

Known Distributions

Bernoulli Distribution: $X \sim \text{Bernoulli}(p)$

$$P(X=0) = 1-p \quad E[X] = p \\ P(X=1) = p \quad \text{Var}[X] = p(1-p) \quad F(x) = \begin{cases} 0 & x < 0 \\ 1-p & 0 \leq x < 1 \\ 1 & x \geq 1 \end{cases}$$

Binomial Distribution: $X \sim \text{Binomial}(n, p)$

$$P(X=i) = \binom{n}{i} p^i (1-p)^{n-i}, \quad i=0,1,2,\dots,n$$

$$E[X] = np \quad \text{Var}[X] = np(1-p)$$

Geometric Distribution: $X \sim \text{Geom}(p)$

$$P(X=i) = (1-p)^{i-1} p, \quad i=1,2,\dots \quad E[X] = \frac{1}{p} \quad \text{Var}[X] = \frac{1-p}{p^2}$$

$$\text{pgf}(x) = \frac{px}{1-(1-p)x}; \quad P(X) = 1 - (1-p)^X; \quad P(X > nm | X > n) = P(X > m) \quad n, m > 0$$

Poisson Distribution: $X \sim \text{Poisson}(\lambda)$

$$P(X=i) = \frac{\lambda^i e^{-\lambda}}{i!}, \quad i=0,1,2,\dots \quad E[X] = \lambda, \quad \text{Var}[X] = \lambda$$

$$X+Y \sim \text{Poisson}(\lambda+\mu); \quad \text{Binomial}(n, \frac{\lambda}{n}) \xrightarrow{n \rightarrow \infty} \text{Poisson}(\lambda)$$

Exponential Distribution: $X \sim \text{Expo}(\lambda)$

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0 & \text{otherwise} \end{cases} \quad E[X] = \frac{1}{\lambda}, \quad \text{Var}[X] = \frac{1}{\lambda^2} \\ E[X^2] = \frac{2}{\lambda^2}, \quad F_X(x) = 1 - e^{-\lambda x}$$

$$X \sim \text{Expo}(\lambda), Y \sim \text{Expo}(\mu) \Rightarrow \min(X, Y) \sim \text{Expo}(\lambda + \mu)$$

$$X_i \sim \text{Expo}(\lambda_i) \Rightarrow P(X_i = \min(X_i)) = \frac{\lambda_i}{\sum \lambda_j}$$

$$\text{Discrete Uniform: } X \sim \text{Uniform}\{a, \dots, b\} \quad P_X(k) = \begin{cases} \frac{1}{b-a+1} \\ 0 \text{ else} \end{cases} \\ E[X] = \frac{a+b}{2} \quad \text{Var}[X] = \frac{(b-a+1)^2 - 1}{12}$$

Continuous Uniform: $X \sim \text{Uniform}(a, b)$

$$E[X] = \frac{a+b}{2}; \quad \text{Var}[X] = \frac{(b-a)^2}{12}; \quad f_X(x) = \frac{1}{b-a}; \quad F_X(x) = \frac{x-a}{b-a}$$

Normal Distribution: $X \sim N(\mu, \sigma^2) \quad Z \sim N(0, 1)$

$$f_X(x) = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}; \quad f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}; \quad Z = \frac{X-\mu}{\sigma} \sim N(0, 1)$$

$$X \sim N(0, \sigma^2) \Rightarrow E[X^{2n}] = (2n-1)!! \sigma^{2n}$$

Erlang Distribution: $X \sim \text{Erlang}(k, \lambda)$ (sum of k iid $\text{Expo}(\lambda)$)

$$E[X] = \frac{k}{\lambda}, \quad \text{Var}[X] = \frac{k}{\lambda^2}; \quad f_X(x) = \frac{\lambda^k x^{k-1} e^{-\lambda x}}{(k-1)!}, \quad X \sim \text{Erlang}(k, \lambda) \Rightarrow aX \sim \text{Erlang}(k, \frac{\lambda}{a})$$

$$F_X(x) = 1 - \sum_{n=0}^{k-1} \frac{1}{n!} e^{-\lambda x} (\lambda x)^n; \quad X \sim \text{Erlang}(k_1, \lambda) \vee X \sim \text{Erlang}(k_2, \lambda) \Rightarrow X+Y \sim \text{Erlang}(k_1+k_2, \lambda)$$

Pascal Distribution

 $X \sim \text{Pascal}(k, p)$ (sum of k iid $\text{Geom}(p)$)

$$E[X] = \frac{k}{p}; \quad \text{Var}[X] = \frac{k(1-p)}{p^2}; \quad E[X^2] = \frac{k^2 + k(1-p)}{p^2}$$

$$P_X(k) = \binom{k-1}{k-1} p^k (1-p)^{k-1}, \quad k=1,2,\dots; \quad x=k, k+1, \dots$$

$$F_X(x) = \sum_{k=x}^{\infty} \binom{n-1}{k-1} p^k (1-p)^{n-k}$$

Joint and Conditional Probabilities

$$P_{X|Y}(x|y) = \frac{P(X=x, Y=y)}{P(Y=y)} = \frac{P_{X,Y}(x,y)}{P_Y(y)}$$

$$P_{X,Y}(x,y) = P_Y(y) P_{X|Y}(x|y)$$

$$f_{X,Y}(x,y) = \frac{\partial F_{X,Y}}{\partial x \partial y}(x,y) \quad \bullet \quad f_{X,Y}(x,y) = f_Y(y) f_{X|Y}(x|y)$$

$$\bullet \quad f_X(x) = \int_{-\infty}^{\infty} f_Y(y) f_{X|Y}(x|y) dy$$

$$\bullet \quad f_X(x) = \sum_{i=1}^n f_{X|A_i}(x) P(A_i)$$

Convolution

$$P[Z=z] = P[X+Y=z] = \sum_x P[X=x] P[Y=z-x]$$

$$f_Z(z) = \int_{-\infty}^{\infty} f_X(x) f_Y(z-x) dx$$

Probabilistic Bounding

$$1. [\text{Markov}] \text{ nonnegative } X, \text{ finite mean } P[X \geq c] \leq \frac{E[X]}{c}, \quad c > 0, c$$

$$2. [\text{Chebyshev}] c > 0. \quad P[|X-\mu| \geq c] \leq \frac{\text{Var}[X]}{c^2} \\ P[|X-\mu| \geq kc] \leq \frac{1}{k^2}$$

$$3. [\text{Generalized Markov}] Y \text{ not necessarily nonnegative, finite mean; } c, r > 0 \quad P[|Y| \geq c] \leq \frac{E[|Y|^r]}{c^r}$$

$$4. [\text{Extended Markov}] X \text{ not necessarily nonnegative, } \Phi(x) \text{ nonnegative function, monotonically increasing for } x > 0 \\ x > 0 \quad P[X \geq a] \leq \frac{E[\Phi(X)]}{\Phi(a)}$$

$$5. [\text{Cantelli}] x > 0. \quad P[X - E[X] \geq x] \leq \frac{\sigma^2}{x^2 + \sigma^2}$$

$$6. [\text{Law of Large Numbers}] X_1, \dots, X_n \text{ i.i.d. RV with } \mu < \infty \\ S_n = X_1 + \dots + X_n. \quad \forall \varepsilon, \lim_{n \rightarrow \infty} P[|\frac{1}{n} S_n - \mu| < \varepsilon] = 1$$

$$7. [\text{Central Limit Theorem}] \text{ Distribution of } \frac{S_n - n\mu}{\sqrt{\sigma^2 n}} \text{ approaches a normal distribution}$$

$$P\left[\frac{S_n - n\mu}{\sqrt{\sigma^2 n}} \leq c\right] \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^c e^{-\frac{x^2}{2}} dx$$

$$\text{i.e. define } Z_n = \frac{X_1 + \dots + X_n - n\mu}{\sqrt{\sigma^2 n}}, \text{ then } (Z_n)_n \xrightarrow{d} N(0, 1) \text{ in distribution.}$$

$$8. [\text{Chernoff}] P[X \geq a] = P[e^{sX} \geq e^{sa}] \leq \frac{M_X(s)}{e^{sa}} \quad \forall s \geq 0$$

$$9. [\text{Jensen}] f(E[X]) \leq E[f(X)] \quad \forall f \text{ convex.}$$

$$10. [\text{WLLN}] \text{ Let } (X_n)_n \text{ be a sequence of i.i.d. RV with mean } \mu. \quad \forall \varepsilon > 0, P\left[\left|\frac{X_1 + \dots + X_n}{n} - \mu\right| \geq \varepsilon\right] \rightarrow 0.$$

$$11. [\text{SLLN}] \text{ Let } (X_n)_n \text{ be a sequence of i.i.d. RV with mean } \mu. \text{ Then } M_n = \frac{X_1 + \dots + X_n}{n} \text{ converges to } \mu \text{ with probability 1} \\ P\left[\lim_{n \rightarrow \infty} \frac{X_1 + \dots + X_n}{n} = \mu\right] = 1. \quad (\text{i.e. } \forall \varepsilon > 0, P(|\mu_n - \mu| > \varepsilon \text{ i.o.}) = 0)$$

$$12. [\text{De Moivre-Laplace Approximation}] \text{ If } S_n \sim \text{Binom}(n, p) \text{ and } n \gg 1, k, \lambda \text{ nonnegative, then} \\ P[k \leq S_n \leq \lambda] \approx \left[\Phi\left(\frac{\lambda + \frac{1}{2} - np}{\sqrt{np(1-p)}}\right) - \Phi\left(\frac{k - \frac{1}{2} - np}{\sqrt{np(1-p)}}\right) \right]$$

Conditional Expectation and Variance

Note: $E[Y|X] = f(X)$ i.e. is a function of X $E[Y|X=x]$ is a real numberTo find $E[Y|X]$, generalize pattern from $E[Y|X=x]$

$$1. (\text{Linearity}) E[a_1 Y_1 + a_2 Y_2 | X] = a_1 E[Y_1 | X] + a_2 E[Y_2 | X]$$

$$2. (\text{Factoring Known Values}) E[h(X) Y | X] = h(X) E[Y | X]$$

$$3. (\text{Independence}) E[X | Y] = E[X] \quad (X, Y \text{ independent})$$

$$E[X] = E[E[X|Y]]; \quad E[XY] = E[E(XY|Y)] = E[Y E(X|Y)]$$

$$\text{Var}[X] = E[\text{Var}[X|Y]] + \text{Var}[E(X|Y)]$$

Conditioning on Event

Conditioning on RV

$$\bullet E[X|A] = \int_{-\infty}^{\infty} x f_{X|A}(x) dx \quad \bullet E[Y] = E[E(Y|X)] = \int_{-\infty}^{\infty} E(Y|X=x) f_X(x) dx$$

$$\bullet E[g(X)|A] = \int_{-\infty}^{\infty} g(x) f_{X|A}(x) dx \quad \bullet E[X|Y=y] = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx$$

$$\bullet E[g(X)|Y=y] = \int_{-\infty}^{\infty} g(x) f_{X|Y}(x|y) dx \quad \bullet E[g(X, Y)|Y=y] = \int_{-\infty}^{\infty} g(x, y) f_{X,Y}(x, y) dx$$

Bayes and Continuous Bayes

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

$$f_{Y|X}(y|x) = \frac{f_X(x)f_{Y|X}(y|x)}{\int_{-\infty}^{\infty} f_X(t)f_{Y|X}(y|t)dt}$$

$$f_Y(y)P(N=n|Y=y) = P(N=n)f_{Y|N}(y|n)$$

$$P(N=n|Y=y) = \frac{P(N=n)f_{Y|N}(y|n)}{\sum_i P(N=i)f_{Y|N}(y|i)}$$

$$f_{Y|N}(y|n) = \frac{f_Y(y)P(N=n|Y=y)}{\int_{-\infty}^{\infty} f_Y(t)P(N=n|Y=t)dt}$$

Differential Probability

$$f_X(x)dx = P(X \leq x \leq x+dx); f_{X|Y}(x|y)dx = \frac{P(X \leq x \leq x+dx, Y \leq y \leq y+dy)}{P(Y \leq y \leq y+dy)}$$

$$\frac{d}{dz} P(z \leq Z|X=x) = f_{Z|X}(z|x)$$

$$f_Y(y) = f_X(f^{-1}(y)) \left| \frac{df^{-1}}{dy}(y) \right|$$

Moment Generating Functions (MGF)

$$\text{Properties: } e^{sX} = 1 + \frac{sX}{1!} + \frac{(sX)^2}{2!} + \frac{(sX)^3}{3!} + \dots$$

$$E[e^{sX}] = M_X(s) = 1 + sE[X] + \frac{s^2}{2!}E[X^2] + \frac{s^3}{3!}E[X^3] + \dots$$

$$\left(\frac{d^n}{ds^n} E[e^{sX}] \right) \Big|_{s=0} = E[X^n]; M_X(0) = 1 \rightarrow \text{independent}$$

$$Y = aX + b \Rightarrow M_Y(s) = e^{sb} M_X(as); Z = \sum X_i \Rightarrow M_Z(s) = \prod M_{X_i}(s)$$

Distributions:

$$\text{Bernoulli}(p): M(s) = (1-p) + pe^s; \text{Binomial}(n, p): M(s) = (1-p + pe^s)^n$$

$$\text{Geometric}(p): M(s) = \frac{pe^s}{1-(1-p)e^s}; \text{Poisson}(\lambda): M(s) = e^{\lambda(e^s-1)}$$

$$\text{Expo}(\lambda): M(s) = \frac{\lambda}{\lambda-s} (\lambda > s); \text{Uniform}(a, b): M(s) = \frac{e^{sb}-e^{sa}}{s(b-a)}$$

$$N(\mu, \sigma^2): M(s) = e^{\mu s + \frac{\sigma^2 s^2}{2}}; \text{Erlang}(k, \lambda): M(s) = \left(\frac{\lambda}{\lambda-s} \right)^k$$

$$\text{Pascal}(k, p): M(s) = \left(\frac{pe^s}{1-(1-p)e^s} \right)^k$$

Convergence and Theorems

Convergence in Distribution: A sequence $(X_n)_n$ converges in distribution to X , denoted as $X_n \xrightarrow{d} X$ if $\forall x$,

$$\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x) \quad (\text{i.e. CDF of } X_n \text{ converges to CDF of } X)$$

Theorem: For integer valued X , $(X_n)_n$, suffices to show

$$\forall x \quad P(X=x) = \lim_{n \rightarrow \infty} P(X_n=x).$$

Convergence in Probability: A sequence $(X_n)_n$ converges in probability to X denoted as $X_n \xrightarrow{p} X$ if $\forall \epsilon > 0$,

$$\lim_{n \rightarrow \infty} P(|X_n - X| \geq \epsilon) = 0$$

Theorem: If $X_n \xrightarrow{d} c$ constant, then $X_n \xrightarrow{p} c$.

Convergence with Probability 1: A sequence $(X_n)_n$ converges to X with probability 1, denoted as $X_n \xrightarrow{a.s.} X$ if under sample space Ω , $P(\{\omega \in \Omega : \lim_{n \rightarrow \infty} X_n(\omega) = X(\omega)\}) = 1$

Theorem: Consider $(X_n)_n$. If $\forall \epsilon > 0 \quad \sum_{n=1}^{\infty} P(|X_n - X| > \epsilon) < \infty$, then $X_n \xrightarrow{a.s.} X$

Theorem: Consider $(X_n)_n$. For $\epsilon > 0$, define $A_n = \{|X_n - X| > \epsilon\}$, then $X_n \xrightarrow{a.s.} X$ if and only if $\lim_{n \rightarrow \infty} P(A_n) = 0$.

Miscellaneous

$$f_{X|N}(x) = n \binom{n-1}{i-1} f(x) F(x)^{i-1} (1-F(x))^{n-i}$$

$$E[X^{(n)}] = \frac{1}{n!}$$

$$E[X] = \sum_{i=1}^{\infty} P(X \geq i); E[Z] = \int_0^{\infty} P(Z \geq z) dz$$

Conditional Expectation and Variance

Series of RV

$$Y = Y_1 + Y_2 + \dots + Y_N$$

$$E[Y|N] = N \cdot E[X]$$

$$E[Y] = E[E[Y|N]] = E[N]E[X]$$

$$Var[Y] = E[N]Var[X] + E[X]^2 Var[N]$$

$$M_Y(s) = \sum_{n=0}^{\infty} (M_X(s))^n P[N=n]$$

Conditioning on RV

$$E[g(X, Y)] = E[E[g(X, Y)|Y]]$$

$$= \int_{-\infty}^{\infty} E[g(X, Y)|Y=y] f_Y(y) dy$$

$$E[X|N > k] = E[E[X|N]|N > k]$$

$$= \sum_{n=k+1}^{\infty} E[X|N=n, N > k] P[N=n|N > k]$$

Covariance and Correlation

Covariance: $Cov[X, Y] = E[(X - \mu_X)(Y - \mu_Y)] = E[XY] - E[X]E[Y]$

$$Cov[X, X] = Var[X]; Var(X+Y) = Var[X] + Var[Y] + 2Cov[X, Y]$$

$$Cov[aX_1 + bX_2, cY_1 + dY_2] = acCov[X_1, Y_1] + adCov[X_1, Y_2] + bcCov[X_2, Y_1] + bdCov[X_2, Y_2]$$

X, Y independent $\Rightarrow Cov(X, Y) = 0$

Correlation: $\rho(X, Y) = \frac{Cov(X, Y)}{\sigma_X \sigma_Y}; X' = \frac{X - \mu_X}{\sigma_X}; Y' = \frac{Y - \mu_Y}{\sigma_Y}$

$$-1 \leq \rho(X, Y) = Cov(X', Y') \leq 1$$

$$\rho(X, Y) = 1 \Rightarrow Y = AX + B, A > 0 (Y' = X')$$

$$\rho(X, Y) = -1 \Rightarrow Y = AX + B, A < 0 (Y' = -X')$$

Borel Cantelli lemma and Continuity

First lemma: Let $(A_n)_n$ be sequence of events. If $\sum_{n=1}^{\infty} P[A_n] < \infty$ then $P[A_n \text{ i.o.}] = 0$ (i.e. probability that infinitely many of A_n occurring is 0).

Second lemma: Let $(A_n)_n$ be sequence of events. If A_n independent and $\sum_{n=1}^{\infty} P[A_n] = \infty$, then $P[A_n \text{ i.o.}] = 1$ (i.e. probability that infinitely many of them occurring is 1).

Let $h: \mathbb{R} \rightarrow \mathbb{R}$ be continuous:

If $X_n \xrightarrow{d} X$, then $h(X_n) \xrightarrow{d} h(X)$. If $X_n \xrightarrow{p} X$, $h(X_n) \xrightarrow{p} h(X)$.

If $X_n \xrightarrow{a.s.} X$, then $h(X_n) \xrightarrow{a.s.} h(X)$.

Classics

[Ballot] Let A, B be players s.t. A scored $n > m$ points and B scored m points. Then $P[A \text{ always ahead of } B] = \frac{n-m}{n+m}$.

[Gambler's Ruin] Let A be a player who start at i and at each step increment by $-1, +1$ with probability $\frac{1}{2}$. Then $P[A \text{ reaches } n] = \frac{i}{n}$.


[Gambler's Ruin (unfair)] If step probability is p , then $P[A \text{ reaches } n] = \frac{1 - (\frac{p}{1-p})^i}{1 - (\frac{p}{1-p})^n}$.

[Secretary] Optimal cutoff is $n e^{-1}$.

[Coupon collector] $n H_n = n \sum_{i=1}^n \frac{1}{i}$.

Last Resort and Final Checks

Indicators! $Var[X] = Cov[X, X]$; Symmetry Bijection

Draw Pictures 

Define all ranges for f_X, F_X . Isolate objects

One off error, conditioning, intuition, momentum.

$$E[X] = \int_0^{\infty} P(X > x) dx - \int_{-\infty}^0 P(X < -x) dx; \int_0^{\infty} P(-Y)^n dy = \frac{n!}{(n+1)!}$$

$$P[X=0] \leq \frac{Var[X]}{E[X]^2}; P(\bigcap A_i) \leq \sum P[A_i]; f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$