

1. Exercise 3.4

- (a) we know  $y = w^{*T}x + \epsilon$  and  $H = X(X^T X)^{-1}X^T$  from (3.6), and we know  $\hat{y} = Hy$  by definition, we want to prove  $\hat{y} = Xw^* + H\epsilon$

$$\begin{aligned}\hat{y} &= H(w^*X + \epsilon) \\ &= X(X^T X)^{-1}X^T(w^*X + \epsilon) \\ &= X(X^T X)^{-1}X^T w^*X + X(X^T X)^{-1}X^T \epsilon \\ &= w^*X + H\epsilon\end{aligned}$$

- (b) for  $\hat{y} - y$ , we have

$$\begin{aligned}\hat{y} - y &= w^*X + H\epsilon - (w^* + \epsilon) \\ &= H\epsilon - \epsilon \\ &= \epsilon(H - I)\end{aligned}$$

where  $I$  denotes the identity matrix

- (c) let  $E_{in}(w) = \frac{1}{N}||\hat{y} - y||^2$

$$\begin{aligned}E_{in}(w) &= \frac{1}{N}||\epsilon(H - I)||^2 \\ &= \frac{1}{N}(\epsilon(H - I))^T(\epsilon(H - I)) \\ &= \frac{1}{N}\epsilon^T(H - I)^T\epsilon(H - I)\end{aligned}$$

We know  $H - I$  is symmetric, so  $(H - I)^T = (H - I)$

$$\begin{aligned}E_{in}(w) &= \frac{1}{N}\epsilon^T\epsilon(H - I)^2 \\ &= \frac{1}{N}\epsilon^T\epsilon(I - H)^2\end{aligned}$$

- (d) We know

$$\begin{aligned}E_D[E_{in}(w_{lin})] &= E_D[\frac{1}{N}(\epsilon^T\epsilon(I - H))] \\ &= \frac{1}{N}(E_D[\epsilon^T\epsilon] - E_D[\epsilon^T\epsilon H])\end{aligned}$$

Given that  $\epsilon$  is a noise term with zero mean and  $\sigma^2$  variance. The variance of each noise

component  $\epsilon$  is  $\sigma^2$ , so

$$\begin{aligned} E_D[E_{in}(w_{lin})] &= \frac{1}{N}(N\sigma^2 - E_D[\epsilon^T \epsilon H]) \\ &= \sigma^2 - \frac{1}{N}E_D[\epsilon^T \epsilon H] \end{aligned}$$

Now we can calculate

$$\begin{aligned} E_D[\epsilon^T \epsilon H] &= E_D\left[\sum_{i=1}^N \epsilon_i^2 H\right] \\ &= H \sum_{i=1}^N E_D[\epsilon_i^2] \end{aligned}$$

By the problem, we know that each component of  $\epsilon$  is a random variable with zero mean and variance  $\sigma^2$ , so this means that  $E_D[\epsilon_i] = 0$  and  $E_D[\epsilon_i^2] = \sigma^2$  for all  $i$ .

$$\begin{aligned} E_D[\epsilon^T \epsilon H] &= H \sum_{i=1}^N \sigma^2 \\ &= HN\sigma^2 \end{aligned}$$

We continue the problem by substituting the result into our original equation

$$\begin{aligned} E_D[E_{in}(w_{lin})] &= \sigma^2 - \frac{1}{N}E_D[\epsilon^T \epsilon H] = \sigma^2 - \frac{1}{N}HN\sigma^2 \\ &= \sigma^2 - H\sigma^2 \end{aligned}$$

Then, we can calculate for the  $trace(H)$

$$\begin{aligned} trace(H) &= trace(X(X^T X)^{-1} X^T) \\ &= trace((X^T X)^{-1} (X^T X)) \end{aligned}$$

Given that  $X^T X$  is a square matrix of size  $(d+1)$ , and it's inverse  $(X^T X)^{-1}$  is also present, then we have

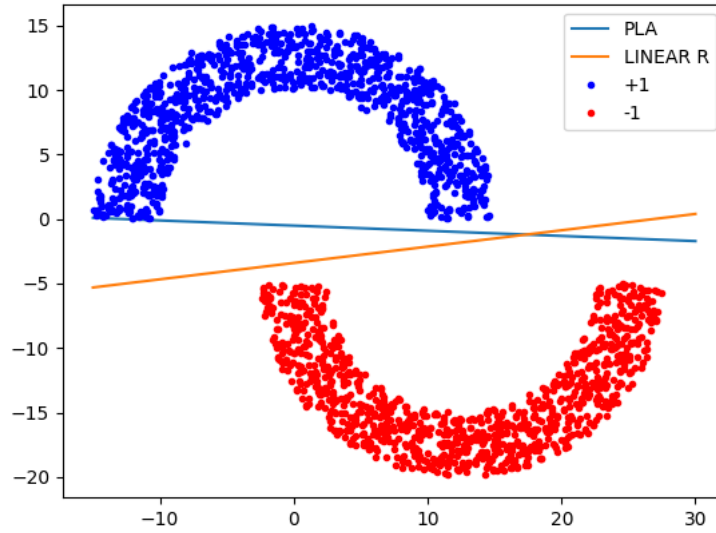
$$\begin{aligned} trace((X^T X)^{-1} (X^T X)) &= d+1 \\ trace(H) &= \frac{d+1}{N} \end{aligned}$$

Then, we have proved that  $E_D[E_{in}(w_{lin})] = \sigma^2(1 - \frac{d+1}{N})$  ■

(e) to do

$$E_{D,\epsilon'}[E_{test}(w_{lin})] = E_{D,\epsilon'}\left[\frac{1}{N}||Xw - y'||^2\right]$$

2. Problem 3.1



(a)

- (b) Both PLA and linear regression found ways to separate this data, however, one could say that the linear regression algorithm found a better way to separate the data as the PLA appears to be closer to the top part of the semicircle, barely missing on misclassifying one of the +1 points. With this, one can predict that linear regression will have a lower  $E_{out}$  than the PLA, however, this isn't guaranteed.