

Let It Be, Let It Move

Regression Discontinuity Insights on Demand-Responsive Transit and Property Value Changes

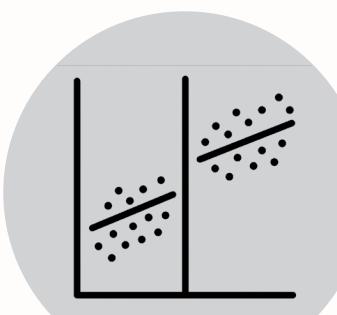
KEYWORDS



DRT



Property



RDD

Hani Jung, Jiyo Shin
Advisor: Professor Up Lim

Department of Urban Planning and Engineering
Yonsei University

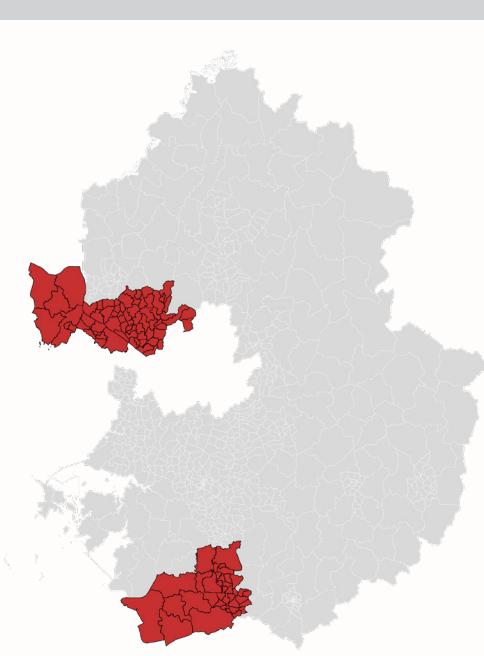
Introduction

Problem Statement

In areas with insufficient transportation infrastructure, limited mobility constrains access to opportunities, thereby influencing local property values. While Demand-Responsive Transit (DRT) has recently been introduced to improve accessibility in such areas, its actual economic impact has yet to be fully understood.

Scope

- The analysis focuses on Goyang, Gimpo, and Pyeongtaek in Gyeonggi Province, analyzing transportation accessibility and property transaction data at the district level.
- This study utilizes the monthly average price per pyeong across various housing types, including apartments, detached housing, row houses, multi-household housing, multi-family housing, and officetels.



Research Objectives

- This study identifies transportation-disadvantaged areas and examines how DRT introduction influences property values.
- In particular, this research places emphasis on clarifying the clear causal link between accessibility improvements and property value shifts through rigorous statistical inference.

Sources

- Administrative Boundaries: District-level shapefiles of Gyeonggi Province provided by GIS Developer
- Bus Stop Locations: Nationwide Bus Stop Location Information provided by the Public Data Portal
- Subway Station Locations: Integrated dataset compiled from multiple subway lines across Gyeonggi Province, including metropolitan, regional, and airport lines, provided by the Public Data Portal
- Demand-Responsive Transit (DRT) Service Areas: Operational areas of the Ddkbus (돌나스) service provided by Gyeonggi Transportation Research Institute
- Property Transaction Data: Housing transaction data provided by the Real Transaction Price Disclosure System

Literature Review

- Yoo and Yeon (2021) calculated an Accessibility Index by analyzing transit variables (e.g. route counts) within grid cells. This index was utilized to identify transportation-disadvantaged areas. Adopting this methodology, this study defined transportation-disadvantaged areas by calculating an Accessibility Index at the administrative dong level, based on accessible subway stations and bus stops.
- Byun and Ko (2019) analyzed the determinants of apartment sales and Jeonse prices in non-capital metropolitan areas, revealing that public transportation accessibility significantly influences price determination. Extending beyond this focus on apartments, this research expands the analytical scope to include a diverse range of housing types to comprehensively examine the market-wide impact of accessibility.
- Kang et al. (2024) applied a Regression Discontinuity Design to examine how the COVID-19 pandemic, real estate regulations, and monetary policies affected apartment prices in Korea. Following this methodological approach, this study uses real transaction data and employs RDD to identify the causal impact of DRT implementation on property values.

Identification of Transportation Disadvantaged Areas

Definition

Transportation-disadvantaged areas were defined as districts with low accessibility to public transit.

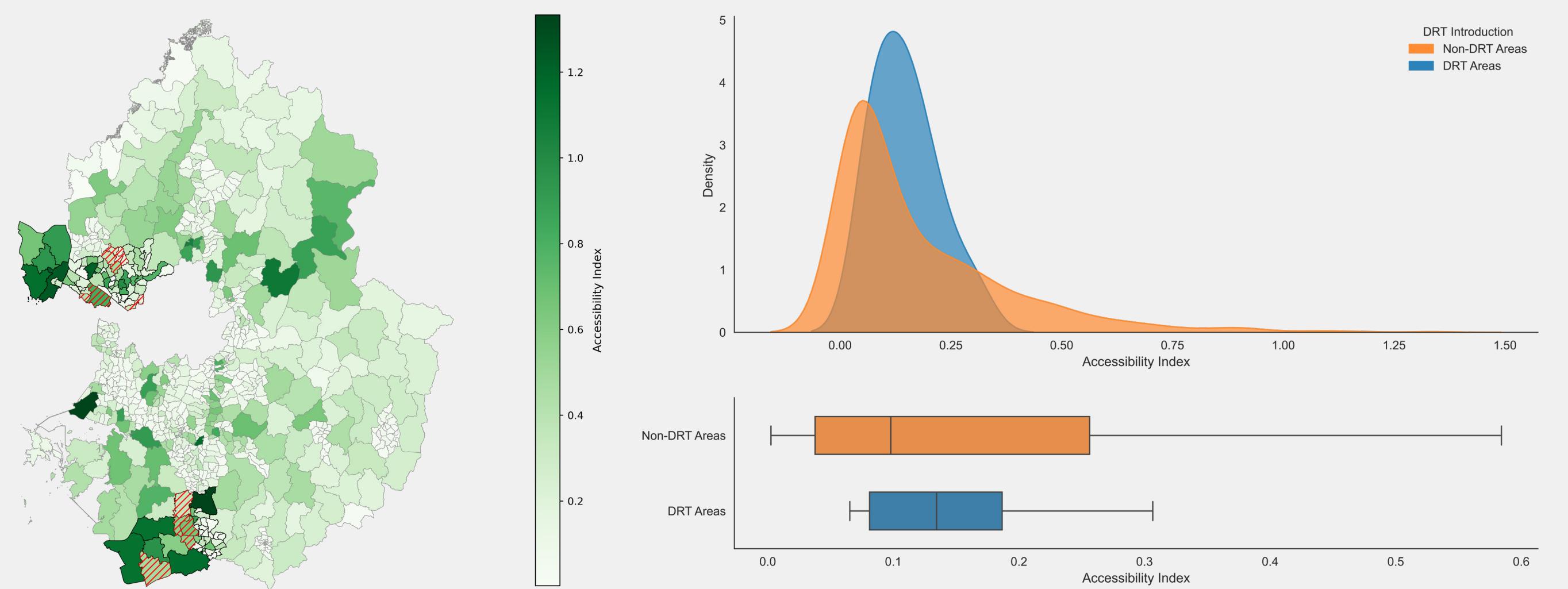
- Accessibility indicates how easily residents can reach nearby bus and subway services. By measuring the spatial distribution of transit facilities, it provides an objective way to identify areas with limited mobility.

$$\text{Accessibility Index}_d = \frac{\text{BusStops}_d}{\max(\text{BusStops})} + \frac{\text{Stations}_d}{\max(\text{Stations})}$$

- The accessibility index was constructed using two factors: number of bus stops, number of subway stations
- Equal importance was assigned to both factors, and values were normalized within each comparison group to ensure interpretability.

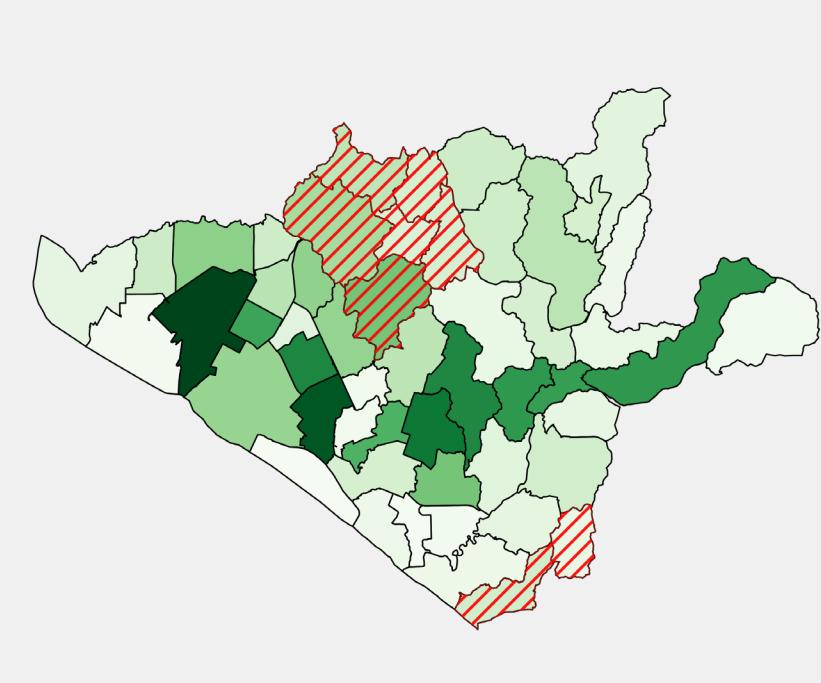
Spatial Distribution of Accessibility

The spatial distribution of accessibility across Gyeonggi Province revealed pronounced geographic disparities. Accessibility was generally high near major transit corridors, while many northwestern and southern peripheral districts exhibited particularly limited access to bus and subway infrastructure. These differences highlight structural imbalances in transit provision across the region.

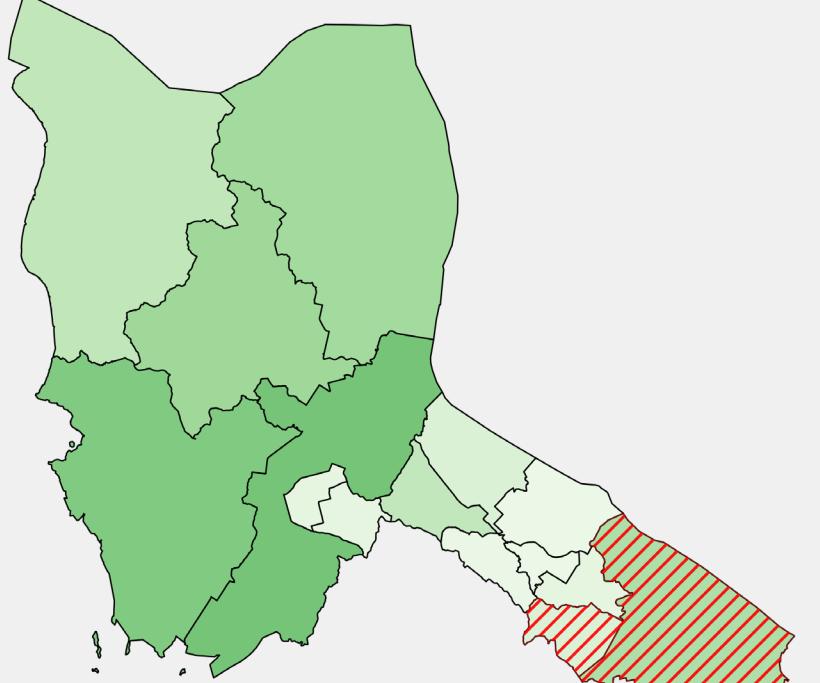


To provide a fair basis for comparison in the subsequent analysis, accessibility indices were recalculated at the city level. Because each city differs in its overall infrastructure density and urban scale, direct comparison of districts across cities could be influenced more by structural characteristics than by the effect of DRT itself. City-level normalization ensured that each district was evaluated relative to others within the same urban context.

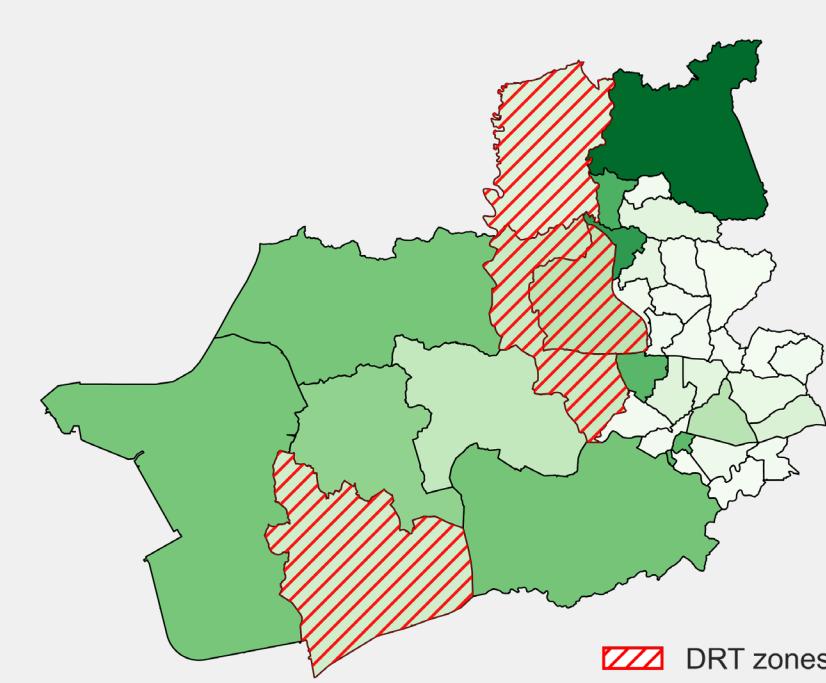
Goyang



Gimpo



Pyeongtaek



The resulting city-specific indices allowed the study to match districts with similar levels of transportation disadvantage. This step minimized confounding differences in baseline accessibility and created a more comparable analytical frame, enabling a clearer examination of how the introduction of DRT influenced subsequent changes in property values.

Research Design and Method

Quasi-Experimental Design

This study employs a quasi-experimental approach to estimate the causal impact of DRT introduction on property values where randomized assignment is not feasible.

- To ensure valid comparisons, administrative districts were matched using the city-level accessibility index, which groups areas with similar baseline levels of transportation disadvantage.

- Under this framework, districts that adopted DRT were classified as the treatment group, while the most comparable non-adopting districts served as the control group.

City	Treatment Group (Implementation Date, Index)	Control Group (Non-Implementation, Index)
Gimpo	Pungmu-dong (Aug 2023, 0.341)	Unyang-dong (0.335)
	Gochon-eup (Jul 2023, 0.654)	Haseong-myeon (0.749)
Pyeongtaek	Godeok-dong (May 2023, 0.434)	Bijeon-dong (0.461)
	Godeok-myeon (May 2023, 0.370)	Yongi-dong (0.274)
Goyang	Deokseon-dong (Nov 2024, 0.413)	Hyangdong-dong (Oct 2024, 0.198)
	Wondang-dong (0.190)	Siksa-dong (Jun 2023, 1.000)
	Pung-dong (1.000)	Sarihyeon-dong (Jun 2023, 0.387)
	Goyang-dong (0.388)	Jiyeong-dong (Jun 2023, 0.373)
	Sinwon-dong (0.371)	Seolmun-dong (Jun 2023, 0.560)
	Daeja-dong (0.570)	Munbong-dong (Jun 2023, 0.347)
	Jeongbalsan-dong (0.293)	Seongseok-dong (Jun 2023, 0.693)
	Jugyo-dong (0.595)	Jugyo-dong (0.595)

Conclusions

Direct Value Appreciation

In Godeok-dong in Pyeongtaek, where market conditions were generally weakening, the introduction of DRT corresponded with a clear shift toward price improvement. While the matched control group, Bijeon-dong, recorded declines in both price and growth trend, the treatment group showed the opposite pattern. After implementation, Godeok-dong registered an increase of 1.59 million KRW in average price and a growth trend of 3.51, reflecting a strong upward response. This contrast suggests that improved mobility from DRT can stimulate housing demand even in soft market conditions.

Market Amplification Effect

In Sarihyeon-dong in Goyang, where the overall housing market was on an upward trend, the treatment area displayed a notably stronger increase than its matched control group. The average transaction price rose by 1.80 million KRW in the DRT-introduced district, which is approximately 1.5 times the increase observed in Goyang-dong, the control group. This difference, despite both districts being exposed to the same broader market conditions, suggests that the added accessibility provided by DRT strengthened local housing demand.

Mitigation of Decline

In Seongseok-dong in Goyang, which experienced a general decline in market prices, the introduction of DRT was associated with a noticeably smaller drop in property values. The treatment group saw a decrease of 1.91 million KRW, whereas the matched control group, Jugyo-dong, recorded a much steeper decline of 4.73 million KRW. This contrast indicates that enhanced mobility provided by DRT can moderate the extent of price depreciation during downturns.

Regression Discontinuity Design

What is RDD?

- Cutoff-based identification: Treatment status changes at a predefined threshold.
- Local comparability: Units near the cutoff share similar characteristics.
- Causal inference: Discontinuities in either the level or slope of the outcome at the threshold indicate the policy's effect.

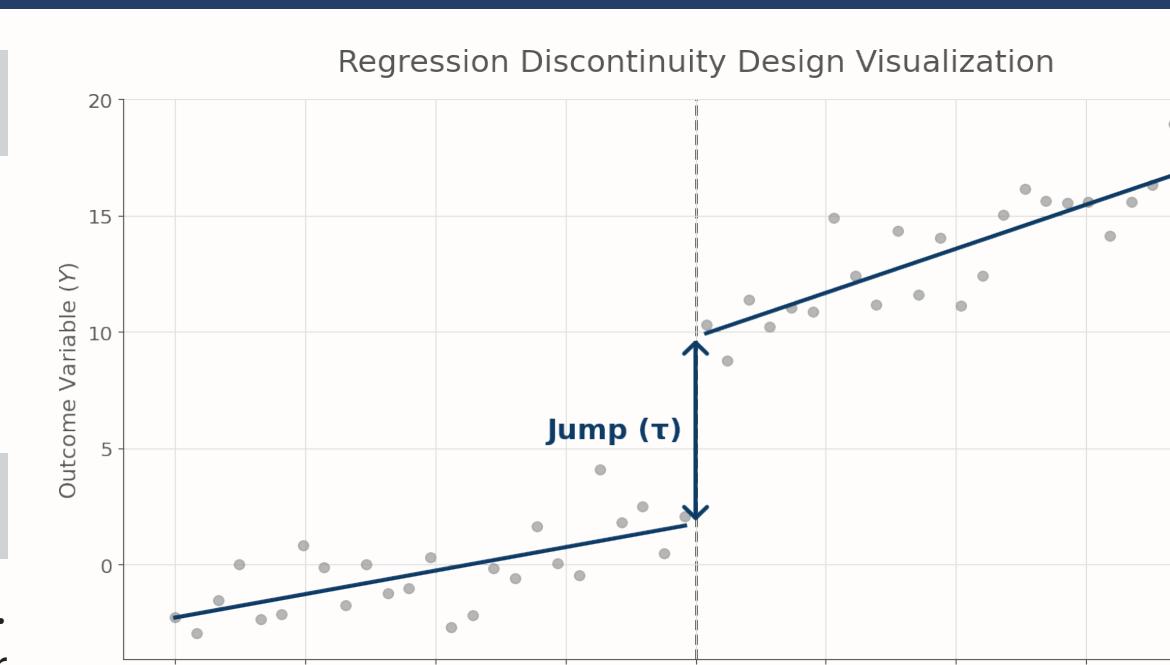
Weighted Least Squares Method

WLS was applied to obtain stable estimates of monthly average transaction prices. Because transaction volumes vary substantially across months, periods with fewer sales are more prone to random fluctuations. By assigning greater weights to months with higher transaction counts, WLS prevents low-volume periods from exerting outsized influence on the results. This approach reduces heteroscedasticity and enhances the overall robustness of the estimation.

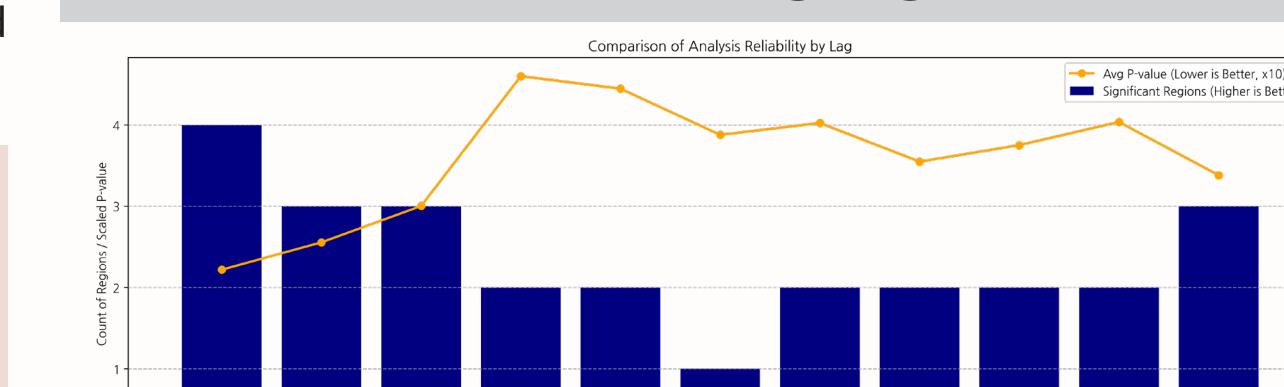
The Basic Estimation Model

- $$\text{Price} = \beta_0 + \beta_1 \cdot \text{timeVar} + \beta_2 \cdot \text{postTreatment} + \beta_3 \cdot (\text{timeVar} \times \text{postTreatment}) + \epsilon$$
- β_0 : property price at the time of cutoff
 - β_1 : price trend before the cutoff (pre-treatment trend)
 - β_2 : immediate change in price at the time of cutoff (jump effect)
 - β_3 : change in price trend after the introduction (post-treatment slope change)

In other words, a significant β_2 indicates a sharp price shift at the time of cutoff, whereas a significant β_3 suggests a change in the rate of price increase or decrease following the introduction.

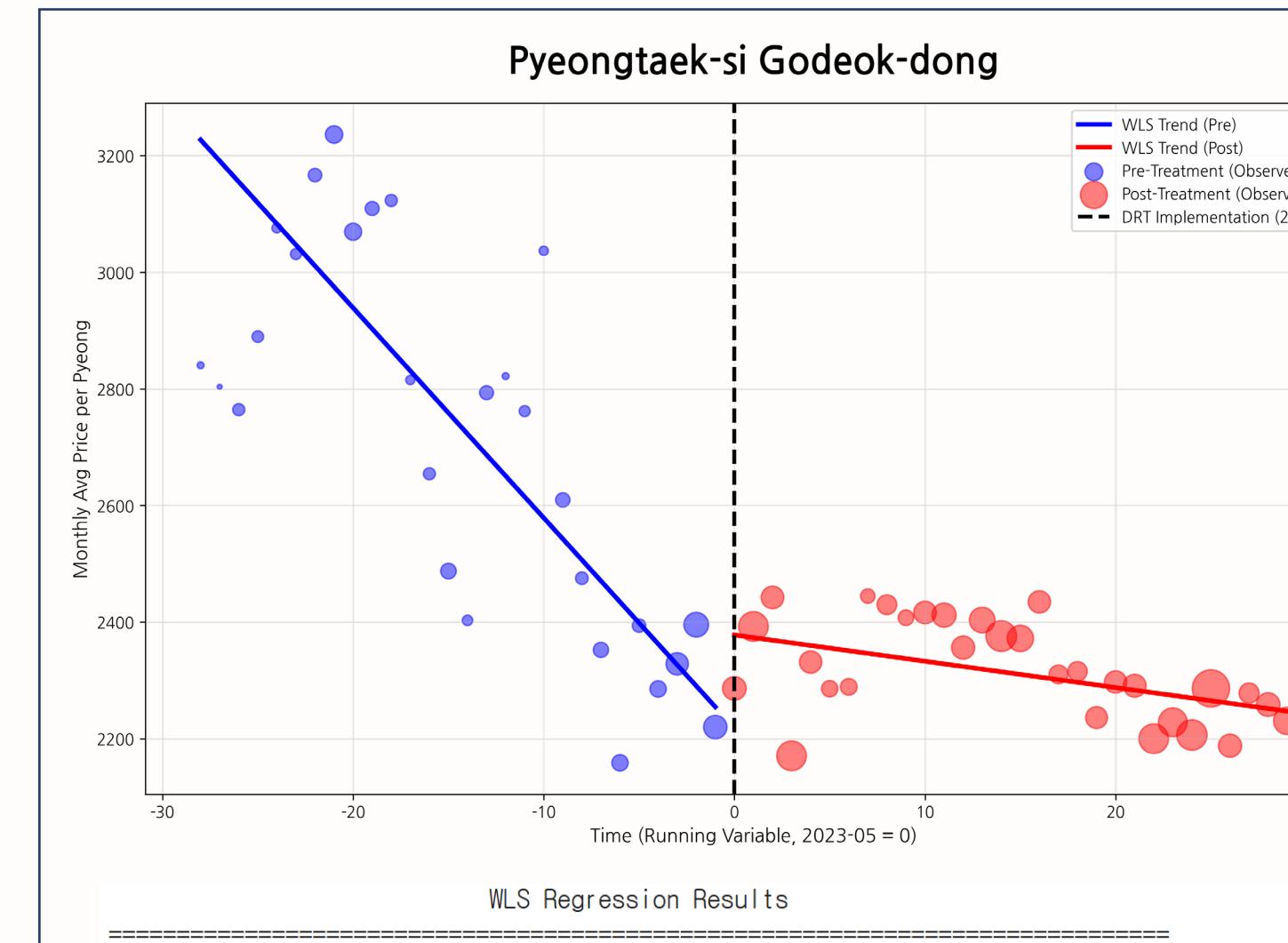


Considering Lag

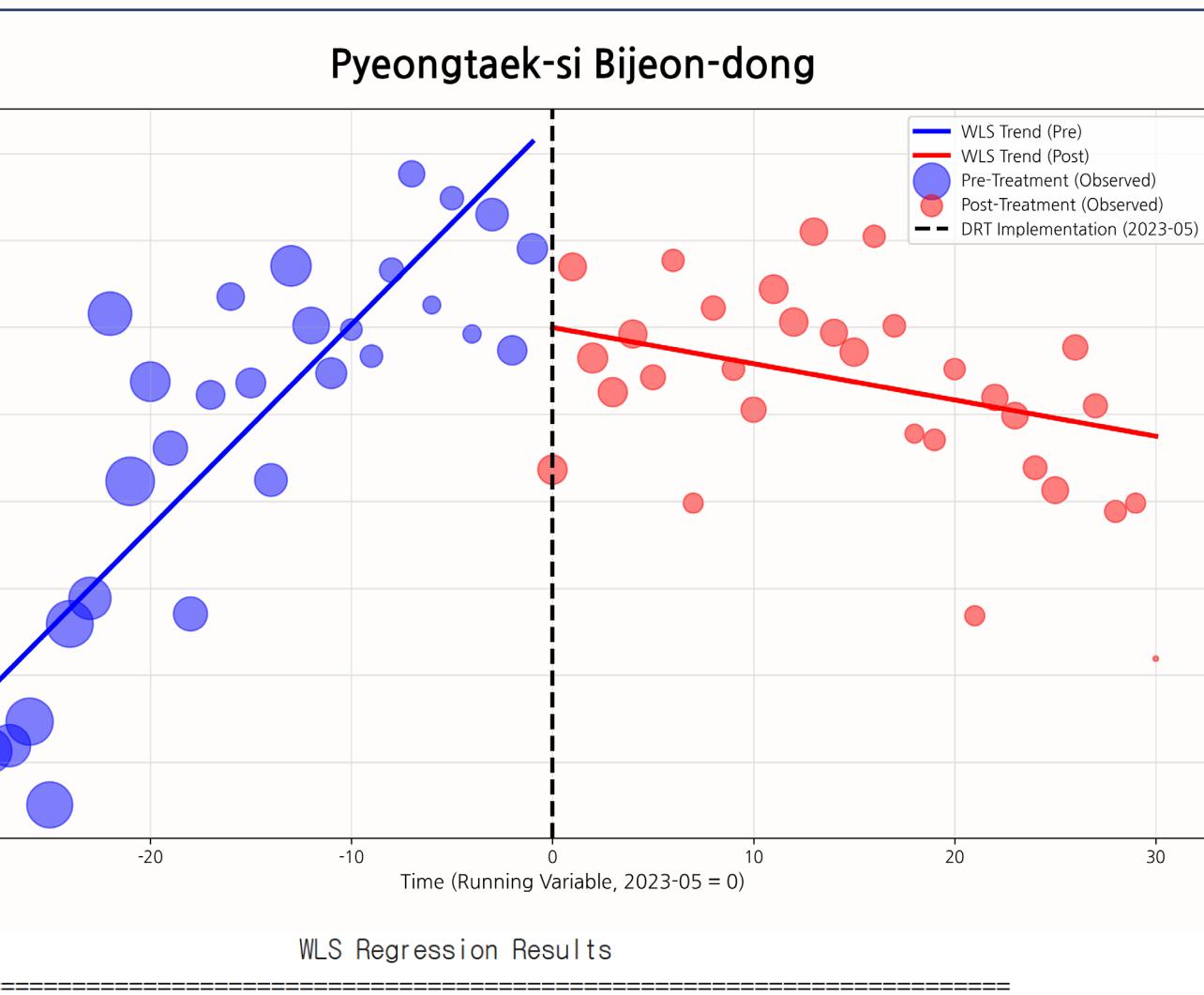


Because the effects of DRT may not appear immediately, property value changes were tested across 0 to 10 months after introduction. Month 0 was selected as the cutoff because it showed the lowest average p-values and the largest number of districts with significant price changes, indicating the clearest treatment response.

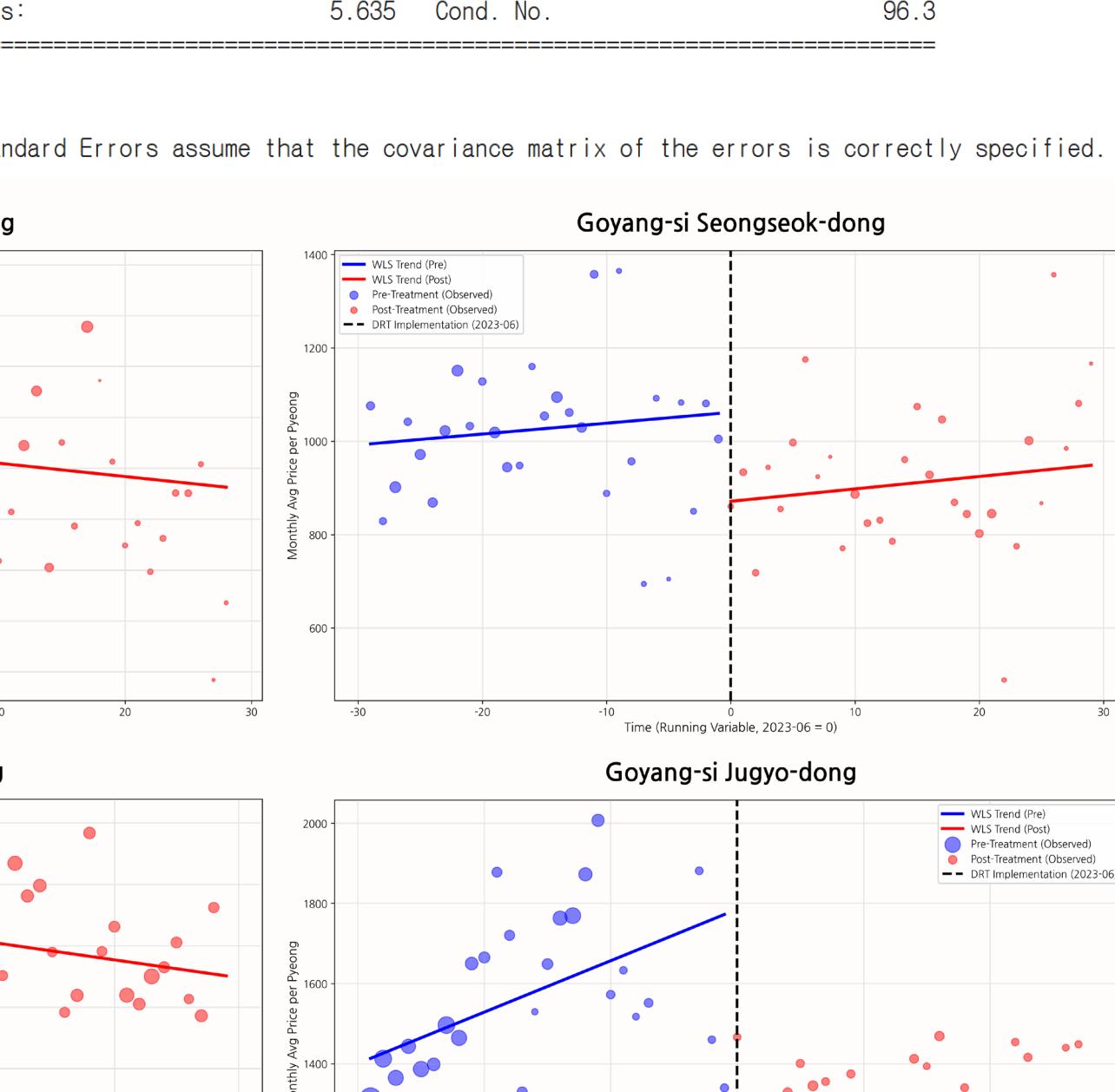
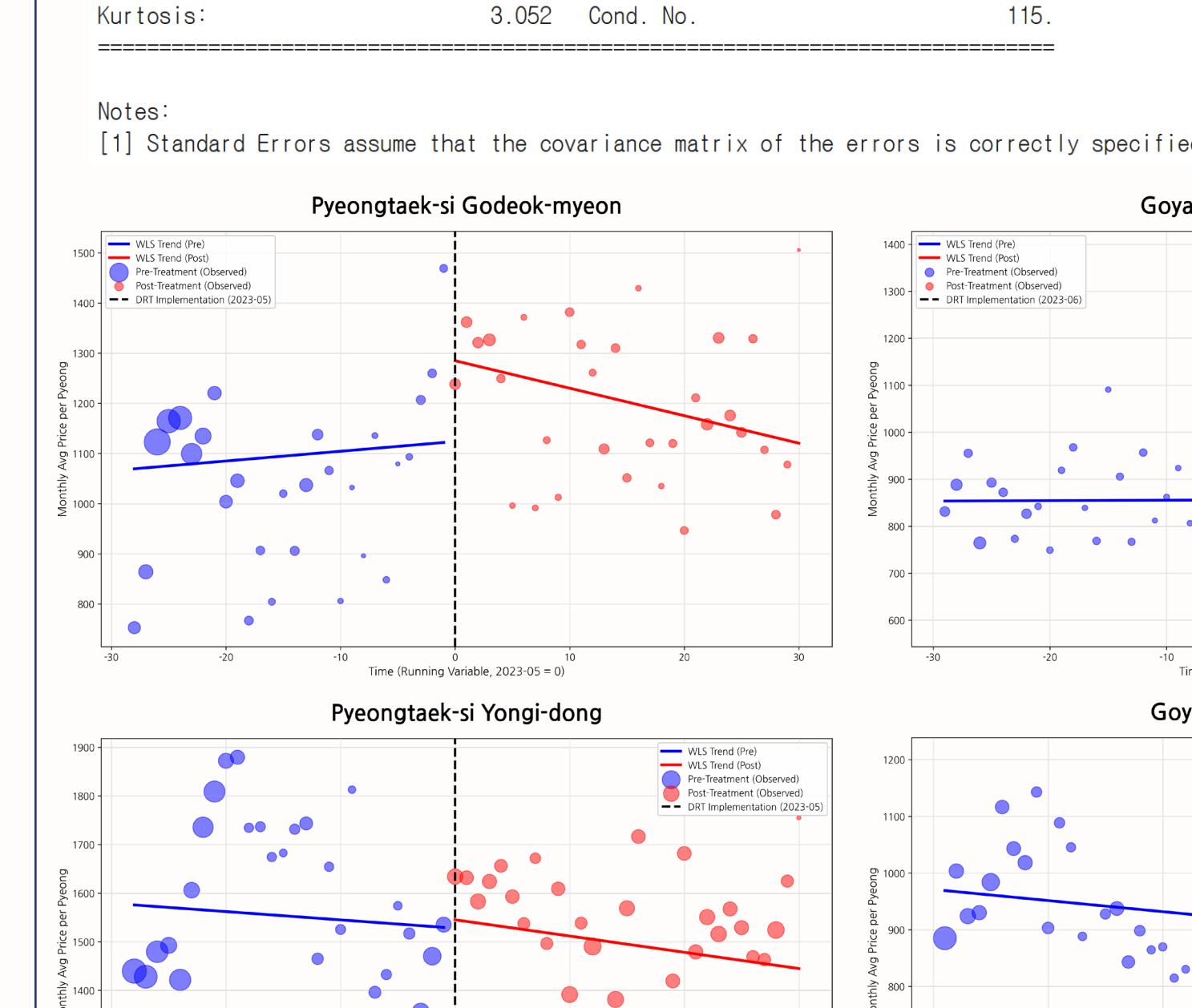
Empirical Results



WLS Regression Results
Dep. Variable: price_per_pyeong R-squared: 0.776
Model: OLS Date: Wed, 19 Nov 2025 Prob (F-statistic): 6.88e-18
Time: 06:08:12 Log-Likelihood: -268.73
No. Observations: 59 AIC: 545.5
Df Residuals: 55 BIC: 753.8
Df Model: 3 Covariance Type: nonrobust
coef std err t P> t [0.025] [0.975]
Intercept 2216.4251 49.038 45.239 0.000 2120.150 2316.700
time_var -36.0182 3.559 -10.120 0.000 -43.151 -28.886
post_treatment 159.3814 59.463 2.680 0.010 40.215 278.547
time_var*post_treatment 31.5101 4.019 7.840 0.000 23.456 39.564
Omnibus: 2.079 Durbin-Watson: 1.075
Prob(Omnibus): 0.354 Jarque-Bera (JB): 1.611
Skew: -0.404 Prob(JB): 0.447
Kurtosis: 3.052 Cond. No.: 15.



WLS Regression Results
Dep. Variable: price_per_pyeong R-squared: 0.667
Model: OLS Date: Wed, 19 Nov 2025 Prob (F-statistic): 3.74e-13
Time: 06:08:12 Log-Likelihood: -327.19
No. Observations: 59 AIC: 662.4
Df Residuals: 55 BIC: 670.7
Df Model: 3 Covariance Type: nonrobust
coef std err t P> t [0.025] [0.975]
Intercept 1268.2109 23.121 53.599 0.000 1220.793 1315.629
time_var 11.6672 3.924 9.624 0.000 10.750 14.097
post_treatment -18.2977 33.300 -0.552 0.001 -185.033 -51.563
time_var*post_treatment -13.7499 1.696 -7.253 0.000 -17.549 -9.951
Omnibus: 12.101 Durbin-Watson: 1.447
Prob(Omnibus): 0.002 Jarque-Bera (JB): 20.750
Skew: 0.611 Prob(JB): 3.12e-05
Kurtosis: 5.635 Cond. No.: 98.3



Notes:
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

WLS Regression Results
Dep. Variable: price_per_pyeong R-squared: 0.667
Model: OLS Date: Wed, 19 Nov 2025 Prob (F-statistic): 3.74e-13
Time: 06:08:12 Log-Likelihood: -327.19
No. Observations