# Generative models for discrete data

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### Outline

- Bayesian concept learning
- The beta-binomial model
- The Dirichlet-multinomial model
- Naive Bayes classifier

### Introduction

$$p(y = c | \mathbf{x}, \boldsymbol{\theta}) = \frac{p(y = c | \boldsymbol{\theta}) p(\mathbf{x} | y = c, \boldsymbol{\theta})}{\sum_{c'} p(y = c' | \boldsymbol{\theta}) p(\mathbf{x} | y = c', \boldsymbol{\theta})}$$

• This is called a **generative classifier**, since it specifies how to generate the data using the **class conditional density** p(x|y = c) and the class prior p(y = c).

• Posterior = 
$$\frac{prior * likelihood}{evidence}$$

### Bayesian concept learning

- Goal: to learn the indicator function f, which just defines which elements are in the set C.
- Example: the number game
- 1. choose some simple arithmetical concept C
- 2. give a series of randomly chosen positive examples  $D = \{x_1, \ldots, x_N\}$  drawn from C
- 3. ask whether some new test case x' belongs to C? eg: p(x'|D)? which is the probability that  $x' \in C$  given the data D for any  $x' \in \{1, ..., 100\}$

### Likelihood

Strong sampling assumption:

Examples are sampled uniformly at random from the extension of a concept.

(eg: the extension of  $h_{even}$  is {2, 4, 6, . . . , 98, 100})

$$p(\mathcal{D}|h) = \left[\frac{1}{\text{size}(h)}\right]^N = \left[\frac{1}{|h|}\right]^N$$
 N is the size of D (**Occam's razor**)

eg: for D={16} 
$$p(D|h_{PowerOfTwo}) = \frac{1}{6}$$
  $p(D|h_{even}) = \frac{1}{50}$ 

### Prior

- Suppose D =  $\{16, 8, 2, 64\}$ . Given this data, the concept h' = "powers of two except 32" is more likely than h = "powers of two"
- conceptually unnatural!
- We can capture such intuition by assigning low prior probability to unnatural concepts. (subjective)

### Posterior

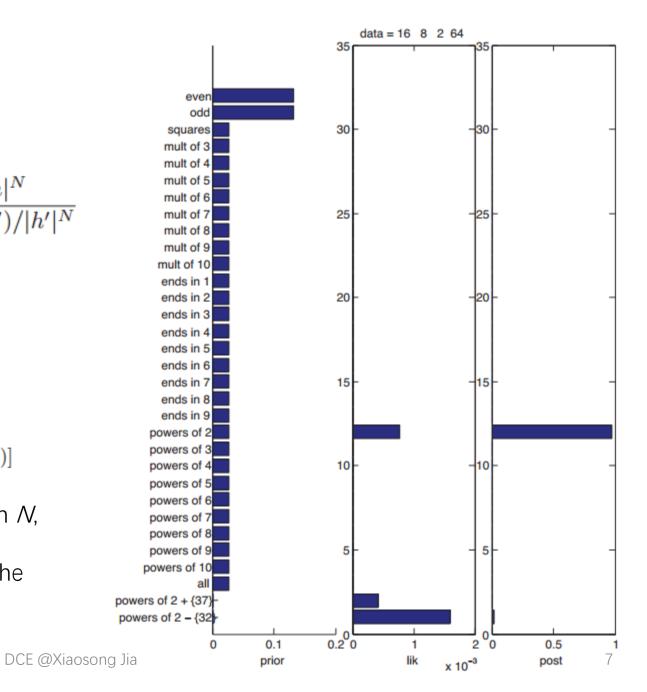
$$p(h|\mathcal{D}) = \frac{p(\mathcal{D}|h)p(h)}{\sum_{h'\in\mathcal{H}} p(\mathcal{D},h')} = \frac{p(h)\mathbb{I}(\mathcal{D}\in h)/|h|^N}{\sum_{h'\in\mathcal{H}} p(h')\mathbb{I}(\mathcal{D}\in h')/|h'|^N}$$

 When we have enough data, the posterior p(h|D) becomes peaked on a single concept, namely the MAP(Maximum A Posteriori Estimation) estimate.

$$\begin{split} p(h|\mathcal{D}) &\to \delta_{\hat{h}^{MAP}}(h) \\ \hat{h}^{MAP} &= \operatorname*{argmax}_{h} p(\mathcal{D}|h) p(h) = \operatorname*{argmax}_{h} \left[ \log p(\mathcal{D}|h) + \log p(h) \right] \end{split}$$

 Since the likelihood term depends exponentially on N, and the prior stays constant, as we get more and more data, the MAP estimate converges towards the MLE(Maximum Likelihood Estimation)

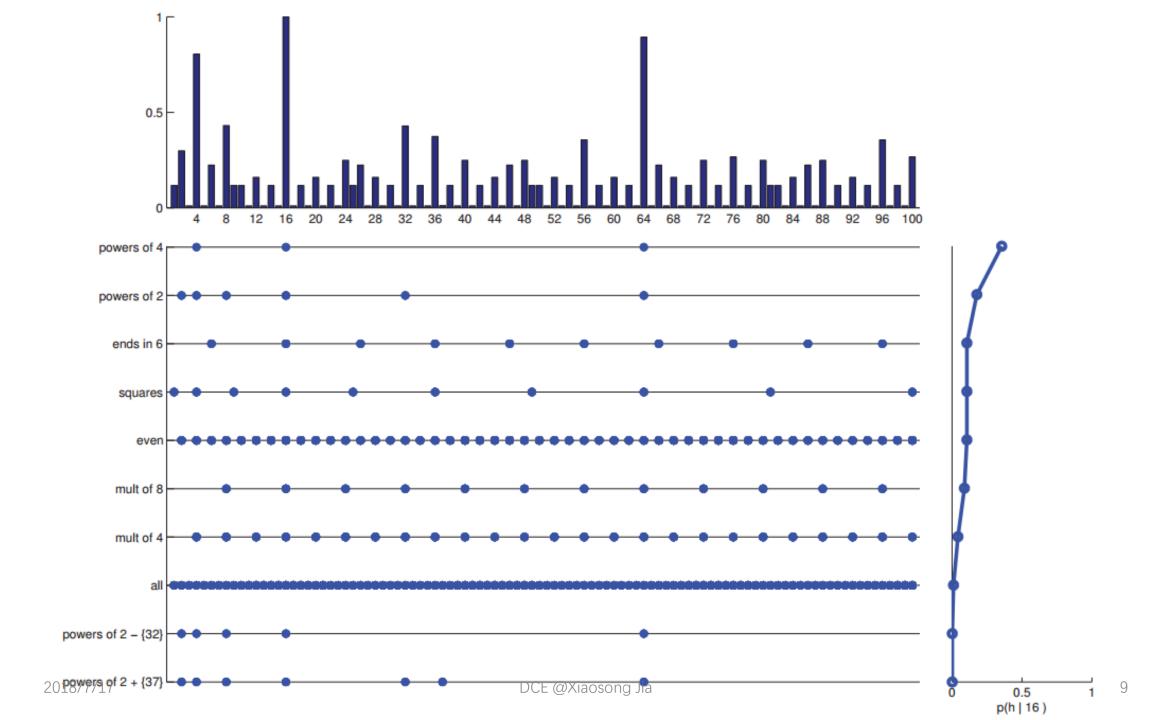
$$\hat{h}^{mle} \triangleq \underset{2018/77_{h}}{\operatorname{argmax}} p(\mathcal{D}|h) = \underset{h}{\operatorname{argmax}} \log p(\mathcal{D}|h)$$



# Posterior predictive distribution

#### Bayes model averaging

$$p(\tilde{x} \in C|\mathcal{D}) = \sum_{h} p(y = 1|\tilde{x}, h)p(h|\mathcal{D})$$



### The beta-binomial model

- the unknown parameters are continuous
- Example:

the problem of inferring the probability that a coin shows up heads, given a series of observed coin tosses.

### Likelihood

• Suppose  $X_i \sim \text{Ber}(\theta)$ , where  $X_i = 1$  represents "heads",  $X_i = 0$  represents "tails", and  $\theta \in [0, 1]$  is the rate parameter (probability of heads).

$$p(\mathcal{D}|\theta) = \theta^{N_1} (1-\theta)^{N_0}$$

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### Prior

 To make the math easier, it would convenient if the prior had the same form as the likelihood.

$$p(\theta) \propto \theta^{\gamma_1} (1 - \theta)^{\gamma_2}$$

Beta
$$(\theta|a,b) \propto \theta^{a-1}(1-\theta)^{b-1}$$

• When a=1 and b=1, it's called a uniform prior.

### Posterior

- $p(\theta|D) \propto Bin(N1|\theta, N0 + N1)Beta(\theta|a, b) \propto Beta(\theta|N1 + a, N0 + b)$
- When the prior and the posterior have the same form, we say that the prior is a **conjugate prior** for the corresponding likelihood.
- MAP MLE

$$\hat{\theta}_{MAP} = \frac{a + N_1 - 1}{a + b + N - 2} \qquad \hat{\theta}_{MLE} = \frac{N_1}{N}$$

Mean

$$\overline{\theta} = \frac{a + N_1}{a + b + N}$$

- The strength of the prior, also known as the **effective sample size** of the prior, is the sum of the pseudo counts a + b;
- Variance

$$var \left[ \theta | \mathcal{D} \right] = \frac{(a + N_1)(b + N_0)}{(a + N_1 + b + N_0)^2 (a + N_1 + b) }$$

### Posterior predictive distribution

$$p(\tilde{x} = 1|\mathcal{D}) = \int_0^1 p(x = 1|\theta)p(\theta|\mathcal{D})d\theta$$
$$= \int_0^1 \theta \operatorname{Beta}(\theta|a, b)d\theta = \mathbb{E}[\theta|\mathcal{D}]$$

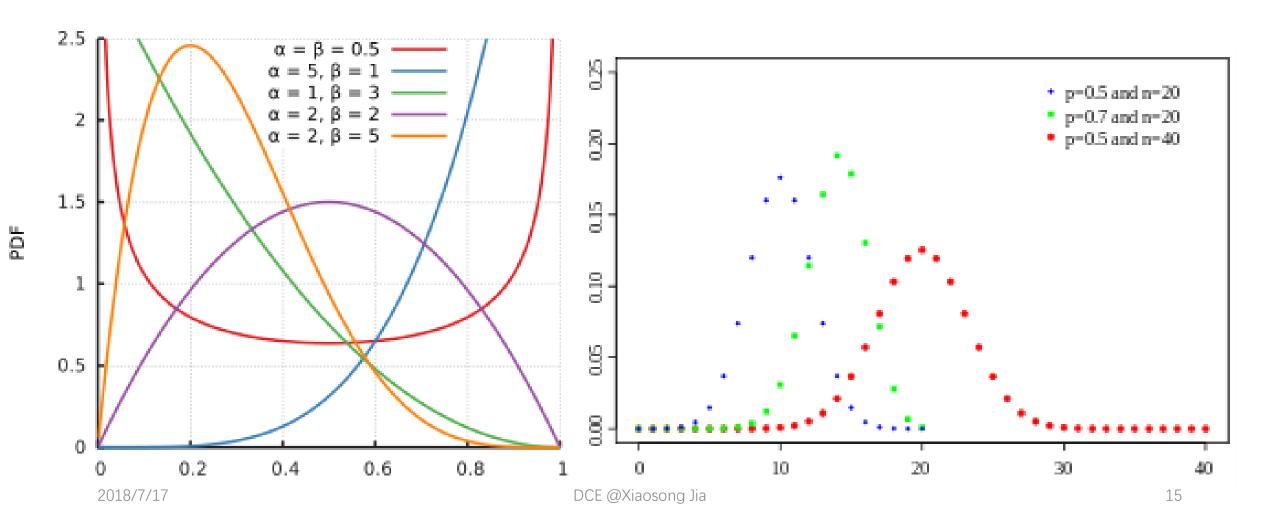
• Predicting the outcome of multiple future trials
Suppose now we were interested in predicting the number of heads, x, in M future trials

#### beta-binomial distribution

$$Bb(x|a,b,M) \triangleq {M \choose x} \frac{B(x+a,M-x+b)}{B(a,b)}$$

$$\mathbb{E}[x] = M \frac{a}{a+b}, \text{ var } [x] = \frac{Mab}{(a+b)^2} \frac{(a+b+M)}{a+b+D\text{LE @Xiaosong Jia}}$$

# Probability Density Graph



# Overfitting and the black swan paradox

- black swan paradox
- zero count problem or the sparse data problem
- Laplace's rule of succession(uniform prior)

$$p(\tilde{x} = 1|\mathcal{D}) = \frac{N_1 + 1}{N_1 + N_0 + 2}$$

### The Dirichlet-multinomial model

- Infer the probability that a dice with K sides comes up as face k.
- Likelihood

Suppose we observe N dice rolls, D =  $\{x_1, \ldots, x_N\}$ , where  $x_i \in \{1, \ldots, K\}$ .

If we assume the data is iid, the likelihood has the form

$$p(\mathcal{D}|\boldsymbol{\theta}) = \prod_{k=1}^{K} \theta_k^{N_k}$$

Prior(conjugate: Dirichlet distribution)

# Posterior and Posterior predictive

#### Posterior

$$p(\boldsymbol{\theta}|\mathcal{D}) \propto p(\mathcal{D}|\boldsymbol{\theta})p(\boldsymbol{\theta})$$

$$\propto \prod_{k=1}^{K} \theta_k^{N_k} \theta_k^{\alpha_k - 1} = \prod_{k=1}^{K} \theta_k^{\alpha_k + N_k - 1}$$

$$= \operatorname{Dir}(\boldsymbol{\theta}|\alpha_1 + N_1, \dots, \alpha_K + N_K)$$

MAP

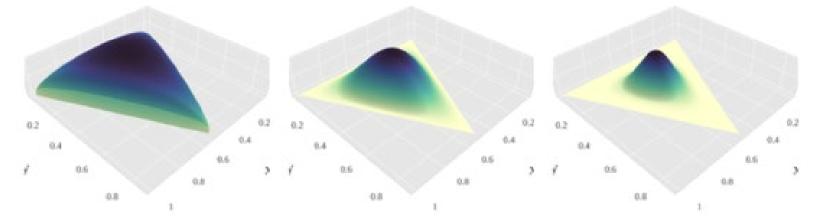
MLE

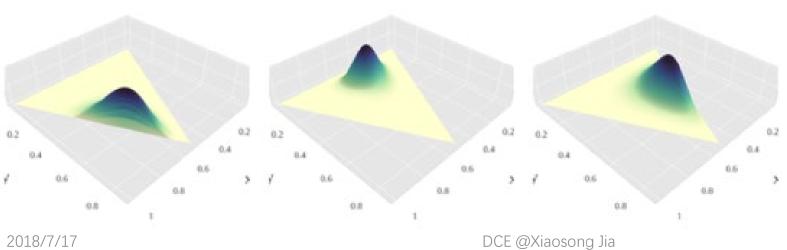
$$\hat{\theta}_{MAP} = \frac{(N_k + \alpha_k - 1)}{N + \alpha_0 - K} \qquad \hat{\theta}_{MLE} = \frac{N_k}{N}$$

Posterior predictive

$$\begin{split} p(X=j|\mathcal{D}) &= \int p(X=j|\theta)p(\theta|\mathcal{D})d\theta \\ &= \int p(X=j|\theta_j) \left[\int p(\theta_{-j},\theta_j|\mathcal{D})d\theta_{-j}\right] d\theta_j \\ &= \int p(X=j|\theta_j) \left[\int p(\theta_{-j},\theta_j|\mathcal{D})d\theta_{-j}\right] d\theta_j \\ &= \int p(X=j|\theta_j) \left[\int p(\theta_{-j},\theta_j|\mathcal{D})d\theta_j \right] = \frac{\alpha_j + N_j}{\sum_k (\alpha_k + N_k)} \exp\left(\frac{\alpha_j + N_j}{\sum_k (\alpha_k + N_k)}\right) \exp\left(\frac{\alpha_j + N_j}{\sum_k (\alpha_k + N_k$$

# Probability Density Graph





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# Naive Bayes classifiers

- Goal: classify vectors of discrete-valued features,  $\mathbf{x} \in \{1, ..., K\}^D$ , where K is the number of values for each feature, and D is the number of features.
- Assumption:

the features are **conditionally independent** given the class label.

$$p(\mathbf{x}|y=c,\boldsymbol{\theta}) = \prod_{j=1}^{D} p(x_j|y=c,\boldsymbol{\theta}_{jc})$$

### Naive Bayes classifiers

- In the case of real-valued features, we can use the Gaussian distribution:  $p(\mathbf{x}|y=c,\boldsymbol{\theta}) = \prod_{j=1}^{D} \mathcal{N}(x_j|\mu_{jc},\sigma_{jc}^2)$ , where  $\mu_{jc}$  is the mean of feature j in objects of class c, and  $\sigma_{jc}^2$  is its variance.
- In the case of binary features,  $x_j \in \{0,1\}$ , we can use the Bernoulli distribution:  $p(\mathbf{x}|y=c,\boldsymbol{\theta}) = \prod_{j=1}^D \mathrm{Ber}(x_j|\mu_{jc})$ , where  $\mu_{jc}$  is the probability that feature j occurs in class c. This is sometimes called the **multivariate Bernoulli naive Bayes** model. We will see an application of this below.

$$p(\mathbf{x}|y=c,\boldsymbol{\theta}) = \prod_{j=1}^{D} p(x_j|y=c,\boldsymbol{\theta}_{jc})$$

# Model fitting

- How to train?
- The probability for a single data(Using assumption)

$$p(\mathbf{x}_i, y_i | \boldsymbol{\theta}) = p(y_i | \boldsymbol{\pi}) \prod_j p(x_{ij} | \boldsymbol{\theta}_j)$$

• the log-likelihood

$$\log p(\mathcal{D}|\boldsymbol{\theta}) = \sum_{c=1}^{C} N_c \log \pi_c + \sum_{j=1}^{D} \sum_{c=1}^{C} \sum_{i:y_i=c} \log p(x_{ij}|\boldsymbol{\theta}_{jc})$$

Method:

let us suppose all features are binary

$$\hat{\pi}_c = \frac{N_c}{N} \qquad \hat{\theta}_{jc} = \frac{N_{jc}}{N_c}$$

#### • Data Set

ゆ。     不好。     矮。     不上进。     不嫁。       不帅。     好。     矮。     上进。     家。       不帅。     好。     高。     上进。     嫁。       小。     不好。     矮。     上进。     不嫁。       不好。     大好。     大力。     不成。
帅→     好→     矮→     上进→     嫁→       不帅→     好→     高→     上进→     嫁→       帅→     不好→     矮→     上进→     不嫁→
不帅。     好。     高。     上进。     嫁。       帅。     不好。     矮。     上进。     不嫁。
<b>帅</b> ↵ 不好↵ 矮↵ 上进↩ 不嫁↩
구대 구녀 본 구나내 구녀
<b>不帅。</b> 不好。    矮。    不上进。    不嫁。
ケッ
<b>不帅</b> 。
<b>帅</b> ₽ 好₽ 高₽ 上进₽ 嫁₽
<b>不帅</b> 。
<b>帅</b> 。
帅。     好。     矮。     不上进。     不嫁。

- 现在给我们的问题是,如果一对男女朋友,男生向女生求婚,男生的四个特点分别是不帅,性格不好,身高矮,不上进,请你判断一下女生是嫁还是不嫁?
- 转换为数学语言
- p(嫁|(不帅、性格不好、身高矮、不上进))与p(不嫁|(不帅、性格不好、身高矮、不上进))哪个大?

• 朴素贝叶斯公式

p(不嫁) = 6/12(总样本数) = 1/2

```
p(g|\text{不帅、性格不好、身高矮、不上进}) = \frac{p(\text{不帅、性格不好、身高矮、不上进}|g)*p(g)}{p(\text{不帅、性格不好、身高矮、不上进})}
= \frac{p(不帅|嫁)*p(性格不好|嫁)*p(身高矮|嫁)*p(不上进|嫁)*p(嫁)}{p(不帅)*p(性格不好)*p(身高矮)*p(不上进)}
      p(嫁) = 6/12(总样本数) = 1/2
```

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<b>帅?</b>	性格好?』	身高?◢	上进?	嫁与否。
不帅↵	好ℯ	高↩	上进↩	嫁↵
不帅↵	好♪ http:/	/l‡og. csdn. net∕yi	z <u>le</u> 进@lp	嫁↩
不帅↵	不好↵	高↩	上进↩	嫁↩

p(不帅|嫁) = 3/6 = 1/2

. . . . . .

帅?。	性格好?	身高?』	上进?』	嫁与否。	
帅↩	不好₽	矮↵	不上进↵	不嫁↵	•
不帅↩	好↩	矮↵	上进↩	不嫁↩	4
帅↩	好↩	矮↵	上进↩	嫁↩	•
不帅↵	好↩	高↩	上进↩	嫁↩	•
帅↩	不好₽	矮↵	上进↩	不嫁↩	4
帅。	不好₽	矮↓ tp://blog.esdn.	上进→	不嫁↩	4
帅↩	好₽	ip:// <u>h</u> log. esuil.	不上进ℯ	嫁↩	4
不帅₽	好↩	<b>中</b> ₽	上进↩	嫁↩	4
<b>冲</b> 。	好↩	中↩	上进↵	嫁↩	•
不帅↩	不好↩	高↩	上进↩	嫁↩	4
帅↩	好↩	矮↵	不上进↵	不嫁↩	4
<b>冲</b> ≈	好↩	矮↩	不上进↵	不嫁↩	4

. . . . . .

```
p(g|| \text{不帅、性格不好、身高矮、不上进}) = \frac{p(\text{不帅、性格不好、身高矮、不上进}|g)*p(f)}{p(\text{不帅、性格不好、身高矮、不上进})}
p(\text{不帅}||g)*p(性格不好||g)*p(身高矮||g)*p(不上进||g)*p(嫁)
```

 $= \frac{p(不帅|嫁)*p(性格不好|嫁)*p(身高矮|嫁)*p(不上进|嫁)*p(嫁)}{p(不帅)*p(性格不好)*p(身高矮)*p(不上进)}$ 

= (1/2\*1/6\*1/6\*1/6\*1/2)/(1/3\*1/3\*7/12\*1/3)

同理可得 p (不嫁|不帅、性格不好、身高矮、不上进)

# Bayesian naive Bayes

- Above method: it can overfit.
- Solution: use a factored prior

$$p(\boldsymbol{\theta}) = p(\boldsymbol{\pi}) \prod_{j=1}^{D} \prod_{c=1}^{C} p(\theta_{jc})$$

We will use a Dir( $\alpha$ ) prior for  $\pi$  and a Beta( $\beta_0$ ,  $\beta_1$ ) prior for each  $\theta$ jc. Often we just take  $\alpha=1$  and  $\beta=1$ , corresponding to add-one or Laplace smoothing

# Bayesian naive Bayes

Posterior

$$p(\boldsymbol{\theta}|\mathcal{D}) = p(\boldsymbol{\pi}|\mathcal{D}) \prod_{j=1}^{D} \prod_{c=1}^{C} p(\theta_{jc}|\mathcal{D})$$

$$p(\boldsymbol{\pi}|\mathcal{D}) = \text{Dir}(N_1 + \alpha_1 \dots, N_C + \alpha_C)$$

$$p(\theta_{jc}|\mathcal{D}) = \text{Beta}((N_c - N_{jc}) + \beta_0, N_{jc} + \beta_1)$$

Using the model for prediction

$$p(y = c | \mathbf{x}, \mathcal{D}) \propto \overline{\pi}_c \prod_{j=1}^D (\overline{\theta}_{jc})^{\mathbb{I}(x_j = 1)} (1 - \overline{\theta}_{jc})^{\mathbb{I}(x_j = 0)}$$

$$\overline{\theta}_{jk} = \frac{N_{jc} + \beta_1}{N_c + \beta_0 + \beta_1}$$

$$\overline{\pi}_c = \frac{N_c + \alpha_c}{N + \alpha_0}$$

### Document classification

- Goal: classify text documents into different categories
- let  $x_i$  be a vector of counts for document i, so  $x_{ij} \in \{0, 1, ..., N_i\}$ , where  $N_i$  is the number of terms in document i (so  $\sum_{j=1}^{D} x_{ij} = N_i$ ).
- For the class conditional densities, we can use a multinomial distribution

$$p(\mathbf{x}_i|y_i = c, \boldsymbol{\theta}) = \operatorname{Mu}(\mathbf{x}_i|N_i, \boldsymbol{\theta}_c) = \frac{N_i!}{\prod_{j=1}^D x_{ij}!} \prod_{j=1}^D \theta_{jc}^{x_{ij}}$$

Here  $\theta_{jc}$  is the probability of generating word j in documents of class c; these parameters satisfythe constraint that  $\sum_{j=1}^{D} \theta_{jc} = 1$  for each class c.

# Quiz

• Calculate p (不嫁|不帅、性格不好、身高矮、不上进) with naïve Bayesian Classifier