

General EM Problem

Given

- Observed data $\{\mathbf{x}_d\}_{d \in D}$
- Unobserved data $\{\mathbf{z}_d\}_{d \in D}$ where $\mathbf{z}_d = \langle z_{d1}, \dots, z_{dM} \rangle$
- Parameterized probability distribution $p(D|h)$ where
 - $D = \{d\}$ is the complete data where $d = \langle \mathbf{x}_d, \mathbf{z}_d \rangle$
 - h comprises the parameters

Determine **ML** hypothesis h' that (locally) maximizes $\mathbb{E}[\ln p(D|h')]$

General EM Algorithm

Define function $Q(h'|h) = \mathbb{E}[\ln p(D|h')|h, \{\mathbf{x}_d\}_{d \in D}]$ given current parameters h and observed data $\{\mathbf{x}_d\}_{d \in D}$ to estimate the latent variables $\{\mathbf{z}_d\}_{d \in D}$

EM Algorithm. Pick random initial h . Then, iterate

- **E Step.** Calculate $Q(h'|h)$ using current hypothesis h and observed data $\{\mathbf{x}_d\}_{d \in D}$ to estimate the latent variables $\{\mathbf{z}_d\}_{d \in D}$:
 $Q(h'|h) \leftarrow \mathbb{E}[\ln p(D|h')|h, \{\mathbf{x}_d\}_{d \in D}]$
- **M Step.** Replace hypothesis h by the hypothesis h' that maximizes this Q function: $h \leftarrow \operatorname{argmax}_{h'} Q(h'|h)$

Applying EM to Estimate M Means

Given

- Instances from X generated by mixture of M Gaussians with the same known variance σ^2
- Unknown means $\langle \mu_1, \dots, \mu_M \rangle$ of the M Gaussians
- Don't know which instance x_d is generated by which Gaussian

Determine **maximum likelihood (ML)** estimates of $\langle \mu_1, \dots, \mu_M \rangle$

Consider full description of each instance as $d = \langle x_d, z_{d1}, \dots, z_{dM} \rangle$ where

- z_{dm} is unobservable and is of value 1 if m -th Gaussian is selected to generate x_d , and 0 otherwise
- x_d is observable

Applying EM to Estimate M Means

$$\begin{aligned} p(d|h') &= p(x_d, z_{d1}, \dots, z_{dM} | h') \\ &= \frac{1}{\sqrt{2\pi}\sigma} \exp \left(-\frac{1}{2\sigma^2} \sum_{m=1}^M z_{dm} (x_d - \mu'_m)^2 \right) \end{aligned}$$

$$\begin{aligned} \ln p(D|h') &= \ln \prod_{d \in D} p(d|h') = \sum_{d \in D} \ln p(d|h') \\ &= \sum_{d \in D} \left(\ln \frac{1}{\sqrt{2\pi}\sigma} - \frac{1}{2\sigma^2} \sum_{m=1}^M z_{dm} (x_d - \mu'_m)^2 \right) \end{aligned}$$

$$\begin{aligned} Q(h'|h) &= \mathbb{E}[\ln p(D|h')] \\ &= \mathbb{E} \left[\sum_{d \in D} \left(\ln \frac{1}{\sqrt{2\pi}\sigma} - \frac{1}{2\sigma^2} \sum_{m=1}^M z_{dm} (x_d - \mu'_m)^2 \right) \right] \\ &= \sum_{d \in D} \left(\ln \frac{1}{\sqrt{2\pi}\sigma} - \frac{1}{2\sigma^2} \sum_{m=1}^M \mathbb{E}[z_{dm}] (x_d - \mu'_m)^2 \right) \end{aligned}$$

$$\text{where } \mathbb{E}[z_{dm}] = \frac{\exp(-\frac{1}{2\sigma^2} (x_d - \mu_m)^2)}{\sum_{\ell=1}^M \exp(-\frac{1}{2\sigma^2} (x_d - \mu_\ell)^2)}$$

Applying EM to Estimate M Means

$$\begin{aligned} & \arg \max_{h'} Q(h'|h) \\ &= \arg \max_{h'} \sum_{d \in D} \left(\ln \frac{1}{\sqrt{2\pi}\sigma} - \frac{1}{2\sigma^2} \sum_{m=1}^M \mathbb{E}[z_{dm}] (x_d - \mu'_m)^2 \right) \\ &= \arg \max_{h'} \sum_{d \in D} - \sum_{m=1}^M \mathbb{E}[z_{dm}] (x_d - \mu'_m)^2 \\ &= \arg \min_{h'} \sum_{d \in D} \sum_{m=1}^M \mathbb{E}[z_{dm}] (x_d - \mu'_m)^2 \\ & \mu'_m \leftarrow \frac{\sum_{d \in D} \mathbb{E}[z_{dm}] x_d}{\sum_{d \in D} \mathbb{E}[z_{dm}]} \end{aligned}$$

Hold on...

Let $X_D = \{\mathbf{x}_d\}_{d \in D}$ and $Z_D = \{\mathbf{z}_d\}_{d \in D}$. Shouldn't h' be selected to maximize log-likelihood of **observed** data $\ln p(X_D|h')$ (and marginalize out the unobserved variables Z_D) instead of $Q(h'|h)$?

Proposition. Log-likelihood of observed data monotonically increases with an increasing number of EM iterations.

Proof.

$$\begin{aligned}
 p(X_D|h') &= p(X_D, Z_D|h') / P(Z_D|h', X_D) \\
 \ln p(X_D|h') &= \ln p(X_D, Z_D|h') - \ln P(Z_D|h', X_D) \\
 \sum_{Z_D} P(Z_D|h, X_D) \ln p(X_D|h') &= \sum_{Z_D} P(Z_D|h, X_D) \ln p(X_D, Z_D|h') - \sum_{Z_D} P(Z_D|h, X_D) \ln P(Z_D|h', X_D) \\
 \ln p(X_D|h') &= \underbrace{\sum_{Z_D} P(Z_D|h, X_D) \ln p(X_D, Z_D|h')}_{=Q(h'|h)} - \underbrace{\sum_{Z_D} P(Z_D|h, X_D) \ln P(Z_D|h', X_D)}_{=R(h'|h)} \\
 &= Q(h'|h) + R(h'|h) \tag{1} \\
 \ln p(X_D|h) &= Q(h|h) + R(h|h) \quad \text{by subst. } h' = h \text{ in (1)} \tag{2} \\
 \ln p(X_D|h') - \ln p(X_D|h) &= Q(h'|h) - Q(h|h) + R(h'|h) - R(h|h) \quad \text{using (1) - (2)} \\
 \ln p(X_D|h') - \ln p(X_D|h) &\geq Q(h'|h) - Q(h|h) \quad \text{by Gibbs' inequality: } R(h'|h) \geq R(h|h)
 \end{aligned}$$