NANYANG TECHNOLOGICAL UNIVERSITY SPMS/DIVISION OF MATHEMATICAL SCIENCES

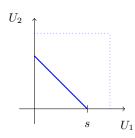
2016/17 Semester 1 MH2500 Probability and Introduction to Statistics Tutorial 8

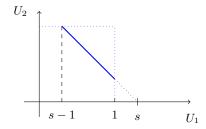
For the tutorial on 13 October, let us discuss

• Ex. 3.8.43, 48, 51, 67, 70, 74

Ex. 3.8.43. Let U_1 and U_2 be independent and uniform on [0,1]. Find and sketch the density function of $S = U_1 + U_2$.

[Solution:]





The joint density function is given by

$$f_{U_1U_2}(x,y) = \begin{cases} 1, & 0 \le x, y \le 1; \\ 0, & \text{otherwise.} \end{cases}$$

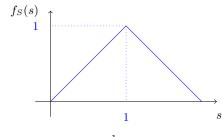
From lecture, the density function is

$$f_S(s) = \int_{-\infty}^{\infty} f(x, s - x) dx$$
$$= \int_0^s dx = s, \qquad (0 \le s < 1),$$

and

$$f_S(s) = \int_{s-1}^{1} f(x, s - x) dx$$

= 2 - s, (1 \le s \le 2),



Ex. 3.8.48. Let T_1 and T_2 be independent exponentials with parameters λ_1 and λ_2 . Find the density function of $T_1 + T_2$.

[Solution:] Let $Z = T_1 + T_2$. From lecture,

$$f_Z(z) = \int_{-\infty}^{\infty} f(x, z - x) dx$$

$$= \int_{0}^{z} \lambda_1 e^{-\lambda_1 x} \lambda_2 e^{-\lambda_2 (z - x)} dx$$

$$= \lambda_1 \lambda_2 \int_{0}^{z} e^{-\lambda_1 x - \lambda_2 (z - x)} dx$$

$$= \lambda_1 \lambda_2 \left[\frac{1}{-\lambda_1 + \lambda_2} e^{-\lambda_1 x - \lambda_2 (z - x)} \right]_{0}^{z}$$

$$= \frac{\lambda_1 \lambda_2}{\lambda_2 - \lambda_1} (e^{-\lambda_1 z} - e^{-\lambda_2 z}), \qquad z > 0.$$

Ex. 3.8.51. Let X and Y have the joint density function f(x, y) and let Z = XY. Show that the density function of Z is

$$f_Z(z) = \int_{-\infty}^{\infty} f\left(y, \frac{z}{y}\right) \frac{1}{|y|} dy.$$

[Solution:]

$$F_Z(z) = \int_{-\infty}^0 \int_{z/x}^\infty f(x,y) \, dy dx + \int_0^\infty \int_{-\infty}^{z/x} f(x,y) \, dy dx.$$

Make a change of variable y = v/x, we find that

$$F_{Z}(z) = \int_{-\infty}^{0} \int_{z}^{-\infty} \frac{1}{x} f\left(x, \frac{v}{x}\right) dv dx + \int_{0}^{\infty} \int_{-\infty}^{z} \frac{1}{x} f\left(x, \frac{v}{x}\right) dv dx$$

$$= \int_{-\infty}^{0} \int_{-\infty}^{z} \frac{1}{|x|} f\left(x, \frac{v}{x}\right) dv dx + \int_{0}^{\infty} \frac{1}{x} \int_{-\infty}^{z} f\left(x, \frac{v}{x}\right) dv dx$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{z} \frac{1}{|x|} f\left(x, \frac{v}{x}\right) dv dx$$

$$= \int_{-\infty}^{z} \int_{-\infty}^{\infty} \frac{1}{|x|} f\left(x, \frac{v}{x}\right) dx dv.$$

Hence

$$f_Z(z) = \int_{-\infty}^{\infty} \frac{1}{|x|} f\left(x, \frac{z}{x}\right) dx.$$

Ex. 3.8.67. A card contains n chips and has an error-correcting mechanism such that the card still functions if a single chip fails but does not function if two or more chips fail. If each

chip has a lifetime that is an independent exponential with parameter λ , find the density function of the card's lifetime.

[Solution:] First, let f denote the density function of an exponential function with parameter λ . Then

$$f(t) = \lambda e^{-\lambda t}, \qquad (t \ge 0),$$

and

$$F(t) = 1 - e^{-\lambda t}, \qquad (t \ge 0).$$

Let $X_1, X_2, \dots X_n$ be the lifetime of the n chips. then the lifetime of the card is the lifetime of $X_{(2)}$. Let $U = X_{(2)}$ and then we require

- $X_{(1)} \le u$, $U \in [u, u + du]$, and $X_{(3)}, \dots, X_{(n)} \ge u$.

Then the density function is

$$f_U(u) = \frac{n!}{1!1!(n-2)!} F(u) f(u) [1 - F(u)]^{n-2}$$

$$= n(n-1)(1 - e^{-\lambda u}) \lambda e^{-\lambda u} [1 - (1 - e^{-\lambda u})^{n-2}]$$

$$= n(n-1) \lambda e^{-(n-1)\lambda u} (1 - e^{-\lambda u})$$

Ex. 3.8.70. If five numbers are chosen at random in the interval [0, 1], what is the probability that they all lie in the middle half of the interval?

[Solution:] Method 1: Using ordered statistics. Let the five random numbers be X_1, X_2, \ldots, X_5 and let $U = X_{(5)}, V = X_{(1)}$. Then all five numbers are in the middle have interval if and only

$$\frac{1}{4} \leq X_{(1)} \leq X_{(2)} \leq X_{(3)} \leq X_{(4)} \leq X_{(5)} \leq \frac{3}{4}.$$

First, we find the joint density function for U and V. By the differential argument, we require

- $X_{(2)}, X_{(3)}, X_{(4)}$ in [v, u], and $U \in [u, u + du]$.

Let f(x) be the uniform density function on [0,1]. Thus

$$f_{U,V}(u,v) = \frac{5!}{1!3!1!} f(v)(u-v)^3 f(u) = 20(u-v)^3 \qquad (0 \le v \le u \le 1).$$

The required probability is

$$20 \int_{1/4}^{3/4} \int_{1/4}^{u} (u - v)^{3} dv du = 20 \int_{1/4}^{3/4} \left[-\frac{1}{4} (u - v)^{4} \right]_{1/4}^{u} du$$

$$= 5 \int_{1/4}^{3/4} \left(u - \frac{1}{4} \right)^{4} du$$

$$= 5 \left[\frac{1}{5} \left(u - \frac{1}{4} \right)^{5} \right]_{1/4}^{3/4}$$

$$= \frac{1}{32}.$$

Method 2: We want each random number to lie in the interval [1/4, 3/4], and since the random numbers are chosen at random, they have the uniform distribution on [0,1]. Thus the probability of any one random number lying in the interval [1/4, 3/4] is $\frac{1}{2}$. Since the random numbers are independent of one another, the probability is

$$\left(\frac{1}{2}\right)^5 = \frac{1}{32}.$$

Ex. 3.8.74. Let U_1, U_2 , and U_3 be independent uniform random variables.

- a. Find the joint density of $U_{(1)}, U_{(2)}$, and $U_{(3)}$.
- b. The locations of three gas stations are independently and randomly placed along a mile of highway. What is the probability that no two gas stations are less than $\frac{1}{3}$ mile apart?

[Solution:]

- a. Let $X=U_{(1)},\,Y=U_{(2)},\,$ and $Z=U_{(3)}.$ Then in order to apply the differential argument, we require
 - $\bullet \ X \in [x, x + dx],$
 - $Y \in [y, y + dy]$, and
 - $Z \in [z, z + dz]$.

Let f be the uniform density function on [0,1]. Then

$$f(x,y,z) = \frac{3!}{1!1!1!} f(x)f(y)f(z) = 6 \qquad (0 \le x \le y \le z).$$

b. The probability is

$$\int_{2/3}^{1} \int_{1/3}^{z-1/3} \int_{0}^{y-1/3} 6 \, dx dy dz = \int_{2/3}^{1} \int_{1/3}^{z-1/3} [6x]_{0}^{y-1/3} \, dy dz$$

$$= \int_{2/3}^{1} \int_{1/3}^{z-1/3} [6y - 2] \, dy dz$$

$$= \int_{2/3}^{1} \left[3y^{2} - 2y \right]_{1/3}^{z-1/3} dz$$

$$= \int_{2/3}^{1} \left[3z^{2} - 4z + \frac{4}{3} \right] dz$$

$$= \left[z^{3} - 2z^{2} + \frac{4}{3}z \right]_{2/3}^{1}$$

$$= 1 - 2 + \frac{4}{3} - \frac{8}{27} + \frac{8}{9} - \frac{8}{9}$$

$$= \frac{1}{27}.$$