

Classics

Conditional Expectations and Random Sum	Branching Process and Generating Functions
<ul style="list-style-type: none"> $\mathbb{P}[X = x] = \sum_i \mathbb{P}[X = x Y = i] \mathbb{P}[Y = i]$ $f_X(x) = \sum_i f_{X N}(x n) \mathbb{P}[N = n]$ $\mathbb{E}[X] = \sum_i \mathbb{E}[X Y] \mathbb{P}[Y = i]$ <ul style="list-style-type: none"> $\mathbb{E}[X] = \mathbb{E}[\mathbb{E}[X Y]]$ $\text{Var}[X] = \text{Var}[\mathbb{E}[X Y]] + \mathbb{E}[\text{Var}[X Y]]$ $\mathbb{P}[\cup_i A_i] \leq \sum_i \mathbb{P}[A_i]$ Let $S = \sum_{i=1}^N X_i$ where N, X_i independent, then $\mathbb{E}[S] = \mathbb{E}[N] \mathbb{E}[X_i]$ If X, Y independent and $Z = X + Y$, then $f_Z(z) = \int_{-\infty}^{\infty} f_X(x) f_Y(z-x) dx$ 	<ul style="list-style-type: none"> $X_{n+1} = \sum_{i=1}^{X_n} \xi_i^{(n)}$ with $X_0 = 1$ $\mathbb{E}[X_{n+1}] = \mu \mathbb{E}[X_n]$ where $\mu = \mathbb{E}[\xi]$ $\mathbb{E}[X_n] = \mu^n$ $\text{Var}[X_{n+1}] = \sigma^2 \mathbb{E}[X_n] + \mu^2 \text{Var}[X_n]$ $\text{Var}[X_n] = \begin{cases} \sigma^2 n, & \mu = 1 \\ \sigma^2 \mu^{n-1} \frac{1-\mu^n}{1-\mu}, & \mu \neq 1 \end{cases}$ [Time of Extinction] $N = \min_n \{X_n = 0\}$ [Extinction Probability] $u_n = \mathbb{P}[N \leq n] = \mathbb{P}[X_n = 0]$ i.e. $u_0 = 0, u_1 = p_0$ $0 \leq u_n \leq u_{n+1} \leq 1 \forall n$ (monotonicity) $u_n = \sum_{k=0}^{\infty} p_k (u_{n-1})^k = \phi_{\xi}(u_{n-1})$ u_{∞} is the smallest solution to $u = \phi_{\xi}(u)$ $\phi_{X_{n+1}}(s) = \phi_{X_n}(\phi_{\xi}(s))$ If $\phi_{X_0}(s) = s, \phi_{X_n}(s) = \phi_{\xi}^n(s)$ [Pure Death Process] $\phi_{\xi}(s) = (1-p) + ps; \phi_{X_n}(s) = 1 - p^n + p^n s$ [Time of extinction] $\mathbb{P}[T = n X_0 = k] = \mathbb{P}[X_n = 0 X_0 = k] - \mathbb{P}[X_{n-1} = 0 X_0 = k] = \phi_{X_n}(0) - \phi_{X_{n-1}}(0)$
Random Walk	
<ul style="list-style-type: none"> [Stirling] $n! \sim n^n e^{-n} \sqrt{2\pi n}$ $[\mathbb{Z}^1] P_{0,0}^{(2n)} = \binom{2n}{n} p^n q^n \sim \frac{(4pq)^n}{\sqrt{\pi n}}$ Only recurrent when $p = q = \frac{1}{2}$ $[\mathbb{Z}^2] P_{0,0}^{(2n)} = \binom{2n}{n} p^n q^n \sim \frac{(4pq)^n}{\sqrt{\pi n}}$ Only recurrent when symmetric $[\mathbb{Z}^3]$ All transient 	
Gambler's Ruin	
<ul style="list-style-type: none"> $[p = q = \frac{1}{2}] \mathbb{P}[X_T = 0 X_0 = k] = \frac{N-k}{N}$ $\mathbb{E}[T X_0 = k] = k(N-k)$ $[p \neq \frac{1}{2}] \mathbb{P}[X_T = 0 X_0 = i] = \left(\frac{q}{p}\right)^k \frac{1 - \left(\frac{q}{p}\right)^{N-k}}{1 - \left(\frac{q}{p}\right)^N}$ 	
Final Checks and Last Resorts	
<ul style="list-style-type: none"> Check you modelled the problem correctly Return what the problem asked for First-step and previous-step analysis Understand dynamics of MC first Method of differences GFs: go back to first principle If have time, rigorously prove MC with limits 	
Markov Chain (Definitions)	Markov Chain (Theorems)
<ul style="list-style-type: none"> [Regularity] P regular if $\exists n$ s.t. $P_{i,j}^n > 0 \forall i, j$ [Stationary Distribution] π s.t. $\pi P = \pi$ [Limiting Distribution] $\lim_{n \rightarrow \infty} \tau P^n = \pi \forall \tau$ [Accessible] $i \rightarrow j \Rightarrow \exists n > 0$ s.t. $P_{i,j}^{(n)} > 0$ [Communicate] $i \leftrightarrow j \Leftrightarrow i \rightarrow j, j \rightarrow i$ 	<ul style="list-style-type: none"> Period and recurrence/transience are class properties i.e. $i \leftrightarrow j \Rightarrow d(i) = d(j)$ $\exists N$ s.t. $\forall n > N, P_{i,i}^{(nd(i))} > 0$ [Finite] Regular finite MC have unique limiting distribution.

- [Irreducible] $\forall i, j \in S, \exists n$ s.t. $P_{i,j}^{(n)} > 0$
- [Period] $d(i) = \gcd\{n | P_{ii}^{(n)} > 0\}$
 - If $P_{ii}^{(n)} = 0 \forall n$, define $d(i) = 0$.
 - [Aperiodic] $d(i) = 1$
- [Probability of return in n steps $f_{i,i}^{(n)}$]
 $f_{i,i}^{(n)} := \mathbb{P}[X_n = i, X_{n-1} \neq i, \dots, X_1 \neq i | X_0 = i]$
- [Probability of return $f_{i,i}$]

$$f_{i,i} = \sum_{k=0}^{\infty} f_{i,i}^{(k)}$$

- [Recurrent] State i recurrent $\Leftrightarrow f_{i,i} = 1$
 - $\Leftrightarrow \mathbb{P}[R_i < \infty | X_0 = i] = 1$
 - $\Leftrightarrow \sum_{m=0}^{\infty} P_{i,i}^{(m)} = \infty$
 - $\Leftrightarrow \lim_{m \rightarrow \infty} \prod_{j=0}^m (1 - P_{i,i}^{(j)}) = 0$
- [Transient] State i transient $\Leftrightarrow f_{i,i} < 1$
 - $\Leftrightarrow \mathbb{P}[R_i < \infty | X_0 = i] < 1$
 - $\Leftrightarrow \sum_{m=0}^{\infty} P_{i,i}^{(m)} < \infty$
 - $\Leftrightarrow \lim_{m \rightarrow \infty} \prod_{j=0}^m (1 - P_{i,i}^{(j)}) > 0$
- $R_i = \min\{n > 0 | X_n = i\}$
- $m_i = \mathbb{E}[R_i | X_0 = i]$
- [Positive Recurrent] $m_i < \infty$
- [Null Recurrent] $m_i = \infty$
- M_i is number of visits to state i
- $M_i = \sum_{k=1}^{\infty} \mathbb{1}\{X_k = i\}$
- $\mathbb{E}[M_i | X_0 = i] = \sum_{k=1}^{\infty} (f_{i,i})^k$

- [Infinite/Mean Return Time] Let $(X_n)_{n \geq 0}$ be an irreducible aperiodic recurrent MC, then $\lim_{n \rightarrow \infty} P_{ii}^{(n)} = \frac{1}{m_i}$ and $\lim_{n \rightarrow \infty} P_{ji}^{(n)} = \frac{1}{m_i}$
- [Infinite/Limiting] In an irreducible, aperiodic, positive recurrent MC, \exists unique limiting distribution π (satisfying $\pi P = \pi$)
- [Infinite/Positive] In a positive recurrent aperiodic class, $\lim_{n \rightarrow \infty} P_{jj}^{(n)} = \pi_j = \sum_{i=0}^{\infty} \pi_i P_{ij}$ and π is uniquely determined by the equations $\pi_i \geq 0$, $\sum_{i=0}^{\infty} \pi_i = 1$ and $\pi_j = \sum_{i=0}^{\infty} \pi_i P_{ij}$
- [Infinite/Transient] In an irreducible, aperiodic, transient MC, $\lim_{n \rightarrow \infty} \mathbb{P}[X_n = j | X_0 = i] = 0 \forall i, j$
- [Infinite/Null] In an irreducible, aperiodic, null recurrent MC, $\lim_{n \rightarrow \infty} \mathbb{P}[X_n = j | X_0 = i] = 0 \forall i, j$
- [Infinite/Periodic] In an irreducible, periodic MC with period d , $P_{i,i}^{(m)} = 0$ if $d \nmid m$. Else, $\lim_{n \rightarrow \infty} P_{i,i}^{nd} = \frac{d}{m_i}$