CS591S1 Homework 2: More on Differential Privacy

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1 Problem 1

1. (Sensitivity)

Given data set \mathbf{x} , for arbitrary $y \in \{1, \dots, R\}$ and adjacent dataset \mathbf{x}' by insertion or deletion of one data point x_k , we have following cases for insertion (deletion will be symmetric):

 \bullet $y = x_k$:

$$q(y; \mathbf{x}') = -|\sum_{i=0}^{n} sign(y-x_i) + sign(y-x_k)| = -|\sum_{i=0}^{n} sign(y-x_i) + 0| = -|\sum_{i=0}^{n} sign(y-x_i)| = q(y; \mathbf{x})$$

• $y < x_k$

$$q(y; \mathbf{x}') = -|\sum_{i=0}^{n} sign(y - x_i) + sign(y - x_k)|$$

= -|\sum_{i=0}^{n} sign(y - x_i) - 1| (\ddot)

By triangle inequality, we have:

$$(\star) \leq -|\sum_{i=0}^{n} sign(y-x_i)| + 1 = q(y; \mathbf{x}) + 1$$

 $(\star) \geq -|\sum_{i=0}^{n} sign(y-x_i)| - 1 = q(y; \mathbf{x}) - 1$

Then we can get:

$$-1 \le q(y; \mathbf{x}') - q(y; \mathbf{x}) \le 1$$

• $y > x_k$

$$q(y; \mathbf{x}') = -|\sum_{i=0}^{n} sign(y - x_i) + sign(y - x_k)|$$

= -|\sum_{i=0}^{n} sign(y - x_i) + 1| (\ddot)

By triangle inequality, we have:

$$\begin{array}{lll} (\star) & \leq & -|\sum_{i=0}^{n} sign(y-x_i)| + 1 = q(y;\mathbf{x}) + 1 \\ (\star) & \geq & -|\sum_{i=0}^{n} sign(y-x_i)| - 1 = q(y;\mathbf{x}) - 1 \end{array}$$

Then we can get:

$$-1 \le q(y; \mathbf{x}') - q(y; \mathbf{x}) \le 1$$

The Deletion is symmetric where we can get: $-1 \le q(y; \mathbf{x}) - q(y; \mathbf{x}') \le 1$ in the same way. Then, we can conclude from all cases, the $|q(y; \mathbf{x}) - q(y; \mathbf{x}')| \le 1$, i.e., the sensitivity be at most 1.

2. Proof. By the definition of $rank_{\mathbf{x}}(y)$, we have:

$$|rank_{\mathbf{x}}(y) - \frac{n}{2}| = -q(y; \mathbf{x}).$$

Then, we know:

$$\begin{array}{ll} Pr_{y \sim A_{\epsilon}(\mathbf{x})}[|rank_{\mathbf{x}}(y) - \frac{n}{2}| > c \cdot \frac{\ln(R) + \ln(1/\beta)}{\epsilon}] & \equiv & Pr_{y \sim A_{\epsilon}(\mathbf{x})}[-q(y; \mathbf{x}) > c \cdot \frac{\ln(R) + \ln(1/\beta)}{\epsilon}] \\ & = & Pr_{y \sim A_{\epsilon}(\mathbf{x})}[q(y; \mathbf{x}) \leq -c \cdot \frac{\ln(R) + \ln(1/\beta)}{\epsilon}] \end{array}$$

By definition of exponential mechanism, we have:

$$\begin{split} Pr_{y \sim A_{\epsilon}(\mathbf{x})}[q(y;\mathbf{x}) \leq -c \cdot \frac{\ln(R) + \ln(1/\beta)}{\epsilon}] &= \sum_{\substack{y \mid q(y;\mathbf{x}) < -c \cdot \frac{\ln(R) + \ln(1/\beta)}{\epsilon} \\ y \mid q(y;\mathbf{x}) < -c \cdot \frac{\ln(R) + \ln(1/\beta)}{\epsilon} \cdot \frac{e}{2S} \\ \end{pmatrix}} \frac{\exp\left(q(y;\mathbf{x})\epsilon/2S\right)}{\sum_{y} \exp\left(q(y;\mathbf{x})\epsilon/2S\right)} \\ &\leq R \frac{\exp\left(-c \cdot \frac{\ln(R) + \ln(1/\beta)}{\epsilon} \cdot \frac{e}{2S}\right)}{\sum_{y} \exp\left(q(y;\mathbf{x})\epsilon/2S\right)} \\ &= R \frac{\exp\left(\frac{c}{2}(\ln(\frac{1}{R}) + \ln(\beta))\right)}{\sum_{y} \exp\left(q(y;\mathbf{x})\epsilon/2S\right)} \left(\star\right) \end{split}$$

Since the only one optimal output candidate is the median value where $q(y, \mathbf{x}) = 0$, so we have:

$$(\star) \leq \frac{\exp\left(\frac{c}{2}(\ln(\frac{1}{R}) + \ln(\beta))\right)}{\exp(0)}$$

In order to have this probability be at most β , we take the equality and get:

$$\frac{\exp\left(\frac{c}{2}(\ln(\frac{1}{R}) + \ln(\beta))\right)}{\exp(0)} = \beta$$

$$R^{1-\frac{c}{2}} = \beta^{1-\frac{c}{2}}$$

Since we have $R \ge 1$ and $\beta \in (0,1)$, there exists c=2 which can make the equation holds. \square

2 Problem 2

Proof. The proof are developed by two symmetric cases: insertion and deletion.

case Insertion

Taking two adjacent data sets \mathbf{x} , \mathbf{x}' where \mathbf{x}' contains one more data point. For any output set S, there are following cases by output space:

 $-S \subseteq E_{bad}$ Inserting one data that makes an empty bin $(k \in \mathcal{X})$ be nonempty and this bin is contained in the output set S.

$$\begin{array}{rcl} Pr[A(\mathbf{x}') = S] & = & Pr[\tilde{c_k'} > \tau] \leq \frac{\delta}{2} \\ Pr[A(\mathbf{x}) = S] & = & 0 \end{array}$$

 $-S \subseteq E_0$ Inserting one data that makes an empty bin $(k \in \mathcal{X})$ be nonempty and this bin is not contained in the output set. The probability ratio

$$\begin{split} 1 > \frac{Pr[A(\mathbf{x}') \in S]}{Pr[A(\mathbf{x}) \in S]} &= \frac{Pr[A(\mathbf{x}) \in S \land \tilde{c_k'} < \tau]}{Pr[A(\mathbf{x}) \in S]} \\ &= \frac{Pr[A(\mathbf{x}) \in S] \cdot Pr[\tilde{c_k'} < \tau]}{Pr[A(\mathbf{x}) \in S]} \\ &= Pr[\tilde{c_k'} = 1 + Lap(\frac{1}{\epsilon}) < \tau] \\ &\geq (1 - \frac{\delta}{2}) \end{split}$$

 $-S \subseteq E_1$ Inserting one data that doesn't change non-empty bins:

$$e^{-\epsilon} \leq \frac{Pr[A(\mathbf{x}') \in S]}{Pr[A(\mathbf{x}) \in S]} = \frac{Pr[A(\mathbf{x} \setminus k) \in S \setminus k \land \tilde{c_k} \in S \cap k]}{Pr[A(\mathbf{x}' \setminus k) \in S \setminus k \land \tilde{c_k} \in S \cap k]}$$

$$= \frac{Pr[A(\mathbf{x} \setminus k) = S \setminus k] \cdot Pr[\tilde{c_k} \in S \cap k]}{Pr[A(\mathbf{x} \setminus k) = S \setminus k] \cdot Pr[\tilde{c_k} \in S \cap k]}$$

$$= \frac{Pr[\tilde{c_k} \in S \cap k]}{Pr[\tilde{c_k} \in S \cap k]}$$

$$\leq e^{\epsilon}$$

By summarization, we have following equations:

$$Pr[A(\mathbf{x}') \in S] = Pr[A(\mathbf{x}') \in S \cap E_0] + Pr[A(\mathbf{x}') \in S \cap E_1] + Pr[A(\mathbf{x}') \in S \cap E_{Bad}]$$

$$\leq e^{\epsilon} Pr[A(\mathbf{x}) \in S \cap E_0] + Pr[A(\mathbf{x}) \in S \cap E_1] + \frac{\delta}{2}$$

$$\leq e^{\epsilon} Pr[A(\mathbf{x}) \in S \cap E_0] + e^{\epsilon} Pr[A(\mathbf{x}) \in S \cap E_1] + \frac{\delta}{2}$$

$$= e^{\epsilon} Pr[A(\mathbf{x}) \in S \cap (E_0 \cup E_1)] + \frac{\delta}{2}$$

$$\leq e^{\epsilon} Pr[A(\mathbf{x}) \in S] + \frac{\delta}{2}$$

On the other side, we have:

$$\begin{array}{ll} Pr[A(\mathbf{x}') \in S] & = & Pr[A(\mathbf{x}') \in S \cap E_0] + Pr[A(\mathbf{x}') \in S \cap E_1] + Pr[A(\mathbf{x}') \in S \cap E_{Bad}] \\ & \geq & e^{-\epsilon} Pr[A(\mathbf{x}) \in S \cap E_0] + (1 - \frac{\delta}{2}) Pr[A(\mathbf{x}) \in S \cap E_1] \\ & \geq & \min(e^{-\epsilon}, 1 - \frac{\delta}{2}) Pr[A(\mathbf{x}) \in S \cap E_0] + \min(e^{-\epsilon}, 1 - \frac{\delta}{2}) Pr[A(\mathbf{x}) \in S \cap E_1] \\ & = & \min(e^{-\epsilon}, 1 - \frac{\delta}{2}) Pr[A(\mathbf{x}) \in S \cap (E_0 \cup E_1)] \\ & = & \min(e^{-\epsilon}, 1 - \frac{\delta}{2}) Pr[A(\mathbf{x}) \in S] \end{array}$$

case deletion.

By deletion, we have exactly the symmetric cases as insertion.

By summarization, the probability of failure would be δ in both cases. So we have the algorithm be $(\max(\epsilon, \ln(\frac{1}{1-\frac{\delta}{2}})), \delta)$ -DP. When δ is small, we have $\ln(\frac{1}{1-\frac{\delta}{2}}) \sim 0$ and $\epsilon > 0$, then $(\max(\epsilon, \ln(\frac{1}{1-\frac{\delta}{2}})), \delta)$ -DP is (ϵ, δ) -DP.