

Prep: bring ID, water, glasses, jacket, pen and the handwritten version of this set of notes

**You got this!**

Distributions	Distribution (Hacks)								
<b>Bernoulli Distribution:</b> $X \sim \text{Bernoulli}(p)$ <table><tr><td><math>\mathbb{P}[X = 0] = (1 - p)</math></td><td><math>\mathbb{E}[X] = p</math></td></tr><tr><td><math>\mathbb{P}[X = 1] = p</math></td><td><math>\text{Var}[X] = p(1 - p)</math></td></tr></table>	$\mathbb{P}[X = 0] = (1 - p)$	$\mathbb{E}[X] = p$	$\mathbb{P}[X = 1] = p$	$\text{Var}[X] = p(1 - p)$	Trick: $\min(X_1, X_2, \dots) \sim \text{Geometric}(1 - \prod(1 - p_i))$				
$\mathbb{P}[X = 0] = (1 - p)$	$\mathbb{E}[X] = p$								
$\mathbb{P}[X = 1] = p$	$\text{Var}[X] = p(1 - p)$								
<b>Binomial Distribution:</b> $X \sim \text{Binom}(n, p)$ <table><tr><td colspan="2"><math>\mathbb{P}[X = i] = \binom{n}{i} p^i (1 - p)^{n-i}, i = 0, 1, \dots, n</math></td></tr><tr><td><math>\mathbb{E}[X] = np</math></td><td><math>\text{Var}[X] = np(1 - p)</math></td></tr></table>	$\mathbb{P}[X = i] = \binom{n}{i} p^i (1 - p)^{n-i}, i = 0, 1, \dots, n$		$\mathbb{E}[X] = np$	$\text{Var}[X] = np(1 - p)$	$\lambda$ is like expected number of success $X \sim \text{Binom}\left(n, \frac{\lambda}{n}\right)$ $\mathbb{P}[X = i] \rightarrow \left(\frac{\lambda^i}{i!}\right) e^{-\lambda}$ as $n \rightarrow \infty$				
$\mathbb{P}[X = i] = \binom{n}{i} p^i (1 - p)^{n-i}, i = 0, 1, \dots, n$									
$\mathbb{E}[X] = np$	$\text{Var}[X] = np(1 - p)$								
<b>Geometric Distribution:</b> $X \sim \text{Geometric}(p)$ <table><tr><td colspan="2"><math>\mathbb{P}[X = i] = (1 - p)^{i-1} p, i = 1, 2, \dots</math></td></tr><tr><td><math>\mathbb{E}[X] = \frac{1}{p}</math></td><td><math>\text{Var}[X] = \frac{1 - p}{p^2}</math></td></tr><tr><td colspan="2"><math>pgf(x) = \frac{px}{1 - (1 - p)x}</math></td></tr></table>	$\mathbb{P}[X = i] = (1 - p)^{i-1} p, i = 1, 2, \dots$		$\mathbb{E}[X] = \frac{1}{p}$	$\text{Var}[X] = \frac{1 - p}{p^2}$	$pgf(x) = \frac{px}{1 - (1 - p)x}$		<b>Memoryless Property</b> $\mathbb{P}[X \geq x + y   X \geq x] = \mathbb{P}[X \geq y]$		
$\mathbb{P}[X = i] = (1 - p)^{i-1} p, i = 1, 2, \dots$									
$\mathbb{E}[X] = \frac{1}{p}$	$\text{Var}[X] = \frac{1 - p}{p^2}$								
$pgf(x) = \frac{px}{1 - (1 - p)x}$									
<b>Poisson Distribution:</b> $X \sim \text{Poisson}(\lambda)$ <table><tr><td colspan="2"><math>\mathbb{P}[X = i] = \frac{\lambda^i}{i!} e^{-\lambda}, i = 0, 1, 2, \dots</math></td></tr><tr><td><math>\mathbb{E}[X] = \lambda</math></td><td><math>\text{Var}[X] = \lambda</math></td></tr><tr><td colspan="2"><math>X + Y \sim \text{Poisson}(\lambda + \mu)</math></td></tr></table>	$\mathbb{P}[X = i] = \frac{\lambda^i}{i!} e^{-\lambda}, i = 0, 1, 2, \dots$		$\mathbb{E}[X] = \lambda$	$\text{Var}[X] = \lambda$	$X + Y \sim \text{Poisson}(\lambda + \mu)$		<b>Useful trick:</b> $\min(X_1, X_2, \dots) \sim \text{Expo}(\sum \lambda_i)$ $\mathbb{P}[X < Y] = \frac{\lambda_X}{\lambda_X + \lambda_Y}$		
$\mathbb{P}[X = i] = \frac{\lambda^i}{i!} e^{-\lambda}, i = 0, 1, 2, \dots$									
$\mathbb{E}[X] = \lambda$	$\text{Var}[X] = \lambda$								
$X + Y \sim \text{Poisson}(\lambda + \mu)$									
<b>Exponential Distribution:</b> $X \sim \text{Expo}(\lambda)$ <table><tr><td colspan="2"><math>f(x) = \begin{cases} \lambda e^{-\lambda x}, &amp; x \geq 0 \\ 0, &amp; \text{otherwise} \end{cases}</math></td></tr><tr><td><math>\mathbb{E}[X] = \frac{1}{\lambda}</math></td><td><math>\text{Var}[X] = \frac{1}{\lambda^2}</math></td></tr><tr><td colspan="2"><math>F(x) = \mathbb{P}[X \leq x] = 1 - e^{-\lambda x}</math></td></tr></table>	$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}$		$\mathbb{E}[X] = \frac{1}{\lambda}$	$\text{Var}[X] = \frac{1}{\lambda^2}$	$F(x) = \mathbb{P}[X \leq x] = 1 - e^{-\lambda x}$		$\lambda$ : success rate per unit time		
$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}$									
$\mathbb{E}[X] = \frac{1}{\lambda}$	$\text{Var}[X] = \frac{1}{\lambda^2}$								
$F(x) = \mathbb{P}[X \leq x] = 1 - e^{-\lambda x}$									
<b>Uniform Continuous:</b> $X \sim \text{Uniform}([a, b])$ <table><tr><td><math>\mathbb{E}[X] = \frac{a + b}{2}</math></td><td><math>\text{Var}[X] = \frac{(b - a)^2}{12}</math></td></tr><tr><td><math>f(x) = \frac{1}{b - a}</math></td><td><math>F(x) = \frac{x - a}{b - a}</math></td></tr></table>	$\mathbb{E}[X] = \frac{a + b}{2}$	$\text{Var}[X] = \frac{(b - a)^2}{12}$	$f(x) = \frac{1}{b - a}$	$F(x) = \frac{x - a}{b - a}$	<b>Functions on Random Variables</b>				
$\mathbb{E}[X] = \frac{a + b}{2}$	$\text{Var}[X] = \frac{(b - a)^2}{12}$								
$f(x) = \frac{1}{b - a}$	$F(x) = \frac{x - a}{b - a}$								
<b>Normal Distribution:</b> $X \sim N(\mu, \sigma^2)$ <table><tr><td colspan="2"><math>X \sim N(\mu, \sigma^2) \rightarrow Z = \frac{X - \mu}{\sigma} \sim N(0, 1)</math></td></tr><tr><td><math>f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}</math></td><td><math>f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}</math></td></tr></table> <p>* No closed formula for CDF.</p>	$X \sim N(\mu, \sigma^2) \rightarrow Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$		$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$	$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$	<b>Covariance (bilinear)</b> <table><tr><td><math>\text{Cov}[X, Y] = E[(X - \mu_X)(Y - \mu_Y)] = E[XY] - E[X]E[Y]</math></td></tr><tr><td><math>\text{Cov}[X, X] = \text{Var}[X]</math></td></tr><tr><td><math>\text{Var}[X + Y] = \text{Var}[X] + \text{Var}[Y] + 2 \text{Cov}[X, Y]</math></td></tr><tr><td><math>\text{Cov}[aX_1 + bX_2, cY_1 + dY_2]</math> <math>= ac \text{Cov}[X_1, Y_1] + ad \text{Cov}[X_2, Y_1]</math> <math>+ bc \text{Cov}[X_2, Y_1] + bd \text{Cov}[X_2, Y_2]</math></td></tr></table>	$\text{Cov}[X, Y] = E[(X - \mu_X)(Y - \mu_Y)] = E[XY] - E[X]E[Y]$	$\text{Cov}[X, X] = \text{Var}[X]$	$\text{Var}[X + Y] = \text{Var}[X] + \text{Var}[Y] + 2 \text{Cov}[X, Y]$	$\text{Cov}[aX_1 + bX_2, cY_1 + dY_2]$ $= ac \text{Cov}[X_1, Y_1] + ad \text{Cov}[X_2, Y_1]$ $+ bc \text{Cov}[X_2, Y_1] + bd \text{Cov}[X_2, Y_2]$
$X \sim N(\mu, \sigma^2) \rightarrow Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$									
$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$	$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$								
$\text{Cov}[X, Y] = E[(X - \mu_X)(Y - \mu_Y)] = E[XY] - E[X]E[Y]$									
$\text{Cov}[X, X] = \text{Var}[X]$									
$\text{Var}[X + Y] = \text{Var}[X] + \text{Var}[Y] + 2 \text{Cov}[X, Y]$									
$\text{Cov}[aX_1 + bX_2, cY_1 + dY_2]$ $= ac \text{Cov}[X_1, Y_1] + ad \text{Cov}[X_2, Y_1]$ $+ bc \text{Cov}[X_2, Y_1] + bd \text{Cov}[X_2, Y_2]$									
<b>Miscellaneous Hacks</b> <div><math>\int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} dx = \sqrt{2\pi}</math><math display="block">\mathbb{P}[A B] = \frac{\mathbb{P}[B A]\mathbb{P}[A]}{\mathbb{P}[B]}</math><b>Discrete Tail Sum (<math>X</math> nonnegative)</b><math display="block">\mathbb{E}[X] = \sum_{i=0}^{\infty} \mathbb{P}[X &gt; i]</math><b>Continuous Tail sum (<math>Z</math> nonnegative)</b><math display="block">\mathbb{E}[Z] = \int_0^{\infty} \mathbb{P}[Z \geq z] dz</math></div>	<b>Probability Density Hacks</b> <div><math display="block">\text{cdf}_X(x) = \mathbb{P}[X \leq x]</math><math display="block">\text{pdf}_X(x) = \frac{d}{dx} \text{cdf}_X \big _x</math><b>Quick hacks:</b><math display="block">\mathbb{P}[X \leq Y] = \int_{-\infty}^{\infty} \mathbb{P}[x \leq X \leq x + dx] \mathbb{P}[x \leq Y]</math><math display="block">\mathbb{E}[g(X)] = \int_{-\infty}^{\infty} g(x) \mathbb{P}[X = x] = \int_{-\infty}^{\infty} g(x) f_X(x) dx</math><math display="block">n! \sim \sqrt{2\pi n} \left(\frac{n}{e}\right)^n</math></div>								

\* No closed formula for CDF.

Properties of Conditional Expectations	Probabilistic Bounding		
$\mathbb{E}[Y X] = f(X)$ <ol style="list-style-type: none"> <li>(Linearity)  <math display="block">\mathbb{E}[a_1 Y_1 + a_2 Y_2   X] = a_1 \mathbb{E}[Y_1   X] + a_2 \mathbb{E}[Y_2   X]</math> </li> <li>(Factoring Known Values)  <math display="block">\mathbb{E}[h(X)Y   X] = h(X)\mathbb{E}[Y   X]</math> </li> <li>(Smoothing)  <math display="block">\mathbb{E}[\mathbb{E}[Y   X]] = \mathbb{E}[Y]</math> </li> <li>(Independence) If <math>X, Y</math> independent:  <math display="block">\mathbb{E}[Y   X] = \mathbb{E}[Y]</math> </li> </ol>	<p><b>Definitions:</b></p> $\mathbb{P}[ \hat{p} - p  \geq \varepsilon] \leq 1 - \delta$ <p><math>\varepsilon</math>: error / accuracy / tolerance; <math>1 - \delta</math>: confidence</p> <p><b>Toolbox:</b></p> <ul style="list-style-type: none"> <li>[Markov] Nonnegative RV <math>X</math>, finite mean  <math display="block">\mathbb{P}[X \geq c] \leq \frac{\mathbb{E}[X]}{c}, \quad c &gt; 0</math> </li> <li>[Generalized Markov] <math>Y</math> not necessarily nonnegative, finite mean; <math>c, r &gt; 0</math>  <math display="block">\mathbb{P}[ Y  \geq c] \leq \frac{\mathbb{E}[ Y ^r]}{c^r}</math> </li> <li>[Extended Markov] <math>X</math> not necessarily nonnegative; <math>\Phi(X)</math> nonnegative function, monotonically increasing for <math>x &gt; 0</math>; <math>\alpha &gt; 0</math>  <math display="block">\mathbb{P}[X \geq \alpha] \leq \frac{\mathbb{E}[\Phi(X)]}{\Phi(\alpha)}</math> </li> <li>[Chebyshev] <math>c &gt; 0</math>  <math display="block">\mathbb{P}[ X - \mu  \geq c] \leq \frac{\text{Var}[X]}{c^2}</math> <math display="block">\mathbb{P}[ X - \mu  \geq k\sigma] \leq \frac{1}{k^2}</math> </li> <li>[Cantelli] <math>\alpha &gt; 0</math>  <math display="block">\mathbb{P}[X - \mathbb{E}[X] \geq \alpha] \leq \frac{\sigma^2}{\alpha^2 + \sigma^2}</math> </li> <li>[Law of Large Numbers] <math>X_1, \dots, X_n</math> i.i.d. RV with <math>\mathbb{E}[X_i] = \mu &lt; \infty</math>. Define <math>S_n = X_1 + \dots + X_n</math>  <math display="block">\forall \varepsilon \quad \lim_{n \rightarrow \infty} \mathbb{P}\left[\left \frac{1}{n}S_n - \mu\right  &lt; \varepsilon\right] = 1</math> </li> <li>[Central Limit Theorem] Distribution of sample average <math>\frac{S_n}{n}</math> approaches a <b>normal distribution</b> with mean <math>\mu</math> and variance <math>\frac{\sigma^2}{n}</math>.  <math display="block">\frac{\frac{S_n}{n} - \mu}{\sqrt{\frac{\sigma^2}{n}}} = \frac{S_n - n\mu}{\sigma\sqrt{n}} \sim N(0, 1)</math> <math display="block">\mathbb{P}\left[\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq c\right] \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^c e^{-\frac{x^2}{2}} dx</math> </li> </ul>		
Estimation and Linear Regression			
<p><b>Case #1:</b> Know the joint distribution  Want to find <math>L[Y X] = g(X) = a + bX</math> that minimizes cost function <math>C(g)</math>  <math display="block">C(g) = \mathbb{E}[ Y - g(X) ^2] = \mathbb{E}[ Y - a - bX ^2]</math> <math display="block">b = \frac{\text{Cov}[X, Y]}{\text{Var}[X]}</math> <math display="block">a = \mathbb{E}[Y] - \mathbb{E}[X] \cdot \frac{\text{Cov}[X, Y]}{\text{Var}[X]}</math> <math display="block">L[Y X] = \mathbb{E}[Y] + \frac{\text{Cov}[X, Y]}{\text{Var}[X]}(X - \mathbb{E}[X])</math> </p> <p><b>Case #2:</b> Linear regression  Observed <math>K</math> samples <math>(X_1, Y_1), (X_2, Y_2), \dots, (X_K, Y_K)</math>  Choose <math>a, b</math> to minimize <math>\frac{1}{K} \sum_{k=1}^K  Y_k - a - bX_k ^2</math>  <math display="block">\mathbb{E}[ Y - a - bX ^2] = \frac{1}{K} \sum_{k=1}^K  Y_k - a - bX_k ^2</math> As <math>K \rightarrow \infty</math>, linear regression approaches LLSE, assuming <math>(X_k, Y_k)</math> are i.i.d.</p>			
MMSE (Minimum Mean Squared Error)			
$g(X) = \mathbb{E}[Y X]$ <p>(Orthogonality)</p> $g(X) = \mathbb{E}[Y X]$ $\Leftrightarrow \mathbb{E}[(Y - g(X))\Phi(X)] = 0 \quad \forall \Phi(X)$			
Markov Chain	Continuous Probability		
<p><math>P</math>: transition matrix, <math>\pi_0</math>: initial distribution vector</p> <table border="1" data-bbox="113 1727 780 1771"> <tr> <td><math>\pi_n = P^n \pi_0</math></td> <td><math>\pi = P\pi</math></td> </tr> </table> <p><b>Irreducibility:</b> can go from every state <math>i</math> to every other state <math>j</math> in finite moves.</p> <p><b>Theorem:</b> For finite irreducible Markov chain, for any <math>\pi_0</math>, exists unique invariant distribution <math>\pi</math> s.t.</p> $\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{m=0}^{n-1} \mathbb{P}[X_m = i] = \pi(i)$	$\pi_n = P^n \pi_0$	$\pi = P\pi$	<p><b>PDF:</b> <math>f: \mathbb{R} \rightarrow \mathbb{R}</math>, nonnegative, normalized</p> $\mathbb{P}[a \leq X \leq b] = \int_a^b f(x) dx$ <p><b>CDF:</b></p> $F(x) = \mathbb{P}[X \leq x] = \int_{-\infty}^x f(z) dz$ <p><b>Joint Density:</b> <math>f: \mathbb{R}^2 \rightarrow \mathbb{R}</math> that is nonnegative and normalized. <math>\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1</math>.</p> $\mathbb{P}[x \leq X \leq x + dx, y \leq Y \leq y + dy] = f(x, y) dx dy$
$\pi_n = P^n \pi_0$	$\pi = P\pi$		

**Definition:**

For  $i \in \mathcal{H}$ , define:

$$d(i) := \gcd\{n > 0 \mid P^n(i, i) = \mathbb{P}[X_n = i \mid X_0 = i] > 0\}$$

An *irreducible* Markov chain is **aperiodic** if  $d = 1$ , else **periodic** with period  $d$ .

**Theorem:** For aperiodic Markov Chain,  $\mathbb{P}[X_n = i] \rightarrow \pi(i)$  as  $n \rightarrow \infty$ .

**Classic Problems:**

Expected Steps	$state_A$ before $state_B$
$\beta(state_1)$ $= 1 + \sum \beta(state_i)$	Set $\alpha(B) = 0$ and $\alpha(A) = 1$ and solve.

**Techniques:**

- Additional start and end state
- Clumping of equivalent states
- Redefine states/transitions
- Check all outgoing edges sum up to 1.

**Countability****Techniques**

1. Bijection
2.  $|S_1| \leq |S_2|$  and  $|S_2| \leq |S_1|$
3. Diagonalization, show  $\in S$
4. Subset of countable
5. Superset of uncountable
6. Reduction (to solving Turing)
7. Self-referencing contradictions

Countable	Uncountable
<ul style="list-style-type: none"> <li>• <math>\mathbb{N}, \mathbb{Q}, \mathbb{Z}, \mathbb{N} \times \mathbb{N}</math></li> <li>• Binary strings</li> <li>• Subset <math>T</math> of countable set <math>S</math></li> </ul>	<ul style="list-style-type: none"> <li>• <math>\mathbb{R}</math></li> <li>• <math>x \in [0, 1], \mathbb{R}</math></li> <li>• <math>\mathcal{P}(\mathbb{N})</math></li> </ul>

**Final Checks**

- Define all random variables.
- Check the domain of PMF, PDF, CDF.
- For PMF diagrams, draw 0 for "elsewhere".

**Last Resorts**

- PIE (Midterm 1 horror)
- Difference method, hockey stick theorem
- If PDF method fails, work with CDF
- Indicator variable approach + algebra
- Consider other forms of indicators
- Union bound

$$\mathbb{P}\left[\bigcup A_i\right] \leq \sum \mathbb{P}[A_i]$$

**Independence:**  $X, Y$  are independent if  $\forall a, b, c, d$ :

$$\mathbb{P}[a \leq X \leq b, c \leq Y \leq d] = \mathbb{P}[a \leq X \leq b] \cdot \mathbb{P}[c \leq Y \leq d]$$

The joint density becomes separable:

$$f(x, y) = f_X(x)f_Y(y)$$

$f(x) = \frac{dF(x)}{dx}$	$\mathbb{E}[X] = \int_{-\infty}^{\infty} xf(x) dx$
$\text{Var}[X] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$ $= \int_{-\infty}^{\infty} x^2 f(x) dx - \left( \int_{-\infty}^{\infty} xf(x) dx \right)^2$	

**Computability**

$\text{HALT}(P, I) \# \text{True if } P(I) \text{ halts, else False}$   
 $\text{TURING}(P)$ :

```
if HALT(P, P) == "halts": ∞
else: "halts"
```

**Un-computable problems:**

- Halting: Does program  $P$  halt on input  $I$ ?
- Variant: Does program  $P$  halt on 0? (or any other input for that matter)

```
HALT(P, x):
def P'(y):
    return P(x)
return VARIANT_HALT(P')
```

- Variant: A program  $P(F, x, y)$  that returns true if  $F(x) = y$

```
HALT(F, x):
def Q(y):
    F(x)
    return 0
return P(Q, x, 0)
```

- Variant: A program  $P(F, G)$  that returns true if  $F, G$  halt on same set of input

```
def HALT(F, x):
def Q(y): loop
def R(y):
    if y == x: return F(x)
    else: loop
return not P(Q, R)
```

- Variant: A program  $E2(P, x)$  that returns true if  $P$  runs an even number of lines on  $x$ .

```
def TuringE(P):
if E2(P, P):
    print("Filler")
return
```

Logic and Function	Graph Theory (Definition)																																
Implies: $P \Rightarrow Q \equiv \neg(P \wedge \neg Q) \equiv \neg P \vee Q$ Converse: $Q \Rightarrow P$ Inverse: $\neg P \Rightarrow \neg Q$ Contrapositive: $\neg Q \Rightarrow \neg P$ De Morgan's Law: $\neg(P \wedge Q) \equiv (\neg P) \vee (\neg Q)$ $\neg(P \vee Q) \equiv (\neg P) \wedge (\neg Q)$ $\neg(\forall x P) \equiv \exists x (\neg P)$ $\neg(\exists x P) \equiv \forall x (\neg P)$	<b>Path</b> : a sequence of edges, vertices distinct. <b>Cycle</b> : a path (distinct vertices) with $v_1 = v_n$ <b>Walk</b> : a path without distinct vertices condition <b>Tour</b> : a walk with $v_1 = v_n$  A cycle is a walk. A tour is a walk.  <b>Eulerian walk</b> : uses all edge exactly once. <b>Eulerian tour</b> : walk that ends at start vertex <b>Hamiltonian walk/cycle</b> : a walk/cycle that visits all vertices <i>exactly once</i> . <b>Hypercube</b> (dim $N$ ): $2^N$ nodes, $N2^{N-1}$ edges																																
Final Checks	Function																																
<ul style="list-style-type: none"><li>Check if it is <i>Stable Matching or Propose-and-Reject problem</i>.</li><li>Polynomials in <math>GF</math> must mod coefficients</li><li>RSA: write <math>N, e, d</math> explicitly to avoid errors</li><li><math>0 \in \mathbb{N}</math> for this class</li><li>Be careful of the bound in vertex coloring</li><li>Be careful of base cases for graph</li><li>Counting: rotations / inversions included?</li></ul>	$f(X) = \{y \mid \exists x \in X \text{ s.t. } y = f(x)\}$ $f^{-1}(Y) = \{x \mid f(x) \in Y\}$																																
Stable Matching	Graph Theory																																
When a candidate does not immediately reject a job, the job is still assumed to “propose” to the candidate on the next day.  [Improvement Lemma] Candidate’s matching can only improve. (exchange argument)  Job-Propose and Reject always terminate with matching (contradiction), gives job-optimal and candidate-pessimal (contradiction).	<b>Lines of Attack</b> : Induction on $ V $ , $ E $ , tree-shaving (removal of leaf node), Eulerian tours, pigeonhole,  <b>Euler’s Theorem</b> : Planar graphs with $v \geq 3$ satisfy $v + f = e + 2$ <b>Corollary</b> : All planar graphs satisfy $e \leq 3v - 6$ $K_{3,3}$ <b>Variant</b> : $e \leq 2v - 4$ <b>Kuratowski’s Theorem</b> : A graph is planar iff it doesn’t contain $K_5$ or $K_{3,3}$  <b>Coloring</b> <ul style="list-style-type: none"><li>A graph with max degree <math>k</math> is <math>k + 1</math> colorable. (induct on <math> V </math>)</li><li>A connected graph of max degree <math>d \geq 2</math> can be vertex colored with <math>d</math> colors so long as <math>\exists</math> vertex with degree <math>&lt; d</math>. (<math> V </math>)</li><li>Graph with max degree <math>d \geq 1</math> can be edge colored in <math>2d - 1</math> colors. (induct <math> E </math>)</li></ul>																																
<table><tr><th>Job</th><th>I</th><th>II</th><th>III</th><th>C</th><th>I</th><th>II</th><th>III</th></tr><tr><td>A</td><td>1</td><td>2</td><td>3</td><td>1</td><td>B</td><td>C</td><td>A</td></tr><tr><td>B</td><td>2</td><td>3</td><td>1</td><td>2</td><td>C</td><td>A</td><td>B</td></tr><tr><td>C</td><td>3</td><td>1</td><td>2</td><td>3</td><td>A</td><td>B</td><td>C</td></tr></table> $\{(A, 1), (B, 2), (C, 3)\}, \{(A, 3), (B, 1), (C, 2)\}, \{(A, 2), (B, 3), (C, 1)\}$ are all stable.	Job	I	II	III	C	I	II	III	A	1	2	3	1	B	C	A	B	2	3	1	2	C	A	B	C	3	1	2	3	A	B	C	
Job	I	II	III	C	I	II	III																										
A	1	2	3	1	B	C	A																										
B	2	3	1	2	C	A	B																										
C	3	1	2	3	A	B	C																										
Stable Matching Trivia	Error Correcting Codes																																
<ul style="list-style-type: none"><li>Always exists a candidate who is not proposed to until the last day.</li><li>Propose-and-reject algorithm must terminate in at most <math>(n - 1)^2 + 1</math> days.</li><li>For even <math>n \geq 2</math>, exists instance of stable matching of <math>n</math> jobs and candidates with at least <math>2^{n/2}</math> distinct stable matching. (induct on <math>n</math>)</li></ul>	Message of $n$ packets $(m_1, m_2, \dots, m_n)$ where $m_i = P(i)$ for some polynomial $P$ of at most degree $n - 1$ .  Bounding of $GF(q)$ , $q$ prime: $q \geq \max(m_i + 1, n + k)$ $q \geq \max(m_i + 1, n + 2k)$  Error Correction: $Q(x) = P(x)E(x)$																																

<ul style="list-style-type: none"> <li>• In a job propose algorithm, jobs can't lie to improve their own outcomes, but can to improve others.</li> <li>• If candidate rejects a job in JPA, there is no stable matching in which the candidate and job is paired.</li> <li>• If a candidate misbehaves (rejects falsely), then it is the only candidate that can be in a rogue couple.</li> </ul>	$Q(x_i) = r_i E(x_i)$ $E(x) = (x - e_1) \dots (x - e_k)$ <p>Fractional variant:</p> $n'(1 - \alpha) = n \Rightarrow n' = \frac{n}{1 - \alpha}$ $n'(1 - 2\alpha) = n \Rightarrow n' = \frac{n}{1 - 2\alpha}$
<b>RSA</b>	<b>Secret Sharing</b>
<p>Key <math>(N, e, d)</math>. <math>(N, e)</math> is public. <math>d</math> is private.  <math>N = pq</math> where <math>p, q</math> are large primes.  <math>p, q</math> must be secret, but if forgotten it's fine. Only requires <math>d</math> to decode.</p> $(e, (p-1)(q-1)) = 1$ $d^{-1} \equiv e \pmod{(p-1)(q-1)}$ $E(x) = x^e \pmod{N}$ $D(x) = x^d \pmod{N}$ <p>Security relies on the computational intractability of obtaining <math>x</math> in <math>y = x^e \pmod{N}</math></p>	<p>Bounding of <math>GF(q)</math>, <math>q</math> prime:</p> <p>Secret sharing among <math>m</math> people (the +1 comes from the secret):</p> $q \geq \max(s + 1, m + 1)$ <p>Can delegate sub-polynomials for hierarchy.</p> <p>Spy variants: spies can corrupt messages.</p>