

# **An extensive comparison of species-abundance distribution models**

Elita Baldridge<sup>1,2</sup>, David J. Harris<sup>3</sup>, Xiao Xiao<sup>1,2,4,5</sup>, Ethan P. White<sup>\*,1,2,3,6</sup>

<sup>1</sup> Department of Biology, Utah State University, Logan, Utah, USA

<sup>2</sup> Ecology Center, Utah State University, Logan, Utah, USA

<sup>3</sup> Department of Wildlife Ecology & Conservation, University of Florida, Gainesville, Florida, USA

<sup>4</sup> School of Biology and Ecology, University of Maine, Orono, Maine, USA

<sup>5</sup> Mitchell Center for Sustainability Solutions, University of Maine, Orono, Maine, USA

<sup>6</sup> Informatics Institute, University of Florida, Gainesville, Florida, USA

## **Abstract**

A number of different models have been proposed as descriptions of the species-abundance distribution (SAD). Most evaluations of these models use only one or two models, focus only a single ecosystem or taxonomic group, or fail to use appropriate statistical methods. We use likelihood and AIC to compare the fit of four of the most widely used models to data on over 16,000 communities from a diverse array of taxonomic groups and ecosystems. Across all datasets combined the log-series, Poisson lognormal, and negative binomial all yield similar overall fits to the data. Therefore, when correcting for differences in the number of parameters the log-series generally provides the best fit to data. Within individual datasets some other distributions performed nearly as well as the log-series even after correcting for the number of parameters. The Zipf distribution is generally a poor characterization of the SAD.

## Introduction

The species abundance distribution (SAD) describes the full distribution of commonness and rarity in ecological systems. It is one of the most fundamental and ubiquitous patterns in ecology, and exhibits a consistent general form with many rare species and few abundant species occurring within a community. The SAD is one of the most widely studied patterns in ecology, leading to a proliferation of models that attempt to characterize the shape of the distribution and identify potential mechanisms for the pattern (see McGill et al. 2007 for a recent review of SADs). These models range from arbitrary distributions that are chosen based on providing a good fit to the data (Fisher et al. 1943), to distributions chosen based on the most likely states of generic random systems (Frank 2011, Harte 2011, Locey and White 2013), to models based more directly on ecological processes (Tokeshi 1993, Hubbell 2001, Volkov et al. 2003, Alroy 2015).

Which model or models provide the best fit to the data, and the resulting implications for the processes structuring ecological systems, is an active area of research (e.g., McGill 2003, Volkov et al. 2003, Ulrich et al. 2010, White et al. 2012, Connolly et al. 2014). However, most comparisons of the different models: 1) use only a small subset of available models (typically two; e.g., McGill 2003, Volkov et al. 2003, White et al. 2012, Connolly et al. 2014); 2) focus on a single ecosystem or taxonomic group (e.g., McGill 2003, Volkov et al. 2003); or 3) fail to use the most appropriate statistical methods (e.g., Ulrich et al. 2010, see Matthews and Whittaker 2014 for discussion of best statistical methods for fitting SADs). This makes it difficult to draw general conclusions about which, if any, models provide the best empirical fit to species abundance distributions.

Here, we evaluate the performance of four of the most widely used models for the species abundance distribution using likelihood-based model selection on data from 16,209 communities and nine major taxonomic groups. This includes data from terrestrial, aquatic, and marine ecosystems representing roughly 50 million individual organisms in total.

## Methods

### Data

We compiled data from citizen science projects, government surveys, and literature mining to produce a dataset with 16,209 communities, from nine taxonomic groups, representing nearly 50 million individual terrestrial, aquatic, and marine organisms. Data for trees, birds, butterflies and mammals was compiled by White et al. (2012) from six data sources: the US Forest Service Forest Inventory and Analysis (FIA; USDA Forest Service 2010), the North American Butterfly Association's North American Butterfly Count (NABC; North American Butterfly Assoc. 2009), the Mammal Community Database (MCDB; Thibault et al. 2011), Alwyn Gentry's Forest Transect Data Set (Gentry; Phillips and Miller 2002), the Audubon Society Christmas Bird Count (CBC; National Audubon Society 2002), and the US Geological Survey's North American Breeding Bird Survey (BBS; Pardieck et al. 2014) (see Table 1 for details). The publicly available datasets (FIA, MCDB, Gentry, and BBS) were acquired using the EcoData Retriever (<http://ecodataretriever.org>; Morris and White 2013). Details of the treatment of these datasets can be found in Appendix A of White et al. (2012), but in general data were analyzed at the level of the site defined in the dataset and a single year of data was selected for each site. We modified the data slightly by removing sites 102 and 179 from the Gentry data due to issues with decimal abundances appearing in raw data due to either data entry or data structure errors. Data on Actinopterygii, Reptilia, Coleoptera, Arachnida, and Amphibia, were mined from literature by Baldrige and are publicly available (Baldrige 2013) (see Table 1 for details). These data were collected at the level of the site defined in the publication if raw data were available at that scale, and at the scale of the entire study otherwise. Time scales of collection for this data depended on the study but was typically one or a few years. All data sources used in the analysis a samples (or censuses) of a taxonomic assemblage, where all individuals of any species seen are recorded. Abundances in the compiled datasets were counts of individuals.

Table 1: Details of datasets used to evaluate the form of the species abundance distribution. Datasets

marked as Private were obtained through data requests to the providers.

Dataset	Dataset code	Availability	Total sites	Citation
Breeding Bird Survey	BBS	Public	2769	Pardieck et al. (2014)
Christmas Bird Count	CBC	Private	1999	National Audubon Society (2002)
Gentry's Forest Transects	Gentry	Public	220	Phillips and Miller (2002)
Forest Inventory Analysis	FIA	Public	10355	USDA Forest Service (2010)
Mammal Community DB	MCDB	Public	103	Thibault et al. (2011)
NA Butterfly Count	NABA	Private	400	North American Butterfly Assoc. (2009)
Actinopterygii	Actinopterygii	Public	161	Baldrige (2013)
Reptilia	Reptilia	Public	129	Baldrige (2013)
Amphibia	Amphibia	Public	43	Baldrige (2013)
Coleoptera	Coleoptera	Public	5	Baldrige (2013)
Arachnida	Arachnida	Public	25	Baldrige (2013)

## Models

We selected models for analysis based on four criteria. First, since the majority of species abundance distributions (SADs) are constructed using counts of individuals (for discussion of alternative approaches see McGill et al. 2007 and @morlon2009) we selected models with discrete distributions (i.e., those that only have non-zero probabilities for positive integer values of abundance). Second, in order to use best practices for comparing species abundance distributions we selected models with analytically defined probability mass functions that allow the calculation of likelihoods (see details in Analysis). Third, McGill et al. (2007) classified species abundance distribution models into five different families: purely statistical, branching process, population dynamics, niche partitioning, and spatial distribution of individuals. We evaluated models from each of these families, with some models having been derived from more than one family of processes. Finally, we selected models

that have been widely used in the ecological literature. Based on these criteria we evaluated the log-series, the Poisson lognormal, the negative binomial, and the Zipf distributions. All distributions were defined to be capable of having non-zero probability at integer values from 1 to infinity.

The log-series is one of the first distributions used to describe the SAD, being derived as a purely statistical distribution by Fisher (1943). It has since been derived as the result of ecological processes, the metacommunity SAD for ecological neutral theory (Hubbell 2001, Volkov et al. 2003), and several different maximum entropy models (Pueyo et al. 2007, Harte et al. 2008).

The lognormal is one of the most commonly used distributions for describing the SAD (McGill 2003) and has been derived as a null form of the distribution resulting from the central limit theorem (May 1975), population dynamics (Engen and Lande 1996), and niche partitioning (Sugihara 1980). We use the Poisson lognormal because it is a discrete form of the distribution appropriate for fitting discrete abundance data (Bulmer 1974).

The negative binomial (which can be derived as a Gamma-distributed mixture of Poisson distributions) provides a good characterization of the SAD predictions for several different ecological neutral models for the purposes of model selection (Connolly et al. 2014). We use it to represent neutral models as a class.

The Zipf (or power law) distribution was derived based on both branching processes and as the outcome of the McGill and Collin's (2003) spatial model. It was one of the best fitting distributions in a recent meta-analysis of SADs (Ulrich et al. 2010). We use the discrete form of the distribution which is appropriate for fitting discrete abundance data (White et al. 2008).

Figure 1 shows three example sites with the empirical distribution and associated models fit to the data. Zipf distributions tend to predict the most rare species followed by the log-series, the negative binomial, and Poisson lognormal.

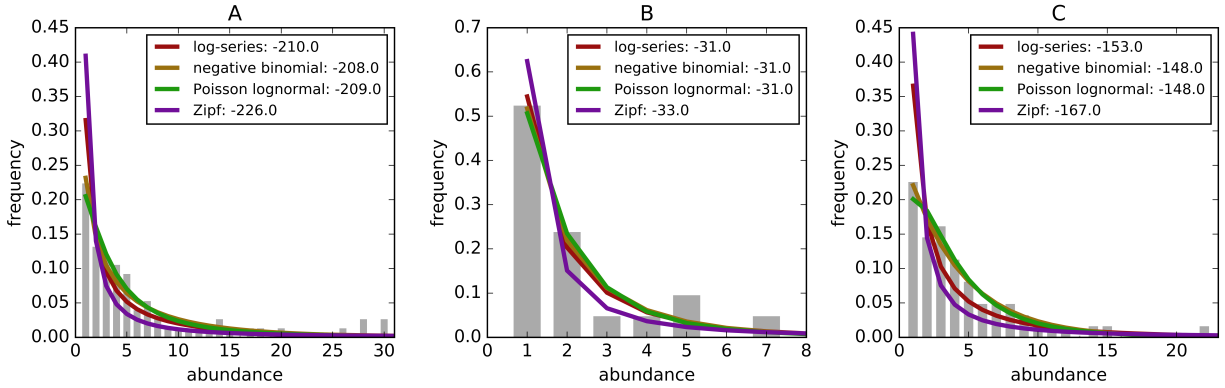


Figure 1: Example species-abundance distributions including the empirical distributions (grey bars) and the best fitting log-series (black line), negative binomial (green line), Poisson lognormal (red line), and Zipf (purple line). Distributions are for (a) Breeding Bird Survey - Route 36 in New York, (b) Forest Inventory and Analysis - Unit 4, County 57, Plot 12 in Alabama, and (c) Gentry - Araracuara High Campina site in Colombia. Log-likelihoods of the models are included in parenthesis in the legend

## Analysis

Following current best practices for fitting distributions to data and evaluating their fit, we used maximum likelihood estimation to fit models to the data (Clark et al. 1999, Newman 2005, White et al. 2008) and likelihood-based model selection to compare the fits of the different models (Burnham and Anderson 2002, Edwards et al. 2007). This approach has recently been affirmed as best practice for species abundance distributions (Connolly et al. 2014, Matthews and Whittaker 2014). This requires that likelihoods for the models can be solved for and therefore we excluded models that lack probability mass functions and associated likelihoods. While methods have been proposed for comparing models without probability mass functions in this context (Alroy 2015), these methods have not been evaluated to determine how well they perform compared to the widely accepted likelihood-based approaches.

For model comparison we used corrected Akaike Information Criterion (AICc) weights to compare the fits of models while correcting for differences in the number of parameters and appropriately handling the small sample sizes (i.e., numbers of species) in some communities (Burnham and Anderson 2002). The Poisson lognormal and the negative binomial each have two fitted parameters,

while the log-series distribution and the Zipf distributions have one fitted parameter each. The model with the greatest AICc weight in each community was considered to be the best fitting model for that community. We also assessed the full distribution of AICc weights to evaluate the similarity of the fits of the different models.

In addition to evaluating AICc of each model, we also examined the log-likelihood values of the models directly. We did this to assess the fit of the model while ignoring corrections for the number of parameters and the influence of similarities to other models in the set of candidate models. This also allows us to make more direct comparisons to previous analyses that have not corrected for the number of parameters (i.e., Ulrich et al. 2010, Alroy 2015)

Model fitting, log-likelihood, and AICc calculations were performed using Python (Van Rossum and Drake 2011) and R (R Core Team 2015). Python packages used for analysis include numpy (Oliphant 2007, Van Der Walt et al. 2011), matplotlib (Hunter and others 2007), sqlalchemy (Bayer 2014), pandas (McKinney and others 2010), macroecotools (Xiao et al. 2016), retriever (Morris and White 2013), R packages used for analysis include ggplot2 (Wickham 2009), magrittr (Bache and Wickham 2014), tidyr (Wickham 2016), dplyr (Wickham and Francois 2016). All of the code and all of the publicly available data necessary to replicate these analyses is available at <https://github.com/weecology/sad-comparison> and archived on Zenodo (Baldrige et al. 2016). The CBC datasets and NABA datasets are not publicly available and therefore are not included.

## Results

Across all datasets, the negative binomial and Poisson lognormal distributions had very similar average log-likelihoods (within 0.01 of one another; Figure 2). The log-likelihoods for each of these distributions averaged 0.8 units higher than for the log-series distribution and 5 units higher than for the Zipf distribution (corresponding to likelihoods that were twice as high and 140 times as high, respectively).

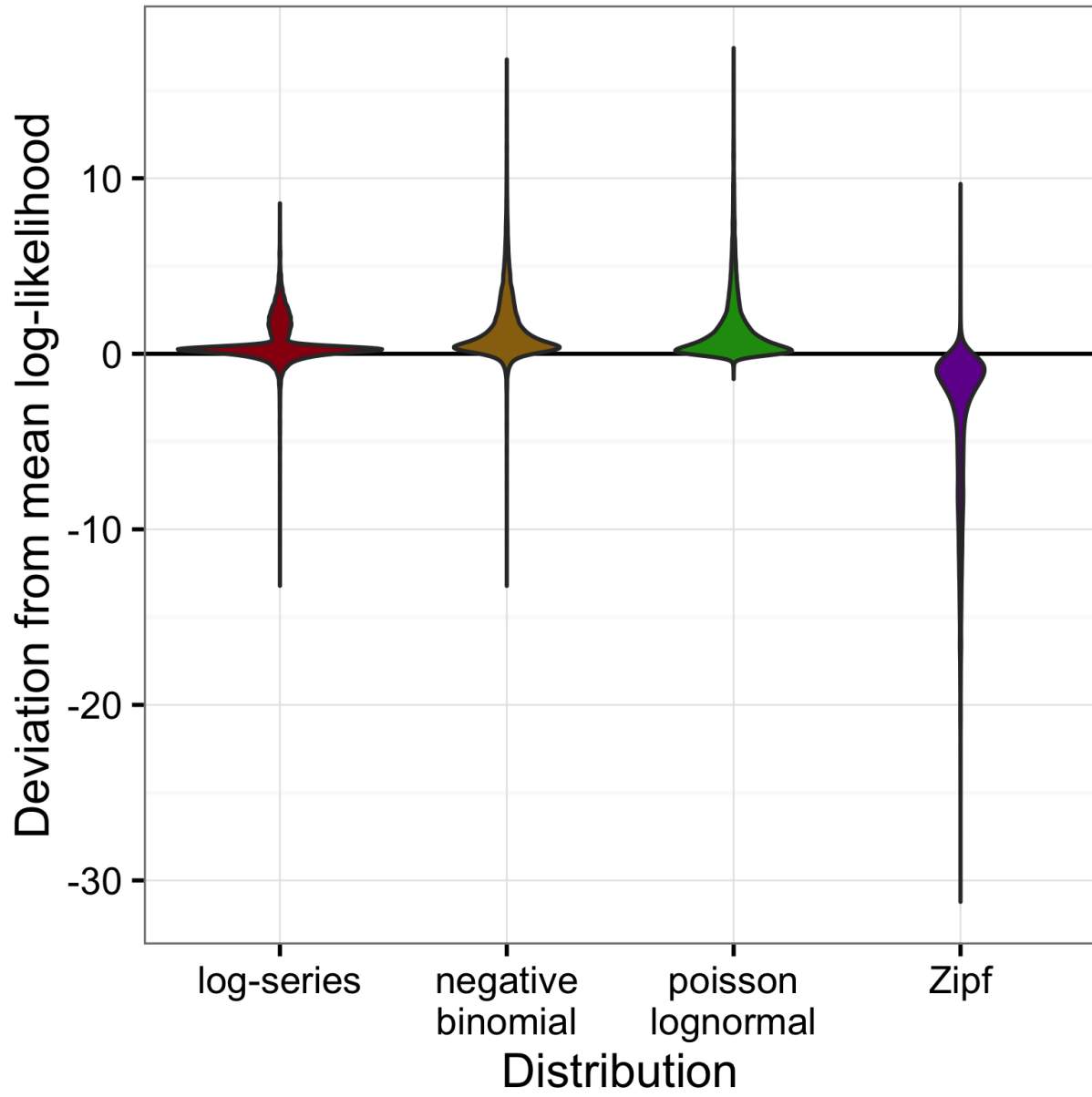


Figure 2: Violin plots of the deviation from the mean log-likelihood for each site for all datasets combined. Positive values indicate that the model fits better than the average fit across the four models.



Although the negative binomial and Poisson lognormal distributions matched the data most closely, the likelihood provides a biased estimate of these distributions' ability to generalize to unobserved species. AICc approximately removes this bias by penalizing models with more degrees of freedom (e.g. the negative binomial and Poisson lognormal distributions, which have two free parameters instead of one like the log-series and Zipf distributions). After applying this penalty, the log-series distribution would be expected to make the best predictions for 69.2% of the sites. The Poisson lognormal and negative binomial distributions were each preferred in about 12% of the sites, and the Zipf distribution was preferred least often (6.0% of sites; Figure 3).

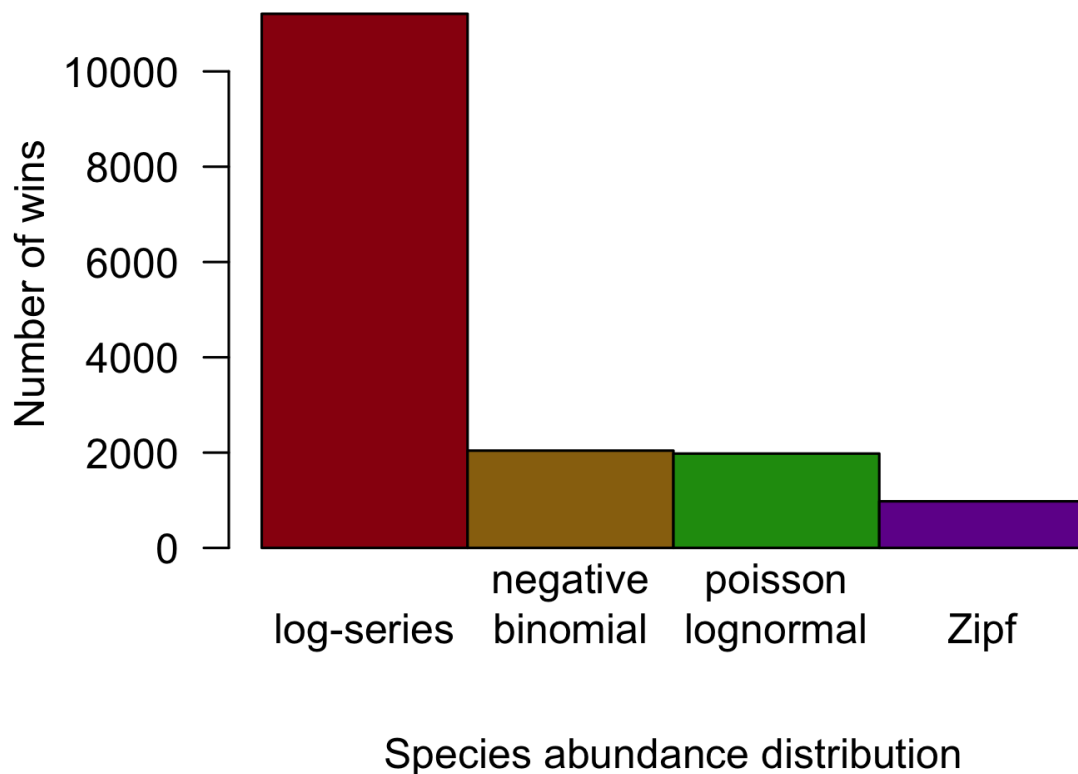


Figure 3: Number of cases in which each model provided the best fit to the data based on AICc for all datasets combined.

Across all datasets and taxonomic groups, the log-series distribution had the highest AICc weights

153 more often than any other model. The negative binomial performed well for BBS, but was almost  
 154 never the best fitting model for plants (FIA and Gentry), butterflies (NABA), Acintopterygii, or  
 155 Coleoptera. The Poisson lognormal performed well for the bird datasets (BBS and CBC) and the  
 156 Gentry tree data, but was almost never best in the FIA and Coleoptera datasets (Figure 4). The  
 157 Zipf distribution only performed consistently well for Arachnida. Because datasets differ in both  
 158 taxonomic groups and sampling methods care should be taken in interpreting these differences.

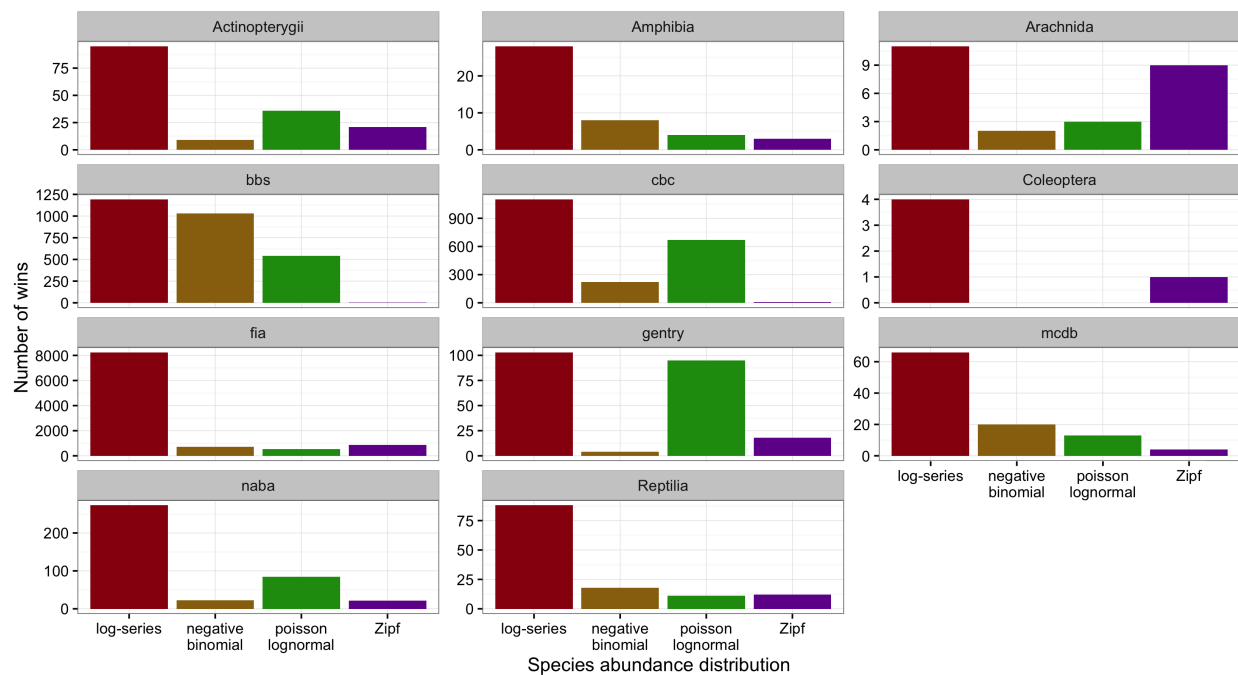


Figure 4: Number of cases in which each model provided the best fit to the data based on AICc for each dataset separately.

159 The full distribution of AICc weights shows separation among models (Figure 5). Although the  
 160 log-series distribution had the best AICc score much more often than the other models, its lead was  
 161 never decisive: across all 16,209 sites, it never had more than about 75% of the AICc weight (Figure  
 162 5). Most of the remaining weight was assigned to the negative binomial and Poisson lognormal  
 163 distributions (each of which usually had at least 12-15% of the weight but was occasionally favored  
 164 very strongly). The Zipf distribution showed a strong mode near zero, and usually had less than 7%  
 165 of the weight.

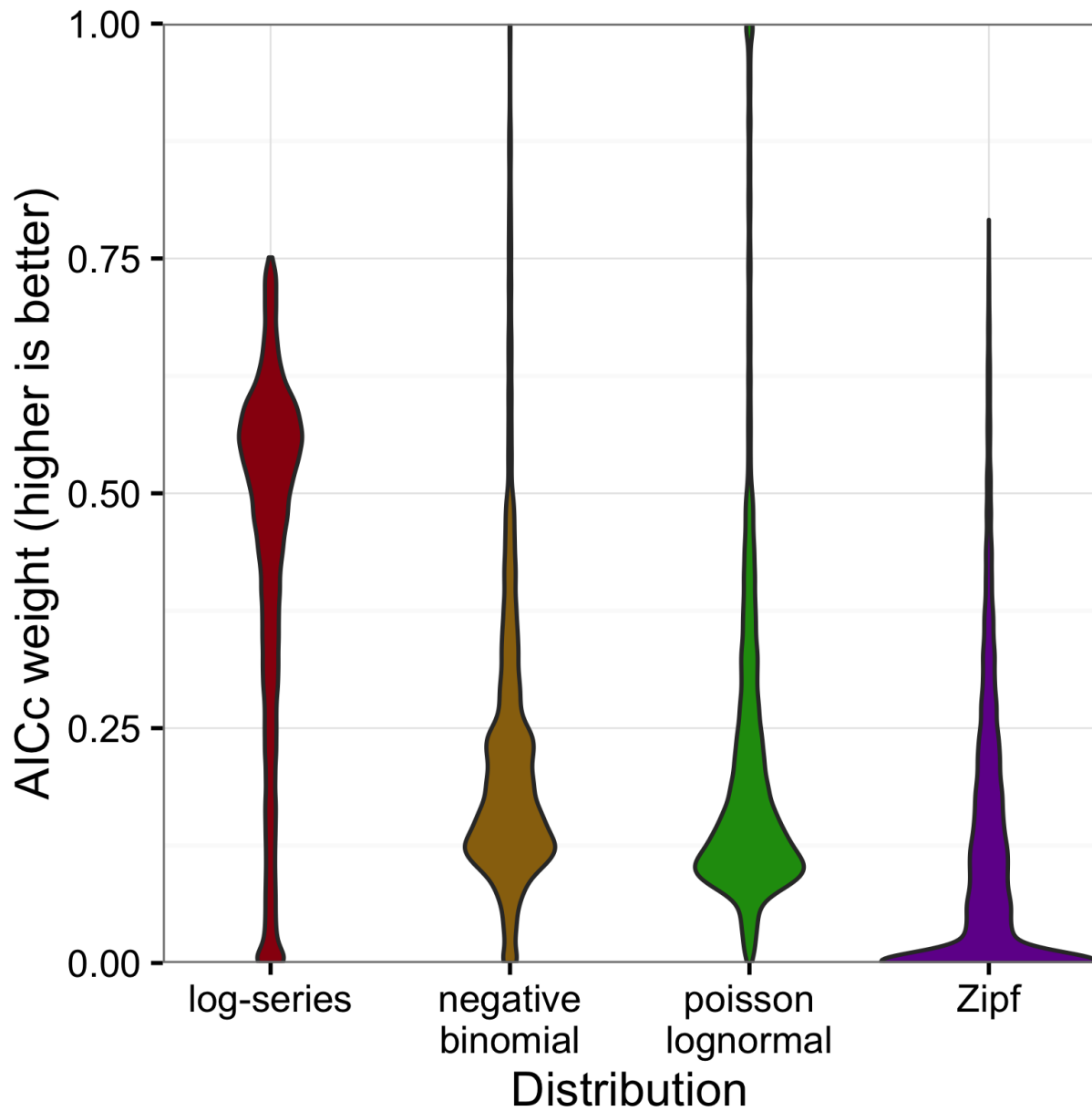


Figure 5: Violin plots of the AICc weights for each model. Weights indicate the probability that the model is the best model for the data

## Discussion

Our extensive comparison of different models for the species abundance distribution (SAD) using rigorous statistical methods demonstrates that several of the most popular existing models provide equivalently good absolute fits to empirical data. Log-series, negative binomial, and Poisson lognormal all had model relative likelihoods between 0.25 and 0.5 suggesting that the three distributions provide roughly equivalent fits in most cases, but with the two-parameter model performing slightly better on average. Because the log-series has only a single parameter but fits the data almost as well as the two-parameter models, the log-series performed better in AICc-based model selection, which penalizes model complexity. These results differ from two other recent analyses of large numbers of species abundance distributions (Ulrich et al. 2010, Connolly et al. 2014) and are generally consistent with a third recent analysis (Alroy 2015).

Ulrich et al. (2010) analyzed ~500 SADs and found support for three major forms of the SAD that changed depending on whether the community had been fully censused or not. They found that “fully censused” communities were best fit by the lognormal, and “incompletely sampled” communities, best fit by the Zipf and log-series (Ulrich et al. 2010). In contrast we find effectively no support for the Zipf across ecosystems and taxonomic groups, including a number of datasets that are incompletely sampled. Our AICc value results also do not support the conclusion that the lognormal outperforms the log-series in fully censused communities. The Gentry and FIA forest inventories both involve large stationary organisms and were collected with the goal of including all trees above a certain stem diameter. Therefore, above the minimum stem diameter, they are as close to fully censused communities as is typically possible. In these communities the log-series provides the best fit to the data most frequently. The discrepancy between our results and those found in (Ulrich et al. 2010) may be due to: 1) their use of binning and fitting curves to rank abundance plots, which deviates from the likelihood-based best practices (Matthews and Whittaker 2014) used in this paper; 2) the statistical methods they use to identify communities as “fully censused”, which tend to exclude communities with large numbers of singletons that would be better fit by distributions like

the log-series; 3) the use of the continuous lognormal instead of the Poisson lognormal; 4) the fact that our censused communities are also a different taxonomic group from our sampled communities, making it difficult to distinguish between taxonomic and sampling differences.

Connolly et al. (2014) use likelihood-based methods to compare the negative binomial distribution (which they call the Poisson gamma) to the Poisson lognormal for a large number of marine communities. They found that the Poisson lognormal provides a substantially better fit than the negative binomial to empirical data and that the negative-binomial provides a better fit to communities simulated using neutral models. They conclude that these analyses of the SAD demonstrate that marine communities are structured by non-neutral processes. Our analysis differs from that in Connolly et al. (2014) in that they aggregate communities at larger spatial scales than those sampled and find the strongest results at large spatial scales. This may explain the difference between the two analyses or there may be differences between the terrestrial systems analyzed here and the marine systems analyzed by Connolly et al. (2014). The explanation for these differences is being explored elsewhere (Connolly et al. unpublished data).

Alroy (2015) compared the fits of the lognormal, log-series, Zipf, geometric series, broken stick, and a new model dubbed the “double geometric”, to over 1000 terrestrial community datasets assembled from the literature. To incorporate the geometric series, broken stick, and the double geometric, this research used non-standard methods for evaluating the fits of the models to the data, however the results were generally consistent with those presented here. The central Kullback-Leibler divergence statistics results showed that: 1) the Zipf, geometric series, and broken stick all perform consistently worse than the other distributions; 2) the double geometric, log-series, and lognormal all provide the best overall fit for at least one taxonomic group; and 3) the lognormal and double geometric fit the data equivalently well and slightly better than the log-series when not controlling for differences in the number of parameters (Alroy’s tables S1, S2, and S3). Penalizing the two-parameter models (lognormal and double geometric) for their complexity, as we do here with AICc, would likewise improve the relative performance of the log-series distribution.

In combination, the results of these three papers suggest that in general the Zipf is a poor characterization of species-abundance distributions and that both the log-series and lognormal distributions provide reasonable fits in many cases. Differences in the performance of the log-series, lognormal, double geometric, and negative binomial, appear to be more minor. How these differences relate to differences in intensity of sampling, spatial scale, taxonomy, and ecosystem type (marine vs. terrestrial) remain open questions. Our analyses suggest that controlling for the number of parameters makes the log-series a slightly better fitting model, at least in the terrestrial systems we studied. Neither of the other papers that include the log-series (Ulrich et al. 2010, Alroy 2015) make this correction and both show that it is still a reasonably competitive model even against those with more parameters.

The relatively similar fit of several commonly used distributions emphasizes the challenge of inferring the processes operating in ecological systems from the form of the abundance distribution. It is already well established that models based on different processes can yield equivalent models of the SAD, i.e., they predict distributions of exactly the same form (Cohen 1968, Boswell and Patil 1971, Pielou 1975, McGill et al. 2007). To the extent that SADs are determined by random statistical processes, one might expect the observed distributions to be compatible with a wide variety of different process-based and process-free models (Frank 2009, 2011, Locey and White 2013). Regardless of the underlying reason that the models performed similarly, our results indicate that the SAD usually does not contain sufficient information to distinguish among the possible statistical processes—let alone biological processes—with any degree of certainty (Volkov et al. 2005), though it is possible that this result differs in marine systems (see Connolly et al. 2014). A more promising way to draw inferences about ecological processes is to evaluate each model's ability to simultaneously explain multiple macroecological patterns, rather than relying on a single pattern like the SAD (McGill 2003, McGill et al. 2006, Newman et al. 2014, Xiao et al. 2015). It has also been suggested that examining second-order effects, such as the scale-dependence of macroecological patterns (Blonder et al. 2014) or how the parameters of the distribution change across gradients (Mac Nally et al. 2014), can provide better inference about process from these

kinds of pattern.

## Acknowledgments

We thank all of the individuals involved in the collection and provision of the data used in this paper, including the citizen scientists who collect the BBS, CBC, and NABC data, the USGS and CWS scientists and managers, the Audubon Society, the North American Butterfly Association, the USDA Forest Service, the Missouri Botanical Garden, and Alwyn H. Gentry. We also thank all of the scientists who published their raw data allowing it to be combined in Baldrige (2013).

## References

- Alroy, J. 2015. The shape of terrestrial abundance distributions. *Science advances* 1:e1500082.
- Bache, S. M., and H. Wickham. 2014. Magrittr: A forward-pipe operator for r.
- Baldrige, E. 2013. Community abundance data.
- Baldrige, E., D. J. Harris, X. Xiao, and E. P. White. 2016, November. weecology/sad-comparison: First submission to PeerJ.
- Bayer, M. 2014. Ssqlalchemy. *The Architecture of Open Source Applications: Elegance, Evolution, and a Few More Fearless Hacks* 2.
- Blonder, B., L. Sloat, B. J. Enquist, and B. McGill. 2014. Separating macroecological pattern and process: Comparing ecological, economic, and geological systems. *PloS one* 9:e112850.
- Boswell, M., and G. Patil. 1971. Chance mechanisms generating the logarithmic series distribution used in the analysis of number of species and individuals. *Statistical ecology* 1:99–130.
- Bulmer, M. 1974. On fitting the poisson lognormal distribution to species-abundance data.

265 Biometrics:101–110.

266 Burnham, K. P., and D. R. Anderson. 2002. Model selection and multimodel inference: A practical  
267 information-theoretic approach. Springer.

268 Clark, R., S. Cox, and G. Laslett. 1999. Generalizations of power-law distributions applicable to  
269 sampled fault-trace lengths: Model choice, parameter estimation and caveats. *Geophysical Journal*  
270 *International* 136:357–372.

271 Cohen, J. E. 1968. Alternate derivations of a species-abundance relation. *American naturalist*:165–  
272 172.

273 Connolly, S. R., M. A. MacNeil, M. J. Caley, N. Knowlton, E. Cripps, M. Hisano, L. M. Thibaut, B.  
274 D. Bhattacharya, L. Benedetti-Cecchi, R. E. Brainard, and others. 2014. Commonness and rarity in  
275 the marine biosphere. *Proceedings of the National Academy of Sciences*:8524–8529.

276 Edwards, A. M., R. A. Phillips, N. W. Watkins, M. P. Freeman, E. J. Murphy, V. Afanasyev, S. V.  
277 Buldyrev, M. G. da Luz, E. P. Raposo, H. E. Stanley, and others. 2007. Revisiting lévy flight search  
278 patterns of wandering albatrosses, bumblebees and deer. *Nature* 449:1044–1048.

279 Engen, S., and R. Lande. 1996. Population dynamic models generating species abundance distribu-  
280 tions of the gamma type. *Journal of Theoretical Biology* 178:325–331.

281 Fisher, R. A., A. S. Corbet, and C. B. Williams. 1943. The relation between the number of species  
282 and the number of individuals in a random sample of an animal population. *The Journal of Animal*  
283 *Ecology*:42–58.

284 Frank, S. A. 2009. The common patterns of nature. *Journal of evolutionary biology* 22:1563–1585.

285 Frank, S. A. 2011. Measurement scale in maximum entropy models of species abundance. *Journal*  
286 *of evolutionary biology* 24:485–496.

287 Harte, J. 2011. Maximum entropy and ecology: A theory of abundance, distribution, and energetics.



288 Oxford University Press.

289 Harte, J., T. Zillio, E. Conlisk, and A. Smith. 2008. Maximum entropy and the state-variable  
290 approach to macroecology. *Ecology* 89:2700–2711.

291 Hubbell, S. P. 2001. The unified neutral theory of biodiversity and biogeography (mPB-32).  
292 Princeton University Press.

293 Hunter, J. D., and others. 2007. Matplotlib: A 2D graphics environment. *Computing in science and*  
294 *engineering* 9:90–95.

295 Locey, K. J., and E. P. White. 2013. How species richness and total abundance constrain the  
296 distribution of abundance. *Ecology letters* 16:1177–1185.

297 Mac Nally, R., C. A. McAlpine, H. P. Possingham, and M. Maron. 2014. The control of rank-  
298 abundance distributions by a competitive despotic species. *Oecologia* 176:849–857.

299 Matthews, T. J., and R. J. Whittaker. 2014. Fitting and comparing competing models of the species  
300 abundance distribution: Assessment and prospect. *Frontiers of Biogeography* 6.

301 May, R. M. 1975. Patterns of species abundance and diversity. *Ecology and evolution of*  
302 *communities*:81–120.

303 McGill, B. J. 2003. A test of the unified neutral theory of biodiversity. *Nature* 422:881–885.

304 McGill, B. J., R. S. Etienne, J. S. Gray, D. Alonso, M. J. Anderson, H. K. Benecha, M. Dornelas,  
305 B. J. Enquist, J. L. Green, F. He, and others. 2007. Species abundance distributions: Moving  
306 beyond single prediction theories to integration within an ecological framework. *Ecology letters*  
307 10:995–1015.

308 McGill, B. J., B. A. Maurer, and M. D. Weiser. 2006. Empirical evaluation of neutral theory.  
309 *Ecology* 87:1411–1423.

310 McGill, B., and C. Collins. 2003. A unified theory for macroecology based on spatial patterns of

311 abundance. *Evolutionary Ecology Research* 5:469–492.

312 McKinney, W., and others. 2010. Data structures for statistical computing in python. Pages 51–56  
 313 *in* Proceedings of the 9th python in science conference.

314 Morlon, H., E. P. White, R. S. Etienne, J. L. Green, A. Ostling, D. Alonso, B. J. Enquist, F. He,  
 315 A. Hurlbert, A. E. Magurran, and others. 2009. Taking species abundance distributions beyond  
 316 individuals. *Ecology Letters* 12:488–501.

317 Morris, B. D., and E. P. White. 2013. The ecoData retriever: Improving access to existing ecological  
 318 data. *PloS one* 8:e65848.

319 Newman, E. A., M. E. Harte, N. Lowell, M. Wilber, and J. Harte. 2014. Empirical tests of within-and  
 320 across-species energetics in a diverse plant community. *Ecology* 95:2815–2825.

321 Newman, M. E. 2005. Power laws, pareto distributions and zipf’s law. *Contemporary physics*  
 322 46:323–351.

323 North American Butterfly Assoc. 2009. NABA butterfly counts: 2009 report. NABA, Morristown,  
 324 New Jersey, USA.

325 Oliphant, T. E. 2007. Python for scientific computing. *Computing in Science & Engineering*  
 326 9:10–20.

327 Pardieck, K. L., D. J. Ziolkowski Jr, and M.-A. Hudson. 2014. North american breeding bird survey  
 328 dataset 1966 - 2013, version 2013.0. U.S. Geological Survey, Patuxent Wildlife Research Center.

329 Phillips, O., and J. S. Miller. 2002. Global patterns of plant diversity: Alwyn h. gentry’s forest  
 330 transect data set. Missouri Botanical Garden Press St., Louis, Missouri.

331 Pielou, E. 1975. *Ecological diversity*. Wiley, New York.

332 Pueyo, S., F. He, and T. Zillio. 2007. The maximum entropy formalism and the idiosyncratic theory

333 of biodiversity. *Ecology Letters* 10:1017–1028.

334 R Core Team. 2015. R: A language and environment for statistical computing. R Foundation for  
 335 Statistical Computing, Vienna, Austria.

336 Society, N. A. 2002. The christmas bird count historical results. National Audobon Society, New  
 337 York, New York, USA.

338 Sugihara, G. 1980. Minimal community structure: An explanation of species abundance patterns.  
 339 *American naturalist*:770–787.

340 Thibault, K. M., S. R. Supp, M. Giffin, E. P. White, and S. M. Ernest. 2011. Species composition  
 341 and abundance of mammalian communities: *Ecological archives* e092-201. *Ecology* 92:2316–2316.

342 Tokeshi, M. 1993. Species abundance patterns and community structure. *Advances in ecological*  
 343 *research* 24:111–186.

344 Ulrich, W., M. Ollik, and K. I. Ugland. 2010. A meta-analysis of species–abundance distributions.  
 345 *Oikos* 119:1149–1155.

346 USDA Forest Service. 2010. Forest inventory and analysis national core field guide (phase 2 and 3).  
 347 version 4.0. USDA Forest Service, Forest Inventory; Analysis.

348 Van Der Walt, S., S. C. Colbert, and G. Varoquaux. 2011. The numPy array: A structure for efficient  
 349 numerical computation. *Computing in Science & Engineering* 13:22–30.

350 Van Rossum, G., and F. L. Drake. 2011. The python language reference manual. Network Theory  
 351 Ltd.

352 Volkov, I., J. R. Banavar, F. He, S. P. Hubbell, and A. Maritan. 2005. Density dependence explains  
 353 tree species abundance and diversity in tropical forests. *Nature* 438:658–661.

354 Volkov, I., J. R. Banavar, S. P. Hubbell, and A. Maritan. 2003. Neutral theory and relative species

355 abundance in ecology. *Nature* 424:1035–1037.

356 White, E. P., B. J. Enquist, and J. L. Green. 2008. On estimating the exponent of power-law  
357 frequency distributions. *Ecology* 89:905–912.

358 White, E. P., K. M. Thibault, and X. Xiao. 2012. Characterizing species abundance distributions  
359 across taxa and ecosystems using a simple maximum entropy model. *Ecology* 93:1772–1778.

360 Wickham, H. 2009. *Ggplot2: Elegant graphics for data analysis*. Springer-Verlag New York.

361 Wickham, H. 2016. *Tidyr: Easily tidy data with ‘spread()’ and ‘gather()’ functions*.

362 Wickham, H., and R. Francois. 2016. *Dplyr: A grammar of data manipulation*.

363 Xiao, X., D. J. McGlinn, and E. P. White. 2015. A strong test of the maximum entropy theory of  
364 ecology. *American Naturalist*.

365 Xiao, X., K. Thibault, D. J. Harris, E. Baldrige, and E. White. 2016. *Weecology/macroecotools*:  
366 V0.4.0.