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Disastrous disappointments: Asset-pricing with disaster risk and disappointment aversion

Presentation for ESEM 2013, Gothenburg

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07/12/13

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 Look at asset-pricing implications of a rare disasters (RD) model where representative agent is endowed with plausible degree of disappointment aversion (DA).

- Two environments in paper:
 - Similar to Rietz (1988 JME), 'normal times' given by Mehra-Prescott Markov chain, single disaster state, preferences collapse to EU-CRRA if no DA. Just to illustrate mechanisms at work.
 - Richer model in the spirit of Barro (2006 QJE) and others, distribution of disaster sizes, partial default on 'risk-free' asset in disaster states, levered equity, fully utilize EZ aggregator.
 - Focus on second in this talk.
- Disappointment aversion as defined by Gul (1991 Ecta), used in asset-pricing literature by Routledge and Zin (2010 JF) or Campanale, Castro and Clementi (2010 RED).

Why would we want to do this?

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- Disaster risk models have recently had considerable success in rationalizing data on consumption and asset returns.
- Literature was dormant after Rietz, jump-started by Barro (2006, QJE).
- Disaster risk is hard to quantify with any precision given limited data.
 - Does a disaster happen once every couple hundred years or a couple times every hundred years?
 - Hard to say with 100 years of data.
- Barro: used cross-country evidence.

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- Adding more countries' experiences gives more observations to estimate disaster risk with, but only works if we assume all countries draw disasters from a common distribution.
- Debatable whether the sizable risks Barro finds—e.g., 1.7 percent chance of average 29% decline—are a good characterization of risk faced by, say, US investors.
- If risks are a lot smaller for the US than indicated by multi-country estimates, does disaster risk no longer explain US asset returns data?
- That's where DA comes in.

Rationale for adding DA to RD

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- We have experimental and field evidence that people over-weight bad or disappointing outcomes.
- If we endow agent in RD model with plausible degree of DA, maybe we can get good results with much smaller disaster risks?
- Paper shows that is in fact the case.
- Empirical discipline imposed by keeping DA in the range of estimates from experimental literature—Choi et al. (2007, AER) and Camerer and Ho (1994, JRU).

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 DA does complement RD: Can get good results with disaster probabilities and sizes one-half or one-fourth the average sizes in Barro's QJE paper.

- Quality of model's fit depends on assumption about EIS.
- For a low EIS (0.10), can match almost exactly the means and standard deviations of risk-free rate and equity return.
- For larger EIS (1 or 1.5), fit is less satisfactory, but still broadly similar to results obtained by Gourio (2010, FRL).
 - The one big departure is that our model doesn't generate enough volatility in the risk-free rate.
 - Not often viewed as a problem.

Roadmap of the talk

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- One of the earliest responses to Mehra-Prescott (Rietz 1988 JME), but lost favor.
 - Mehra-Prescott's response (1988 JME): disasters are too big; risk aversion, discount factor too high; risk-free asset not likely risk-free in a crisis.
 - Merits just a footnote mention in Kocherlakota's 1996 JEL survey.
- Interest revived by empirical work of Barro (2006 *QJE*), theoretical contributions by Gabaix, Gourio, others.
- Disaster risk provides potential resolution to equity premium puzzle by providing a channel for holding down risk-free rate.

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- As with Epstein-Zin preferences, uncertain future utility reduced to a certainty equivalent value. Disappointment aversion enters through the form of the certainty equivalent operator.
 - Disappointing outcomes get extra weight.
- But what outcomes are disappointing?
- Gul (1991 Ecta), certainty equivalent obeys a consistency requirement—threshold for disappointment is certainty equivalent itself.
 - So certainty equivalent only defined implicitly.

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- Very roughly paraphrased: disappointment averse agents take expectations over-weighting outcomes that would fall below expectations.
- Used by Routledge and Zin (2010 JF) in a Mehra-Prescott framework; Campanale, Castro, Clementi (2010 RED) in a model with capital accumulation.
 - Routledge and Zin use a 'generalized' form of DA that allows them some control over which states are disappointing.

Disappointment aversion

Disappointment aversion: Formally

 More formally, the agent's intertemporal preferences have the Epstein-Zin form

$$U_t = [(1-\beta)c_t^{1-\rho} + \beta\mu_t(U_{t+1})^{1-\rho}]^{1/(1-\rho)}$$

for $\rho \geq 0$, $\rho \neq 1$.

- The agent's EIS is ρ^{-1} .
- Conditionality in μ_t eventually will come from Markov structure of consumption growth; for now, just ignore it.
 - I'll drop the t for now to avoid confusion.

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• Let z be a discrete random variable taking the value z_i with probability p_i (given the state at t). The certainty equivalent of z is defined implicitly by

$$\mu(z)^{1-\alpha} = \frac{\sum_{i} p_{i} z_{i}^{1-\alpha} [1 + \theta I (z_{i} < \mu(z))]}{1 + \theta \sum_{i} p_{i} I (z_{i} < \mu(z))}$$

- $I(\cdot)$ is an indicator function, taking the value one when $z_i < \mu(z)$ (z_i disappoints), and zero otherwise.
- $\alpha \geq 0$, $\alpha \neq 1$ and $\theta \geq 0$ are parameters.

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• Suppose state i = 1 is the only disappointing state. Then:

$$\mu(z)^{1-\alpha} = \frac{1+\theta}{1+\theta p_1} p_1 z_1^{1-\alpha} + \frac{1}{1+\theta p_1} \sum_{i \neq 1} p_i z_i^{1-\alpha}$$

• This is precisely the form of a standard, Epstein-Zin certainty equivalent, with distorted probabilities \hat{p} ,

• Think of state 1 as a rare disaster. Then, $p_1 = .01$ and $\theta = 1$ gives same c.e. as $p_1 = .02$ and $\theta = 0$.

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• What does the SDF look like?

- Assume log consumption growth, x_{t+1} , follows a Markov chain, $\{x_i, P_{i,j}\}$.
- The SDF has the form:

$$m_{i,j} = \beta e^{-\rho x_j} \left(\frac{\phi_j e^{x_j}}{\mu_i (\phi e^x)} \right)^{\rho - \alpha} \frac{1 + \theta I_{i,j}}{1 + \theta \sum_j P_{i,j} I_{i,j}},$$

which is similar to the version given by Campanale et al.

• ϕ is lifetime utility scaled by consumption (v_t/c_t) , and $I_{i,j}=1$ if state j is disappointing conditional on state i today.

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• Starting point is Mehra and Prescott's Markov chain—two states, with gross consumption growth y_L and y_H , and symmetric transition matrix Q.

- Add a distribution of N disaster states D_i with relative frequencies f_i and relative sizes z_i.
 - $\sum_i f_i = 1$ and $\sum_i f_i z_i = 1$.
- If average disaster size is b, then size of disaster in state D_i is bz_i.
- If overall disaster probability is p, then probability of state D_i is
 pf_i.

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• Modify transition matrix as follows:

- If the state today is either L or H, then the D_i state occurs with probability pf_i.
- If today's state is a disaster state, growth returns next period to $\{y_L, y_H\}$ with probabilities $\{1/2, 1/2\}$.
- There is zero probability of remaining in a disaster state or transitioning from one disaster state to another.
- Unlike Rietz, I assume Mehra-Prescott Markov chain describes "normal" (non-disaster) times.
 - Keep y_L, y_H, Q fixed as we vary p, b.

► Markov chain

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 I consider claims to a one-period 'riskless' asset, paying one unit of consumption next period (if no disaster), and two forms of 'equity.'

 Riskless asset is 'riskless' because I allow partial default in the disaster state. Call its return the bill rate to avoid confusion.

- Two equity claims:
 - One is standard claim to aggregate consumption
 - The other (a 'dividend' claim) is a claim to a leveraged version of aggregate consumption.
- Returns are defined in the usual way. PRETURNS

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• Vary the environment by varying p, disaster probability, and b, the average disaster size. Find parameters of certainty equivalent— θ , strength of DA, and α , curvature—that get us close to first and second moments of returns data (from Gourio 2008 *FRL*).

- I use discrete grids for θ and α .
- $\theta \in \{0, 0.1, \dots, 3\}, \ \alpha \in \{0, 0.1, \dots 5\}.$
- The $\theta \leq 3$ restriction keeps us in the empirically plausible range. Choosing θ
- Model moments I compare to data are for non-disaster samples (average behavior in 'normal times').
- Focus on bill rate and dividend (leveraged consumption) return.

Structure

Starting point is p and average b from Barro's QJE paper:

p = 0.017

b = 0.29

 Consider in turn (and in combinations) disaster probabilities and disaster sizes that are 1/2 or 1/4 the averages from Barro's data. I.e., p and b such that

$$p \in \{0.017, 0.0085, 0.00425\}$$

 $b \in \{0.29, 0.145, 0.0725\}.$

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 Some parameters fixed throughout experiments—discounting, leverage, default fraction, EIS, relative sizes and frequencies of disaster states.

- In all the environments, data prefer a very low EIS, around 0.1. But we know there are also good reasons to consider EIS bigger than one. Simulate model for EIS of 0.1, 1.0, 1.5.
- I follow Gourio (2008 *FRL*), in setting fraction of bill payoff defaulted on in disaster (η_i) to be 40% of the disaster size.

Basic parameters

$$\frac{\rho^{-1}}{\{0.10, 1.0, 1.5\}}$$
 $\frac{\beta}{0.97}$ $\frac{\lambda}{3}$ $\frac{\eta_i}{0.4bz_i}$ Derived from Barro

Results: Low EIS case

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• The model's best fits are for the low EIS case.

Results for $(\rho^{-1}=0.10)$											
	p (%)	b (%)	θ	α	$E(R^f)$	$E(R^d)$	$\sigma(R^f)$	$\sigma(R^d)$			
	1.70	29.0	0.3	2.9	1.02	9.58	4.37	15.17			
	0.85	29.0	0.9	2.8	1.01	9.33	3.97	14.81			
	0.43	29.0	1.6	2.1	1.04	9.09	3.71	14.55			
	1.70	14.5	1.2	3.3	1.04	9.32	3.75	14.64			
	1.70	7.25	1.9	1.5	1.06	9.04	3.56	14.42			
	0.85	14.5	1.7	2.6	1.04	9.21	3.62	14.51			
	0.43	7.25	2.4	0.3	1.03	8.97	3.52	14.37			
	0	0	2.5	0.5	1.04	9.07	3.50	14.36			
	Data	_	_	_	1.03	8.91	4.36	15.04			

 Data are from Gourio (2008 FRL), based on 1926–2004 sample used in Cochrane (2008 RFS).

Results: Large EIS case

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Results f	or (ho^{-1})	= 1.5	5)					
p (%)	b (%)	θ	α	$E(R^f)$	$E(R^d)$	$\sigma(R^f)$	$\sigma(R^d)$	
1.70	29.0	0.0	4.0	1.11	10.23	0.25	10.58	
0.85	29.0	0.0	4.8	1.07	9.74	0.24	10.52	
0.43	29.0	8.0	5.0	1.14	10.89	0.14	10.86	
1.70	14.5	3.0	5.0	2.14	12.20	0.12	11.33	
1.70	7.25	3.0	5.0	2.48	10.84	0.16	11.08	
0.85	14.5	3.0	5.0	2.35	11.21	0.14	11.14	
0.43	7.25	3.0	5.0	2.55	10.32	0.17	10.97	
Data	_	_	_	1.03	8.91	4.36	15.04	

- Results are less satisfactory—especially volatility of the bill rate.
- Even so, apart from standard deviation of bill rate, results are broadly similar to those obtained by Gourio (even as we greatly reduce disaster risk).
- Bang-bang nature of optimal θ choice also interesting—either very small or at upper bound.

Conclusions, directions

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- Smaller disaster risks can still generate realistic asset returns, when agent is endowed with empirically plausible degree of disappointment aversion.
- · Obvious directions for further work:
 - Time-varying disaster probability (useful for getting times series properties of the data, but problems imposing empirical discipline).
 - Generalized disappointment aversion (as in Routledge-Zin); can more control over which states disappointment help with the large-EIS results?
 - Changing the shape of the distribution of disasters, rather than
 just joint disaster probability p, average size b.

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The current version of the paper and Matlab code for solving the model are available at:

www.jimdolmas.net/economics/current-work

I would value your feedback; email me at jim@jimdolmas.net

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• Recall that in representative agent, CRRA, lognormal case, risk-free rate obeys:

$$R^f = -\log(\beta) + \alpha\gamma - \frac{\alpha^2\sigma^2}{2}$$

where α is the CRRA coefficient, γ and σ are mean and standard deviation of log consumption growth.

- Well-known that for typical estimates of γ and σ , the first-order term $\alpha\gamma$ dominates as α increases.
 - A high value of α, necessary to get a non-negligible equity risk premium, thus produces a counterfactually high risk-free rate.

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• Let a disaster add $\log(1-b) < 0$ to $\log(c_{t+1}/c_t)$, and occur with probability p (independently of the log-normal component). Then, the log risk-free rate becomes

$$R^f = -\log(\beta) + \alpha\gamma - \frac{\alpha^2\sigma^2}{2} - \log(1 - p + p(1 - b)^{-\alpha}).$$

• For typical values of γ and σ , and for small values of p, we can generate low values of R^f even for large values of α , and with $\beta < 1$.

Supplemental slides

 Mehra and Prescott's two-state Markov chain for gross consumption growth y_{t+1} is:

$$y_{t+1} \in \begin{bmatrix} y_L \\ y_H \end{bmatrix} = \begin{bmatrix} 0.982 \\ 1.054 \end{bmatrix}$$

with the transition matrix

$$Q = \begin{bmatrix} Q_{LL} & Q_{LH} \\ Q_{HL} & Q_{HH} \end{bmatrix} = \begin{bmatrix} 0.43 & 0.57 \\ 0.57 & 0.43 \end{bmatrix}$$

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• The modified Markov chain is given by the set of states

$$y_{t+1} \in \begin{bmatrix} \mathbf{1} - b\mathbf{z} \\ 0.982 \\ 1.054 \end{bmatrix}$$

and the transition matrix

$$P = \begin{bmatrix} \mathbf{0} & \frac{1}{2}[\mathbf{1}, \mathbf{1}] \\ p[\mathbf{f}, \mathbf{f}]^{\top} & (1 - p)Q \end{bmatrix}$$

z denotes the z_i's, f denotes the f_i's, 0 a matrix of 0's, 1 a vector of 1's.

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• The 'riskless' asset here is a one-period zero-coupon bond (in zero net supply) whose price at t is q_t and whose payoff at t+1 is F_{t+1} , where

$$F_{t+1} = \begin{cases} 1 - \eta_i & \text{if } y_{t+1} = y_{D,i} \\ 1 & \text{if } y_{t+1} \in \{y_L, y_H\} \end{cases}$$

• In the experiments, I'll follow Gourio (2010 *FRL*) in making η_i a fraction of the disaster size in state D_i .

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• The 'dividend' claim is a claim to a process ('dividends') whose log growth rate is a multiple of log consumption growth.

• Letting y_{t+1}^d denote gross dividend growth from t to t+1,

$$y_{t+1}^d = (y_{t+1})^\lambda,$$

where $\lambda \geq 1$ is the "leverage" parameter.

• This dividend formulation is now pretty standard—used, for example, in Routledge and Zin, Gourio, Bansal-Yaron, etc.

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• Returns are defined in the standard way.

• If q is the price of the 'riskless asset,' and w^c and w^d denote the price-dividend ratios for the two equity claims, then the key pricing equations are

$$\begin{aligned} q_t &= \mathsf{E}_t[m_{t+1} F_{t+1}] \\ w^c_t &= \mathsf{E}_t \left[m_{t+1} y_{t+1} \left(1 + w^c_{t+1} \right) \right] \\ w^d_t &= \mathsf{E}_t \left[m_{t+1} y^d_{t+1} \left(1 + w^d_{t+1} \right) \right], \end{aligned}$$

Asset returns

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• Then, for Markov states *i* today and *j* tomorrow:

$$\begin{split} R_{i,j}^f &= \frac{F_j}{q_i} & \text{(Risk-free rate)} \\ R_{i,j}^c &= \frac{y_j(w_j^c+1)}{w_i^c} & \text{(Consumption claim return)} \\ R_{i,j}^d &= \frac{y_j^d(w_j^d+1)}{w_i^d} & \text{(Dividend claim return)} \end{split}$$

where F_j is $1 - \eta_j$ if $y_j = y_{D,j}$, 1 otherwise.

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- Experimental evidence on size of θ :
 - Choi, Fisman, Gale and Kariv (2007 AER) estimate mean around 0.3, with ninety percent of estimates for individual subjects between 0 and around 1.4.
 - Camerer and Ho (1994, JRU) synthesize data from nine experimental choice studies, estimate $\theta=3$ treating the choice data as coming from a single representative agent.
- How does our $\theta \in [0,3]$ compare with calibrations in other work on disappointment aversion and asset pricing?
 - Campanale et al. rely on Choi et al., consider $\theta \in [0, 0.4]$.
 - Routledge and Zin use θ's from 9 (mostly) to 24 (in one case).
 Given their generalized form of DA, hard to say what experimental evidence implies for their θ.

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