TMA4300: Exercise 1

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Problem A

A.1

The exponential distribution has cumulative density function (CDF)

$$F(x) = 1 - e^{-\lambda x},$$

with rate parameter λ . By defining u := F(x), we can express x as

$$x = -\frac{1}{\lambda} \ln(1 - u) =: F^{-1}(u).$$

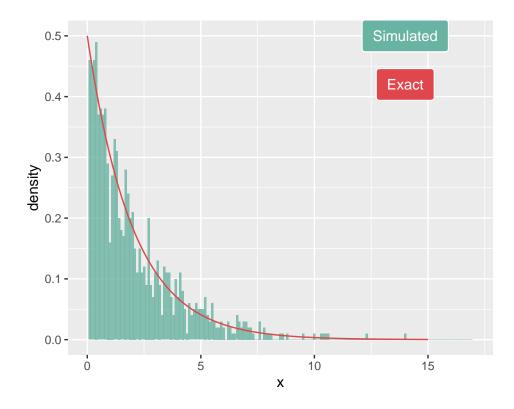
This means that we can use the *inversion method* to simulate from the exponential distribution. I.e., we let $U \sim \mathcal{U}_{[0,1]}$ and calculate $X = F^{-1}(U)$. Then, $X \sim \text{Exp}(\lambda)$. The function which simulates the exponential distribution is given below.

```
sim.exp <- function(rate, n){
  u <- runif(n,0,1)
  return(-1/rate * log(1 - u))
}</pre>
```

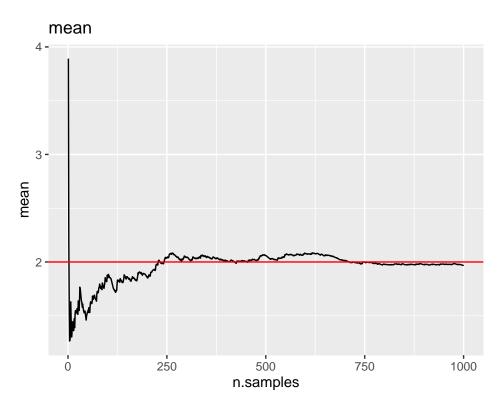
Next, we need to check if this gives reasonable results by comparing our simulated values to the theoretical knowledge.

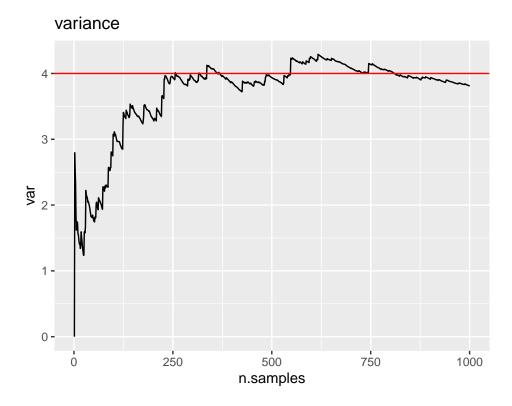
```
label.size = 0.35,
  color = "white",
  fill= "#69b3a2"
) +
geom_label(
  label="Exact",
    x=14,
    y=0.42,
  label.padding = unit(0.55, "lines"), # Rectangle size around label
  label.size = 0.35,
  color = "white",
  fill = "#e0474c"
) +
xlim(0,17)
```

Warning: Removed 2 rows containing missing values (geom_bar).



We also compare the estimated mean and variance to first and second central moments of the exponential distribution





We observe that the sampled mean and variance approach the theoretical values as the number of samples grows larger.

A.3

a)

To find the value of c, we integrate the density over the entire domain and equate the result to 1:

$$1 = \int_{-\infty}^{\infty} f(x)dx = c \int_{-\infty}^{\infty} \frac{e^{\alpha x}}{(1 + e^{\alpha x})^2} dx.$$

To progress from here, we introduce the substitution, $v = e^{\alpha x}$, which gives

$$1 = \frac{c}{\alpha} \int_0^\infty \frac{dv}{(1+v)^2} = \frac{c}{\alpha} \left(-\frac{1}{1+v}\right) \Big|_0^\infty = \frac{c}{\alpha}.$$

Consequently, $c = \alpha$.

b)

The CDF is defined as follows,

$$F(x) = \int_{-\infty}^{x} f(z)dz = \int_{0}^{\exp(\alpha x)} \frac{dv}{(1+v)^2}$$
$$= 1 - \frac{1}{1 + \exp(\alpha x)} = \frac{\exp(\alpha x)}{1 + \exp(\alpha x)},$$

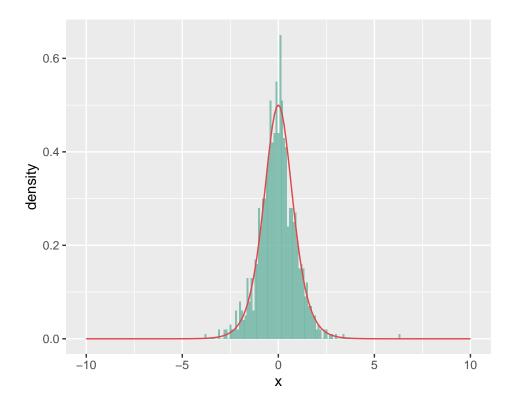
where we have used the same substitution as earlier, namely $v = e^{\alpha z}$. We notice that F(x) is the Sigmoid function, which has the well known logit-function as its inverse. I.e.,

$$F^{-1}(x) = \frac{1}{\alpha} \ln \left(\frac{x}{1-x} \right).$$

c)

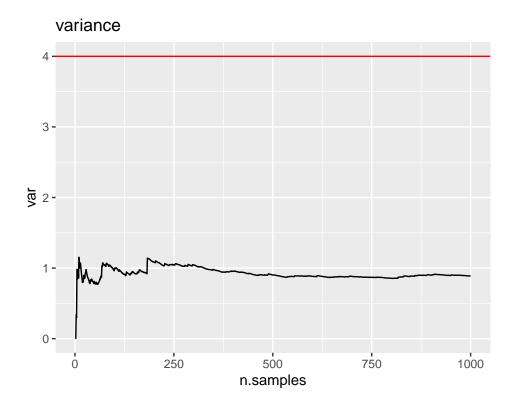
Since we have an analytic expression for the inverse, we can again use the *inversion method* to sample from f. The sampling-function is given below.

```
sim.sigm <- function(alpha, n){
  u <- runif(n,0,1)
  return(1/alpha * log(u/(1-u)))
}</pre>
```



To compare the simulated values with the theoretical mean and variance, we first need to compute the expression of these moments. Think maybe the mean is undefined??

mean 2.0 1.5 0.5 0.0 250 0.5 000 0.5 1000 0.5 1000

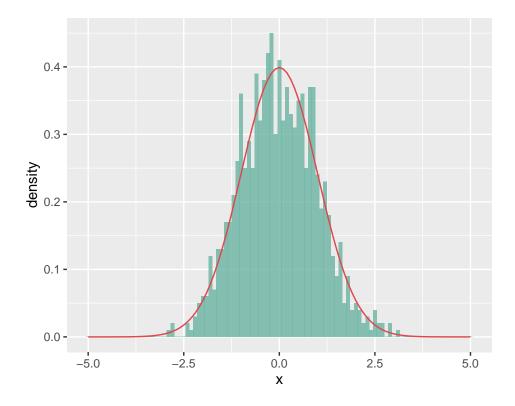


A.4

Below is our implementation of the Box-Muller algorithm.

```
box.mul <- function(n){</pre>
  odd <- FALSE
  if(n %% 2 != 0){
    odd <- TRUE
    n \leftarrow n + 1
  }
  x1 <- runif(n/2, 0, 2*pi)
  x2 <- sim.exp(0.5, n/2)
  y1 <- sqrt(x2)*cos(x1)</pre>
  y2 <- sqrt(x2)*sin(x1)</pre>
  concat <- c(y1,y2)</pre>
  if(odd){
    return(head(concat, -1))
  }
  else{
    return(concat)
  }
```

As usual, we compare the results of a simulation against the theoretical distribution.



Finally, we compare the mean and variance resulting from the simulation with the theoretical values.

