spredning
\overline{y}_{si}	682	75.688
\overline{y}_R	$=682*\frac{58437}{610}=65334$	7.676
\overline{y}_{reg}	=682+1,1469*(58437-610)=67004	7.124

Proportional og optimal

$$n_k = nW_k$$

$$\overline{y}_p = \sum_{k=1}^K W_k \overline{y_k} = \frac{1}{n} \sum_{i=1}^n y_i$$
 dvs. det "oprindelige" estimat, selvvejende

$$V(\overline{y}_p) = \frac{N-n}{N} \frac{1}{n} \sum_{k=1}^{K} W_k S_k^2$$

Sammenlignet med simpel tilfældig

$$V(\overline{y}_{si}) = \frac{(N-n)}{N} \frac{1}{n} S^2$$