

R Notebook for Chapter 6: Mortality surface modeling

Companion Code to Gaussian Process Models for Quantitative Finance

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This RMarkdown file presents an illustrative use of Gaussian Process surrogates for mortality surface modeling, linking to **Section 6.4** of the book.

Accessing Mortality Data

We make use of the following packages

- **DiceKriging**: kriging methods for single-output GP (SOGP).
- **rgenoud**: genetic optimization solver used for Maximum Likelihood Estimation of GP hyperparameters.
- **data.table**: extension of **data.frame** to enhance data manipulation.

The script in the `createMortData.R` loads the csv files containing the Deaths and Exposures for the respective population.

```
library(dplyr)
library(rgenoud)
library(DiceKriging)
library(data.table)
library(plot3D)
source("createMortData.R")
```

The input data is an R dataframe with factors **age**, **year**, **deaths** and **exposures**. The corresponding log mortality rates are computed as

$$y^n = \log(D^n/E^n)$$

where D^n and E^n are, respectively, the number of deaths and midyear count of exposed lives for the n th age/year pair $\mathbf{x}^n = (x_{ag}^n, x_{yr}^n)$.

For the illustrations below we work with the dataset for Denmark Males.

```
mortData = createMortData(year_start=1990, year_end=2016,
age_start=60, age_end=80, sex="m", sex_cat="no") # M=males
mortData = mortData[country %in% c("Denmark")]
head(mortData)
```

##	age	year	country	rate	y
## 1:	60	1990	Denmark	0.01561052	-4.159810
## 2:	60	1991	Denmark	0.01602873	-4.133373
## 3:	60	1992	Denmark	0.01762344	-4.038525
## 4:	60	1993	Denmark	0.01486386	-4.208822
## 5:	60	1994	Denmark	0.01553909	-4.164396
## 6:	60	1995	Denmark	0.01599507	-4.135475

Fitting the baseline GP model

For our Gaussian Process Regression models, we utilize the package **DiceKriging** available on CRAN.

The function **km()** is used to fit a GP model based on input-output set (x, y) and the following parameters:

- *formula* determines the mean function $\mu(x)$; **formula = ~x.age** corresponds to a prior mean function that is linear in Age $\mu(x) = \beta_0 + \beta_{ag}^1 x_{ag}$
- *covtype* refers to kernel type, taken to be Matern-5/2, separable in Age and Year.
- *nugget.estim=TRUE* tells **km()** to infer the intrinsic (homoskedastic, input-independent) noise variance σ_ϵ^2 as part of the model.
- *optim.method="gen"* tells **km()** to utilize a genetic optimization algorithm from library **rgenoud**, which produces more reliable hyperparameter estimates and is recommended by **DiceKriging** authors.
- *control=...* are internal (recommended) parameters of the above optimization algorithm
- See package manual <https://cran.r-project.org/web/packages/DiceKriging/DiceKriging.pdf> for a more detailed explanation of **km()** options.

```
xMort <- data.frame(age = mortData$age, year = mortData$year)
yMort <- mortData$y
mortModel_nug <- km(formula = ~x.age,
                     design = data.frame(x = xMort), response = yMort,
                     nugget.estim=TRUE,
                     covtype="matern5_2",
                     optim.method="gen",
                     # the "control" parameters below handle speed versus risk of
                     # converging to local minima. See "rgenoud" package for details
                     control=list(max.generations=100,pop.size=100,wait.generations=8,
                                   solution.tolerance=1e-5,trace=F))
```

```
mortModel_nug
```

```
##
## Call:
## km(formula = ~x.age, design = data.frame(x = xMort), response = yMort,
##     covtype = "matern5_2", nugget.estim = TRUE, optim.method = "gen",
##     control = list(max.generations = 100, pop.size = 100, wait.generations = 8,
##                     solution.tolerance = 1e-05, trace = F))
##
## Trend  coeff.:
##              Estimate
## (Intercept)  -9.8521
##      x.age    0.0919
##
## Covar. type  : matern5_2
## Covar. coeff.:
##              Estimate
## theta(x.age)  24.2729
## theta(x.year)  9.8827
##
## Variance estimate: 0.03687721
##
```

```
## Nugget effect estimate: 0.001787395
```

The above code output shows estimates of (in order) $\beta_0, \beta_{ag}^1, \theta_{ag}, \theta_{yr}, \eta^2$ and σ_ϵ^2 . Despite the assumption of a nugget effect, `km()` by default treats the observations as without having noise variance. To be consistent with our setting, we re-enter the model using the above fitted parameters along with inputting `noise.var`:

```
nug <- mortModel_nug@covariance@nugget
mortModel <- km(formula = ~x.age, design = mortModel_nug@X, response = mortModel_nug@y,
  noise.var = rep(nug, mortModel_nug@n),
  coef.trend = mortModel_nug@trend.coef, # re-use obtained hyperparameters
  coef.cov = mortModel_nug@covariance@range.val,
  coef.var = mortModel_nug@covariance@sd2,
  covtype = mortModel_nug@covariance@name)
```

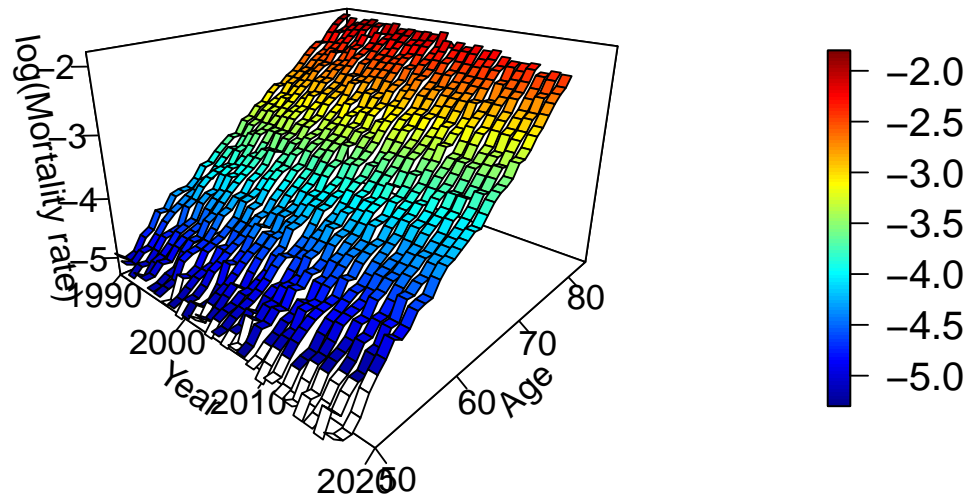
Figure 6.2: raw and predicted mortality surfaces

We use the `plot3D` library to visualize the raw log-mortality surface (the ribbons emphasize that data is collected annually) for Males in Denmark.

```
agesObserved <- 50:84
yearsObserved <- 1990:2016
mortData_ext = createMortData(year_start=min(yearsObserved), year_end=max(yearsObserved),
  age_start=min(agesObserved), age_end=max(agesObserved), sex="m", sex_cat="no")
mortData_ext = mortData_ext[country %in% c("Denmark")]

rateObserved <- dplyr::filter(mortData_ext, age %in% agesObserved, year %in% yearsObserved)$rate
rateObserved <- matrix(log(rateObserved), nrow=length(yearsObserved));

ribbon3D(yearsObserved, agesObserved, z=rateObserved, along="y", expand=0.67,
  border="black", zlim=c(-5.3, -1.8), phi=22, theta=42, xlim=c(1990, 2020),
  xlab="Year", ylab="Age", zlab="log(Mortality rate)", space=0.1,
  colkey = list(plot=FALSE), clim=c(-5.3, -1.8),
  ticktype="detailed", nticks=4, cex.axis=1.2, cex.lab=1.2)
colkey(side=4, length = 0.75, width = 0.6, dist=-0.09, cex.axis=1.3,
  clim=c(-5.3, -1.8), add=TRUE)
```



We then show the smooth surface based on GP projections from the previous **mortModel**. The predictive surface extends beyond the range of observed data: covering Ages [55, 85] in the plot (beyond the training range of [60, 80]) and covering the future years of 2017–2024 relative to the training range of 1990–2016.

```
agesForecast <- 55:85
yearsForecast <- 1990:2024

# build data frame for desired forecasts, then call predict
xPredImage <- data.frame(expand.grid(agesForecast, yearsForecast))
colnames(xPredImage) <- c("age", "year")

mortPredImage <- predict(mortModel, newdata=data.frame(x=xPredImage),
                        cov.compute=TRUE,
                        se.compute=TRUE, type="UK")
mortPredImage <- matrix(mortPredImage$m, nrow=length(yearsForecast), byrow=TRUE)

persp3D(yearsForecast, agesForecast, z=mortPredImage, border="black",
        xlab="Year", ylab="Age", zlab="log(Mortality rate)", zlim=c(-5.3, -1.8),
        colkey = FALSE, clim=c(-5.3, -1.8),
        ticktype="detailed", phi=22, theta=42, expand=0.67, cex.axis=1.2,
        cex.lab=1.2, nticks=4 )
colkey(side=4, length = 0.75, width = 0.6, dist=-0.09, cex.axis=1.3,
        clim=c(-5.3, -1.8), add=TRUE)
```

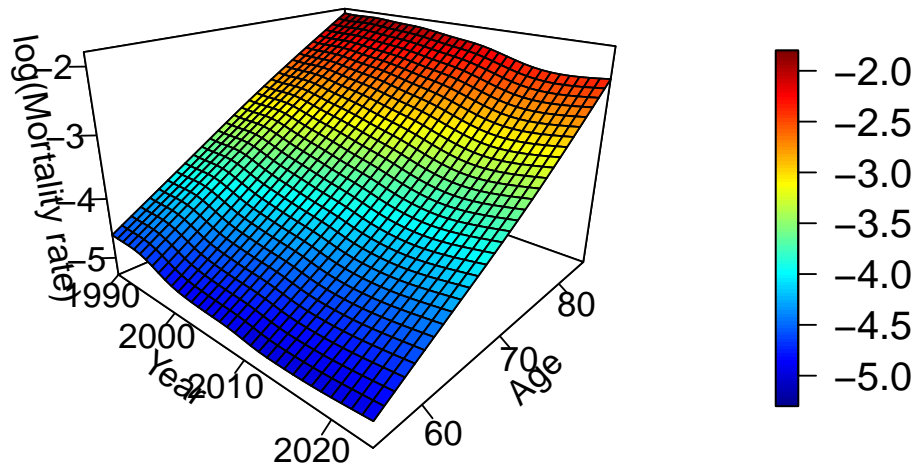


Fig 6.3 left: Mortality over time

After constructing a **km** object, the main workhorse is the **predict()** command. Among its outputs are the posterior mean, posterior standard deviation, and 95% credible bands. The **predict()** command allows for *any* test set with the same number of covariates as the training set. Our implementation of building a model with noise variance equal to the nugget causes **predict()** to return credible bands for the latent mortality surface $f(\cdot)$ and not the observation process $Y(\cdot)$. Thus we need to add σ_ϵ^2 to achieve the bands for the observed mortality y that would be relevant for coverage tests.

As illustration, we show evolution of mortality across time for a fixed Age. Decreasing trend corresponds to mortality improvement. Below, we additionally show the 95% credible intervals that quantify the (statistical) credibility of the smoothing. We use the library **dplyr** for easier data handling.

```
agesForecast <- 70
yearsForecast <- 2000:2026
yearsObserved <- 1999:2016

nYr <- length(yearsForecast)
nAg <- length(agesForecast)
xPred <- data.frame(age=rep(agesForecast,each=nYr), year=rep(yearsForecast,nAg))

mortPred <- predict(mortModel, newdata=data.frame(x=xPred), cov.compute=TRUE,
                    se.compute=TRUE, type="UK")

xPred$m <- mortPred$mean
y_sd = sqrt((mortPred$sd ^2) + nug)
xPred$lower95 = mortPred$mean - 1.96 * y_sd # predictive CI for y at 95% using z_0.975=1.96
xPred$upper95 = mortPred$mean + 1.96 * y_sd
```

```

for(ageObs in agesForecast){
  rateObserved <- dplyr::filter(mortData, age == ageObs, year %in% yearsObserved)$rate
  dataPred <- dplyr::filter(xPred, age == ageObs, year %in% yearsForecast)

  # exponentiate to show actual mortality rates (not logged)
  ratePred <- exp(dataPred$m);
  upper95Pred <- exp(dataPred$upper95); lower95Pred <- exp(dataPred$lower95)

  plot(yearsObserved,rateObserved,
       pch=18, main = paste("Age", ageObs),
       xlab="Year", ylab="Mortality Rate", cex.axis=1.4, cex.lab=1.4, cex=1.4,
       xlim=c(2000.5,2024), ylim=c(min(lower95Pred),max(upper95Pred)), bty="n",
       col="darkblue")
  lines(yearsForecast,ratePred, col=2, lwd=3)
  polygon(c(yearsForecast,rev(yearsForecast)),c(lower95Pred,rev(upper95Pred)),
        col="skyblue", density=25,lwd=1)

  legend("topright",c("Observed","Predicted","95% Credible Band"),
        lwd=c(2,2,2), lty=c(0,1,2), pch=c(18,NA,NA),
        col=c("darkblue",2,"skyblue"), cex=1.3, bty="n")
}

```

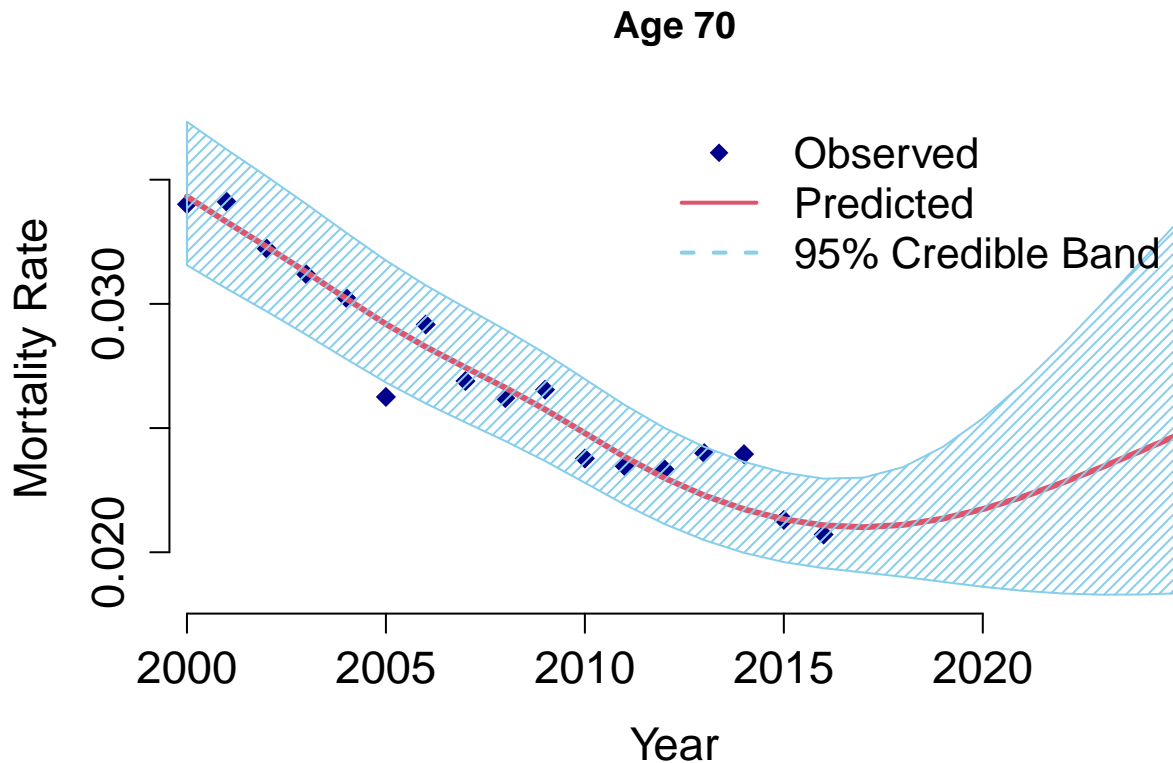


Figure 6.3 right: Mortality Improvement

It is common to interpret a mortality surface via the (annual) mortality *improvement factors* which measure longevity changes year-over-year. The raw annual percentage mortality improvement is:

$$MI^{raw}(\mathbf{x}) = 1 - \frac{\exp(y(x_{ag}, x_{yr}))}{\exp(y(x_{ag}, x_{yr} - 1))}.$$

The smoothed improvement factors are obtained by replacing y 's by the GP model posterior $m(\mathbf{x})$'s:

$$MI^{GP}(\mathbf{x}) := 1 - \frac{\exp(m_*(x_{ag}, x_{yr}))}{\exp(m_*(x_{ag}, x_{yr} - 1))}.$$

The following code produces improvement rates $MI^{GP}(\mathbf{x})$ for every odd calendar year and plots them.

```
agesForecast <- 60:80
yStart = 2008; yEnd = 2018;
yearsForecast <- yStart:yEnd

nYr <- length(yearsForecast)
nAg <- length(agesForecast)
xPred <- data.frame(age=rep(agesForecast,each=nYr), year=rep(yearsForecast,nAg))
mortPred <- predict(mortModel, newdata=data.frame(x=xPred), cov.compute=TRUE,
                  se.compute=TRUE, type="UK")

xPred$m <- mortPred$mean
forwardPred <- dplyr::filter(xPred, age %in% agesForecast, year %in% (yStart+1):yEnd)$m
backwardPred <- dplyr::filter(xPred, age %in% agesForecast, year %in% yStart:(yEnd-1))$m
mibackgp <- 1-exp(forwardPred)/exp(backwardPred) # GP-smoothed improvement rate
xPred$mibackgp <- NA
xPred$mibackgp[which(xPred$year > yStart)] <- mibackgp

miPlot <- dplyr::filter(xPred, age %in% agesForecast, year == yEnd)$mibackgp
plot(c(0,100), c(0,0), type="l", lwd=1, xlab="Age", xaxs="i",
     ylab="YoY Mortality Improvement Rate", cex.axis=1.4, cex.lab=1.4, cex=1.5,
     ylim=c(-0.01, 0.04), xlim=c(60,80), bty='n')
for(yr in seq(yStart+1,yEnd-1,2)){
  # plot the MI for that year
  miPlot <- dplyr::filter(xPred, age %in% agesForecast, year == yr)$mibackgp
  lines(agesForecast, miPlot, col=(yr-2007)/2, lty=(yr-2007)/2, lwd=2)
}
legend("bottomright", legend=seq(yStart+1,yEnd-1,by=2), col=c(1:5,1),
      lwd=rep(2,5), lty=c(1:5,1), cex=1.5, bty='n')
```

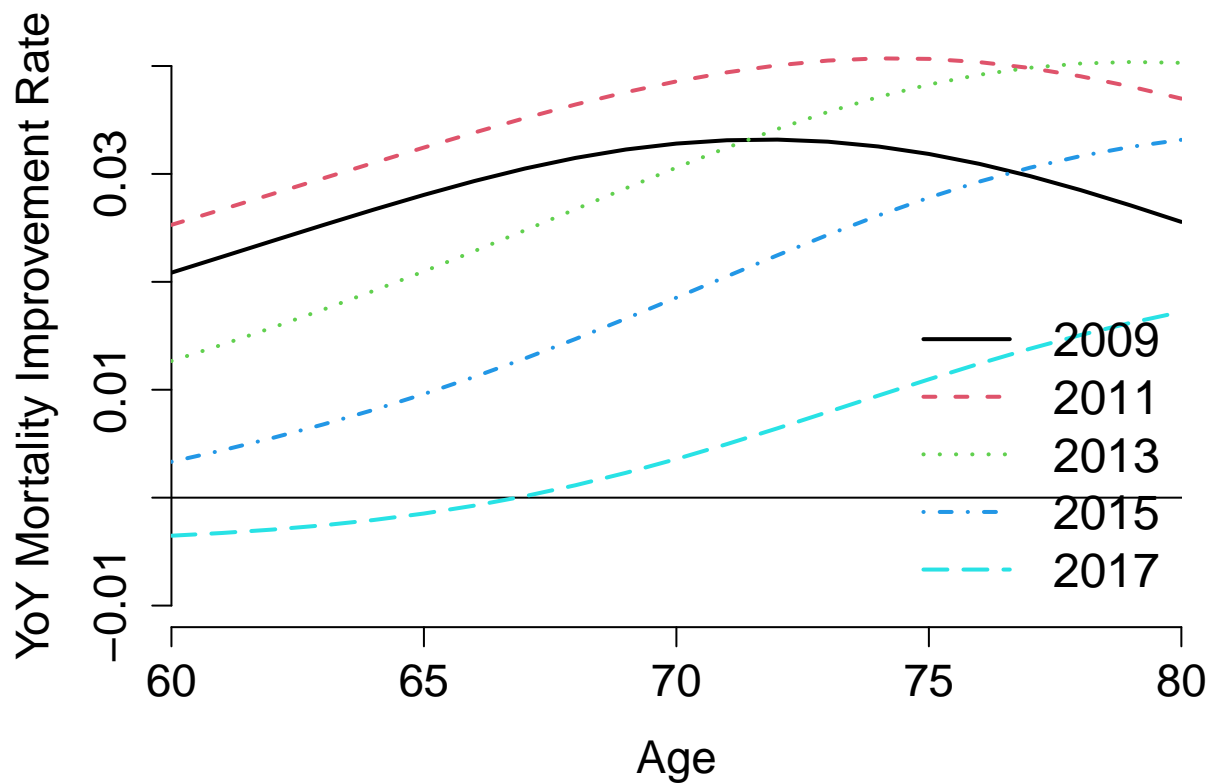


Figure 6.4 left: Life Expectancy Boxplot

To generate scenarios for period life expectancy, we train a model on a wider range of Ages. In the Danish dataset, observations are reliable up to age 103, so we use that for training the new `mortDataE`. Due to the larger training set (of size $27 \times 44 = 1188$ training inputs), this takes a couple of minutes.

```
mortDataE = createMortData(year_start=1990, year_end=2016,
                           age_start=60, age_end=103, sex="m", sex_cat="no")
mortDataE = mortDataE[country %in% c("Denmark")]
control.params = list(max.generations=100, pop.size=100,
                      wait.generations=8, solution.tolerance=1e-5, trace=F)

# fit linear age-year trend model
xMortE <- data.frame(age = mortDataE$age, year = mortDataE$year)
yMortE <- mortDataE$y
mortModelE_nug <- km(formula = ~x.age+x.year, #linear in age, linear in year
                     design = data.frame(x = xMortE), response = yMortE,
                     nugget.estim=TRUE,
                     covtype="matern5_2",
                     optim.method="gen",
                     control=control.params)

nug_E <- mortModelE_nug@covariance@nugget
mortModelE <- km(formula = ~x.age+x.year,
```



```

design = mortModelE_nug@X, response = mortModelE_nug@y,
noise.var = rep(nug_E,mortModelE_nug@n),
coef.trend = mortModelE_nug@trend.coef,
coef.cov = mortModelE_nug@covariance@range.val,
coef.var = mortModelE_nug@covariance@sd2,
covtype = mortModelE_nug@covariance@name)

```

We next generate 200 posterior samples (i.e., scenarios $\omega_1, \dots, \omega_n$) based on the GP posterior mortality curves in Age for the period Life Expectancy (LE) e_{65} at age 65 from the point of view of Year 2017 through 2026. After aggregating across samples, we visualize the resulting boxplots of simulated future e_{65} , computed as

$$e_{65}(t; \omega) = \sum_{x=66}^{103} \exp(f_*(x, t; \omega)) \cdot \left(\prod_{u=66}^{x-1} (1 - e^{f_*(u, t; \omega)}) \right) \cdot (x - 65).$$

We find that the GP model projects an average gain in LE of about 1 year between 2016 and 2026, although the range is from 0 (no gain) to 2+ years of gain.

```

nsim <- 200
maxAge <- 105
agesForecast <- seq(65,maxAge, 1)
lifeExp65 <- array(0, dim=c(10,nsim)) # period life expectancy for each scenario
for (j in 1:10) {
  yearsForecast <- 2016+j #:2026

  xPredE <- data.frame(age=agesForecast,year=yearsForecast)
  mortPredE <- predict(mortModelE, newdata=data.frame(x=xPredE),
                      cov.compute=TRUE,
                      se.compute=TRUE,type="UK")

  sim <- simulate(mortModelE, nsim=nsim, newdata = data.frame(x=xPredE), cond=TRUE,
                 nugget.sim=nug_E/100)

  # first two terms: dying in the first year, dying in exactly the second year
  lifeExp65[j,] = exp(sim[,1]) *1 + exp(sim[,2])*(1-exp(sim[,1])) *2
  for (ag in 3:(maxAge-65+1)) # add probability of dying at age 'ag'
    lifeExp65[j,] = lifeExp65[j,] + exp(sim[,ag])*apply( 1-exp(sim[,1:(ag-1)]),1,prod) * ag
}

# Display a boxplot for the distribution of future e_{65}
par(bty="n")
boxplot(t(lifeExp65), pch=18, main="", names=2017:2026, xlab="Year",
        ylab="Life Expectancy at 65", cex.axis=1.3, cex.lab=1.2, cex=1.2,col="lightblue3")

```

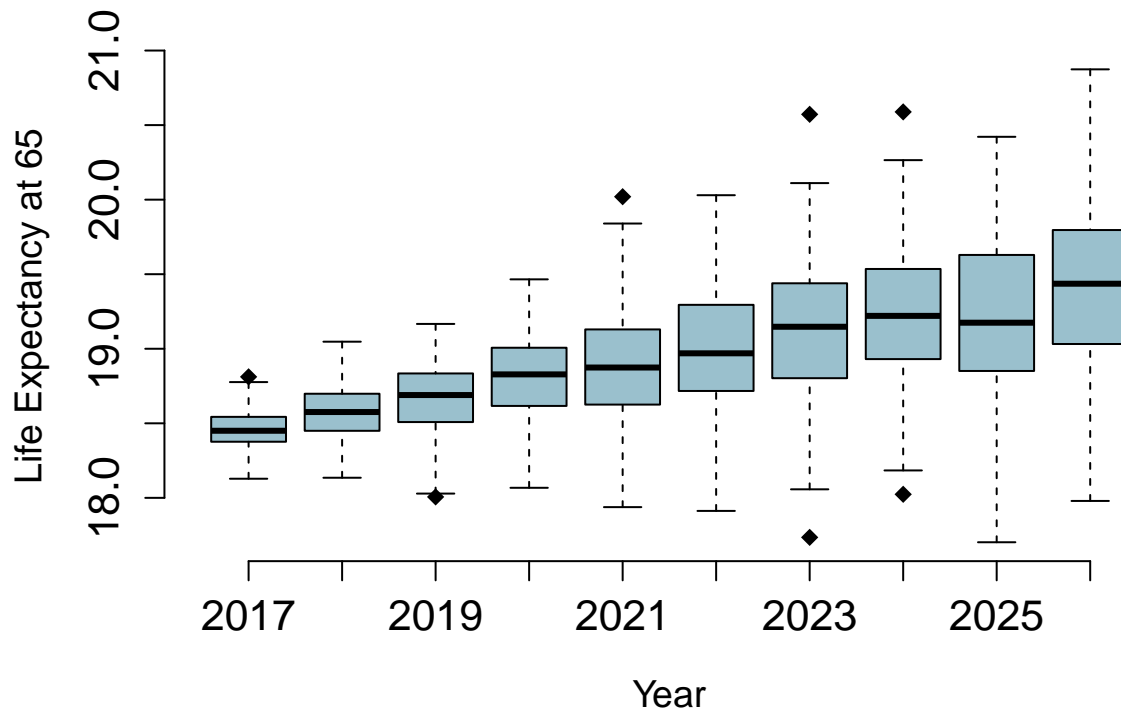


Figure 6.4 right: Comparison of Mean-function specifications

`km()` takes in an object of class **formula** to specify the prior mean function $\mu(\cdot)$, just like with linear OLS model `lm()` and generalized linear model `glm()` fitting. The following code produces models with constant, linear and quadratic prior means.

```
mortData3 <- dplyr::filter(mortData, age %in% 60:80, year %in% 1999:2016)
xMort3 <- data.frame(age = mortData3$age, year = mortData3$year)
yMort3 <- mortData3$y

# fit intercept only model
mortModel3int_nug <- km(formula = ~1,
  design = data.frame(x = xMort3), response = yMort3,
  nugget.estim=TRUE,
  covtype="matern5_2",
  optim.method="gen",
  control=control.params)

nug_int <- mortModel3int_nug@covariance@nugget
mortModelint <- km(formula = ~1,
  design = mortModel3int_nug@X, response = mortModel3int_nug@y,
  noise.var = rep(nug_int,mortModel3int_nug@n),
  coef.trend = mortModel3int_nug@trend.coef,
```

```

coef.cov = mortModel3int_nug@covariance@range.val,
coef.var = mortModel3int_nug@covariance@sd2,
covtype = mortModel3int_nug@covariance@name)

# fit linear trend model in both age and year
mortModel3lin_nug <- km(formula = ~x.age+x.year, # linear in age, linear in year
  design = data.frame(x = xMort3), response = yMort3,
  nugget.estim=TRUE,
  covtype="matern5_2",
  optim.method="gen",
  control=control.params)

nug_lin <- mortModel3lin_nug@covariance@nugget
mortModellin <- km(formula = ~x.age+x.year,
  design = mortModel3lin_nug@X, response = mortModel3lin_nug@y,
  noise.var = rep(nug_lin,mortModel3lin_nug@n),
  coef.trend = mortModel3lin_nug@trend.coef,
  coef.cov = mortModel3lin_nug@covariance@range.val,
  coef.var = mortModel3lin_nug@covariance@sd2,
  covtype = mortModel3lin_nug@covariance@name)

# fit quadratic-age trend model
mortModel3quad_nug <- km(formula = ~x.age+I(x.age^2)+x.year, # quadratic in age, linear in year
  design = data.frame(x = xMort3), response = yMort3,
  nugget.estim=TRUE,
  covtype="matern5_2",
  optim.method="gen",
  control=control.params)

nug_quad <- mortModel3quad_nug@covariance@nugget
mortModelquad <- km(formula = ~x.age+I(x.age^2)+x.year,
  design = mortModel3quad_nug@X, response = mortModel3quad_nug@y,
  noise.var = rep(nug_quad,mortModel3quad_nug@n),
  coef.trend = mortModel3quad_nug@trend.coef,
  coef.cov = mortModel3quad_nug@covariance@range.val,
  coef.var = mortModel3quad_nug@covariance@sd2,
  covtype = mortModel3quad_nug@covariance@name)

```

Using `show()` we can compare the fitted coefficients of the trends, as well as the GP hyperparameters, such as the lengthscales θ_{ag}, θ_{yr} .

```

show(mortModelint)

##
## Call:
## km(formula = ~1, design = mortModel3int_nug@X, response = mortModel3int_nug@y,
##     covtype = mortModel3int_nug@covariance@name, coef.trend = mortModel3int_nug@trend.coef,
##     coef.cov = mortModel3int_nug@covariance@range.val, coef.var = mortModel3int_nug@covariance@sd2,
##     noise.var = rep(nug_int, mortModel3int_nug@n))
##
## Trend  coeff.:
##
## (Intercept)    -3.5476
##
## Covar. type   : matern5_2

```

```

## Covar. coeff.:
##
##   theta(x.age)    32.6842
##   theta(x.year)   34.0000
##
## Variance: 1.261979
show(mortModellin)

##
## Call:
## km(formula = ~x.age + x.year, design = mortModel3lin_nug@X, response = mortModel3lin_nug@y,
##     covtype = mortModel3lin_nug@covariance@name, coef.trend = mortModel3lin_nug@trend.coef,
##     coef.cov = mortModel3lin_nug@covariance@range.val, coef.var = mortModel3lin_nug@covariance@sd2,
##     noise.var = rep(nug_lin, mortModel3lin_nug@n))
##
## Trend  coeff.:
##
##   (Intercept)    40.8071
##         x.age      0.0965
##         x.year    -0.0255
##
## Covar. type  : matern5_2
## Covar. coeff.:
##
##   theta(x.age)    5.1058
##   theta(x.year)    4.4593
##
## Variance: 0.001748755
show(mortModelquad)

##
## Call:
## km(formula = ~x.age + I(x.age^2) + x.year, design = mortModel3quad_nug@X,
##     response = mortModel3quad_nug@y, covtype = mortModel3quad_nug@covariance@name,
##     coef.trend = mortModel3quad_nug@trend.coef, coef.cov = mortModel3quad_nug@covariance@range.val,
##     coef.var = mortModel3quad_nug@covariance@sd2, noise.var = rep(nug_quad,
##         mortModel3quad_nug@n))
##
## Trend  coeff.:
##
##   (Intercept)    44.2132
##         x.age      0.0292
##   I(x.age^2)      0.0005
##         x.year    -0.0260
##
## Covar. type  : matern5_2
## Covar. coeff.:
##
##   theta(x.age)    4.9115
##   theta(x.year)    3.2769
##
## Variance: 0.001201927

```

We can also print out the respective log-likelihood scores that can be converted into BIC scores and ultimately

into Bayes Factors for comparing goodness-of-fit across models. As expected, the quadratic prior mean has the highest log-likelihood because it has the most degrees of freedom. However, after applying the BIC penalties, the linear-mean model ends up with the highest BIC.

```
c(mortModel3int_nug@logLik,mortModel3lin_nug@logLik,mortModel3quad_nug@logLik)
```

```
## [1] 616.2822 633.9895 635.7808
```

Next, we use these three GP models to forecast temporal mortality trends at ages 60, 70, 80, to see how well they can extrapolate. In the plots below we compare their Age-specific forecasts for years 2005–2030, where 2017 and onward is an extrapolation compared to the training data.

Note: we plot all the GP predictions on the same plot since we have the natural ordering that (log-)mortality is much lower at Age 60 compared to Age 70 compared to Age 80.

```
# Setup the test data.frame to be displayed
agesForecast <- c(60,70,80)
yearsForecast <- 2005:2030
nYr <- length(yearsForecast)
nAg <- length(agesForecast)
xPred <- data.frame(age=rep(agesForecast,each=nYr), year=rep(yearsForecast,nAg))

# Setup the raw training data that was used
agesObserved <- 60:80
yearsObserved <- 1999:2016
rateObserved <- dplyr::filter(mortData,age %in% agesObserved, year %in% yearsForecast)$rate

# predict using the 3 GP models above on the same test set
mortPredInt <- predict(mortModelint, newdata=data.frame(x=xPred), cov.compute=TRUE,
  se.compute=TRUE,type="UK")
mortPredLin <- predict(mortModellin, newdata=data.frame(x=xPred), cov.compute=TRUE,
  se.compute=TRUE,type="UK")
mortPredQuad <- predict(mortModelquad, newdata=data.frame(x=xPred), cov.compute=TRUE,
  se.compute=TRUE,type="UK")

xPred$mInt <- (mortPredInt$mean)
xPred$mLin <- (mortPredLin$mean)
xPred$mQuad <- (mortPredQuad$mean)

xPredag <- dplyr::filter(xPred,age==80)
rateObserved <- log(dplyr::filter(mortData, year %in% yearsObserved, age == 80)$rate)
# set-up the overall plot layout
plot(yearsObserved, rateObserved,
  bty='n', main="", pch=19, cex.axis=1.2, cex.lab=1.2, cex=1.3,
  xlab="Year", ylab="Log Mortality Rate", ylim=c(-5, -2), xlim=range(yearsForecast))

for(ag in agesForecast){
  rateObserved <- log(dplyr::filter(mortData, year %in% yearsObserved, age == ag)$rate)
  # add points representing training data of observed log-mortality
  points(yearsObserved,rateObserved,col="darkblue",
    pch=19,cex=1.3)
  text(2008, rateObserved[2]+0.1*(ag != 70), paste("Age ",ag), cex=1.15)

  # add lines representing GP posterior mean
```

```

xPredag <- dplyr::filter(xPred,age==ag)
lines(yearsForecast, xPredag$mInt, col=2, lty=3, lwd=3)
lines(yearsForecast, xPredag$mLin, col=3, lty=2, lwd=3)
lines(yearsForecast, xPredag$mQuad, col=4, lty=1, lwd=3)
}
legend("topright",legend=c("Intercept Only","Linear Trend","Quadratic Trend"),
      lwd=rep(3,3), col=2:4, lty=3:1, cex=1.2, bty="n")

```

