# Package 'gap'

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Type Package

Title Genetic Analysis Package

**Version** 1.5-2 **Date** 2023-2-18

**Description** 

As first reported [Zhao, J. H. 2007. ``gap: Genetic Analysis Package". J Stat Soft 23(8):1-18. <doi:10.18637/jss.v023.i08>], it is designed as an integrated package for genetic data analysis of both population and family data. Currently, it contains functions for sample size calculations of both population-based and family-based designs, probability of familial disease aggregation, kinship calculation, statistics in linkage analysis, and association analysis involving genetic markers including haplotype analysis with or without environmental covariates. Over years, the package has been developed in-between many projects hence also in line with the name (gap).

**License** GPL (>= 2)

URL https://github.com/jinghuazhao/R

BugReports https://github.com/jinghuazhao/R/issues

**Depends** R (>= 2.10), gap.datasets

Imports dplyr, ggplot2, plotly, Rdpack

Suggests BradleyTerry2, DiagrammeR, DOT, MASS, Matrix, MCMCglmm,

R2jags, bdsmatrix, bookdown, calibrate, circlize, coda, cowplot, coxme, foreign, forestplot, genetics, grid, haplo.stats, htmlwidgets, jsonlite, kinship2, knitr, lattice, magic, manhattanly, matrixStats, meta, metafor, nlme, pedigree, pedigreemm, plotrix, readr, reshape, rmarkdown, rmeta, rms, survival

Enhances shiny

LazyData Yes

LazyLoad Yes

**NeedsCompilation** yes

**Encoding** UTF-8

VignetteBuilder knitr

**Roxygen** list(markdown = TRUE)

RoxygenNote 7.2.3

RdMacros Rdpack

2 R topics documented:

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a2g	Allele-to-genotype conversion	

# Description

Allele-to-genotype conversion

# Usage

a2g(a1, a2)

# Arguments

a1 first allele.

a2 second allele.

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ab

Test/Power calculation for mediating effect

# Description

Test/Power calculation for mediating effect

# Usage

```
ab(

type = "power",

n = 25000,

a = 0.15,

sa = 0.01,

b = log(1.19),

sb = 0.01,

alpha = 0.05,

fold = 1
```

# **Arguments**

type	string option: "test", "power".
n	default sample size to be used for power calculation.
a	regression coefficient from indepdendent variable to mediator.
sa	SE(a).
b	regression coefficient from mediator variable to outcome.
sb	SE(b).
alpha	size of siginficance test for power calculation.
fold	fold change for power calculation, as appropriate for a range of sample sizes.

#### **Details**

This function tests for or obtains power of mediating effect based on estimates of two regression coefficients and their standard errors. Note that for binary outcome or mediator, one should use log-odds ratio and its standard error.

# Value

The returned value are z-test and significance level for significant testing or sample size/power for a given fold change of the default sample size.

# Author(s)

Jing Hua Zhao

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#### References

Freathy RM, Timpson NJ, Lawlor DA, Pouta A, Ben-Shlomo Y, Ruokonen A, Ebrahim S, Shields B, Zeggini E, Weedon MN, Lindgren CM, Lango H, Melzer D, Ferrucci L, Paolisso G, Neville MJ, Karpe F, Palmer CN, Morris AD, Elliott P, Jarvelin MR, Smith GD, McCarthy MI, Hattersley AT, Frayling TM (2008). "Common variation in the FTO gene alters diabetes-related metabolic traits to the extent expected given its effect on BMI." *Diabetes*, **57**(5), 1419-26. ISSN 0012-1797 (Print) 0012-1797, doi:10.2337/db071466.

Kline RB. Principles and practice of structural equation modeling, Second Edition. The Guilford Press 2005.

MacKinnon DP. Introduction to Statistical Mediation Analysis. Taylor & Francis Group 2008.

Preacher KJ, Leonardelli GJ. Calculation for the Sobel Test-An interactive calculation tool for mediation tests https://quantpsy.org/sobel/sobel.htm

#### See Also

ccsize

# **Examples**

```
## Not run:
ab()
n <- power <- vector()
for (j in 1:10)
{
    z <- ab(fold=j*0.01)
    n[j] <- z[1]
    power[j] <- z[2]
}
plot(n,power,xlab="Sample size",ylab="Power")
title("SNP-BMI-T2D association in EPIC-Norfolk study")
## End(Not run)</pre>
```

AE3

AE model using nuclear family trios

### **Description**

AE model using nuclear family trios

# Usage

```
AE3(model, random, data, seed = 1234, n.sim = 50000, verbose = TRUE)
```

# **Arguments**

model a linear mixed model formula, see example below.

random random effect, see exampe below.

data data to be analyzed.

allele.recode 7

random number seed.n.simnumber of simulations.verbosea flag for printing out results.

#### **Details**

This function is adapted from example 7.1 of Rabe-Hesketh et al. (2008). It also procides heritability estimate and confidence intervals.

# Value

The returned value is a list containing:

- lme.result the linear mixed model result.
- h2 the heritability estimate.
- CL confidence intervals.

#### Note

Adapted from f.mbf.R from the paper.

# Author(s)

Jing Hua Zhao

#### References

Rabe-Hesketh S, Skrondal A, Gjessing HK (2008). "Biometrical modeling of twin and family data using standard mixed model software." *Biometrics*, **64**(1), 280-8. ISSN 0006-341X (Print) 0006-341x, doi:10.1111/j.15410420.2007.00803.x.

# **Examples**

allele.recode

Allele recoding

# Description

Allele recoding

### Usage

```
allele.recode(a1, a2, miss.val = NA)
```

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### **Arguments**

a1	first allele.
a2	second allele.
miss.val	missing value.

asplot

Regional association plot

# **Description**

Regional association plot

### Usage

```
asplot(
  locus,
  map,
  genes,
  flanking = 1000,
  best.pval = NULL,
  sf = c(4, 4),
  logpmax = 10,
  pch = 21
)
```

# Arguments

locus	Data frame with columns c("CHR", "POS", "NAME", "PVAL", "RSQR") con-
	taining association results.

map Genetic map, i.e, c("POS","THETA","DIST").

genes Gene annotation with columns c("START", "STOP", "STRAND", "GENE").

flanking Flanking length.

best.pval Best p value for the locus of interest.

sf scale factors for p values and recombination rates, smaller values are necessary

for gene dense regions.

logpmax Maximum value for -log10(p).

pch Plotting character for the SNPs to be highlighted, e.g., 21 and 23 refer to circle

and diamond.

#### **Details**

This function obtains regional association plot for a particular locus, based on the information about recombinatino rates, linkage disequilibria between the SNP of interest and neighbouring ones, and single-point association tests p values.

Note that the best p value is not necessarily within locus in the original design.

### Author(s)

Paul de Bakker, Jing Hua Zhao, Shengxu Li

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#### References

Saxena R, Voight BF, Lyssenko V, Burtt NP, de Bakker PI, Chen H, Roix JJ, Kathiresan S, Hirschhorn JN, Daly MJ, Hughes TE, Groop L, Altshuler D, Almgren P, Florez JC, Meyer J, Ardlie K, Bengtsson Boström K, Isomaa B, Lettre G, Lindblad U, Lyon HN, Melander O, Newton-Cheh C, Nilsson P, Orho-Melander M, Råstam L, Speliotes EK, Taskinen MR, Tuomi T, Guiducci C, Berglund A, Carlson J, Gianniny L, Hackett R, Hall L, Holmkvist J, Laurila E, Sjögren M, Sterner M, Surti A, Svensson M, Svensson M, Tewhey R, Blumenstiel B, Parkin M, Defelice M, Barry R, Brodeur W, Camarata J, Chia N, Fava M, Gibbons J, Handsaker B, Healy C, Nguyen K, Gates C, Sougnez C, Gage D, Nizzari M, Gabriel SB, Chirn GW, Ma Q, Parikh H, Richardson D, Ricke D, Purcell S (2007). "Genome-wide association analysis identifies loci for type 2 diabetes and triglyceride levels." *Science*, 316(5829), 1331-6. ISSN 0036-8075, doi:10.1126/science.1142358.

```
## Not run:
require(gap.datasets)
asplot(CDKNlocus, CDKNmap, CDKNgenes)
title("CDKN2A/CDKN2B Region")
asplot(CDKNlocus, CDKNmap, CDKNgenes, best.pval=5.4e-8, sf=c(3,6))
## NCBI2R
options(stringsAsFactors=FALSE)
p <- with(CDKNlocus,data.frame(SNP=NAME,PVAL))</pre>
hit <- subset(p,PVAL==min(PVAL,na.rm=TRUE))$SNP</pre>
library(NCBI2R)
# LD under build 36
chr_pos <- GetSNPInfo(with(p,SNP))[c("chr","chrpos")]</pre>
1 <- with(chr_pos,min(as.numeric(chrpos),na.rm=TRUE))</pre>
u <- with(chr_pos,max(as.numeric(chrpos),na.rm=TRUE))</pre>
LD <- with(chr_pos,GetLDInfo(unique(chr),1,u))
# We have complaints; a possibility is to get around with
# https://ftp.ncbi.nlm.nih.gov/hapmap/
hit_LD <- subset(LD,SNPA==hit)</pre>
hit_LD <- within(hit_LD,{RSQR=r2})</pre>
info <- GetSNPInfo(p$SNP)</pre>
haldane <- function(x) 0.5*(1-exp(-2*x))
locus <- with(info, data.frame(CHR=chr,POS=chrpos,NAME=marker,</pre>
                     DIST=(chrpos-min(chrpos))/1000000,
                     THETA=haldane((chrpos-min(chrpos))/100000000)))
locus <- merge.data.frame(locus,hit_LD,by.x="NAME",by.y="SNPB",all=TRUE)</pre>
locus <- merge.data.frame(locus,p,by.x="NAME",by.y="SNP",all=TRUE)</pre>
locus <- subset(locus,!is.na(POS))</pre>
ann <- AnnotateSNPList(p$SNP)</pre>
genes <- with(ann,data.frame(ID=locusID,CLASS=fxn_class,PATH=pathways,</pre>
                               START=GeneLowPoint, STOP=GeneHighPoint,
                               STRAND=ori,GENE=genesymbol,BUILD=build,CYTO=cyto))
attach(genes)
ugenes <- unique(GENE)</pre>
ustart <- as.vector(as.table(by(START,GENE,min))[ugenes])</pre>
ustop <- as.vector(as.table(by(STOP,GENE,max))[ugenes])</pre>
ustrand <- as.vector(as.table(by(as.character(STRAND),GENE,max))[ugenes])</pre>
detach(genes)
genes <- data.frame(START=ustart,STOP=ustop,STRAND=ustrand,GENE=ugenes)</pre>
genes <- subset(genes,START!=0)</pre>
```

b2r

```
rm(l,u,ugenes,ustart,ustop,ustrand)
# Assume we have the latest map as in CDKNmap
asplot(locus,CDKNmap,genes)
## End(Not run)
```

b2r

Obtain correlation coefficients and their variance-covariances

#### **Description**

Obtain correlation coefficients and their variance-covariances

#### Usage

```
b2r(b, s, rho, n)
```

# **Arguments**

b	the vector of linear regression coefficients.
S	the corresponding vector of standard errors.
rho	triangular array of between-SNP correlation.
n	the sample size.

### **Details**

This function converts linear regression coefficients of phenotype on single nucleotide polymorphisms (SNPs) into Pearson correlation coefficients with their variance-covariance matrix. It is useful as a preliminary step for meta-analyze SNP-trait associations at a given region. Between-SNP correlations (e.g., from HapMap) are required as auxiliary information.

### Value

The returned value is a list containing:

- r the vector of correlation coefficients.
- V the variance-covariance matrix of correlations.

# Author(s)

Jing Hua Zhao

# References

Becker BJ (2004). Multivariate meta-analysis. in Tinsley HEA, Brown SD (Ed.) Handbook of Applied Multivariate Statistics and Mathematical Modeling (Chapter 17, pp499-525). Academic Press.

Casella G, Berger RL (2002). Statistical Inference, 2nd Edition, Duxbury.

Elston RC (1975). "On the correlation between correlations." *Biometrika*, **62**(1), 133-140. ISSN 0006-3444, doi:10.1093/biomet/62.1.133.

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#### See Also

```
mvmeta, LD22
```

# **Examples**

```
## Not run:
n <- 10
r <- c(1,0.2,1,0.4,0.5,1)
b <- c(0.1,0.2,0.3)
s <- c(0.4,0.3,0.2)
bs <- b2r(b,s,r,n)
## End(Not run)</pre>
```

**BFDP** 

Bayesian false-discovery probability

# Description

Bayesian false-discovery probability

#### Usage

```
BFDP(a, b, pi1, W, logscale = FALSE)
```

# Arguments

a parameter value at which the power is to be evaluated.

b the variance for a, or the uppoer point  $(RR_{hi})$  of a 95%CI if logscale=FALSE.

pi1 the prior probability of a non-null association.

W the prior variance.

logscale FALSE=the orginal scale, TRUE=the log scale.

### **Details**

This function calculates BFDP, the approximate  $P(H_0|\hat{\theta})$ , given an estiamte of the log relative risk,  $\hat{\theta}$ , the variance of this estimate, V, the prior variance, W, and the prior probability of a non-null association. When logscale=TRUE, the function accepts an estimate of the relative risk,  $\hat{RR}$ , and the upper point of a 95% confidence interval  $RR_{hi}$ .

# Value

The returned value is a list with the following components: PH0. probability given a,b). PH1. probability given a,b,W). BF. Bayes factor,  $P_{H_0}/P_{H_1}$ . BFDP. Bayesian false-discovery probability. ABF. approxmiate Bayes factor. ABFDP. approximate Bayesian false-discovery probability.

### Note

Adapted from BFDP functions by Jon Wakefield on 17th April, 2007.

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#### Author(s)

Jon Wakefield, Jing Hua Zhao

#### References

Wakefield J (2007). "A Bayesian measure of the probability of false discovery in genetic epidemiology studies." *Am J Hum Genet*, **81**(2), 208-27. ISSN 0002-9297 (Print) 0002-9297, doi:10.1086/519024.

#### See Also

**FPRP** 

```
## Not run:
# Example from BDFP.xls by Jon Wakefield and Stephanie Monnier
# Step 1 - Pre-set an BFDP-level threshold for noteworthiness: BFDP values below this
#
           threshold are noteworthy
\# The threshold is given by R/(1+R) where R is the ratio of the cost of a false
# non-discovery to the cost of a false discovery
T <- 0.8
# Step 2 - Enter up values for the prior that there is an association
pi0 <- c(0.7, 0.5, 0.01, 0.001, 0.00001, 0.6)
\# Step 3 - Enter the value of the OR that is the 97.5% point of the prior, for example
           if we pick the value 1.5 we believe that the prior probability that the
           odds ratio is bigger than 1.5 is 0.025.
ORhi <- 3
W <- (\log(ORhi)/1.96)^2
# Step 4 - Enter OR estimate and 95% confidence interval (CI) to obtain BFDP
OR <- 1.316
OR_L <- 1.10
OR_U <- 2.50
logOR <- log(OR)
selogOR \leftarrow (log(OR_U)-log(OR))/1.96
r <- W/(W+selogOR^2)
z <- logOR/selogOR
ABF <- \exp(-z^2*r/2)/sqrt(1-r)
ABF
FF <- (1-pi0)/pi0
FF
BFDPex <- FF*ABF/(FF*ABF+1)
BFDPex
pi0[BFDPex>T]
```

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```
## now turn to BFDP

pi0 <- c(0.7,0.5,0.01,0.001,0.00001,0.6)

ORhi <- 3

OR <- 1.316

OR_U <- 2.50

W <- (log(ORhi)/1.96)^2

z <- BFDP(OR,OR_U,pi0,W)

z

## End(Not run)
```

bt

Bradley-Terry model for contingency table

# Description

Bradley-Terry model for contingency table

# Usage

bt(x)

# **Arguments**

Χ

the data table.

# **Details**

This function calculates statistics under Bradley-Terry model.

# Value

The returned value is a list containing:

- y A column of 1.
- count the frequency count/weight.
- allele the design matrix.
- bt.glm a glm.fit object.
- etdt.dat a data table that can be used by ETDT.

# Note

Adapted from a SAS macro for data in the example section.

# Author(s)

Jing Hua Zhao

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#### References

Bradley RA, Terry ME (1952) Rank analysis of incomplete block designs I. the method of paired comparisons. Biometrika 39:324–345

Sham PC, Curtis D (1995) An extended transmission/disequilibrium test (TDT) for multi-allelic marker loci. Ann. Hum. Genet. 59:323-336

#### See Also

mtdt

#### **Examples**

```
## Not run:
# Copeman JB, Cucca F, Hearne CM, Cornall RJ, Reed PW,
# Ronningen KS, Undlien DE, Nistico L, Buzzetti R, Tosi R, et al.
\# (1995) Linkage disequilibrium mapping of a type 1
# diabetes susceptibility gene (IDDM7) to chromosome 2q31-q33.
# Nat Genet 9: 80-5
0,0, 1, 3, 0,0, 0, 2, 3, 0, 0, 0,
             2,3,26,35, 7,0, 2,10,11, 3, 4, 1,
             2,3,22,26, 6,2, 4, 4,10, 2, 2, 0,
             0,1, 7,10, 2,0, 0, 2, 2, 1, 1, 0,
             0,0, 1, 4, 0,1, 0, 1, 0, 0, 0, 0,
             0,2, 5, 4, 1,1, 0, 0, 0, 2, 0, 0,
             0,0, 2, 6, 1,0, 2, 0, 2, 0, 0, 0,
             0,3, 6,19, 6,0, 0, 2, 5, 3, 0, 0,
             0,0, 3, 1, 1,0, 0, 0, 1, 0, 0, 0,
             0,0, 0, 2, 0,0, 0, 0, 0, 0, 0, 0,
             0,0, 1, 0, 0,0, 0, 0, 0, 0, 0, 0), nrow=12)
# Bradley-Terry model, only deviance is available in glm
# (SAS gives score and Wald statistics as well)
bt.ex<-bt(x)
anova(bt.ex$bt.glm)
summary(bt.ex$bt.glm)
## End(Not run)
```

ccsize

Power and sample size for case-cohort design

# **Description**

Power and sample size for case-cohort design

# Usage

```
ccsize(n, q, pD, p1, theta, alpha, beta = 0.2, power = FALSE, verbose = FALSE)
```

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### **Arguments**

n	the total number of subjects in the cohort.
q	the sampling fraction of the subcohort.
pD	the proportion of the failures in the full cohort.
p1	proportions of the two groups (p2=1-p1).
theta	log-hazard ratio for two groups.
alpha	type I error – significant level.
beta	type II error.
power	if specified, the power for which sample size is calculated.
verbose	error messages are explicitly printed out.

# **Details**

The power of the test is according to

$$\Phi\left(Z_{\alpha} + m^{1/2}\theta\sqrt{\frac{p_1p_2p_D}{q + (1 - q)p_D}}\right)$$

where  $\alpha$  is the significance level,  $\theta$  is the log-hazard ratio for two groups,  $p_j$ , j=1, 2, are the proportion of the two groups in the population. m is the total number of subjects in the subcohort,  $p_D$  is the proportion of the failures in the full cohort, and q is the sampling fraction of the subcohort.

Alternatively, the sample size required for the subcohort is

$$m = nBp_D/(n - B(1 - p_D))$$

where  $B = (Z_{1-\alpha} + Z_{\beta})^2/(\theta^2 p_1 p_2 p_D)$ , and n is the size of cohort.

When infeaisble configurations are specified, a sample size of -999 is returned.

# Value

The returned value is a value indicating the power or required sample size.

# Note

Programmed for EPIC study. keywords misc

# Author(s)

Jing Hua Zhao

#### References

Cai J, Zeng D (2004). "Sample size/power calculation for case-cohort studies." *Biometrics*, **60**(4), 1015-24. ISSN 0006-341X (Print) 0006-341x, doi:10.1111/j.0006341X.2004.00257.x.

# See Also

pbsize

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```
## Not run:
# Table 1 of Cai & Zeng (2004).
outfile <- "table1.txt"</pre>
cat("n","pD","p1","theta","q","power\n",file=outfile,sep="\t")
alpha <- 0.05
n <- 1000
for(pD in c(0.10, 0.05))
   for(p1 in c(0.3, 0.5))
   {
       for(theta in c(0.5,1.0))
       {
           for(q in c(0.1,0.2))
              power <- ccsize(n,q,pD,p1,alpha,theta)</pre>
              \mathtt{cat}(\texttt{n}, \texttt{"}\texttt{t"}, \texttt{pD}, \texttt{"}\texttt{t"}, \texttt{p1}, \texttt{"}\texttt{t"}, \texttt{theta}, \texttt{"}\texttt{t"}, \texttt{q}, \texttt{"}\texttt{t"}, \texttt{signif}(\texttt{power}, 3), \texttt{"}\texttt{'n"},
                    file=outfile,append=TRUE)
           }
       }
   }
}
n <- 5000
for(pD in c(0.05, 0.01))
{
   for(p1 in c(0.3, 0.5))
       for(theta in c(0.5,1.0))
           for(q in c(0.01, 0.02))
              power <- ccsize(n,q,pD,p1,alpha,theta)</pre>
              cat(n,"\t",pD,"\t",p1,"\t",theta,"\t",q,"\t",signif(power,3),"\n",
                    file=outfile,append=TRUE)
           }
       }
   }
}
table1<-read.table(outfile,header=TRUE,sep="\t")</pre>
unlink(outfile)
# ARIC study
outfile <- "aric.txt"</pre>
n <- 15792
pD <- 0.03
p1 <- 0.25
alpha <- 0.05
theta <-c(1.35,1.40,1.45)
beta1 <- 0.8
s_nb <- c(1463,722,468)
cat("n","pD","p1","hr","q","power","ssize\n",file=outfile,sep="\t")
for(i in 1:3)
{
  q \leftarrow s_nb[i]/n
  power <- ccsize(n,q,pD,p1,alpha,log(theta[i]))</pre>
  ssize <- ccsize(n,q,pD,p1,alpha,log(theta[i]),beta1)</pre>
  cat(n,"\t",pD,"\t",p1,"\t",theta[i],"\t",q,"\t",signif(power,3),"\t",ssize,"\n",
```

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```
file=outfile,append=TRUE)
}
aric<-read.table(outfile,header=TRUE,sep="\t")</pre>
unlink(outfile)
# EPIC study
outfile <- "epic.txt"</pre>
n <- 25000
alpha <- 0.00000005
power <- 0.8
s_pD \leftarrow c(0.3, 0.2, 0.1, 0.05)
s_p1 < - seq(0.1, 0.5, by=0.1)
s_hr <- seq(1.1, 1.4, by=0.1)
cat("n","pD","p1","hr","alpha","ssize\n",file=outfile,sep="\t")
# direct calculation
for(pD in s_pD)
   for(p1 in s_p1)
   {
      for(hr in s_hr)
         ssize <- ccsize(n,q,pD,p1,alpha,log(hr),power)</pre>
         if (ssize>0) cat(n,"\t",pD,"\t",p1,"\t",hr,"\t",alpha,"\t",ssize,"\n",
                            file=outfile,append=TRUE)
      }
   }
}
epic<-read.table(outfile,header=TRUE,sep="\t")</pre>
unlink(outfile)
# exhaustive search
outfile <- "search.txt"</pre>
s_q < - seq(0.01, 0.5, by=0.01)
cat("n","pD","p1","hr","nq","alpha","power\n",file=outfile,sep="\t")
for(pD in s_pD)
{
   for(p1 in s_p1)
      for(hr in s_hr)
      {
         for(q in s_q)
            power <- ccsize(n,q,pD,p1,alpha,log(hr))</pre>
             cat(n,"\t",pD,"\t",p1,"\t",hr,"\t",q*n,"\t",alpha,"\t",power,"\n",
                 file=outfile,append=TRUE)
         }
      }
   }
search<-read.table(outfile,header=TRUE,sep="\t")</pre>
unlink(outfile)
## End(Not run)
```

18 chow.test

#### **Description**

Chow's test for heterogeneity in two regressions

### Usage

```
chow.test(y1, x1, y2, x2, x = NULL)
```

#### **Arguments**

y1	a vector of dependent variable.
x1	a matrix of independent variables.
y2	a vector of dependent variable.
x2	a matrix of independent variables.
Х	a known matrix of independent variables.

#### **Details**

Chow's test is for differences between two or more regressions. Assuming that errors in regressions 1 and 2 are normally distributed with zero mean and homoscedastic variance, and they are independent of each other, the test of regressions from sample sizes  $n_1$  and  $n_2$  is then carried out using the following steps. 1. Run a regression on the combined sample with size  $n=n_1+n_2$  and obtain within group sum of squares called  $S_1$ . The number of degrees of freedom is  $n_1+n_2-k$ , with k being the number of parameters estimated, including the intercept. 2. Run two regressions on the two individual samples with sizes  $n_1$  and  $n_2$ , and obtain their within group sums of square  $S_2+S_3$ , with  $n_1+n_2-2k$  degrees of freedom. 3. Conduct an  $F_{(k,n_1+n_2-2k)}$  test defined by

$$F = \frac{[S_1 - (S_2 + S_3)]/k}{[(S_2 + S_3)/(n_1 + n_2 - 2k)]}$$

If the F statistic exceeds the critical F, we reject the null hypothesis that the two regressions are equal.

In the case of haplotype trend regression, haplotype frequencies from combined data are known, so can be directly used.

# Value

The returned value is a vector containing (please use subscript to access them):

- F the F statistic.
- df1 the numerator degree(s) of freedom.
- df2 the denominator degree(s) of freedom.
- p the p value for the F test.

### Note

adapted from chow.R.

### Author(s)

Shigenobu Aoki, Jing Hua Zhao

chr\_pos\_a1\_a2

#### References

Chow GC (1960). "Tests of Equality Between Sets of Coefficients in Two Linear Regressions." *Econometrica*, **28**(3), 591-605. ISSN 00129682, 14680262, doi:10.2307/1910133.

#### See Also

htr

# **Examples**

```
## Not run:
dat1 <- matrix(c(</pre>
    1.2, 1.9, 0.9,
     1.6, 2.7, 1.3,
     3.5, 3.7, 2.0,
     4.0, 3.1, 1.8,
     5.6, 3.5, 2.2,
     5.7, 7.5, 3.5,
     6.7, 1.2, 1.9,
     7.5, 3.7, 2.7,
     8.5, 0.6, 2.1,
     9.7, 5.1, 3.6), byrow=TRUE, ncol=3)
dat2 <- matrix(c(</pre>
    1.4, 1.3, 0.5,
     1.5, 2.3, 1.3,
     3.1, 3.2, 2.5,
     4.4, 3.6, 1.1,
     5.1, 3.1, 2.8,
     5.2, 7.3, 3.3,
     6.5, 1.5, 1.3,
     7.8, 3.2, 2.2,
     8.1, 0.1, 2.8,
     9.5, 5.6, 3.9), byrow=TRUE, ncol=3)
y1<-dat1[,3]
y2<-dat2[,3]
x1<-dat1[,1:2]
x2<-dat2[,1:2]
chow.test.r<-chow.test(y1,x1,y2,x2)</pre>
# from http://aoki2.si.gunma-u.ac.jp/R/
## End(Not run)
```

chr\_pos\_a1\_a2

SNP id by chr:pos+a1/a2

# Description

SNP id by chr:pos+a1/a2

20 ci2ms

### Usage

```
chr_pos_a1_a2(
    chr,
    pos,
    a1,
    a2,
    prefix = "chr",
    seps = c(":", "_", "_"),
    uppercase = TRUE
)
```

# **Arguments**

chr	Chromosome.
pos	Position.
a1	Allele 1.
a2	Allele 2.
prefix	Prefix of the identifier.
seps	Delimiters.
uppercase	A flag to return in upper case.

# **Details**

This function generates unique identifiers for variants

# Value

Identifier.

# **Examples**

```
# rs12075
chr_pos_a1_a2(1,159175354,"A","G",prefix="chr",seps=c(":","_","_"),uppercase=TRUE)
```

ci2ms

Effect size and standard error from confidence interval

# **Description**

Effect size and standard error from confidence interval

# Usage

```
ci2ms(ci, logscale = TRUE, alpha = 0.05)
```

# **Arguments**

ci confidence interval (CI). The delimiter between lower and upper limit is either a hyphen (-) or en dash (-).

logscale a flag indicating the confidence interval is based on a log-scale.

alpha Type 1 error.

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#### **Details**

Effect size is a measure of strength of the relationship between two variables in a population or parameter estimate of that population. Without loss of generality, denote m and s to be the mean and standard deviation of a sample from  $N(\mu, \sigma^2)$ ). Let  $z \sim N(0,1)$  with cutoff point  $z_{\alpha}$ , confidence limits L, U in a CI are defined as follows,

$$L = m - z_{\alpha}s$$
$$U = m + z_{\alpha}s$$

 $\Rightarrow U + L = 2m, U - L = 2z_{\alpha}s$ . Consequently,

$$m = \frac{U+L}{2}$$
 
$$s = \frac{U-L}{2z_{\alpha}}$$

Effect size in epidemiological studies on a binary outcome is typically reported as odds ratio from a logistic regression or hazard ratio from a Cox regression,  $L \equiv \log(L)$ ,  $U \equiv \log(U)$ .

#### Value

Based on CI, the function provides a list containing estimates

- m effect size (log(OR))
- · s standard error
- direction a decrease/increase (-/+) sign such that sign(m)=-1, 0, 1, is labelled "-", "0", "+", respectively as in PhenoScanner.

# **Examples**

```
# rs3784099 and breast cancer recurrence/mortality
ms <- ci2ms("1.28-1.72")
print(ms)
# Vector input
ci2 <- c("1.28-1.72","1.25-1.64")
ms2 <- ci2ms(ci2)
print(ms2)</pre>
```

```
circos.cis.vs.trans.plot
```

circos plot of cis/trans classification

# **Description**

circos plot of cis/trans classification

# Usage

```
circos.cis.vs.trans.plot(hits, panel, id, radius = 1e+06)
```

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# **Arguments**

hits A text file as input data with varibles named "CHR", "BP", "SNP", "prot".

Protein panel with prot(ein), uniprot (id) and "chr", "start", "end", "gene".

id Identifier.

radius The flanking distance as cis.

#### **Details**

The function implements a circos plot at the early stage of SCALLOP-INF meta-analysis.

# Value

None.

# **Examples**

```
## Not run:
    circos.cis.vs.trans.plot(hits="INF1.clumped", panel=inf1, id="uniprot")
## End(Not run)
```

circos.cnvplot

circos plot of CNVs.

# Description

circos plot of CNVs.

# Usage

```
circos.cnvplot(data)
```

# **Arguments**

data

CNV data containing chromosome, start, end and freq.

### **Details**

The function plots frequency of CNVs.

# Value

None.

```
## Not run:
circos.cnvplot(cnv)
## End(Not run)
```

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circos.mhtplot

circos Manhattan plot with gene annotation

# Description

circos Manhattan plot with gene annotation

# Usage

```
circos.mhtplot(data, glist)
```

### **Arguments**

data Data to be used. glist A gene list.

#### **Details**

The function generates circos Manhattan plot with gene annotation.

#### Value

None.

# Examples

circos.mhtplot2

Another circos Manhattan plot

# Description

Another circos Manhattan plot

# Usage

```
circos.mhtplot2(dat, labs, species = "hg18", ticks = 0:3 * 10, y = 20)
```

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# **Arguments**

dat Data to be plotted with variables chr, pos, log10p.
labs Data on labels.
species Genome build.
ticks Tick positions.
y Starting position of y-axis label.

#### **Details**

This is adapted from work for a recent publication. It enables a y-axis to the -log10(P) for association statistics

#### Value

There is no return value but a plot.

### **Examples**

```
## Not run:
require(gap.datasets)
library(dplyr)
testdat <- mhtdata[c("chr","pos","p","gene","start","end")] %>%
         rename(log10p=p) %>%
         mutate(chr=paste0("chr",chr),log10p=-log10(log10p))
dat <- mutate(testdat,start=pos,end=pos) %>%
      select(chr,start,end,log10p)
labs <- subset(testdat,gene %in% glist) %>%
       group_by(gene,chr,start,end) %>%
       summarize() %>%
       mutate(cols="blue") %>%
       select(chr,start,end,gene,cols)
circos.mhtplot2(dat,labs,ticks=0:2*10)
# https://www.rapidtables.com/web/color/RGB_Color.html
## End(Not run)
```

```
{\it cis.vs.trans.classification} \\ {\it A cis/trans classifier}
```

# **Description**

A cis/trans classifier

# Usage

```
cis.vs.trans.classification(hits, panel, id, radius = 1e+06)
```

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# **Arguments**

hits Data to be used, which contains prot, Chr, bp, id and/or other information such

as SNPid.

panel Panel data. id Identifier.

radius The flanking distance for variants.

#### **Details**

The function classifies variants into cis/trans category according to a panel which contains id, chr, start, end, gene variables.

#### Value

The cis/trans classification.

#### Author(s)

James Peters

# **Examples**

cnvplot

genomewide plot of CNVs

### **Description**

genomewide plot of CNVs

# Usage

```
cnvplot(data)
```

### **Arguments**

data

Data to be used.

26 comp.score

#### **Details**

The function generates a plot containing genomewide copy number variants (CNV) chr, start, end, freq(uencies).

#### Value

None.

# **Examples**

```
knitr::kable(cnv,caption="A CNV dataset")
cnvplot(cnv)
```

comp.score

score statistics for testing genetic linkage of quantitative trait

#### **Description**

score statistics for testing genetic linkage of quantitative trait

### Usage

```
comp.score(
  ibddata = "ibd_dist.out",
  phenotype = "pheno.dat",
  mean = 0,
  var = 1,
  h2 = 0.3
)
```

#### **Arguments**

ibddata The output file from GENEHUNTER using command "dump ibd". The default

file name is  $ibd_dist.out$ .

phenotype The file of pedigree structure and trait value. The default file name is "pheno.dat".

Columns (no headings) are: family ID, person ID, father ID, mother ID, gender, trait value, where Family ID and person ID must be numbers, not characters.

Use character "NA" for missing phenotypes.

mean (population) mean of the trait, with a default value of 0.
var (population) variance of the trait, with a default value of 1.

heritability of the trait, with a default value of 0.3.

# **Details**

The function empirically estimate the variance of the score functions. The variance-covariance matrix consists of two parts: the additive part and the part for the individual-specific environmental effect. Other reasonable decompositions are possible.

This program has the following improvement over "score.r":

1. It works with selected nuclear families

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- 2. Trait data on parents (one parent or two parents), if available, are utilized.
- 3. Besides a statistic assuming no locus-specific dominance effect, it also computes a statistic that allows for such effect. It computes two statistics instead of one.

Function "merge" is used to merge the IBD data for a pair with the transformed trait data (i.e.,  $w_k w_l$ ).

#### Value

a matrix with each row containing the location and the statistics and their p-values.

#### Note

Adapt from score2.r.

#### Author(s)

Yingwei Peng, Kai Wang

#### References

Kruglyak L, Daly MJ, Reeve-Daly MP, Lander ES (1996). "Parametric and nonparametric linkage analysis: a unified multipoint approach." *Am J Hum Genet*, **58**(6), 1347-63. ISSN 0002-9297 (Print) 0002-9297.

Kruglyak L, Lander ES (1998). "Faster multipoint linkage analysis using Fourier transforms." *J Comput Biol*, **5**(1), 1-7. ISSN 1066-5277 (Print) 1066-5277, doi:10.1089/cmb.1998.5.1.

Wang K (2005). "A likelihood approach for quantitative-trait-locus mapping with selected pedigrees." *Biometrics*, **61**(2), 465-73. ISSN 0006-341X (Print) 0006-341x, doi:10.1111/j.15410420.2005.031213.x.

```
## Not run:
# An example based on GENEHUNTER version 2.1, with quantitative trait data in file
# "pheno.dat" generated from the standard normal distribution. The following
# exmaple shows that it is possible to automatically call GENEHUNTER using R
# function "system".

cwd <- getwd()
cs.dir <- file.path(find.package("gap"),"tests/comp.score")
setwd(cs.dir)
dir()
# system("gh < gh.inp")
cs.default <- comp.score()
setwd(cwd)

## End(Not run)</pre>
```

28 cs

cs Credible set

# Description

Credible set

# Usage

```
cs(tbl, b = "Effect", se = "StdErr", log_p = NULL, cutoff = 0.95)
```

# Arguments

tbl	Input data.
b	Effect size.
se	Standard error.
log_p	if not NULL it will be used to derive z-statistic
cutoff	Threshold for inclusion.

# **Details**

The function implements credible set as in fine-mapping.

# Value

Credible set.

```
## Not run:
\preformatted{
    zcat METAL/4E.BP1-1.tbl.gz | \
    awk 'NR==1 || ($1==4 && $2 >= 187158034 - 1e6 && $2 < 187158034 + 1e6)' > 4E.BP1.z
}
    tbl <- within(read.delim("4E.BP1.z"),{logp <- logp(Effect/StdErr)})
    z <- cs(tbl)
    l <- cs(tbl,log_p="logp")
## End(Not run)</pre>
```

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**ESplot** 

Effect-size plot

### **Description**

Effect-size plot

#### Usage

```
ESplot(ESdat, alpha = 0.05, fontsize = 12)
```

### **Arguments**

ESdat A data frame consisting of model id, parameter estimates and standard errors.

alpha Type-I error rate used to construct 100(1-alpha) confidence interval.

fontsize size of font.

#### **Details**

The function accepts parameter estimates and their standard errors for a range of models.

#### Value

A high resolution plot object.

# Author(s)

Jing Hua Zhao

```
rs12075 <- data.frame(id=c("CCL2","CCL7","CCL8","CCL11","CCL13","CXCL6","Monocytes"),
                       b=c(0.1694, -0.0899, -0.0973, 0.0749, 0.189, 0.0816, 0.0338387),
                       se=c(0.0113,0.013,0.0116,0.0114,0.0114,0.0115,0.00713386))
ESplot(rs12075)
# The function replaces an older implementation.
within(data.frame(
       id=c("Basic model","Adjusted","Moderately adjusted","Heavily adjusted","Other"),
       b=log(c(4.5,3.5,2.5,1.5,1)),
       se=c(0.2,0.1,0.2,0.3,0.2)
), {
   lcl <- exp(b-1.96*se)
   ucl \leftarrow exp(b+1.96*se)
   x \leftarrow seq(-2,8,length=length(id))
   y <- 1:length(id)
   plot(x,y,type="n",xlab="",ylab="",axes=FALSE)
   points((lcl+ucl)/2,y,pch=22,bg="black",cex=3)
   segments(lcl,y,ucl,y,lwd=3,lty="solid")
   axis(1,cex.axis=1.5,lwd=0.5)
   par(las=1)
   abline(v=1)
   axis(2,labels=id,at=y,lty="blank",hadj=0.2,cex.axis=1.5)
```

30 fbsize

```
title("A fictitious plot")
})
```

fbsize

Sample size for family-based linkage and association design

# Description

Sample size for family-based linkage and association design

### Usage

```
fbsize(
  gamma,
  p,
  alpha = c(1e-04, 1e-08, 1e-08),
  beta = 0.2,
  debug = 0,
  error = 0
)
```

# **Arguments**

gamma genotype relative risk assuming multiplicative model.

p frequency of disease allele.

alpha Type I error rates for ASP linkage, TDT and ASP-TDT.

beta Type II error rate.

debug verbose output.

error 0=use the correct formula,1=the original paper.

# **Details**

This function implements Risch and Merikangas (1996) statistics evaluating power for family-based linkage (affected sib pairs, ASP) and association design. They are potentially useful in the prospect of genome-wide association studies.

The function calls auxiliary functions sn() and strlen; sn() contains the necessary thresholds for power calculation while strlen() evaluates length of a string (generic).

# Value

The returned value is a list containing:

- gamma input gamma.
- p input p.
- n1 sample size for ASP.
- n2 sample size for TDT.
- n3 sample size for ASP-TDT.
- · lambdao lambda o.
- lambdas lambda s.

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#### Note

extracted from rm.c.

#### Author(s)

Jing Hua Zhao

#### References

Risch, N. and K. Merikangas (1996). The future of genetic studies of complex human diseases. Science 273(September): 1516-1517.

Risch, N. and K. Merikangas (1997). Reply to Scott el al. Science 275(February): 1329-1330.

Scott, W. K., M. A. Pericak-Vance, et al. (1997). Genetic analysis of complex diseases. Science 275: 1327.

#### See Also

pbsize

```
models <- matrix(c(</pre>
   4.0, 0.01,
   4.0, 0.10,
   4.0, 0.50,
   4.0, 0.80,
   2.0, 0.01,
   2.0, 0.10,
   2.0, 0.50,
   2.0, 0.80,
   1.5, 0.01,
   1.5, 0.10,
   1.5, 0.50,
   1.5, 0.80), ncol=2, byrow=TRUE)
outfile <- "fbsize.txt"</pre>
cat("gamma","p","Y","N_asp","P_A","H1","N_tdt","H2","N_asp/tdt","L_o","L_s\n",
    file=outfile, sep="\t")
for(i in 1:12) {
  g <- models[i,1]</pre>
  p \leftarrow models[i,2]
  z <- fbsize(g,p)
  \texttt{cat}(z\$\texttt{gamma},z\$\texttt{p},z\$\texttt{n1},z\$\texttt{pA},z\$\texttt{h1},z\$\texttt{n2},z\$\texttt{h2},z\$\texttt{n3},z\$\texttt{lambdao},z\$\texttt{lambdas},\texttt{file=outfile},
       append=TRUE,sep="\t")
  cat("\n",file=outfile,append=TRUE)
table1 <- read.table(outfile,header=TRUE,sep="\t")</pre>
nc <- c(4,7,9)
table1[,nc] <- ceiling(table1[,nc])</pre>
dc <- c(3,5,6,8,10,11)
table1[,dc] <- round(table1[,dc],2)</pre>
unlink(outfile)
# APOE-4, Scott WK, Pericak-Vance, MA & Haines JL
# Genetic analysis of complex diseases 1327
g <- 4.5
p < -0.15
```

32 FPRP

```
cat("\nAlzheimer's:\n\n") fbsize(g,p) # note to replicate the Table we need set alpha=9.961139e-05,4.910638e-08 and # beta=0.2004542 or reset the quantiles in fbsize.R
```

**FPRP** 

False-positive report probability

#### **Description**

False-positive report probability

### Usage

```
FPRP(a, b, pi0, ORlist, logscale = FALSE)
```

#### **Arguments**

a parameter value at which the power is to be evaluated.

b the variance for a, or the uppoer point of a 95%CI if logscale=FALSE.

pi0 the prior probability that  $H_0$  is true. ORlist a vector of ORs that is most likely.

logscale FALSE=a,b in orginal scale, TRUE=a, b in log scale.

#### **Details**

The function calculates the false positive report probability (FPRP), the probability of no true association between a genetic variant and disease given a statistically significant finding, which depends not only on the observed P value but also on both the prior probability that the association is real and the statistical power of the test. An associate result is the false negative reported probability (FNRP). See example for the recommended steps.

The FPRP and FNRP are derived as follows. Let  $H_0$ =null hypothesis (no association),  $H_A$ =alternative hypothesis (association). Since classic frequentist theory considers they are fixed, one has to resort to Bayesian framework by introduing prior,  $\pi = P(H_0 = TRUE) = P(association)$ . Let T=test statistic, and  $P(T > z_\alpha|H_0 = TRUE) = P(rejecting \ H_0|H_0 = TRUE) = \alpha$ ,  $P(T > z_\alpha|H_0 = FALSE) = P(rejecting \ H_0|H_A = TRUE) = 1 - \beta$ . The joint probability of test and truth of hypothesis can be expressed by  $\alpha$ ,  $\beta$  and  $\pi$ .

Joint probability of significance of test and truth of hypothesis

Truth of $H_A$	significant	nonsignificant	Total
TRUE	$(1-\beta)\pi$	$eta\pi$	$\pi$
FALSE	$\alpha(1-\pi)$	$(1-\alpha)(1-\pi)$	$1-\pi$
Total	$(1-\beta)\pi + \alpha(1-\pi)$	$\beta\pi + (1-\alpha)(1-\pi)$	1

We have 
$$FPRP = P(H_0 = TRUE | T > z_{\alpha}) = \alpha(1 - \pi)/[\alpha(1 - \pi) + (1 - \beta)\pi] = \{1 + \pi/(1 - \pi)][(1 - \beta)/\alpha]\}^{-1}$$
 and similarly  $FNRP = \{1 + [(1 - \alpha)/\beta][(1 - \pi)/\pi]\}^{-1}$ .

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#### Value

The returned value is a list with components, p p value corresponding to a,b. power the power corresponding to the vector of ORs. FPRP False-positive report probability. FNRP False-negative report probability.

#### Author(s)

Jing Hua Zhao

### References

Wacholder S, Chanock S, Garcia-Closas M, El Ghormli L, Rothman N (2004). "Assessing the probability that a positive report is false: an approach for molecular epidemiology studies." *J Natl Cancer Inst*, **96**(6), 434-42. ISSN 0027-8874 (Print) 0027-8874, doi:10.1093/jnci/djh075.

# See Also

**BFDP** 

```
# Example by Laure El ghormli & Sholom Wacholder on 25-Feb-2004
# Step 1 - Pre-set an FPRP-level criterion for noteworthiness
T < -0.2
# Step 2 - Enter values for the prior that there is an association
pi0 <- c(0.25,0.1,0.01,0.001,0.0001,0.00001)
# Step 3 - Enter values of odds ratios (OR) that are most likely, assuming that
             there is a non-null association
ORlist <- c(1.2, 1.5, 2.0)
# Step 4 - Enter OR estimate and 95% confidence interval (CI) to obtain FPRP
OR <- 1.316
ORlo <- 1.08
ORhi <- 1.60
logOR <- log(OR)
selogOR <- abs(logOR-log(ORhi))/1.96</pre>
p <- ifelse(logOR>0,2*(1-pnorm(logOR/selogOR)),2*pnorm(logOR/selogOR))\\
q \leftarrow qnorm(1-p/2)
POWER <- ifelse(log(ORlist)>0,1-pnorm(q-log(ORlist)/selogOR),
                   pnorm(-q-log(ORlist)/selogOR))
POWER
\label{eq:fprex} \mathsf{FPRPex} \ \leftarrow \ \mathsf{t}(\mathsf{p}*(\mathsf{1}\mathsf{-}\mathsf{pi0})/(\mathsf{p}*(\mathsf{1}\mathsf{-}\mathsf{pi0})\mathsf{+}\mathsf{POWER} \setminus
row.names(FPRPex) <- pi0</pre>
colnames(FPRPex) <- ORlist</pre>
FPRPex
FPRPex>T
```

gc.em

```
## now turn to FPRP
OR <- 1.316
ORhi <- 1.60
ORlist <- c(1.2,1.5,2.0)
pi0 <- c(0.25,0.1,0.01,0.001,0.0001,0.00001)
z <- FPRP(OR,ORhi,pi0,ORlist,logscale=FALSE)
z
## End(Not run)</pre>
```

g2a

Conversion of a genotype identifier to alleles

# Description

Conversion of a genotype identifier to alleles

# Usage

g2a(g)

# **Arguments**

g

a genotype identifier.

gc.em

Gene counting for haplotype analysis

# Description

Gene counting for haplotype analysis

# Usage

```
gc.em(
  data,
  locus.label = NA,
  converge.eps = 1e-06,
  maxiter = 500,
  handle.miss = 0,
  miss.val = 0,
  control = gc.control()
)
```

gc.em 35

### **Arguments**

data Matrix of alleles, such that each locus has a pair of adjacent columns of alleles,

and the order of columns corresponds to the order of loci on a chromosome. If there are K loci, then ncol(data) = 2\*K. Rows represent alleles for each subject.

locus.label Vector of labels for loci, of length K (see definition of data matrix).

converge.eps Convergence criterion, based on absolute change in log likelihood (lnlike).

maxiter Maximum number of iterations of EM.

handle.miss a flag for handling missing genotype data, 0=no, 1=yes.

miss.val missing value.

control a function, see genecounting.

#### **Details**

Gene counting for haplotype analysis with missing data, adapted for hap score

#### Value

List with components:

- converge Indicator of convergence of the EM algorithm (1=converged, 0 = failed).
- niter Number of iterations completed in the EM alogrithm.
- locus.info A list with a component for each locus. Each component is also a list, and the items of a locus-specific list are the locus name and a vector for the unique alleles for the locus.
- locus.label Vector of labels for loci, of length K (see definition of input values).
- haplotype Matrix of unique haplotypes. Each row represents a unique haplotype, and the number of columns is the number of loci.
- hap.prob Vector of mle's of haplotype probabilities. The ith element of hap.prob corresponds to the ith row of haplotype.
- hap.prob.noLD Similar to hap.prob, but assuming no linkage disequilibrium.
- Inlike Value of Inlike at last EM iteration (maximum Inlike if converged).
- lr Likelihood ratio statistic to test no linkage disequilibrium among all loci.
- indx.subj Vector for index of subjects, after expanding to all possible pairs of haplotypes for each person. If indx=i, then i is the ith row of input matrix data. If the ith subject has n possible pairs of haplotypes that correspond to their marker phenotype, then i is repeated n times.
- nreps Vector for the count of haplotype pairs that map to each subject's marker genotypes.
- hap1code Vector of codes for each subject's first haplotype. The values in hap1code are the row numbers of the unique haplotypes in the returned matrix haplotype.
- hap2code Similar to hap1code, but for each subject's second haplotype.
- post Vector of posterior probabilities of pairs of haplotypes for a person, given thier marker phenotypes.
- htrtable A table which can be used in haplotype trend regression.

# Note

Adapted from GENECOUNTING.

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# Author(s)

Jing Hua Zhao

# References

Zhao JH, Lissarrague S, Essioux L, Sham PC (2002). "GENECOUNTING: haplotype analysis with missing genotypes." *Bioinformatics*, **18**(12), 1694-5. ISSN 1367-4803 (Print) 1367-4803, doi:10.1093/bioinformatics/18.12.1694.

Zhao JH, Sham PC (2003). "Generic number systems and haplotype analysis." *Comput Methods Programs Biomed*, **70**(1), 1-9. ISSN 0169-2607 (Print) 0169-2607, doi:10.1016/s01692607(01)00193-6.

#### See Also

```
genecounting, LDkl
```

# **Examples**

```
## Not run:
data(hla)
gc.em(hla[,3:8],locus.label=c("DQR","DQA","DQB"),control=gc.control(assignment="t"))
## End(Not run)
```

gc.lambda

Estionmation of the genomic control inflation statistic (lambda)

# Description

Estionmation of the genomic control inflation statistic (lambda)

# Usage

```
gc.lambda(p)
```

# **Arguments**

р

A vector of p values.

#### Value

Estimate of inflation factor.

```
set.seed(12345)
p <- runif(100)
lambda <- gc.lambda(p)</pre>
```

gcontrol 37

# **Description**

genomic control

# Usage

```
gcontrol(
  data,
  zeta = 1000,
  kappa = 4,
  tau2 = 1,
  epsilon = 0.01,
  ngib = 500,
  burn = 50,
  idum = 2348
)
```

# Arguments

data	the data matrix.
zeta	program constant with default value 1000.
kappa	multiplier in prior for mean with default value 4.
tau2	multiplier in prior for variance with default value 1.
epsilon	prior probability of marker association with default value 0.01.
ngib	number of Gibbs steps, with default value 500.
burn	number of burn-ins with default value 50.
idum	seed for pseudorandom number sequence.

## **Details**

The Bayesian genomic control statistics with the following parameters,

n	number of loci under consideration
11	number of foct under consideration
lambdahat	median(of the n trend statistics)/0.46
	Prior for noncentrality parameter Ai is
	Normal(sqrt(lambdahat)kappa,lambdahat*tau2)
kappa	multiplier in prior above, set at 1.6 * sqrt(log(n))
tau2	multiplier in prior above
epsilon	prior probability a marker is associated, set at 10/n
ngib	number of cycles for the Gibbs sampler after burn in
burn	number of cycles for the Gibbs sampler to burn in

Armitage's trend test along with the posterior probability that each marker is associated with the disorder is given. The latter is not a p-value but any value greater than 0.5 (pout) suggests association.

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#### Value

The returned value is a list containing:

- deltot the probability of being an outlier.
- x2 the  $\chi^2$  statistic.
- A the A vector.

### Note

Adapted from gcontrol by Bobby Jones and Kathryn Roeder, use -Dexecutable for standalone program, function getnum in the original code needs \

### Author(s)

Bobby Jones, Jing Hua Zhao

#### **Source**

```
https://www.cmu.edu/dietrich/statistics-datascience/index.html
```

#### References

Devlin B, Roeder K (1999). "Genomic control for association studies." *Biometrics*, **55**(4), 997-1004. ISSN 0006-341X (Print) 0006-341x, doi:10.1111/j.0006341x.1999.00997.x.

# **Examples**

 ${\tt gcontrol2}$ 

genomic control based on p values

## **Description**

genomic control based on p values

# Usage

```
gcontrol2(p, col = palette()[4], lcol = palette()[2], ...)
```

## Arguments

```
    p a vector of observed p values.
    col colour for points in the Q-Q plot.
    lcol colour for the diagonal line in the Q-Q plot.
    other options for plot.
```

gcp 39

#### **Details**

The function obtains 1-df  $\chi^2$  statistics (observed) according to a vector of p values, and the inflation factor (lambda) according to medians of the observed and expected statistics. The latter is based on the empirical distribution function (EDF) of 1-df  $\chi^2$  statstics.

It would be appropriate for genetic association analysis as of 1-df Armitage trend test for case-control data; for 1-df additive model with continuous outcome one has to consider the compatibility with p values based on z-/t- statistics.

#### Value

A list containing:

- x the expected  $\chi^2$  statistics.
- y the observed  $\chi^2$  statistics.
- lambda the inflation factor.

#### Author(s)

Jing Hua Zhao

#### References

Devlin B, Roeder K (1999) Genomic control for association studies. Biometrics 55:997-1004

# **Examples**

```
## Not run:
x2 <- rchisq(100,1,.1)
p <- pchisq(x2,1,lower.tail=FALSE)
r <- gcontrol2(p)
print(r$lambda)
## End(Not run)</pre>
```

gcp

Permutation tests using GENECOUNTING

## **Description**

Permutation tests using GENECOUNTING

## Usage

```
gcp(
   y,
   cc,
   g,
   handle.miss = 1,
   miss.val = 0,
   n.sim = 0,
```

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```
locus.label = NULL,
quietly = FALSE
)
```

## **Arguments**

y A column of 0/1 indicating cases and controls.

cc analysis indicator, 0 = marker-marker, 1 = case-control.

g the multilocus genotype data.

handle.miss a flag with value 1 indicating missing data are allowed.

miss.val missing value.

n.sim the number of permutations.

locus.label label of each locus.

quietly a flag if TRUE will suppress the screen output.

## **Details**

This function is a R port of the GENECOUNTING/PERMUTE program which generates EHPLUS-type statistics including z-tests for individual haplotypes

#### Value

The returned value is a list containing (p.sim and ph when n.sim > 0):

- x2obs the observed chi-squared statistic.
- pobs the associated p value.
- zobs the observed z value for individual haplotypes.
- p.sim simulated p value for the global chi-squared statistic.
- ph simulated p values for individual haplotypes.

## Note

Built on gcp.c.

#### Author(s)

Jing Hua Zhao

### References

Zhao JH, Curtis D, Sham PC (2000). "Model-free analysis and permutation tests for allelic associations." *Hum Hered*, **50**(2), 133-9. ISSN 0001-5652 (Print) 0001-5652, doi:10.1159/000022901.

Zhao JH (2004). "2LD. GENECOUNTING and HAP: computer programs for linkage disequilibrium analysis." *Bioinformatics*, **20**(8), 1325-6. ISSN 1367-4803 (Print) 1367-4803, doi:10.1093/bioinformatics/bth071.

Zhao JH, Qian WD (2003) Association analysis of unrelated individuals using polymorphic genetic markers – methods, implementation and application, Royal Statistical Society, Hassallt-Diepenbeek, Belgium.

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#### See Also

```
genecounting
```

#### **Examples**

```
## Not run:
data(fsnps)
y<-fsnps$y
cc<-1
g<-fsnps[,3:10]
gcp(y,cc,g,miss.val="Z",n.sim=5)
hap.score(y,g,method="hap",miss.val="Z")
## End(Not run)</pre>
```

genecounting

Gene counting for haplotype analysis

## **Description**

Gene counting for haplotype analysis

### Usage

```
genecounting(data, weight = NULL, loci = NULL, control = gc.control())
```

## **Arguments**

data genotype table.

weight a column of frequency weights.

loci an array containing number of alleles at each locus.

control is a function with the following arguments:

- xdata. a flag indicating if the data involves X chromosome, if so, the first column of data indicates sex of each subject: 1=male, 2=female. The marker data are no different from the autosomal version for females, but for males, two copies of the single allele present at a given locus.
- convll. set convergence criteria according to log-likelihood, if its value set to 1
- handle.miss. to handle missing data, if its value set to 1
- eps. the actual convergence criteria, with default value 1e-5
- tol. tolerance for genotype probabilities with default value 1e-8
- maxit. maximum number of iterations, with default value 50
- pl. criteria for trimming haplotypes according to posterior probabilities
- assignment. filename containing haplotype assignment
- verbose. If TRUE, yields print out from the C routine

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#### **Details**

Gene counting for haplotype analysis with missing data.

#### Value

The returned value is a list containing:

- h haplotype frequency estimates under linkage disequilibrium (LD).
- h0 haplotype frequency estimates under linkage equilibrium (no LD).
- prob genotype probability estimates.
- 10 log-likelihood under linkage equilibrium.
- 11 log-likelihood under linkage disequilibrium.
- hapid unique haplotype identifier (defunct, see gc.em).
- npusr number of parameters according user-given alleles.
- npdat number of parameters according to observed.
- htrtable design matrix for haplotype trend regression (defunct, see gc.em).
- iter number of iterations used in gene counting.
- · converge a flag indicating convergence status of gene counting.
- di0 haplotype diversity under no LD, defined as  $1 \sum (h_0^2)$ .
- dil haplotype diversity under LD, defined as  $1 \sum (h^2)$ .
- resid residuals in terms of frequency weights = o e.

#### Note

adapted from GENECOUNTING.

#### Author(s)

Jing Hua Zhao

#### References

Zhao JH, Lissarrague S, Essioux L, Sham PC (2002). "GENECOUNTING: haplotype analysis with missing genotypes." *Bioinformatics*, **18**(12), 1694-5. ISSN 1367-4803 (Print) 1367-4803, doi:10.1093/bioinformatics/18.12.1694.

Zhao JH, Sham PC (2003). "Generic number systems and haplotype analysis." *Comput Methods Programs Biomed*, **70**(1), 1-9. ISSN 0169-2607 (Print) 0169-2607, doi:10.1016/s01692607(01)00193-6

Zhao JH (2004). "2LD. GENECOUNTING and HAP: computer programs for linkage disequilibrium analysis." *Bioinformatics*, **20**(8), 1325-6. ISSN 1367-4803 (Print) 1367-4803, doi:10.1093/bioinformatics/bth071.

## See Also

gc.em, LDkl

geno.recode 43

## **Examples**

```
## Not run:
require(gap.datasets)
# HLA data
data(hla)
hla.gc <- genecounting(hla[,3:8])</pre>
summary(hla.gc)
hla.gc$10
hla.gc$l1
# ALDH2 data
data(aldh2)
control <- gc.control(handle.miss=1,assignment="ALDH2.out")</pre>
aldh2.gc <- genecounting(aldh2[,3:6],control=control)</pre>
summary(aldh2.gc)
aldh2.gc$10
aldh2.gc$11
# Chromosome X data
# assuming allelic data have been extracted in columns 3-13
# and column 3 is sex
filespec <- system.file("tests/genecounting/mao.dat")</pre>
mao2 <- read.table(filespec)</pre>
dat <- mao2[,3:13]</pre>
loci <- c(12,9,6,5,3)
contr <- gc.control(xdata=TRUE, handle.miss=1)</pre>
mao.gc <- genecounting(dat,loci=loci,control=contr)</pre>
mao.gc$npusr
mao.gc$npdat
## End(Not run)
```

geno.recode

Genotype recoding

## **Description**

Genotype recoding

# Usage

```
geno.recode(geno, miss.val = 0)
```

# Arguments

```
geno genotype.
miss.val missing value.
```

get\_pve\_se

get\_b\_se

Get b and se from AF, n, and z

### **Description**

The function obtains effect size and its standard error.

### Usage

```
get_b_se(f, n, z)
```

#### **Arguments**

```
f Allele frequency.
n Sample size.
z z-statistics.
```

#### Value

b and se.

# **Examples**

```
## Not run:
 library(dplyr)
 # eQTLGen
 cis_pQTL <- merge(read.delim('eQTLGen.lz') %>%
             filter(GeneSymbol=="LTBR"),read.delim("eQTLGen.AF"),by="SNP") %>%
             mutate(data.frame(get_b_se(AlleleB_all,NrSamples,Zscore)))
 head(cis_pQTL,1)
       SNP
             Pvalue SNPChr SNPPos AssessedAllele OtherAllele Zscore
 rs1003563 2.308e-06 12 6424577
                                                           G 4.7245
                                               Α
            Gene GeneSymbol GeneChr GenePos NrCohorts NrSamples
                                                                      FDR
 ENSG00000111321
                     LTBR
                                12 6492472 34
                                                        23991 0.006278872
 BonferroniP hg19_chr hg19_pos AlleleA AlleleB allA_total allAB_total
                  12 6424577
                                           G
                                                   2574
                                                               8483
 allB_total AlleleB_all
                                b
                                           se
       7859 0.6396966 0.04490488 0.009504684
## End(Not run)
```

get\_pve\_se

Get pve and its standard error from n, z

## **Description**

Get pve and its standard error from n, z

### Usage

```
get_pve_se(n, z, correction = TRUE)
```

get\_sdy 45

# Arguments

n Sample size.

z z-statistic, i.e., b/se when they are available instead.
correction if TRUE an correction based on t-statistic is applied.

# **Details**

This function obtains proportion of explained variance of a continuous outcome.

## Value

pve and its se.

get\_sdy

Get sd(y) from AF, n, b, se

# Description

```
Get sd(y) from AF, n, b, se
```

## Usage

```
get_sdy(f, n, b, se, method = "mean", ...)
```

# Arguments

f Allele frequency.

n Sample size.

b effect size.

se standard error.

method of averaging: "mean" or "median".

... argument(s) passed to method

# **Details**

This function obtains standard error of a continuous outcome.

## Value

sd(y).

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#### **Examples**

```
## Not run:
set.seed(1)
X1 <- matrix(rbinom(1200,1,0.4),ncol=2)</pre>
X2 <- matrix(rbinom(1000,1,0.6),ncol=2)</pre>
colnames(X1) \leftarrow colnames(X2) \leftarrow c("f1","f2")
Y1 <- rnorm(600,apply(X1,1,sum),2)
Y2 <- rnorm(500,2*apply(X2,1,sum),5)
summary(lm1 <- lm(Y1~f1+f2,data=as.data.frame(X1)))</pre>
summary(lm2 <- lm(Y2~f1+f2,data=as.data.frame(X2)))</pre>
b1 <- coef(lm1)
b2 <- coef(1m2)
v1 \leftarrow vcov(lm1)
v2 \leftarrow vcov(1m2)
require(coloc)
## Bayesian approach, esp. when only p values are available
abf \leftarrow coloc.abf(list(beta=b1, varbeta=diag(v1), N=nrow(X1), sdY=sd(Y1), type="quant"),
                                                     list(beta=b2, varbeta=diag(v2), N=nrow(X2), sdY=sd(Y2), type="quant"))
abf
# sdY
 {\rm cat}("{\rm sd}(Y)=",{\rm sd}(Y1),"==>Estimates:",{\rm sqrt}({\rm diag}({\rm var}(X1)*b1[-1]^2+{\rm var}(X1)*v1[-1,-1]*nrow(X1))),"\\ {\rm ''}(X1)={\rm cat}("{\rm sd}(Y1),"==>Estimates:",{\rm sqrt}({\rm diag}({\rm var}(X1)*b1[-1]^2+{\rm var}(X1)*v1[-1,-1]*nrow(X1))),"\\ {\rm ''}(X1)={\rm cat}(X1)={\rm cat}(X1)
for(k in 1:2)
{
      k1 <- k + 1
      {\sf cat("Based\ on\ b",k,"\ sd(Y1) = ",sqrt(var(X1[,k])*(b1[k1]^2+nrow(X1)*v1[k1,k1])),"\ n",sep="") } 
}
 {\sf cat}("sd(Y)=",sd(Y2),"==> Estimates:",sqrt(diag(var(X2)*b2[-1]^2+var(X2)*v2[-1,-1]*nrow(X2))),"\\ {\sf ''}n") 
for(k in 1:2)
{
      k1 < - k + 1
     cat("Based on b", k," sd(Y2) = ", sqrt(var(X2[,k])*(b2[k1]^2+nrow(X2)*v2[k1,k1])),"n", sep="")
get_sdy(0.6396966,23991,0.04490488,0.009504684)
## End(Not run)
```

gif

Kinship coefficient and genetic index of familiality

## **Description**

Kinship coefficient and genetic index of familiality

# Usage

```
gif(data, gifset)
```

## Arguments

```
data the trio data of a pedigree.

gifset a subgroup of pedigree members.
```

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#### **Details**

The genetic index of familality is defined as the mean kinship between all pairs of individuals in a set multiplied by 100,000. Formally, it is defined in (Gholami and Thomas 1994) as

$$100,000 \times \frac{2}{n(n-1)} \sum_{i=1}^{n-1} \sum_{j=i+1}^{n} k_{ij}$$

where n is the number of individuals in the set and  $k_{ij}$  is the kinship coefficient between individuals i and j.

The scaling is purely for convenience of presentation.

### Value

The returned value is a list containing:

• gifval the genetic index of familiarity.

#### Note

Adapted from gif.c, testable with -Dexecutable as standalone program, which can be use for any pair of indidivuals

#### Author(s)

Alun Thomas, Jing Hua Zhao

#### References

Gholami K, Thomas A (1994). "A linear time algorithm for calculation of multiple pairwise kinship coefficients and the genetic index of familiality." *Comput Biomed Res*, **27**(5), 342-50. ISSN 0010-4809 (Print) 0010-4809, doi:10.1006/cbmr.1994.1026.

#### See Also

pfc

#### Examples

```
## Not run:
test<-c(
 5,
 1,
 9,
10,
          9,
15,
          9,
21,
         10,
                   15,
3,
          0,
18,
          3,
                   15,
         21,
23,
 2,
          0,
 7,
          0,
```

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```
12,
          9,
                  6,
13,
                   6,
14,
         5,
                   8,
16,
         14,
                   6,
17,
         10,
                  2,
19,
         9,
                 11,
20,
         10,
                 13,
22,
         21,
                 20)
test<-matrix(test,ncol=3,byrow=TRUE)</pre>
gif(test,gifset=c(20,21,22))
# all individuals
gif(test,gifset=1:23)
## End(Not run)
```

grid2d

Two-dimensional grid

# Description

This function build 2-d grids

## Usage

```
grid2d(
  chrlen,
  plot = TRUE,
  cex.labels = 0.6,
  xlab = "QTL position",
  ylab = "Gene position")
```

# Arguments

chrlen Lengths of chromosomes; e.g., hg18, hg19 or hg38.

plot A flag for plot.

cex.labels A scaling factor for labels.

xlab X-axis title.

ylab Y-axis title.

### Value

A list with two variables:

- n Number of chromosomes.
- CM Cumulative lengths starting from 0.

h2.jags 49

h2. jags Heritability estimation based on genomic relationship m JAGS	natrix using
--	--------------

# Description

Heritability estimation based on genomic relationship matrix using JAGS

# Usage

```
h2.jags(
   y,
   x,
   G,
   eps = 1e-04,
   sigma.p = 0,
   sigma.r = 1,
   parms = c("b", "p", "r", "h2"),
   ...
)
```

## **Arguments**

```
y outcome vector.

x covariate matrix.

G genomic relationship matrix.

eps a positive diagonal perturbation to G.

sigma.p initial parameter values.

sigma.r initial parameter values.

parms monitored parmeters.

... parameters passed to jags, e.g., n.chains, n.burnin, n.iter.
```

## **Details**

This function performs Bayesian heritability estimation using genomic relationship matrix.

### Value

The returned value is a fitted model from jags().

## Author(s)

Jing Hua Zhao keywords htest

# References

Zhao JH, Luan J, Congdon P (2018). "Bayesian Linear Mixed Models with Polygenic Effects." *Journal of Statistical Software*, **85**(6), 1 - 27. doi:10.18637/jss.v085.i06.

50 h2G

#### **Examples**

```
## Not run:
require(gap.datasets)
set.seed(1234567)
meyer <- within(meyer,{</pre>
    y[is.na(y)] \gets rnorm(length(y[is.na(y)]), mean(y, na.rm=TRUE), sd(y, na.rm=TRUE))
    g1 <- ifelse(generation==1,1,0)</pre>
    g2 \leftarrow ifelse(generation==2,1,0)
    id <- animal
    animal <- ifelse(!is.na(animal),animal,0)</pre>
    dam <- ifelse(!is.na(dam),dam,0)</pre>
    sire <- ifelse(!is.na(sire),sire,0)</pre>
})
G \leftarrow kin.morgan(meyer)kin.matrix*2
library(regress)
r <- regress(y~-1+g1+g2,~G,data=meyer)</pre>
with(r,h2G(sigma,sigma.cov))
eps <- 0.001
y <- with(meyer,y)</pre>
x <- with(meyer,cbind(g1,g2))</pre>
ex <- h2.jags(y,x,G,sigma.p=0.03,sigma.r=0.014)
print(ex)
## End(Not run)
```

h2G

Heritability and its variance

# Description

Heritability and its variance

# Usage

```
h2G(V, VCOV, verbose = TRUE)
```

# **Arguments**

V Variance estimates.

VCOV Variance-covariance matrix.

verbose Detailed output.

# Value

A list of phenotypic variance/heritability estimates and their variances.

h2GE 51

h2GE

Heritability and its variance when there is an environment component

## **Description**

Heritability and its variance when there is an environment component

# Usage

```
h2GE(V, VCOV, verbose = TRUE)
```

## **Arguments**

V Variance estimates.

VCOV Variance-covariance matrix.

verbose Detailed output.

#### Value

A list of phenotypic variance/heritability/GxE interaction esimates and their variances.

h21

Heritability under the liability threshold model

# **Description**

Heritability under the liability threshold model

# Usage

```
h21(K = 0.05, P = 0.5, h2, se, verbose = TRUE)
```

# Arguments

K Disease prevalence.
 P Phenotypeic variance.
 h2 Heritability estimate.
 se Standard error.
 verbose Detailed output.

## Value

A list of the input heritability estimate/standard error and their counterpart under liability threshold model, the normal deviate..

 $h2_mzdz$ 

h2\_mzdz

Heritability estimation according to twin correlations

## **Description**

Heritability estimation according to twin correlations

# Usage

```
h2_mzdz(
  mzDat = NULL,
  dzDat = NULL,
  rmz = NULL,
  rdz = NULL,
  nmz = NULL,
  ndz = NULL,
  selV = NULL
)
```

## **Arguments**

```
mzDat a data frame for monzygotic twins (MZ).

dzDat a data frame for dizygotic twins (DZ).

rmz correlation for MZ twins.

rdz correlation for DZ twins.

nmz sample size for MZ twins.

ndz sample size for DZ twins.

selV names of variables for twin and cotwin.
```

# Details

The example section shows how to obtain bootstrap 95%CI. Heritability and variance estimation according to twin pair correlations.

#### Value

The returned value is a matrix containing heritability and their variance estimations for "h2", "c2", "e2", "vb", "vc", "ve".

## Author(s)

Jing Hua Zhao

## References

Keeping ES. Introduction to Statistical Inference, Dover Pulications, Inc. 1995

hap 53

#### **Examples**

```
## Not run:
ACE_CI <- function(mzData,dzData,n.sim=5,selV=NULL,verbose=TRUE)
ACEr_twinData <- h2(mzDat=mzData,dzDat=dzData,selV=selV)
print(ACEr_twinData)
nmz <- dim(mzData)[1]</pre>
ndz <- dim(dzData)[1]</pre>
a <- ar <- vector()
set.seed(12345)
for(i in 1:n.sim)
  cat("\rRunning # ",i,"/", n.sim,"\r",sep="")
  sampled_mz <- sample(1:nmz, replace=TRUE)</pre>
  sampled_dz <- sample(1:ndz, replace=TRUE)</pre>
  mzDat <- mzData[sampled_mz,]</pre>
  dzDat <- dzData[sampled_dz,]</pre>
  ACEr_i <- h2(mzDat=mzDat,dzDat=dzDat,selV=selV)
  if(verbose) print(ACEr_i)
  ar <- rbind(ar,ACEr_i)</pre>
cat("\n\ner itability according to correlations\n'")
ar <- as.data.frame(ar)</pre>
m <- mean(ar,na.rm=TRUE)</pre>
s <- sd(ar,na.rm=TRUE)</pre>
allr <- data.frame(mean=m,sd=s,lcl=m-1.96*s,ucl=m+1.96*s)
print(allr)
}
selVars <- c('bmi1','bmi2')</pre>
library(mvtnorm)
n.sim <- 500
cat ("\nThe first study\n")
mzm <- as.data.frame(rmvnorm(195, c(22.75,22.75),</pre>
                       matrix(2.66<sup>2</sup>*c(1, 0.67, 0.67, 1), 2)))
dzm <- as.data.frame(rmvnorm(130, c(23.44,23.44),</pre>
                       matrix(2.75<sup>2</sup>*c(1, 0.32, 0.32, 1), 2)))
mzw <- as.data.frame(rmvnorm(384, c(21.44,21.44),</pre>
                       matrix(3.08<sup>2</sup>*c(1, 0.72, 0.72, 1), 2)))
dzw <- as.data.frame(rmvnorm(243, c(21.72,21.72),</pre>
                       matrix(3.12<sup>2</sup>*c(1, 0.33, 0.33, 1), 2)))
names(mzm) \leftarrow names(dzm) \leftarrow names(mzw) \leftarrow names(dzw) \leftarrow c("bmi1", "bmi2")
ACE_CI(mzm,dzm,n.sim,selV=selVars,verbose=FALSE)
ACE_CI(mzw,dzw,n.sim,selV=selVars,verbose=FALSE)
## End(Not run)
```

54 hap

# **Description**

Haplotype reconstruction

## Usage

```
hap(
   id,
   data,
   nloci,
   loci = rep(2, nloci),
   names = paste("loci", 1:nloci, sep = ""),
   control = hap.control()
)
```

## **Arguments**

```
id a column of subject id.

data genotype table.

nloci number of loci.

loci number of alleles at all loci.

names locus names.

control is a call to hap.control().
```

#### **Details**

Haplotype reconstruction using sorting and trimming algorithms.

The package can hande much larger number of multiallelic loci. For large sample size with relatively small number of multiallelic loci, genecounting should be used.

### Value

The returned value is a list containing:

- 11 log-likelihood assuming linkage disequilibrium.
- converge convergence status, 0=failed, 1=succeeded.
- niter number of iterations.

#### Note

adapted from hap.

### References

Clayton DG (2001) SNPHAP. https://github.com/chr1swallace/snphap.

Zhao JH and W Qian (2003) Association analysis of unrelated individuals using polymorphic genetic markers. RSS 2003, Hassalt, Belgium

Zhao JH (2004). "2LD. GENECOUNTING and HAP: computer programs for linkage disequilibrium analysis." *Bioinformatics*, **20**(8), 1325-6. ISSN 1367-4803 (Print) 1367-4803, doi:10.1093/bioinformatics/bth071.

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#### See Also

```
genecounting
```

## **Examples**

```
## Not run:
require(gap.datasets)
# 4 SNP example, to generate hap.out and assign.out alone
data(fsnps)
hap(id=fsnps[,1],data=fsnps[,3:10],nloci=4)
dir()

# to generate results of imputations
control <- hap.control(ss=1,mi=5,hapfile="h",assignfile="a")
hap(id=fsnps[,1],data=fsnps[,3:10],nloci=4,control=control)
dir()

## End(Not run)</pre>
```

hap.control

Control for haplotype reconstruction

# Description

Control for haplotype reconstruction

# Usage

```
hap.control(
  mb = 0,
  pr = 0,
  po = 0.001,
  to = 0.001,
  th = 1,
  maxit = 100,
  n = 0,
  ss = 0,
  rs = 0,
  rp = 0,
  ro = 0,
  rv = 0,
  sd = 0,
  mm = 0,
  mi = 0,
  mc = 50,
  ds = 0.1,
  de = 0,
  q = 0,
  hapfile = "hap.out",
  assignfile = "assign.out"
```

56 hap.em

## **Arguments**

mb	Maximum dynamic storage to be allocated, in Mb.
pr	Prior (ie population) probability threshold.
ро	Posterior probability threshold.
to	Log-likelihood convergence tolerance.
th	Posterior probability threshold for output.
maxit	Maximum EM iteration.
n	Force numeric allele coding (1/2) on output (off).
ss	Tab-delimited speadsheet file output (off).
rs	Random starting points for each EM iteration (off).
rp	Restart from random prior probabilities.
ro	Loci added in random order (off).
rv	Loci added in reverse order (off).
sd	Set seed for random number generator (use date+time).
mm	Repeat final maximization multiple times.
mi	Create multiple imputed datasets. If set >0.
mc	Number of MCMC steps between samples.
ds	Starting value of Dirichlet prior parameter.
de	Finishing value of Dirichlet prior parameter.
q	Quiet operation (off).
hapfile	a file for haplotype frequencies.
assignfile	a file for haplotype assignment.

# Value

A list containing the parameter specifications to the function.

hap.em Gene counting for haplotype analysis

# Description

Gene counting for haplotype analysis

# Usage

```
hap.em(
  id,
  data,
  locus.label = NA,
  converge.eps = 1e-06,
  maxiter = 500,
  miss.val = 0
)
```

hap.em 57

### **Arguments**

id a vector of individual IDs.

data Matrix of alleles, such that each locus has a pair of adjacent columns of alleles,

and the order of columns corresponds to the order of loci on a chromosome. If there are K loci, then ncol(data) = 2\*K. Rows represent alleles for each subject.

locus.label Vector of labels for loci, of length K (see definition of data matrix).

converge.eps Convergence criterion, based on absolute change in log likelihood (Inlike).

maxiter Maximum number of iterations of EM.

miss.val missing value.

#### **Details**

Gene counting for haplotype analysis with missing data, adapted for hap. score.

#### Value

List with components:

- converge Indicator of convergence of the EM algorithm (1=converged, 0 = failed).
- niter Number of iterations completed in the EM alogrithm.
- locus.info A list with a component for each locus. Each component is also a list, and the items of a locus-specific list are the locus name and a vector for the unique alleles for the locus.
- locus.label Vector of labels for loci, of length K (see definition of input values).
- haplotype Matrix of unique haplotypes. Each row represents a unique haplotype, and the number of columns is the number of loci.
- hap.prob Vector of mle's of haplotype probabilities. The ith element of hap.prob corresponds to the ith row of haplotype.
- Inlike Value of Inlike at last EM iteration (maximum Inlike if converged).
- indx.subj Vector for index of subjects, after expanding to all possible pairs of haplotypes for each person. If indx=i, then i is the ith row of input matrix data. If the ith subject has n possible pairs of haplotypes that correspond to their marker phenotype, then i is repeated n times.
- nreps Vector for the count of haplotype pairs that map to each subject's marker genotypes.
- hap1code Vector of codes for each subject's first haplotype. The values in hap1code are the row numbers of the unique haplotypes in the returned matrix haplotype.
- hap2code Similar to hap1code, but for each subject's second haplotype.
- post Vector of posterior probabilities of pairs of haplotypes for a person, given thier marker phenotypes.

# Note

Adapted from HAP.

### Author(s)

Jing Hua Zhao

### See Also

hap, LDkl

hap.score

## **Examples**

```
## Not run:
data(hla)
hap.em(id=1:length(hla[,1]),data=hla[,3:8],locus.label=c("DQR","DQA","DQB"))
## End(Not run)
```

hap.score

Score statistics for association of traits with haplotypes

# Description

Score statistics for association of traits with haplotypes

# Usage

```
hap.score(
  у,
  geno,
  trait.type = "gaussian",
  offset = NA,
  x.adj = NA,
  skip.haplo = 0.005,
  locus.label = NA,
  miss.val = 0,
  n.sim = 0,
  method = "gc",
  id = NA,
  handle.miss = 0,
  mloci = NA,
  sexid = NA
)
```

# Arguments

У	Vector of trait values. For trait.type = "binomial", y must have values of 1 for event, 0 for no event.
geno	Matrix of alleles, such that each locus has a pair of adjacent columns of alleles, and the order of columns corresponds to the order of loci on a chromosome. If there are K loci, then $ncol(geno) = 2*K$ . Rows represent alleles for each subject.
trait.type	Character string defining type of trait, with values of "gaussian", "binomial", "poisson", "ordinal".
offset	Vector of offset when trait.type = "poisson".
x.adj	Matrix of non-genetic covariates used to adjust the score statistics. Note that intercept should not be included, as it will be added in this function.
skip.haplo	Skip score statistics for haplotypes with frequencies < skip.haplo.
locus.label	Vector of labels for loci, of length K (see definition of geno matrix).
miss.val	Vector of codes for missing values of alleles.

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Number of simulations for empirical p-values. If n.sim=0, no empirical p-values are computed.

method method of haplotype frequency estimation, "gc" or "hap".

id an added option which contains the individual IDs.

handle.miss flag to handle missing genotype data, 0=no, 1=yes.

mloci maximum number of loci/sites with missing data to be allowed in the analysis.

sexid flag to indicator sex for data from X chromosome, i=male, 2=female.

#### Details

Compute score statistics to evaluate the association of a trait with haplotypes, when linkage phase is unknown and diploid marker phenotypes are observed among unrelated subjects. For now, only autosomal loci are considered. This package haplo.score which this function is based is greatly acknowledged.

This is a version which substitutes haplo.em.

#### Value

List with the following components:

- score.global Global statistic to test association of trait with haplotypes that have frequencies >= skip.haplo.
- df Degrees of freedom for score.global.
- score.global.p P-value of score.global based on chi-square distribution, with degrees of free-dom equal to df.
- score.global.p.sim P-value of score.global based on simulations (set equal to NA when n.sim=0).
- score.haplo Vector of score statistics for individual haplotypes that have frequencies >= skip.haplo.
- score.haplo.p Vector of p-values for score.haplo, based on a chi-square distribution with 1 df.
- score.haplo.p.sim Vector of p-values for score.haplo, based on simulations (set equal to NA when n.sim=0).
- score.max.p.sim P-value of maximum score.haplo, based on simulations (set equal to NA when n.sim=0).
- haplotype Matrix of hapoltypes analyzed. The ith row of haplotype corresponds to the ith item of score.haplo, score.haplo.p, and score.haplo.p.sim.
- hap.prob Vector of haplotype probabilies, corresponding to the haplotypes in the matrix haplotype.
- locus.label Vector of labels for loci, of length K (same as input argument).
- n.sim Number of simulations.
- n.val.global Number of valid simulated global statistics.
- n.val.haplo Number of valid simulated score statistics (score.haplo) for individual haplotypes.

#### References

Schaid DJ, Rowland CM, Tines DE, Jacobson RM, Poland GA (2002). "Score tests for association between traits and haplotypes when linkage phase is ambiguous." *Am J Hum Genet*, **70**(2), 425-34. ISSN 0002-9297 (Print) 0002-9297, doi:10.1086/338688.

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#### **Examples**

```
## Not run:
data(hla)
y<-hla[,2]
geno<-hla[,3:8]
# complete data
hap.score(y,geno,locus.label=c("DRB","DQA","DQB"))
# incomplete genotype data
\verb|hap.score(y,geno,locus.label=c("DRB","DQA","DQB"), \verb|handle.miss=1,mloci=1|| \\
unlink("assign.dat")
### note the differences in p values in the following runs
data(aldh2)
# to subset the data since hap doesn't handle one allele missing
deleted < -c(40,239,256)
aldh2[deleted,]
aldh2<-aldh2[-deleted,]
y<-aldh2[,2]
geno<-aldh2[,3:18]</pre>
# only one missing locus
hap.score(y,geno,handle.miss=1,mloci=1,method="hap")
# up to seven missing loci and with 10,000 permutations
hap.score(y,geno,handle.miss=1,mloci=7,method="hap",n.sim=10000)
# hap.score takes considerably longer time and does not handle missing data
hap.score(y,geno,n.sim=10000)
## End(Not run)
```

hg18

Chromosomal lengths for build 36

# Description

Data are used in other functions.

# Usage

hg18

## **Format**

A vector containing lengths of chromosomes.

### **Details**

generated from GRCh.R.

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hg19

Chromosomal lengths for build 37

# Description

Data are used in other functions.

# Usage

hg19

## **Format**

A vector containing lengths of chromosomes.

hg38

Chromosomal lengths for build 38

# Description

Data are used in other functions.

# Usage

hg38

# **Format**

A vector containing lengths of chromosomes.

hmht.control

Controls for highlights

# Description

Specification of highlights

# Usage

```
hmht.control(
  data = NULL,
  colors = NULL,
  yoffset = 0.25,
  cex = 1.5,
  boxed = FALSE
```

htr

# **Arguments**

data Data.
colors Colors.
yoffset Y offset.

cex Scaling factor. boxed Label in box.

#### Value

A list as above.

htr

Haplotype trend regression

# **Description**

Haplotype trend regression

## Usage

```
htr(y, x, n.sim = 0)
```

# **Arguments**

y a vector of phenotype.x a haplotype table.

n.sim the number of permutations.

# **Details**

Haplotype trend regression (with permutation)

# Value

The returned value is a list containing:

- f the F statistic for overall association.
- p the p value for overall association.
- fv the F statistics for individual haplotypes.
- pi the p values for individual haplotypes.

### Note

adapted from emgi.cpp, a pseudorandom number seed will be added on.

## Author(s)

Dimitri Zaykin, Jing Hua Zhao

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#### References

Zaykin DV, Westfall PH, Young SS, Karnoub MA, Wagner MJ, Ehm MG (2002). "Testing association of statistically inferred haplotypes with discrete and continuous traits in samples of unrelated individuals." *Hum Hered*, **53**(2), 79-91. ISSN 0001-5652 (Print) 0001-5652, doi:10.1159/000057986.

Xie R, Stram DO (2005). "Asymptotic equivalence between two score tests for haplotype-specific risk in general linear models." *Genetic Epidemiology*, **29**(2), 166-170. ISSN 0741-0395, doi:10.1002/gepi.20087.

#### See Also

hap.score

## **Examples**

```
## Not run:
# 26-10-03
# this is now part of demo
test2<-read.table("test2.dat")</pre>
y<-test2[,1]
x<-test2[,-1]
y<-as.matrix(y)
x<-as.matrix(x)
htr.test2<-htr(y,x)</pre>
htr.test2
htr.test2<-htr(y,x,n.sim=10)</pre>
htr.test2
# 13-11-2003
require(gap.datasets)
data(apoeapoc)
apoeapoc.gc<-gc.em(apoeapoc[,5:8])</pre>
v<-apoeapoc$v
for(i in 1:length(y)) if(y[i]==2) y[i]<-1
htr(y,apoeapoc.gc$htrtable)
# 20-8-2008
# part of the example from useR!2008 tutorial by Andrea Foulkes
# It may be used beyond the generalized linear model (GLM) framework
HaploEM <- haplo.em(Geno,locus.label=SNPnames)</pre>
HapMat <- HapDesign(HaploEM)</pre>
m1 <- lm(Trait~HapMat)
m2 <- lm(Trait^1)
anova(m2,m1)
## End(Not run)
```

hwe

Hardy-Weinberg equlibrium test for a multiallelic marker

### **Description**

Hardy-Weinberg equlibrium test for a multiallelic marker

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### Usage

```
hwe(data, data.type = "allele", yates.correct = FALSE, miss.val = 0)
```

#### **Arguments**

data A rectangular data containing the genotype, or an array of genotype counts.

data. type An option taking values "allele", "genotype", "count" if data is alleles, genotype

or genotype count.

yates.correct A flag indicating if Yates' correction is used for Pearson  $\chi^2$  statistic.

miss.val A list of missing values.

#### **Details**

Hardy-Weinberg equilibrium test.

This function obtains Hardy-Weinberg equilibrium test statistics. It can handle data coded as allele numbers (default), genotype identifiers (by setting data.type="genotype") and counts corresponding to individual genotypes (by setting data.type="count") which requires that genotype counts for all n(n+1) possible genotypes, with n being the number of alleles.

For highly polymorphic markers when asymptotic results do not hold, please resort to hwe.hardy.

#### Value

The returned value is a list containing:

- allele.freq Frequencies of alleles.
- x2 Pearson  $\chi^2$ .
- p.x2 p value for  $\chi^2$ .
- lrt Log-likelihood ratio test statistic.
- p.lrt p value for lrt.
- df Degree(s) of freedom.
- rho  $\sqrt{\chi^2/N}$  the contingency table coefficient.

#### Author(s)

Jing Hua Zhao

# See Also

hwe.hardy

# **Examples**

```
## Not run:
a <- c(3,2,2)
a.out <- hwe(a,data.type="genotype")
a.out
a.out <- hwe(a,data.type="count")
a.out
require(haplo.stats)
data(hla)
hla.DQR <- hwe(hla[,3:4])</pre>
```

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```
summary(hla.DQR)
# multiple markers
s <- vector()
for(i in seq(3,8,2))
{
   hwe_i <- hwe(hla[,i:(i+1)])
   s <- rbind(s,hwe_i)
}
s</pre>
## End(Not run)
```

hwe.cc

A likelihood ratio test of population Hardy-Weinberg equilibrium for case-control studies

### **Description**

A likelihood ratio test of population Hardy-Weinberg equilibrium for case-control studies

#### Usage

```
hwe.cc(model, case, ctrl, k0, initial1, initial2)
```

#### **Arguments**

model	model specification, dominant, recessive.
case	a vector of genotype counts in cases.
ctrl	a vector of genotype counts in controls.
k0	prevalence of disease in the population.
initial1	initial values for beta, gamma, and q.
initial2	initial values for logit(p) and log(gamma).

#### **Details**

A likelihood ratio test of population Hardy-Weinberg equilibrium for case-control studies

This is a collection of utility functions. The null hypothesis declares that the proportions of genotypes are according to Hardy-Weinberg law, while under the alternative hypothesis, the expected genotype counts are according to the probabilities that particular genotypes are obtained conditional on the prevalence of disease in the population. In so doing, Hardy-Weinberg equilibrium is considered using both case and control samples but pending on the disease model such that 2-parameter multiplicative model is built on baseline genotype  $\alpha$ ,  $\alpha\beta$  and  $\alpha\gamma$ .

#### Value

The returned value is a list with the following components.

- Cox statistics under a general model.
- t2par under the null hypothesis.
- t3par under the alternative hypothesis.
- lrt.stat the log-likelihood ratio statistic.
- pval the corresponding p value.

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#### Author(s)

Chang Yu, Li Wang, Jing Hua Zhao

#### References

Yu C, Zhang S, Zhou C, Sile S (2009). "A likelihood ratio test of population Hardy-Weinberg equilibrium for case-control studies." *Genet Epidemiol*, **33**(3), 275-80. ISSN 0741-0395 (Print) 0741-0395, doi:10.1002/gepi.20381.

#### See Also

hwe

## **Examples**

```
## Not run:
### Saba Sile, email of Jan 26, 2007, data always in order of GG AG AA, p=Pr(G),
### q=1-p=Pr(A)
case=c(155,27,4)
ctrl=c(408,55,15)
k0=.2
initial1=c(1.0,0.94,0.0904)
initial2=c(logit(1-0.0904),log(0.94))
hwe.cc("recessive",case,ctrl,k0, initial1, initial2)
### John Phillips III, TGFb1 data codon 10: TT CT CC, CC is abnormal and increasing
### TGFb1 activity
case=c(29,78,13)
ctrl=c(17,28,6)
k0 <- 1e-5
initial1 <- c(2.45, 2.45, 0.34)
initial2 <- c(logit(1-0.34), log(2.45))
hwe.cc("dominant",case,ctrl,k0,initial1,initial2)
## End(Not run)
```

hwe.hardy

Hardy-Weinberg equilibrium test using MCMC

# Description

Hardy-Weinberg equilibrium test using MCMC

# Usage

```
hwe.hardy(a, alleles = 3, seed = 3000, sample = c(1000, 1000, 5000))
```

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## **Arguments**

a an array containing the genotype counts, as integer.

alleles number of allele at the locus, greater than or equal to 3, as integer.

seed pseudo-random number seed, as integer.

sample optional, parameters for MCMC containing number of chunks, size of a chunk

and burn-in steps, as integer.

#### **Details**

Hardy-Weinberg equilibrium test by MCMC

## Value

The returned value is a list containing:

- method Hardy-Weinberg equilibrium test using MCMC.
- data.name name of used data if x is given.
- p.value Monte Carlo p value.
- p.value.se standard error of Monte Carlo p value.
- switches percentage of switches (partial, full and altogether).

### Note

Codes are commented for taking x a genotype object, as genotype to prepare a and alleles on the fly.

Adapted from HARDY, testable with -Dexecutable as standalone program.

keywords htest

### Author(s)

Sun-Wei Guo, Jing Hua Zhao, Gregor Gorjanc

#### **Source**

https://sites.stat.washington.edu/thompson/Genepi/pangaea.shtml

## References

Guo SW, Thompson EA (1992). "Performing the exact test of Hardy-Weinberg proportion for multiple alleles." *Biometrics*, **48**(2), 361-72. ISSN 0006-341X (Print) 0006-341x.

#### See Also

hwe, genetics::HWE.test, genetics::genotype

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### **Examples**

```
## Not run:
 # example 2 from hwe.doc:
   a<-c(
   3,
   4, 2,
   2, 2, 2,
   3, 3, 2, 1,
   0, 1, 0, 0, 0,
   0, 0, 0, 0, 0, 1,
   0, 0, 1, 0, 0, 0, 0,
   0, 0, 0, 2, 1, 0, 0, 0)
   ex2 <- hwe.hardy(a=a,alleles=8)
   # example using HLA
   data(hla)
   x <- hla[,3:4]
   y <- pgc(x,handle.miss=0,with.id=1)</pre>
   n.alleles <- max(x,na.rm=TRUE)</pre>
   z <- vector("numeric",n.alleles*(n.alleles+1)/2)</pre>
   z[y$idsave] <- y$wt</pre>
   hwe.hardy(a=z,alleles=n.alleles)
   # with use of class 'genotype'
   # this is to be fixed
   library(genetics)
   hlagen <- genotype(a1=x$DQR.a1, a2=x$DQR.a2,</pre>
                       alleles=sort(unique(c(x$DQR.a1, x$DQR.a2))))
   hwe.hardy(hlagen)
   # comparison with hwe
   hwe(z,data.type="count")
   # to create input file for HARDY
   print.tri<-function (xx,n) {</pre>
       cat(n,"\n")
       for(i in 1:n) {
           for(j in 1:i) {
               cat(xx[i,j]," ")
           }
       cat("\n")
       }
       cat("100 170 1000\n")
   xx<-matrix(0,n.alleles,n.alleles)</pre>
   xxx<-lower.tri(xx,diag=TRUE)</pre>
   xx[xxx]<-z
   sink("z.dat")
   print.tri(xx,n.alleles)
   sink()
   # now call as: hwe z.dat z.out
## End(Not run)
```

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hwe.jags

Hardy-Weinberg equlibrium test for a multiallelic marker using JAGS

#### **Description**

Hardy-Weinberg equlibrium test for a multiallelic marker using JAGS

### Usage

```
hwe.jags(
    k,
    n,
    delta = rep(1/k, k),
    lambda = 0,
    lambdamu = -1,
    lambdasd = 1,
    parms = c("p", "f", "q", "theta", "lambda"),
    ...
)
```

## **Arguments**

#### **Details**

Hardy-Weinberg equilibrium test.

This function performs Bayesian Hardy-Weinberg equilibrium test, which mirrors hwe.hardy, another implementation for highly polymorphic markers when asymptotic results do not hold.

### Value

The returned value is a fitted model from jags().

# Author(s)

Jing Hua Zhao, Jon Wakefield

### References

Wakefield J (2010). "Bayesian methods for examining Hardy-Weinberg equilibrium." *Biometrics*, **66**(1), 257-65. ISSN 0006-341X (Print) 0006-341x, doi:10.1111/j.15410420.2009.01267.x.

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#### See Also

hwe.hardy

# **Examples**

```
## Not run:
ex1 <- hwe.jags(4,c(5,6,1,7,11,2,8,19,26,15))
print(ex1)
ex2 <- hwe.jags(2,c(49,45,6))
print(ex2)
ex3 <- hwe.jags(4,c(0,3,1,5,18,1,3,7,5,2),lambda=0.5,lambdamu=-2.95,lambdasd=1.07)
print(ex3)
ex4 <- hwe.jags(9,c(1236,120,3,18,0,0,982,55,7,249,32,1,0,12,0,2582,132,20,1162,29,
                      1312,6,0,0,4,0,4,0,2,0,0,0,0,0,0,115,5,2,53,1,149,0,0,4),
                 delta=c(1,1,1,1,1,1,1,1,1),lambdamu=-4.65,lambdasd=0.21)
print(ex4)
ex5 <- hwe.jags(8,n=c(
         3,
         4, 2,
         2, 2, 2,
         3, 3, 2, 1,
         0, 1, 0, 0, 0,
         0, 0, 0, 0, 0, 1,
         0, 0, 1, 0, 0, 0, 0,
         0, 0, 0, 2, 1, 0, 0, 0))
print(ex5)
\mbox{\tt\#} Data and code accordining to the following URL,
# http://darwin.eeb.uconn.edu/eeb348-notes/testing-hardy-weinberg.pdf
hwe.jags.ABO <- function(n,...)</pre>
{
  hwe <- function() {</pre>
     # likelihood
     pi[1] <- p.a*p.a + 2*p.a*p.o
     pi[2] <- 2*p.a*p.b
     pi[3] \leftarrow p.b*p.b + 2*p.b*p.o
     pi[4] <- p.o*p.o
     n[1:4] \sim dmulti(pi[],N)
     # priors
     a1 \sim dexp(1)
     b1 \sim dexp(1)
     o1 \sim dexp(1)
     p.a <- a1/(a1 + b1 + o1)
     p.b \leftarrow b1/(a1 + b1 + o1)
     p.o <- o1/(a1 + b1 + o1)
  hwd <- function() {</pre>
     # likelihood
     pi[1] \leftarrow p.a*p.a + f*p.a*(1-p.a) + 2*p.a*p.o*(1-f)
     pi[2] \leftarrow 2*p.a*p.b*(1-f)
     pi[3] \leftarrow p.b*p.b + f*p.b*(1-p.b) + 2*p.b*p.o*(1-f)
     pi[4] \leftarrow p.o*p.o + f*p.o*(1-p.o)
     n[1:4] \sim dmulti(pi[],N)
     # priors
     a1 \sim dexp(1)
     b1 \sim dexp(1)
```

invnormal 71

```
o1 ~ dexp(1)
    p.a <- a1/(a1 + b1 + o1)
    p.b <- b1/(a1 + b1 + o1)
    p.o <- o1/(a1 + b1 + o1)
    f ~ dunif(0,1)
}

N <- sum(n)
ABO.hwe <- R2jags::jags(list(n=n,N=N),,c("pi","p.a","p.b","p.o"),hwe,...)
ABO.hwd <- R2jags::jags(list(n=n,N=N),,c("pi","p.a","p.b","p.o","f"),hwd,...)
invisible(list(hwe=ABO.hwe,hwd=ABO.hwd))
}
hwe.jags.ABO.results <- hwe.jags.ABO(n=c(862, 131, 365, 702))
hwe.jags.ABO.results</pre>
## End(Not run)
```

invnormal

Inverse normal transformation

## **Description**

Inverse normal transformation

## Usage

invnormal(x)

# Arguments

Х

Data with missing values.

### Value

Transformed value.

# **Examples**

```
x <- 1:10
z <- invnormal(x)
plot(z,x,type="b")</pre>
```

72

inv\_chr\_pos\_a1\_a2

Retrieval of chr:pos+a1/a2 according to SNP id

## **Description**

This function obtains information embedded in unique identifiers.

## Usage

```
inv_chr_pos_a1_a2(chr_pos_a1_a2, prefix = "chr", seps = c(":", "_", "_"))
```

## **Arguments**

```
chr_pos_a1_a2 SNP id.
```

prefix Prefix of the identifier.
seps Delimiters of fields.

#### Value

A data.frame with the following variables:

- · chr Chromosome.
- pos Position.
- a1 Allele 1.
- a2 Allele 2.

# **Examples**

```
# rs12075
inv_chr_pos_a1_a2("chr1:159175354_A_G",prefix="chr",seps=c(":","_","_"))
```

ixy

Conversion of chrosome name from strings

# Description

This function converts 1:22, X, Y back to 1:24.

# Usage

ixy(x)

# Arguments

Х

Chromosome name in strings

## Value

As indicated.

KCC 73

KCC Disease prevalences in cases and controls
---

## Description

Disease prevalences in cases and controls

## Usage

```
KCC(model, GRR, p1, K)
```

### **Arguments**

model disease model (one of "multiplicative", "additive", "recessive", "dominant", "overdominant").

GRR genotype relative risk.

p1 disease allele frequency.

K disease prevalence in the whole population.

## **Details**

KCC calculates disease prevalences in cases and controls for a given genotype relative risk, allele frequency and prevalencen of the disease in the whole population. It is used by tscc and pbsize2.

### Value

A list of two elements:

- pprime prevlence in cases.
- p prevalence in controls.

kin.morgan	kinship matrix for simple pedigree	
------------	------------------------------------	--

## **Description**

kinship matrix for simple pedigree

## Usage

```
kin.morgan(ped, verbose = FALSE)
```

# Arguments

ped individual's id, father's id and mother's id.
verbose an option to print out the original pedigree.

## **Details**

kinship matrix according to Morgan v2.1.

74 kin.morgan

## Value

The returned value is a list containing:

- kin the kinship matrix in vector form.
- kin.matrix the kinship matrix.

## Note

The input data is required to be sorted so that parents preceed their children.

## Author(s)

Morgan development team, Jing Hua Zhao

## References

```
Morgan\ V2.1\ https://sites.stat.washington.edu/thompson/Genepi/MORGAN/Morgan.shtml
```

## See Also

gif

```
## Not run:
# Werner syndrome pedigree
werner<-c(
1, 0, 0, 1,
2, 0, 0, 2,
3, 0, 0, 2,
4, 1, 2, 1,
5, 0, 0, 1,
6, 1, 2, 2,
 7, 1, 2, 2,
8, 0, 0, 1,
9, 4, 3,
           2,
10, 5,
11, 5, 6, 2,
12, 8, 7, 1,
13,10, 9, 2,
14,12, 11, 1,
15,14, 13, 1)
werner<-t(matrix(werner,nrow=4))</pre>
kin.morgan(werner[,1:3])
## End(Not run)
```

klem 75

klem	Haplotype frequency estimation based on a genotype table of two multiallelic markers

## Description

Haplotype frequency estimation using expectation-maximization algorithm based on a table of genotypes of two multiallelic markers.

## Usage

```
klem(obs, k = 2, 1 = 2)
```

## Arguments

obs a table of genotype counts.

k number of alleles at marker 1.

1 number of alleles at marker 2.

The dimension of the genotype table should be k\*(k+1)/2 x l\*(l+1)/2.

Modified from 2ld.c.

## Value

The returned value is a list containing:

- h haplotype Frequencies.
- 10 log-likelihood under linkage equilibrium.
- 11 log-likelihood under linkage disequilibrium.

## Author(s)

Jing Hua Zhao

# See Also

```
genecounting
```

```
## Not run:
# an example with known genotype counts
z <- klem(obs=1:9)
# an example with imputed genotypes at SH2B1
source(file.path(find.package("gap"),"scripts","SH2B1.R"),echo=TRUE)
## End(Not run)</pre>
```

76 labelManhattan

labelManhattan

Annotate Manhattan or Miami Plot

# Description

Annotate Manhattan or Miami Plot

# Usage

```
labelManhattan(
  chr,
  pos,
  name,
  gwas,
  gwasChrLab = "chr",
  gwasPosLab = "pos",
  gwasPLab = "p",
  gwasZLab = "NULL",
  chrmaxpos,
  textPos = 4,
  angle = 0,
  miamiBottom = FALSE
)
```

# Arguments

 ${\tt miamiBottom}$ 

chr	A vector of chromosomes for the markers to be labelled.
pos	A vector of positions on the chromosome for the markers to be labelled. These must correspond to markers in the GWAS dataset used to make the manhattan plot.
name	A vector of labels to be added next to the points specified by chr and pos.
gwas	The same GWAS dataset used to plot the existing Manhattan or Miami plot to be annotated.
gwasChrLab	The name of the column in gwas containing chromosome number. Defaults to ""chr".
gwasPosLab	The name of the column in gwas containing position. Defaults to "pos".
gwasPLab	The name of the column in gwas containing p-values. Defaults to "p".
gwasZLab	The name of the column in gwas containing z-values. Defaults to "NULL".
chrmaxpos	Data frame containing $x$ coordinates for chromosome start positions, generated by labelManhattan.
textPos	An integer or vector dictating where the label should be plotted relative to each point. Good for avoiding overlapping labels. Provide an integer to plot all points in the same relative position or use a vector to specify position for each label. Passed to the 'pos' option of graphics::text. Defaults to '4'.
angle	An integer or vector dictating the plot angle of the label for each point. rovide an integer to plot all points in the same relative position or use a vector to specify position for each label. Passed to the 'srt' option of graphics::text. Defaults to '0'.

If 'TRUE', labels will be plotted on the lower region of a Miami plot. If 'FALSE',

labels will be plotted on the upper region. Defaults to 'FALSE'.

LD22 77

#### **Details**

Add labels beside specified points on a Manhattan or Miami plot. Ideal for adding locus names to peaks. Currently only designed to work with miamiplot2.

### Value

Adds annotation to existing Manhattan or Miami plot

#### Note

Extended to handle extreme P values.

#### Author(s)

Jonathan Marten

## **Examples**

LD22

LD statistics for two diallelic markers

## **Description**

LD statistics for two diallelic markers

# Usage

```
LD22(h, n)
```

# Arguments

h a vector of haplotype frequencies.

n number of haplotypes.

# Details

It is possible to perform permutation test of  $r^2$  by re-ordering the genotype through R's sample function, obtaining the haplotype frequencies by gc.em or genecounting, supplying the estimated haplotype frequencies to the current function and record x2, and comparing the observed x2 and that from the replicates.

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#### Value

The returned value is a list containing:

- h the original haplotype frequency vector.
- n the number of haplotypes.
- D the linkage disequilibrium parameter.
- VarD the variance of D.
- Dmax the maximum of D.
- VarDmax the variance of Dmax.
- Dprime the scaled disequilibrium parameter.
- VarDprime the variance of Dprime.
- x2 the Chi-squared statistic.
- lor the log(OR) statistic.
- vlor the var(log(OR)) statistic.

### Note

extracted from 2ld.c, worked 28/6/03, tables are symmetric do not fix, see kbyl below

## Author(s)

Jing Hua Zhao

#### References

Zabetian CP, Buxbaum SG, Elston RC, Köhnke MD, Anderson GM, Gelernter J, Cubells JF (2003). "The structure of linkage disequilibrium at the DBH locus strongly influences the magnitude of association between diallelic markers and plasma dopamine beta-hydroxylase activity." *Am J Hum Genet*, **72**(6), 1389-400. ISSN 0002-9297 (Print) 0002-9297, doi:10.1086/375499.

Zapata C, Alvarez G, Carollo C (1997). "Approximate variance of the standardized measure of gametic disequilibrium D'." *Am J Hum Genet*, **61**(3), 771-4. ISSN 0002-9297 (Print) 0002-9297, doi:10.1016/s00029297(07)643420.

## See Also

LDk1

```
## Not run:
h <- c(0.442356,0.291532,0.245794,0.020319)
n <- 481*2
t <- LD22(h,n)
## End(Not run)</pre>
```

LDkl 79

LDk1

LD statistics for two multiallelic markers

## **Description**

LD statistics for two multiallelic loci. For two diallelic makers, the familiar  $r^2$  has standard error seX2.

# Usage

```
LDkl(n1 = 2, n2 = 2, h, n, optrho = 2, verbose = FALSE)
```

## **Arguments**

n1 number of alleles at marker 1.
 n2 number of alleles at marker 2.
 h a vector of haplotype frequencies.

n number of haplotypes.

optrho type of contingency table association, 0=Pearson, 1=Tschuprow, 2=Cramer (de-

fault).

verbose detailed output of individual statistics.

#### Value

The returned value is a list containing:

- n1 the number of alleles at marker 1.
- n2 the number of alleles at marker 2.
- h the haplotype frequency vector.
- n the number of haplotypes.
- Dp D'.
- VarDp variance of D'.
- Dijtable table of Dij.
- VarDijtable table of variances for Dij.
- Dmaxtable table of Dmax.
- Dijptable table of Dij'.
- VarDijptable table of variances for Dij'.
- X2table table of Chi-squares (based on Dij).
- ptable table of p values.
- x2 the Chi-squared statistic.
- seX2 the standard error of x2/n.
- rho the measure of association.
- seR the standard error of rho.
- optrho the method for calculating rho.
- klinfo the Kullback-Leibler information.

LDkl

#### Note

adapted from 2ld.c.

## Author(s)

Jing Hua Zhao

### References

Bishop YMM, Fienberg SE, Holland PW (1975) Discrete Multivariate Analysis – Theory and Practice, The MIT press

Cramer H (1946) Mathematical Methods of Statistics. Princeton Univ. Press

Zapata C, Carollo C, Rodriguez S (2001). "Sampling variance and distribution of the D' measure of overall gametic disequilibrium between multiallelic loci." *Ann Hum Genet*, **65**(Pt 4), 395-406. ISSN 0003-4800 (Print) 0003-4800, doi:10.1017/s0003480001008697.

Zhao JH (2004). "2LD. GENECOUNTING and HAP: computer programs for linkage disequilibrium analysis." *Bioinformatics*, **20**(8), 1325-6. ISSN 1367-4803 (Print) 1367-4803, doi:10.1093/bioinformatics/bth071.

## See Also

LD22

```
## Not run:
# two examples in the C program 2LD:
# two SNPs as in 2by2.dat
# this can be compared with output from LD22

h <- c(0.442356,0.291532,0.245794,0.020319)
n <- 481*2
t <- LDk1(2,2,h,n)
t

# two multiallelic markers as in kbyl.dat
# the two-locus haplotype vector is in file "kbyl.dat"
# The data is now available from 2ld in Haplotype-Analysis

filespec <- system.file("kbyl.dat")
h <- scan(filespec,skip=1)
t <- LDk1(9,5,h,213*2,verbose=TRUE)

## End(Not run)</pre>
```

log10p

log10p

log10(p) for a normal deviate z

# Description

log10(p) for a normal deviate z

## Usage

log10p(z)

## Arguments

Z

normal deviate.

## Value

log10(P)

# Author(s)

James Peters

# **Examples**

log10p(100)

log10pvalue

log10(p) for a P value including its scientific format

# Description

log10(p) for a P value including its scientific format

## Usage

```
log10pvalue(p = NULL, base = NULL, exponent = NULL)
```

# Arguments

p value.

base base part in scientific format.

exponent exponent part in scientific format.

# Value

log10(P)

```
log10pvalue(1e-323)
log10pvalue(base=1,exponent=-323)
```

82 makeped

logp

log(p) for a normal deviate z

## **Description**

log(p) for a normal deviate z

## Usage

logp(z)

## **Arguments**

Z

normal deviate.

## Value

log(P)

## **Examples**

logp(100)

makeped

A function to prepare pedigrees in post-MAKEPED format

## Description

A function to prepare pedigrees in post-MAKEPED format

## Usage

```
makeped(
  pifile = "pedfile.pre",
  pofile = "pedfile.ped",
  auto.select = 1,
  with.loop = 0,
  loop.file = NA,
  auto.proband = 1,
  proband.file = NA
)
```

## **Arguments**

pifile input filename.
pofile output filename.

 $\hbox{auto.select} \qquad \hbox{no loops in pedigrees and probands are selected automatically? 0=no, 1=yes.}$ 

with.loop input data with loops? 0=no, 1=yes.

loop.file filename containing pedigree id and an individual id for each loop, set if with.loop=1.

```
auto.proband probands are selected automatically? 0=no, 1=yes.

proband.file filename containing pedigree id and proband id, set if auto.proband=0 (not implemented).
```

#### Details

Many computer programs for genetic data analysis requires pedigree data to be in the so-called "post-MAKEPED" format. This function performs this translation and allows for some inconsistences to be detected.

The first four columns of the input file contains the following information:

```
pedigree ID, individual ID, father's ID, mother's ID, sex
```

Either father's or mother's id is set to 0 for founders, i.e. individuals with no parents. Numeric coding for sex is 0=unknown, 1=male, 2=female. These can be followed by satellite information such as disease phenotype and marker information.

The output file has extra information extracted from data above.

Before invoking makeped, input file, loop file and proband file have to be prepared.

By default, auto.select=1, so translation proceeds without considering loops and proband statuses. If there are loops in the pedigrees, then set auto.select=0, with.loop=1, loop.file="filespec".

There may be several versions of makeped available, but their differences with this port should be minor.

## Note

adapted from makeped.c by W Li and others. keywords datagen

### **Examples**

```
## Not run:
cwd <- getwd()
cs.dir <- file.path(find.package("gap.examples"),"tests","kinship")
setwd(cs.dir)
dir()
makeped("ped7.pre","ped7.ped",0,1,"ped7.lop")
setwd(cwd)
# https://lab.rockefeller.edu/ott/
## End(Not run)</pre>
```

masize

Sample size calculation for mediation analysis

## **Description**

The function computes sample size for regression problems where the goal is to assess mediation of the effects of a primary predictor by an intermediate variable or mediator.

### Usage

```
masize(model, opts, alpha = 0.025, gamma = 0.2)
```

#### **Arguments**

model "lineari", "logisticj", "poissonk", "coxl", where i,j,k,l range from 1 to 4,5,9,9,

respectively.

opts A list specific to the model

b1 regression coefficient for the primary predictor X1

b2 regression coefficient for the mediator X2

rho correlation between X1 and X2

sdx1, sdx2 standard deviations (SDs) of X1 and X2

f1, f2 prevalence of binary X1 and X2

sdy residual SD of the outcome for the linear model

p marginal prevalence of the binary outcome in the logistic model

m marginal mean of the count outcome in a Poisson model f proportion of uncensored observations for the Cox model

fc proportion of observations censored early

alpha one-sided type-I error rate

gamma type-II error rate

ns number of observations to be simulated

seed random number seed

For linear model, the arguments are b2, rho, sdx2, sdy, alpha, and gamma. For cases CpBm and BpBm, set  $sdx2 = \sqrt{f2(1-f2)}$ . Three alternative functions are included for the linear model. These functions make it possible to supply other combinations of input parameters affecting mediation:

b1\* coefficient for the primary predictor

in the reduced model excluding the mediator (b1star)

b1 coefficient for the primary predictor

in the full model including the mediator

PTE proportion of the effect of the primary predictor

explained by the mediator, defined as (b1\*-b1)/b1\*

These alternative functions for the linear model require specification of an extra parameter, but are provided for convenience, along with two utility files for computing PTE and b1\* from the other parameters. The required arguments are explained in comments within the R code.

alpha Type-I error rate, one-sided.

gamma Type-II error rate.

#### **Details**

Mediation has been thought of in terms of the proportion of effect explained, or the relative attenuation of b1, the coefficient for the primary predictor X1, when the mediator, X2, is added to the model. The goal is to show that b1\*, the coefficient for X1 in the reduced model (i.e., the model with only X1, differs from b1, its coefficient in the full model (i.e., the model with both X1 and the mediator X2. If X1 and X2 are correlated, then showing that b2, the coefficient for X2, differs from zero is equivalent to showing b1\* differs from b1. Thus the problem reduces to detecting an effect of X2, controlling for X1. In short, it amounts to the more familiar problem of inflating sample size to account for loss of precision due to adjustment for X1.

The approach here is to approximate the expected information matrix from the regression model including both X1 and X2, to obtain the expected standard error of the estimate of b2, evaluated at the MLE. The sample size follows from comparing the Wald test statistic (i.e., the ratio of the estimate of b2 to its SE) to the standard normal distribution, with the expected value of the numerator and denominator of the statistic computed under the alternative hypothesis. This reflects the Wald test for the statistical significance of a coefficient implemented in most regression packages.

The function provides methods to calculate sample sizes for the mediation problem for linear, logistic, Poisson, and Cox regression models in four cases for each model:

CpCm continuous primary predictor, continuous mediator
BpCm binary primary predictor, continuous mediator
CpBm continuous primary predictor, binary mediator
BpBm binary primary predictor, binary mediator

The function is also generally applicable to the analogous problem of calculating sample size adequate to detect the effect of a primary predictor in the presence of confounding. Simply treat X2 as the primary predictor and consider X1 the confounder.

For linear model, a single function, linear, implements the analytic solution for all four cases, based on Hsieh et al., is to inflate sample size by a variance inflation factor,  $1/(1-rho^2)$ , where rho is the correlation of X1 and X2. This also turns out to be the analytic solution in cases CpCm and BpCm for the Poisson model, and underlies approximate solutions for the logistic and Cox models. An analytic solution is also given for cases CpBm and BpBm for the Poisson model. Since analytic solutions are not available for the logistic and Cox models, a simulation approach is used to obtain the expected information matrix instead.

For logistic model, the approximate solution due to Hsieh is implemented in the function logistic approx, and can be used for all four cases. Arguments are p, b2, rho, sdx2, alpha, and gamma. For a binary mediator with prevalence f2, sdx2 should be reset to  $\sqrt{f2(1-f2)}$ . Simulating the information matrix of the logistic model provides somewhat more accurate sample size estimates than the Hsieh approximation. The functions for cases CpCm, BpCm, CpBm, and BpBm are respectively logistic.ccs, logistic.bcs, logistic.cbs, and logistic.bbs, as for the Poisson and Cox models. Arguments for these functions include p, b1, sdx1 or f1, b2, sdx2 or f2, rho, alpha, gamma, and ns. As in other functions, sdx1, sdx2, alpha, and gamma are set to the defaults listed above. These four functions call two utility functions, getb0 (to calculate the intercept parameter from the others) and antilogit, which are supplied.

For Poisson model, The function implementing the approximate solution based on the variance inflation factor is poisson.approx, and can be used for all four cases. Arguments are EY (the marginal mean of the Poisson outcome), b2, sdx2, rho, alpha and gamma, with sdx2, alpha and gamma set to the usual defaults; use  $sdx2=\sqrt{f2(1-f2)}$  for a binary mediator with prevalence f2 (cases CpBm and BpBm). For cases CpCm and BpCm (continuous mediators), the approximate formula is also the analytic solution. For these cases, we supply redundant functions poisson.cc and poisson.bc, with the same arguments and defaults as for poisson.approx (it's the same function). For the two cases with binary mediators, the functions are poisson.cb and poisson.bb. In addition to m, b2, f2, rho, alpha, and gamma, b1 and sdx1 or f1 must be specified. Defaults are as usual. Functions using simulation for the Poisson model are available: poisson.ccs, poisson.bcs, poisson.cbs, and poisson.bbs. As in the logistic case, these require arguments b1 and sdx1 or f1. For this case, however, the analytic functions are faster, avoid simulation error, and should be used. We include these functions as templates that could be adapted to other joint predictor distributions. For Cox model, the function implementing the approximate solution, using the variance inflation factor and derived by Schmoor et al., is cox.approx, and can be used for all four cases. Arguments are b2, sdx2, rho, alpha, gamma, and f. For binary X2 set  $sdx2 = \sqrt{f2(1-f2)}$ . The approximation works very well for cases CpCm and BpCm (continuous mediators), but is a bit less accurate for cases CpBm

and BpBm (binary mediators). We get some improvement for those cases using the simulation approach. This approach is implemented for all four, as functions cox.ccs, cox.bcs, cox.cbs, and cox.bbs. Arguments are b1, sdx1 or f1, b2, sdx2 or f2, rho, alpha, gamma, f, and ns, with defaults as described above. Slight variants of these functions, cox.ccs2, cox.bcs2, cox.cbs2, and cox.bbs2, make it possible to allow for early censoring of a fraction fc of observations; but in our experience this has virtually no effect, even with values of fc of 0.5. The default for fc is 0.

A summary of the argumentss is as follows, noting that additional parameter seed can be supplied for simulation-based method.

model	arguments	description
linear1 linear2 linear3 linear4	b2, rho, sdx2, sdy b1star, PTE, rho, sdx1, sdy b1star, b2, PTE, sdx1, sdx2, sdy b1star, b1, b2, sdx1, sdx2, sdy	linear lineara linearb linearc
logistic1 logistic2 logistic3 logistic4 logistic5	p, b2, rho, sdx2 p, b1, b2, rho, sdx1, sdx2, ns p, b1, f1, b2, rho, sdx2, ns p, b1, b2, f2, rho, sdx1, ns p, b1, f1, b2, f2, rho, ns	logistic.approx logistic.ccs logistic.bcs logistic.cbs logistic.bbs
poisson1 poisson2 poisson3 poisson4 poisson5 poisson6 poisson7 poisson8 poisson9	m, b2, rho, sdx2 m, b2, rho, sdx2 m, b2, rho, sdx2 m, b1, b2, f2, rho, sdx1 m, b1, f1, b2, f2, rho m, b1, b2, rho, sdx1, sdx2, ns m, b1, f1, b2, rho, sdx2, ns m, b1, b2, f2, rho, sdx1, ns m, b1, b2, f2, rho, ns	poisson.approx poisson.cc poisson.cb poisson.cb poisson.ccs poisson.bcs poisson.cbs poisson.bs
cox1 cox2 cox3 cox4 cox5 cox6 cox7 cox8	b2, rho, f, sdx2 b1, b2, rho, f, sdx1, sdx2, ns b1, f1, b2, rho, f, sdx2, ns b1, b2, f2, rho, f, sdx1, ns b1, b2, f2, rho, f, ns b1, b2, rho, f, fc, sdx1, sdx2, ns b1, f1, b2, rho, f, fc, sdx2, ns b1, b2, f2, rho, f, fc, sdx1, ns b1, b2, f2, rho, f, fc, ns	cox.approx cox.ccs cox.bcs cox.cbs cox.bbs cox.ccs2 cox.bcs2 cox.cbs2

#### Value

A short description of model (desc, b=binary, c=continuous, s=simulation) and sample size (n). In the case of Cox model, number of events (d) is also indicated.

# References

Hsieh FY, Bloch DA, Larsen MD (1998). "A simple method of sample size calculation for linear and logistic regression." *Stat Med*, **17**(14), 1623-34. ISSN 0277-6715 (Print) 0277-6715, doi:10.1002/(sici)10970258(19980730)17:14<1623::aidsim871>3.0.co;2s.

Schmoor C, Sauerbrei W, Schumacher M (2000). "Sample size considerations for the evaluation of prognostic factors in survival analysis." *Stat Med*, **19**(4), 441-52. ISSN 0277-6715 (Print) 0277-6715, doi:10.1002/(sici)10970258(20000229)19:4<441::aidsim349>3.0.co;2n.

Vittinghoff E, Sen S, McCulloch CE (2009). "Sample size calculations for evaluating mediation." *Stat Med*, **28**(4), 541-57. ISSN 0277-6715 (Print) 0277-6715, doi:10.1002/sim.3491.

#### See Also

ab

```
## Not run:
## linear model
# CpCm
opts <- list(b2=0.5, rho=0.3, sdx2=1, sdy=1)
masize("linear1",opts)
opts <- list(b2=0.75, rho=0.3, f2=0.25, sdx2=sqrt(0.25*0.75), sdy=3)
masize("linear1",opts,gamma=0.1)
## logistic model
# CpBm
opts <- list(p=0.25, b2=log(0.5), rho=0.5, sdx2=0.5)
masize("logistic1",opts)
opts <- list(p=0.25, b1=log(1.5), sdx1=1, b2=log(0.5), f2=0.5, rho=0.5, ns=10000,
             seed=1234)
masize("logistic4",opts)
opts <- list(p=0.25, b1=log(1.5), sdx1=1, b2=log(0.5), f2=0.5, rho=0.5, ns=10000,
             seed=1234)
masize("logistic4",opts)
opts <- list(p=0.25, b1=log(1.5), sdx1=4.5, b2=log(0.5), f2=0.5, rho=0.5, ns=50000,
             seed=1234)
masize("logistic4",opts)
## Poisson model
opts <- list(m=0.5, b2=log(1.25), rho=0.3, sdx2=sqrt(0.25*0.75))
masize("poisson1",opts)
opts <- list(m=0.5, b1=log(1.4), f1=0.25, b2=log(1.25), f2=0.25, rho=0.3)
masize("poisson5",opts)
opts <- c(opts,ns=10000, seed=1234)
masize("poisson9",opts)
## Cox model
# BpBm
opts <- list(b2=log(1.5), rho=0.45, f=0.2, sdx2=sqrt(0.25*0.75))
masize("cox1",opts)
opts <- list(b1=log(2), f1=0.5, b2=log(1.5), f2=0.25, rho=0.45, f=0.2, seed=1234)
masize("cox5",c(opts, ns=10000))
masize("cox5",c(opts, ns=50000))
## End(Not run)
```

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MCMCgrm

Mixed modeling with genetic relationship matrices

### **Description**

Mixed modeling with genetic relationship matrices

## Usage

```
MCMCgrm(
  model,
  prior,
  data,
  GRM,
  eps = 0,
  n.thin = 10,
  n.burnin = 3000,
  n.iter = 13000,
  ...
)
```

### **Arguments**

```
model
                   statistical model.
                   a list of priors for parameters in the model above.
prior
data
                   a data.frame containing outcome and covariates.
GRM
                   a relationship matrix.
eps
                   a small number added to the diagonal of the a nonpositive definite GRM.
n.thin
                   thinning parameter in the MCMC.
                   the number of burn-in's.
n.burnin
n.iter
                   the number of iterations.
                   other options as appropriate for MCMCglmm.
. . .
```

## Details

Mixed modeling with genomic relationship matrix. This is appropriate with relationship matrix derived from family structures or unrelated individuals based on whole genome data.

The function was created to address a number of issues involving mixed modelling with family data or population sample with whole genome data. First, the implementation will shed light on the uncertainty involved with polygenic effect in that posterior distributions can be obtained. Second, while the model can be used with the MCMCglmm package there is often issues with the specification of pedigree structures but this is less of a problem with genetic relationship matrices. We can use established algorithms to generate kinship or genomic relationship matrix as input to the MCMCglmm function. Third, it is more intuitive to specify function arguments in line with other packages such as R2OpenBUGS, R2jags or glmmBUGS. In addition, our experiences of tuning the model would help to reset the input and default values.

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#### Value

The returned value is an object as generated by MCMCglmm.

### Author(s)

Jing Hua Zhao

#### References

Hadfield JD (2010). "MCMC Methods for Multi-Response Generalized Linear Mixed Models: The MCMCglmm R Package." *Journal of Statistical Software*, **33**(2), 1 - 22. doi:10.18637/jss.v033.i02.

```
## Not run:
### with kinship
# library(kinship)
# fam <- with(151,makefamid(id,fid,mid))</pre>
# s <-with(151, makekinship(fam, id, fid, mid))</pre>
# K <- as.matrix(s)*2
### with gap
s \leftarrow kin.morgan(151)
K <- with(s,kin.matrix*2)</pre>
\label{eq:prior}  \mbox{prior} <- \mbox{list(R=list(V=1, nu=0.002), G=list(G1=list(V=1, nu=0.002)))} 
m <- MCMCgrm(qt~1,prior,151,K)</pre>
save(m,file="151.m")
pdf("151.pdf")
plot(m)
dev.off()
# A real analysis on bats
## data
bianfu.GRM <- read.table("bianfu.GRM.txt", header = TRUE)</pre>
bianfu.GRM[1:5,1:6]
Data <- read.table(file = "PHONE.txt", header = TRUE,</pre>
                     colClasses=c(rep("factor",3),rep("numeric",7)))
## MCMCgrm
library("MCMCglmm")
GRM <- as.matrix(bianfu.GRM[,-1])</pre>
colnames(GRM) <- rownames(GRM) <- bianfu.GRM[,1]</pre>
library(gap)
names(Data)[1] <- "id"</pre>
prior \leftarrow list(G = list(G1 = list(V = 1, nu = 0.002)), R = list(V = 1, nu = 0.002))
model1.1 <- MCMCgrm(WEIGTHT ~ 1, prior, Data, GRM, n.burnin=100, n.iter=1000, verbose=FALSE)</pre>
## an alternative
names(Data)[1] <- "animal"</pre>
N <- nrow(Data)
i <- rep(1:N, rep(N, N))
j <- rep(1:N, N)</pre>
s <- Matrix::spMatrix(N, N, i, j, as.vector(GRM))</pre>
Ginv <- Matrix::solve(s)</pre>
class(Ginv) <- "dgCMatrix"</pre>
rownames(Ginv) <- Ginv@Dimnames[[1]] <- with(Data, animal)</pre>
```

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```
model1.2 <- MCMCglmm(WEIGTHT ~ 1, random= ~ animal, data = Data,
    ginverse=list(animal=Ginv), prior = prior, burnin=100, nitt=1000, verbose=FALSE)
## without missing data
model1.3 <- MCMCglmm(Peak_Freq ~ WEIGTHT, random = ~ animal,
    data=subset(Data,!is.na(Peak_Freq)&!is.na(WEIGTHT)),
    ginverse=list(animal=Ginv), prior = prior, burnin=100, nitt=1000, verbose=FALSE)
## End(Not run)</pre>
```

METAL\_forestplot

forest plot as R/meta's forest for METAL outputs

## **Description**

forest plot as R/meta's forest for METAL outputs

### Usage

```
METAL_forestplot(tbl, all, rsid, package = "meta", split = FALSE, ...)
```

## Arguments

tbl	Meta-anslysis summary statistics.
all	statistics from all contributing studies.
rsid	SNPID-rsid mapping file.
package	style of plot as in meta, rmeta or forestplot.
split	when TRUE, individual prot-MarkerName.pdf will be generated.
	Additional arguments to meta::forest.

### **Details**

This functions takes a meta-data from METAL (tbl) and data from contributing studies (all) for forest plot. It also takes a SNPID-rsid mapping (rsid) as contributing studies often involve discrepancies in rsid so it is appropriate to use SNPID, i.e., chr:pos\_A1\_A2 (A1<=A2).

The study-specific and total sample sizes (N) can be customised from METAL commands. By default, the input triplets each contain a MarkerName variable which is the unique SNP identifier (e.g., chr:pos:a1:a2) and the tbl argument has variables A1 and A2 as produced by METAL while the all argument has EFFECT\_ALLELE and REFERENCE\_ALLELE as with a study variable indicating study name. Another variable common the tbl and all is prot variable as the function was developed in a protein based meta-analysis. From these all information is in place for generation of a list of Forest plots through a batch run.

CUSTOMVARIABLE N LABEL N as N WEIGHTLABEL N

## Value

It will generate a forest plot specified by pdf for direction-adjusted effect sizes.

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#### Author(s)

Jing Hua Zhao

### References

Scharzer G. (2007). meta: An R package for meta-analysis. R News, 7:40-5, https://cran.r-project.org/doc/Rnews/Rnews\_2007-3.pdf, https://CRAN.R-project.org/package=meta.

Willer CJ, Li Y, Abecasis GR. (2010). METAL: fast and efficient meta-analysis of genomewideassociation scans. Bioinformations. 26:2190-1, https://github.com/statgen/METAL, https://genome.sph.umich.edu/wiki/METAL

## **Examples**

```
## Not run:
    require(gap.datasets)
    data(OPG)
    METAL_forestplot(OPGtbl,OPGall,OPGrsid,width=8.75,height=5,digits.TE=2,digits.se=2)
## End(Not run)
```

metap

Meta-analysis of p values

# Description

Meta-analysis of p values

# Usage

```
metap(data, N, verbose = "Y", prefixp = "p", prefixn = "n")
```

# **Arguments**

data data frame.

N Number of studies.

verbose Control of detailed output.

 $\label{eq:prefix} \text{Prefix of p value, with default value "p"}.$ 

prefixn Preifx of sample size, with default value "n".

### **Details**

This function is the method of meta-analysis used in the Genetic Investigation of ANThropometric Traits (GIANT) consortium, which is based on normal approximation of p values and weighted by sample sizes from individual studies.

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### Value

- x2 Fisher's chi-squared statistics.
- p P values from Fisher's method according to chi-squared distribution with 2\*N degree(s) of freedom.
- z Combined z value.
- p1 One-sided p value.
- p2 Two-sided p value.

### Author(s)

Jing Hua Zhao

### See Also

metareg

```
## Not run:
s \leftarrow data.frame(p1=0.1^rep(8:2,each=7,times=1),n1=rep(32000,49),
                  p2=0.1^rep(8:2,each=1,times=7),n2=rep(8000,49))
cbind(s,metap(s,2))
# Speliotes, Elizabeth K., M.D. [ESPELIOTES@PARTNERS.ORG]
# 22-2-2008 MRC-Epid JHZ
np <- 7
p <- 0.1<sup>((np+1):2)</sup>
z \leftarrow qnorm(1-p/2)
n <- c(32000,8000)
n1 <- n[1]
s1 <- s2 <- vector("numeric")</pre>
for (i in 1:np)
   a <- z[i]
   for (j in 1:np)
   {
        b \leftarrow z[j]
        metaz1 \leftarrow (sqrt(n1)*a+sqrt(n[1])*b)/sqrt(n1+n[1])
        metap1 <- pnorm(-abs(metaz1))</pre>
        metaz2 <- (sqrt(n1)*a+sqrt(n[2])*b)/sqrt(n1+n[2])
        metap2 <- pnorm(-abs(metaz2))</pre>
        k <- (i-1)*np+j
        \verb|cat(k,"\t",p[i],"\t",p[j],"\t",metap1,metaz1,"\t",metap2,metaz2,"\n")|
        s1[k] \leftarrow metap1
        s2[k] \leftarrow metap2
  }
}
q <- -log10(sort(p,decreasing=TRUE))</pre>
t1 <- matrix(-log10(sort(s1,decreasing=TRUE)),np,np)</pre>
t2 <- matrix(-log10(sort(s2,decreasing=TRUE)),np,np)</pre>
```

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```
par(mfrow=c(1,2),bg="white",mar=c(4.2,3.8,0.2,0.2))
persp(q,q,t1)
persp(q,q,t2)
## End(Not run)
```

metareg

Fixed and random effects model for meta-analysis

### **Description**

Fixed and random effects model for meta-analysis

### Usage

```
metareg(data, N, verbose = "Y", prefixb = "b", prefixse = "se")
```

## **Arguments**

data Data frame to be used.

N Number of studies.

verbose A control for screen output.

prefixb Prefix of estimate; default value is "b".

prefixse Prefix of standard error; default value is "se". The function accepts a wide for-

mat data with estimates as b1,...,bN and standard errors as se1,...,seN. More

generally, they can be specified by prefixes in the function argument.

#### **Details**

Given k=n studies with  $b_1,...,b_N$  being  $\beta$ 's and  $se_1,...,se_N$  standard errors from regression, the fixed effects model uses inverse variance weighting such that  $w_1=1/se_1^2,...,w_N=1/se_N^2$  and the combined  $\beta$  as the weighted average,  $\beta_f=(b_1*w_1+...+b_N*w_N)/w$ , with  $w=w_1+...+w_N$  being the total weight, the se for this estimate is  $se_f=\sqrt{1/w}$ . A normal z-statistic is obtained as  $z_f=\beta_f/se_f$ , and the corresponding p value  $p_f=2*pnorm(-abs(z_f))$ . For the random effects model, denote  $q_w=w_1*(b_1-\beta_f)^2+...+w_N*(b_N-\beta_f)^2$  and  $dl=max(0,(q_w-(k-1))/(w-(w_1^2+...+w_N^2)/w))$ , corrected weights are obtained such that  $w_{1c}=1/(1/w_1+dl),...,w_{Nc}=1/(1/w_N+dl)$ , totaling  $w_c=w_{1c}+...+w_{Nc}$ . The combined  $\beta$  and se are then  $\beta_r=(b_1*w_{1c}+...+b_N*w_{Nc})/w_c$  and  $se_r=\sqrt{(1/w_c)}$ , leading to a z-statistic  $z_r=\beta_r/se_r$  and a p-value  $p_r=2*pnorm(-abs(z_r))$ . Moreover, a p-value testing for heterogeneity is  $p_{heter}=pchisq(q_w,k-1,lower.tail=FALSE)$ .

### Value

The returned value is a data frame with the following variables:

- p\_f P value (fixed effects model).
- p\_r P value (random effects model).
- beta\_f regression coefficient.

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- beta\_r regression coefficient.
- · se\_f standard error.
- se\_r standard error.
- z\_f z value.
- z\_r z value.
- p\_heter heterogeneity test p value.
- i2  $I^2$  statistic.
- k No of tests used.
- eps smallest double-precision number.

#### Note

Adapted from a SAS macro, 23-7-2009 MRC-Epid JHZ

## Author(s)

Shengxu Li, Jing Hua Zhao

#### References

Higgins JP, Thompson SG, Deeks JJ, Altman DG (2003). "Measuring inconsistency in meta-analyses." *Bmj*, **327**(7414), 557-60. ISSN 0959-8138 (Print) 0959-8138, doi:10.1136/bmj.327.7414.557.

## **Examples**

mht.control

Controls for mhtplot

## Description

Parameter specification through function

## Usage

```
mht.control(
  type = "p",
  usepos = FALSE,
  logscale = TRUE,
  base = 10,
  cutoffs = NULL,
```

```
colors = NULL,
labels = NULL,
srt = 45,
gap = NULL,
cex = 0.4,
yline = 3,
xline = 3
```

## **Arguments**

type Type of plot. usepos A flag.

logscale A flag for log-scale.

base Base of log.

cutoffs Cutoffs of P-value, etc.
colors Colours for chromosomes.
labels Labels for chromosomes.

srt Rotation degrees.

gap Gap between data points.
cex Scaling factor of data points.

yline Vertical adjustment. xline Horiztonal adjustment.

#### Value

A list as above.

mhtplot Manhattan plot

## **Description**

Manhattan plot

### Usage

```
mhtplot(data, control = mht.control(), hcontrol = hmht.control(), ...)
```

## Arguments

data

a data frame with three columns representing chromosome, position and p values.

control

A control function named mht.control() with the following arguments:

- type a flag with value "p" or "I" indicating if points or lines are to be drawn.
- usepos a flag to use real chromosomal positions as composed to ordinal positions with default value FALSE.

- logscale a flag to indicate if p value is to be log-transformed with default value TRUE.
- base the base of the logarithm with default value 10.
- cutoffs the cut-offs where horizontal line(s) are drawn with default value NULL.
- colors the color for different chromosome(s), and random if unspecified with default values NULL.
- labels labels for the ticks on x-axis with default value NULL.
- srt degree to which labels are rotated with default value of 45.
- gap gap between chromosomes with default value NULL.
- cex cex for the data points.
- yline Margin line position.
- xline Margin line position.

hcontrol

A control function named hmht.control() with the following arguments:

- data. chunk of data to be highlighted with default value NULL.
- colors colors for annotated genes.
- yoffset. offset above the data point showing most significant p value with default value 0.5.
- cex shrinkage factor for data points with default value 1.5.
- boxed if the label for the highlited region with default value FALSE.

other options in compatible with the R plot function.

### **Details**

. . .

To generate Manhattan plot, e.g., of genomewide significance (p values) and a random variable that is uniformly distributed. By default, a log10-transformation is applied. Note that with real chromosomal positions, it is also appropriate to plot and some but not all chromosomes.

It is possible to specify options such as xlab and ylim when the plot is requested for data in other context.

### Value

The plot is shown on or saved to the appropriate device.

# Author(s)

Jing Hua Zhao

## See Also

qqunif

```
## Not run:
# foo example
test <- matrix(c(1,1,4,1,1,6,1,10,3,2,1,5,2,2,6,2,4,8),byrow=TRUE,6)
mhtplot(test)
mhtplot(test,mht.control(logscale=FALSE))
# fake example with Affy500k data</pre>
```

```
affy <-c(40220, 41400, 33801, 32334, 32056, 31470, 25835, 27457, 22864, 28501, 26273,
         24954, 19188, 15721, 14356, 15309, 11281, 14881, 6399, 12400, 7125, 6207)
CM <- cumsum(affy)
n.markers <- sum(affy)</pre>
n.chr <- length(affy)</pre>
test <- data.frame(chr=rep(1:n.chr,affy),pos=1:n.markers,p=runif(n.markers))</pre>
# to reduce size of the plot
# bitmap("mhtplot.bmp",res=72*5)
oldpar <- par()
par(cex=0.6)
colors <- rep(c("blue", "green"), 11)</pre>
# other colors, e.g.
# colors <- c("red", "blue", "green", "cyan", "yellow", "gray", "magenta", "red", "blue", "green",
            "cyan", "yellow", "gray", "magenta", "red", "blue", "green", "cyan", "yellow", "gray",
               "magenta", "red")
\verb|mhtplot(test,control=mht.control(colors=colors),pch=19,srt=0||
title("A simulated example according to EPIC-Norfolk QCed SNPs")
axis(1,pos=0,labels=FALSE,tick=FALSE)
abline(0,0)
# dev.off()
par(oldpar)
mhtplot(test,control=mht.control(usepos=TRUE,colors=colors,gap=10000),pch=19,bg=colors)
title("Real positions with a gap of 10000 bp between chromosomes")
box()
png("manhattan.png",height=3600,width=6000,res=600)
opar <- par()
par(cex=0.4)
ops <- mht.control(colors=rep(c("lightgray", "lightblue"), 11), srt=0, yline=2.5, xline=2)
require(gap.datasets)
mhtplot(mhtdata[,c("chr","pos","p")],ops,xlab="",ylab="",srt=0)
axis(2,at=1:16)
title("An adaptable plot as .png")
par(opar)
dev.off()
data <- with(mhtdata,cbind(chr,pos,p))</pre>
glist <- c("IRS1", "SPRY2", "FTO", "GRIK3", "SNED1", "HTR1A", "MARCH3", "WISP3", "PPP1R3B",
          "RP1L1", "FDFT1", "SLC39A14", "GFRA1", "MC4R")
hdata <- subset(mhtdata,gene\
color <- rep(c("lightgray", "gray"),11)</pre>
glen <- length(glist)</pre>
hcolor <- rep("red",glen)</pre>
par(las=2, xpd=TRUE, cex.axis=1.8, cex=0.4)
ops <- mht.control(colors=color,yline=1.5,xline=3,labels=paste("chr",1:22,sep=""),</pre>
                    srt=270)
hops <- hmht.control(data=hdata,colors=hcolor)</pre>
mhtplot(data,ops,hops,pch=19)
axis(2,pos=2,at=1:16)
title("Manhattan plot with genes highlighted",cex.main=1.8)
mhtplot(data,mht.control(cutoffs=c(4,6,8,16)),pch=19)
title("Another plain Manhattan plot")
```

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```
# Miami plot

test <- within(test, {pr=1-p})
miamiplot(test,chr="chr",bp="pos",p="p",pr="pr")
## End(Not run)</pre>
```

mhtplot.trunc

Truncated Manhattan plot

## **Description**

Truncated Manhattan plot

## Usage

```
mhtplot.trunc(
  chr = "CHR",
  bp = "BP",
  p = NULL,
  log10p = NULL,
  z = NULL,
  snp = "SNP",
  col = c("gray10", "gray60"),
  chrlabs = NULL,
  suggestiveline = -\log 10(1e-05),
  genomewideline = -\log 10(5e-08),
  highlight = NULL,
  annotatelog10P = NULL,
  annotateTop = FALSE,
  cex.mtext = 1.5,
  cex.text = 0.7,
  mtext.line = 2,
  y.ax.space = 5,
  y.brk1 = NULL,
  y.brk2 = NULL,
  trunc.yaxis = TRUE,
  cex.axis = 1.2,
  delta = 0.05,
)
```

## **Arguments**

```
x A data.frame.
chr Chromosome.
bp Position.
p p values, e.g., "1.23e-600".
log10p log10(p).
```

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z z statistic, i.e., BETA/SE.

snp SNP. Pending on the setup it could either of variant or gene ID(s).

col Colours.

chrlabs Chromosome labels, 1,2,...22,23,24,25.

suggestiveline Suggestive line. genomewideline Genomewide line.

highlight A list of SNPs to be highlighted. annotatelog 10P Threshold of  $-\log 10(P)$  to annotate.

annotateTop Annotate top.

cex.mtext axis label extension factor.
cex.text SNP label extension factor.

mtext.line position of the y lab.

y.ax.space interval of ticks of the y axis. y.brk1 lower -log10(P) break point. y.brk2 upper -log10(P) break point.

trunc.yaxis do not truncate y-axisx when FALSE.

cex.axis extension factor for x-, y-axis.

delta a value to enable column(s) of red points.

... other options.

## **Details**

To generate truncated Manhattan plot, e.g., of genomewide significance (P values) or a random variable that is uniformly distributed.

The rationale of this function is to extend mhtplot() to handle extremely small p values as often seen from a protein GWAS. Optionally, the function also draws an ordinary Manhattan plot.

### Value

The plot is shown on or saved to the appropriate device.

## Author(s)

James Peters, Jing Hua Zhao

## See Also

mhtplot.

```
# giant/images/c/c8/Meta-analysis_Locke_et_al%2BUKBiobank_2018_UPDATED.txt.gz
gz <- gzfile("work/Meta-analysis_Locke_et_al+UKBiobank_2018_UPDATED.txt.gz")</pre>
BMI <- within(read.delim(gz,as.is=TRUE), {Z <- BETA/SE})
print(subset(BMI[c("CHR", "POS", "SNP", "P")], CHR!=16 & P<=1e-150))</pre>
library(Rmpfr)
print(within(subset(BMI, P==0, select=c(CHR,POS,SNP,Z)),
             {P <- format(2*pnorm(mpfr(abs(Z),100),lower.tail=FALSE));</pre>
              Pvalue <- pvalue(Z); log10P <- -log10p(Z)}))</pre>
png("BMI.png", res=300, units="in", width=9, height=6)
par(oma=c(0,0,0,0), mar=c(5,6.5,1,1))
mhtplot.trunc(BMI, chr="CHR", bp="POS", z="Z", snp="SNP",
              suggestiveline=FALSE, genomewideline=-log10(1e-8),
              cex.mtext=1.2, cex.text=1.2,
              annotatelog10P=156, annotateTop = FALSE,
              highlight=c("rs13021737","rs17817449","rs6567160"),
              mtext.line=3, y.brk1=200, y.brk2=280, trunc.yaxis=TRUE,
              cex.axis=1.2, cex=0.5,
              y.ax.space=20,
              col = c("blue4", "skyblue")
dev.off()
## End(Not run)
```

mhtplot2

Manhattan plot with annotations

# Description

Manhattan plot with annotations

### Usage

```
mhtplot2(data, control = mht.control(), hcontrol = hmht.control(), ...)
```

## **Arguments**

data

a data frame with three columns representing chromosome, position and p values.

control

A control function named mht.control() with the following arguments:

- type a flag with value "p" or "l" indicating if points or lines are to be drawn.
- usepos a flag to use real chromosomal positions as composed to ordinal positions with default value FALSE.
- logscale a flag to indicate if p value is to be log-transformed with default value TRUE.
- base the base of the logarithm with default value 10.
- cutoffs the cut-offs where horizontal line(s) are drawn with default value NULL.
- colors the color for different chromosome(s), and random if unspecified with default values NULL.
- labels labels for the ticks on x-axis with default value NULL.

- srt degree to which labels are rotated with default value of 45.
- gap gap between chromosomes with default value NULL.
- cex cex for the data points.
- yline Margin line position.
- xline Margin line position.

hcontrol

A control function named hmht.control() with the following arguments:

- data chunk of data to be highlighted with default value NULL.
- colors colors for annotated genes.
- yoffset offset above the data point showing most significant p value with default value 0.5.
- cex shrinkage factor for data points with default value 1.5.
- boxed if the label for the highlited region with default value FALSE.

other options in compatible with the R plot function.

#### **Details**

To generate Manhattan plot with annotations. The function is generic and for instance could be used for genomewide p values or any random variable that is uniformly distributed. By default, a log10-transformation is applied. Note that with real chromosomal positions, it is also appropriate to plot and some but not all chromosomes.

It is possible to specify options such as xlab, ylim and font family when the plot is requested for data in other context.

To maintain back compatibility options as in mhtplot are used. The positions of the horizontal labels are now in the middle rather than at the beginning of their bands in the plot.

## Value

The plot is shown on or saved to the appropriate device.

#### Author(s)

Jing Hua Zhao

#### References

den Hoed M, Eijgelsheim M, Esko T, Brundel BJ, Peal DS, Evans DM, Nolte IM, Segrè AV, Holm H, Handsaker RE, Westra HJ, Johnson T, Isaacs A, Yang J, Lundby A, Zhao JH, Kim YJ, Go MJ, Almgren P, Bochud M, Boucher G, Cornelis MC, Gudbjartsson D, Hadley D, van der Harst P, Hayward C, den Heijer M, Igl W, Jackson AU, Kutalik Z, Luan J, Kemp JP, Kristiansson K, Ladenvall C, Lorentzon M, Montasser ME, Njajou OT, O'Reilly PF, Padmanabhan S, St Pourcain B, Rankinen T, Salo P, Tanaka T, Timpson NJ, Vitart V, Waite L, Wheeler W, Zhang W, Draisma HH, Feitosa MF, Kerr KF, Lind PA, Mihailov E, Onland-Moret NC, Song C, Weedon MN, Xie W, Yengo L, Absher D, Albert CM, Alonso A, Arking DE, de Bakker PI, Balkau B, Barlassina C, Benaglio P, Bis JC, Bouatia-Naji N, Brage S, Chanock SJ, Chines PS, Chung M, Darbar D, Dina C, Dörr M, Elliott P, Felix SB, Fischer K, Fuchsberger C, de Geus EJ, Goyette P, Gudnason V, Harris TB, Hartikainen AL, Havulinna AS, Heckbert SR, Hicks AA, Hofman A, Holewijn S, Hoogstra-Berends F, Hottenga JJ, Jensen MK, Johansson A, Junttila J, Kääb S, Kanon B, Ketkar S, Khaw KT, Knowles JW, Kooner AS, others (2013). "Identification of heart rate-associated loci and their effects on cardiac conduction and rhythm disorders." *Nat Genet*, **45**(6), 621-31. ISSN 1061-4036 (Print) 1061-4036, doi:10.1038/ng.2610.

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#### **Examples**

```
## Not run:
The following example uses only chromosomes 14 and 20 of the Nat Genet paper.
mdata <- within(hr1420,{</pre>
     c1<-colour==1
     c2<-colour==2
     c3<-colour==3
     colour[c1] <- 62
     colour[c2] <- 73</pre>
     colour[c3] <- 552</pre>
})
mdata <- mdata[,c("CHR","POS","P","gene","colour")]</pre>
ops <- mht.control(colors=rep(c("lightgray","gray"),11),yline=1.5,xline=2,srt=0)</pre>
hops <- hmht.control(data=subset(mdata,!is.na(gene)))</pre>
v <- "Verdana"
ifelse(Sys.info()['sysname']=="Windows", windowsFonts(ffamily=windowsFont(v)),
                   ffamily <- v)
tiff("mh.tiff", width=.03937*189, height=.03937*189/2, units="in", res=1200,
             compress="lzw")
par(las=2, xpd=TRUE, cex.axis=1.8, cex=0.4)
mhtplot2(with(mdata,cbind(CHR,POS,P,colour)),ops,hops,pch=19,
                        ylab=expression(paste(plain("-"),log[10],plain("p-value"),sep=" ")),
                        family="ffamily")
axis(2,pos=2,at=seq(0,25,5),family="ffamily",cex=0.5,cex.axis=1.1)
dev.off()
# To exemplify the use of chr, pos and p without gene annotation
# in response to query from Vallejo, Roger <Roger.Vallejo@ARS.USDA.GOV>
opar <- par()
par(cex=0.4)
ops <- mht.control(colors=rep(c("lightgray","lightblue"),11),srt=0,yline=2.5,xline=2)</pre>
\label{lem:mhtplot2} \\ \texttt{mhtplot2}(\texttt{data.frame(mhtdata[,c("chr","pos","p")],gene=NA,color=NA),ops,xlab="",ylab="",srt=0)} \\ \\ \texttt{mhtplot2}(\texttt{data.frame(mhtdata[,c("chr","pos","p")],gene=NA,color=NA),ops,xlab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",ylab="",y
axis(2,at=1:16)
title("data in mhtplot used by mhtplot2")
par(opar)
## End(Not run)
```

mia

Multiple imputation analysis for hap

### **Description**

Multiple imputation analysis for hap

## Usage

```
mia(
  hapfile = "hap.out",
  assfile = "assign.out",
  miafile = "mia.out",
  so = 0,
```

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```
ns = 0,
mi = 0,
allsnps = 0,
sas = 0
)
```

## **Arguments**

hapfile hap haplotype output file name. hap assignment output file name. assfile miafile mia output file name. to generate results according to subject order. so do not sort in subject order. ns number of multiple imputations used in hap. mi all loci are SNPs. allsnps produce SAS data step program. sas

#### **Details**

This command reads outputs from hap session that uses multiple imputations, i.e. -mi# option. To simplify matters it assumes -ss option is specified together with -mi option there.

This is a very naive version of MIANALYZE, but can produce results for PROC MIANALYZE of SAS.

It simply extracts outputs from hap.

### Value

The returned value is a list.

### Note

adapted from hap, in fact cline.c and cline.h are not used. keywords utilities

#### References

Zhao JH and W Qian (2003) Association analysis of unrelated individuals using polymorphic genetic markers. RSS 2003, Hassalt, Belgium

Clayton DG (2001) SNPHAP. https://github.com/chr1swallace/snphap.

# See Also

hap

```
## Not run:
# 4 SNP example, to generate hap.out and assign.out alone
data(fsnps)
hap(id=fsnps[,1],data=fsnps[,3:10],nloci=4)
# to generate results of imputations
```

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```
control <- hap.control(ss=1,mi=5)
hap(id=fsnps[,1],data=fsnps[,3:10],nloci=4,control=control)

# to extract information from the second run above
mia(so=1,ns=1,mi=5)
file.show("mia.out")

## commands to check out where the output files are as follows:
## Windows
# system("command.com")
## Unix
# system("csh")

## End(Not run)</pre>
```

miamiplot

Miami plot

## Description

Miami plot

### Usage

```
miamiplot(
  х,
  chr = "CHR",
  bp = "BP",
  p = "P",
  pr = "PR"
  snp = "SNP",
  col = c("midnightblue", "chartreuse4"),
  col2 = c("royalblue1", "seagreen1"),
  ymax = NULL,
  highlight = NULL,
  highlight.add = NULL,
  pch = 19,
  cex = 0.75,
  cex.lab = 1,
  xlab = "Chromosome",
  ylab = "-log10(P) [y>0]; log10(P) [y<0]",
  lcols = c("red", "black"),
  1wds = c(5, 2),
  ltys = c(1, 2),
  main = "",
)
```

### **Arguments**

Χ

Input data.

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chr Chromsome. Position. bp P value. р P value of the other GWAS. pr Marker. snp col Colors. Colors. col2 ymax Max y. highlight Highlight flag. Highlight meta-data. highlight.add Symbol. pch cex cex. cex.lab cex for labels. xlab Label for x-axis. ylab Label for y-axis. lcols Colors. lwds lwd. ltys lty. Main title. main

Additional options.

## **Details**

The function allows for contrast of genomewide P values from two GWASs. It is conceptually simpler than at the first sight since it involves only one set of chromosomal positions.

### Value

None.

```
## Not run:
    mhtdata <- within(mhtdata,{pr=p})
    miamiplot(mhtdata,chr="chr",bp="pos",p="p",pr="pr",snp="rsn")
## End(Not run)</pre>
```

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miamiplot2

Miami Plot

# Description

Miami Plot

## Usage

```
miamiplot2(
  gwas1,
  gwas2,
  name1 = "GWAS 1",
  name2 = "GWAS 2",
  chr1 = "chr",
  chr2 = "chr",
  pos1 = "pos",
pos2 = "pos",
p1 = "p",
  p2 = "p",
  z1 = NULL,
  z2 = NULL,
  sug = 1e-05,
  sig = 5e-08,
  pcutoff = 0.1,
  topcols = c("green3", "darkgreen"),
  botcols = c("royalblue1", "navy"),
  yAxisInterval = 5
)
```

# Arguments

gwas1	The first of two GWAS datasets to plot, in the upper region.
gwas2	The second of two GWAS datasets to plot, in the lower region.
name1	The name of the first dataset, plotted above the upper plot region. Defaults to '"GWAS 1"'.
name2	The name of the second dataset, plotted below the lower plot region. Defaults to ""GWAS 2"'.
chr1	The name of the column containing chromosome number in gwas1. Defaults to '"chr".
chr2	The name of the column containing chromosome number in gwas2. Defaults to '"chr".
pos1	The name of the column containing SNP position in gwas1. Defaults to "pos".
pos2	The name of the column containing SNP position in gwas2. Defaults to "pos".
p1	The name of the column containing p-values in gwas1. Defaults to "p".
p2	The name of the column containing p-values in gwas2. Defaults to "p".
<b>z</b> 1	The name of the column containing z-values in gwas1. Defaults to "NULL".
z2	The name of the column containing z-values in gwas2. Defaults to "NULL".

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sug	The threshold for suggestive significance, plotted as a light grey dashed line.
sig	The threshold for genome-wide significance, plotted as a dark grey dashed line.
pcutoff	The p-value threshold below which SNPs will be ignored. Defaults to 0.1. It is not recommended to set this higher as it will narrow the central gap between the two plot region where the chromosome number is plotted.
topcols	A vector of two colours to plot alternating chromosomes in for the upper plot. Defaults to green3 and darkgreen.
botcols	A vector of two colours to plot alternating chromosomes in for the lower plot. Defaults to royalblue1 and navy.
yAxisInterval	The interval between tick marks on the y-axis. Defaults to 5, 2 may be more suitable for plots with larger minimum p-values.

#### **Details**

Creates a Miami plot to compare results from two genome-wide association analyses.

### Value

In addition to creading a Miami plot, the function returns a data frame containing x coordinates for chromosome start positions (required for labelManhattan)

### Note

Extended to handle extreme P values.

### Author(s)

Jonathan Marten

## **Examples**

mr

Mendelian randomization analysis

## **Description**

Mendelian randomization analysis

### Usage

```
mr(data, X, Y, alpha = 0.05, other_plots = FALSE)
```

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#### **Arguments**

data	Data to be used.
X	Exposure.
Υ	Outcome.
alpha	type I error rate for confidence intervals.
other_plots	To add funnel and forest plots.

#### **Details**

The function initially intends to rework on GSMR outputs, but it would be appropriate for general use.

### Value

The result and plots.

### **Examples**

```
library(ggplot2)
library(gap)
mrdat <-
rs188743906 0.6804 0.1104 0.00177 0.01660
  rs10793962 0.2098 0.0212 0.00378 0.00536 -0.014567 0.0138196
  rs11562629 0.1155 0.0181 0.00960 0.00378 -0.010040 0.0151460
v <- c("SNP", "b.LIF.R", "SE.LIF.R", "b.FEV1", "SE.FEV1", "b.CAD", "SE.CAD")
mrdat <- setNames(as.data.frame(scan(file=textConnection(mrdat),</pre>
                          what=list("",0,0,0,0,0,0))), v)
knitr::kable(mrdat,caption="Table 2. LIF.R and CAD/FEV1")
res <- mr(mrdat, "LIF.R", c("CAD", "FEV1"), other_plots=TRUE)</pre>
s <- res$r[-1,]
colnames(s) <- res$r[1,]</pre>
r \leftarrow matrix(as.numeric(s[,-1]),nrow(s),dimnames=list(res*r[-1,1],res*r[1,-1]))
p <- sapply(c("IVW", "EGGER", "WM", "PWM"), function(x)</pre>
   format(2*pnorm(-abs(r[,paste0("b",x)]/r[,paste0("seb",x)])),digits=3,scientific=TRUE))
rp <- t(data.frame(round(r,3),p))</pre>
knitr::kable(rp,align="r",caption="Table 3. LIFR variant rs635634 and CAD/FEV1")
```

 ${\tt mr\_forestplot}$ 

Mendelian Randomization forest plot

## **Description**

Mendelian Randomization forest plot

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### Usage

```
mr_forestplot(dat, sm = "", title = "", ...)
```

### **Arguments**

dat A data.frame with outcome id, effect size and standard error.

sm Summary measure such as OR, RR, MD.

title Title of the meta-analysis.

Other options for meta::forest().

### **Details**

This is a wrapper of meta::forest() for multi-outcome Mendelian Randomization. It allows for the flexibility of both binary and continuous outcomes with and without summary level statistics.

```
## Not run:
tnfb <-
            "multiple sclerosis" 0.69058600 0.059270400
   "systemic lupus erythematosus" 0.76687500 0.079000500
         "sclerosing cholangitis" 0.62671500 0.075954700
 "rheumatoid arthritis" -0.00378072 0.000625160
     "inflammatory bowel disease" -0.14334200 0.025272500
         "ankylosing spondylitis" -0.00316852 0.000626225
                "hypothyroidism" -0.00432054 0.000987324
             "allergic rhinitis" 0.00393075 0.000926002
         "IgA glomerulonephritis" -0.32696600 0.105262000
                 "atopic eczema" -0.00204018 0.000678061
 require(dplyr)
 tnfb <- as.data.frame(scan(file=textConnection(tnfb),what=list("",0,0))) %>%
        setNames(c("outcome","Effect","StdErr")) %>%
        mutate(outcome=gsub("\b(^[a-z])","\U\1",outcome,perl=TRUE))
# default output
mr_forestplot(tnfb, colgap.forest.left="0.05cm", fontsize=14, leftlabs=c("Outcome", "b", "SE"),
           common=FALSE, random=FALSE, print.I2=FALSE, print.pval.Q=FALSE, print.tau2=FALSE,
              spacing=1.6,digits.TE=2,digits.se=2)
 # no summary level statistics
mr_forestplot(tnfb, colgap.forest.left="0.05cm", fontsize=14,
           leftcols="studlab", leftlabs="Outcome", plotwidth="3inch", sm="OR", rightlabs="ci",
              sortvar=tnfb[["Effect"]],
           common=FALSE, random=FALSE, print.I2=FALSE, print.pval.Q=FALSE, print.tau2=FALSE,
              backtransf=TRUE, spacing=1.6)
# with P values
mr_forestplot(tnfb,colgap.forest.left="0.05cm", fontsize=14,
              leftcols=c("studlab"), leftlabs=c("Outcome"),
              plotwidth="3inch", sm="OR", sortvar=tnfb[["Effect"]],
              rightcols=c("effect","ci","pval"), rightlabs=c("OR","95%CI","P"),
              digits=3, digits.pval=2, scientific.pval=TRUE,
           common=FALSE, random=FALSE, print.I2=FALSE, print.pval.Q=FALSE, print.tau2=FALSE,
              addrow=TRUE, backtransf=TRUE, spacing=1.6)
```

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```
## End(Not run)
```

mtdt

Transmission/disequilibrium test of a multiallelic marker

### **Description**

Transmission/disequilibrium test of a multiallelic marker

### Usage

```
mtdt(x, n.sim = 0)
```

#### **Arguments**

x the data table.

n.sim the number of simulations.

#### **Details**

This function calculates transmission-disequilibrium statistics involving multiallelic marker. Inside the function are tril and triu used to obtain lower and upper triangular matrices.

#### Value

It returned list contains the following components:

- SE Spielman-Ewens Chi-square from the observed data.
- ST Stuart or score Statistic from the observed data.
- pSE the simulated p value.
- sSE standard error of the simulated p value.
- pST the simulated p value.
- sST standard error of the simulated p value.

### Author(s)

Mike Miller, Jing Hua Zhao

#### References

Miller MB (1997). "Genomic scanning and the transmission/disequilibrium test: analysis of error rates." *Genet Epidemiol*, **14**(6), 851-6. ISSN 0741-0395 (Print) 0741-0395, doi:10.1002/(sici)1098-2272(1997)14:6<854::aidgepi48>3.3.co;27.

Sham PC, Curtis D (1995). "An extended transmission/disequilibrium test (TDT) for multi-allele marker loci." *Ann Hum Genet*, **59**(3), 323-36. ISSN 0003-4800 (Print) 0003-4800, doi:10.1111/j.14691809.1995.tb00751.x.

Spielman RS, Ewens WJ (1996). "The TDT and other family-based tests for linkage disequilibrium and association." *Am J Hum Genet*, **59**(5), 983-9. ISSN 0002-9297 (Print) 0002-9297.

Zhao JH, Sham PC, Curtis D (1999). "A program for the Monte Carlo evaluation of significance of the extended transmission/disequilibrium test." *Am J Hum Genet*, **64**, 1484-1485.

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### See Also

bt

# **Examples**

```
## Not run:
# Copeman et al (1995) Nat Genet 9: 80-5
0,0, 1, 3, 0,0, 0, 2, 3, 0, 0, 0,
             2,3,26,35, 7,0, 2,10,11, 3, 4, 1,
             2,3,22,26, 6,2, 4, 4,10, 2, 2, 0,
            0,1, 7,10, 2,0, 0, 2, 2, 1, 1, 0,
             0,0, 1, 4, 0,1, 0, 1, 0, 0, 0, 0,
             0,2, 5, 4, 1,1, 0, 0, 0, 2, 0, 0,
             0,0, 2, 6, 1,0, 2, 0, 2, 0, 0, 0,
             0,3, 6,19, 6,0, 0, 2, 5, 3, 0, 0,
             0,0, 3, 1, 1,0, 0, 0, 1, 0, 0, 0,
            0,0,0,2,0,0,0,0,0,0,0,0,0,
            0,0, 1, 0, 0,0, 0, 0, 0, 0, 0, 0), nrow=12)
# See note to bt for the score test obtained by SAS
mtdt(x)
## End(Not run)
```

mtdt2

Transmission/disequilibrium test of a multiallelic marker by Bradley-Terry model

# **Description**

Transmission/disequilibrium test of a multiallelic marker by Bradley-Terry model

### Usage

```
mtdt2(x, verbose = TRUE, n.sim = NULL, ...)
```

# **Arguments**

x the data table.
verbose To print out test statistics if TRUE.
n.sim Number of simulations.
other options compatible with the BTm function.

# **Details**

This function calculates transmission-disequilibrium statistics involving multiallelic marker according to Bradley-Terry model.

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#### Value

It returned list contains the following components:

- c2b A data frame in four-column format showing transmitted vs nontransmitted counts.
- BTm A fitted Bradley-Terry model object.
- X2 Allele-wise, genotype-wise and goodness-of-fit Chi-squared statistics.
- df Degrees of freedom.
- p P value.
- pn Monte Carlo p values when n.sim is specified.

### Author(s)

Jing Hua Zhao keywords models keywords htest

#### References

Firth, D. (2005). Bradley-terry models in R. Journal of Statistical Software 12(1):1-12

Sham PC, Curtis D (1995). "An extended transmission/disequilibrium test (TDT) for multi-allele marker loci." *Ann Hum Genet*, **59**(3), 323-36. ISSN 0003-4800 (Print) 0003-4800, doi:10.1111/j.14691809.1995.tb00751.x.

Turner H, Firth D (2010) Bradley-Terry models in R: The BradleyTerry2 package. https://cran.r-project.org/web/packages/BradleyTerry2/vignettes/BradleyTerry.pdf.

Zhao JH, Sham PC, Curtis D (1999). "A program for the Monte Carlo evaluation of significance of the extended transmission/disequilibrium test." *Am J Hum Genet*, **64**, 1484-1485.

### See Also

mtdt

```
## Not run:
# Copeman et al (1995) Nat Genet 9: 80-5
0,0, 1, 3, 0,0, 0, 2, 3, 0, 0, 0,
             2,3,26,35, 7,0, 2,10,11, 3, 4, 1,
             2,3,22,26, 6,2, 4, 4,10, 2, 2, 0,
             0,1, 7,10, 2,0, 0, 2, 2, 1, 1, 0,
             0,0, 1, 4, 0,1, 0, 1, 0, 0, 0, 0,
             0,2, 5, 4, 1,1, 0, 0, 0, 2, 0, 0,
             0,0, 2, 6, 1,0, 2, 0, 2, 0, 0, 0,
             0,3, 6,19, 6,0, 0, 2, 5, 3, 0, 0,
             0,0, 3, 1, 1,0, 0, 0, 1, 0, 0, 0,
             0,0,0,2,0,0,0,0,0,0,0,0,0,
             0,0, 1, 0, 0,0, 0, 0, 0, 0, 0, 0), nrow=12)
xx <- mtdt2(x,refcat="12")</pre>
## End(Not run)
```

muvar 113

muvar

Means and variances under 1- and 2- locus (biallelic) QTL model

# **Description**

Means and variances under 1- and 2- locus (biallelic) QTL model

# Usage

```
muvar(
    n.loci = 1,
    y1 = c(0, 1, 1),
    y12 = c(1, 1, 1, 1, 1, 0, 0, 0, 0),
    p1 = 0.99,
    p2 = 0.9
)
```

# Arguments

n.loci	number of loci, 1=single locus, 2=two loci.
y1	the genotypic means of aa, Aa and AA.
y12	the genotypic means of aa, Aa and AA at the first locus and bb, Bb and BB at the second locus.
p1	the frequency of the lower allele, or the that for the first locus under a 2-locus model.
p2	the frequency of the lower allele at the second locus.

### **Details**

Function muvar() gives means and variances under 1-locus and 2-locus QTL model (simple); in the latter case it gives results from different avenues. This function is included for experimental purpose and yet to be generalized.

# Value

Currently it does not return any value except screen output; the results can be kept via R's sink() command or via modifying the C/R codes.

# Note

Adapted from an earlier C program written for the above book.

### Author(s)

Jing Hua Zhao

# References

Sham P (1998). Statistics in Human Genetics. Arnold

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### **Examples**

```
## Not run:
# the default 1-locus model
muvar(n.loci=1,y1=c(0,1,1),p1=0.5)
# the default 2-locus model
muvar(n.loci=2,y12=c(1,1,1,1,1,0,0,0,0),p1=0.99,p2=0.9)
## End(Not run)
```

mvmeta

Multivariate meta-analysis based on generalized least squares

# Description

Multivariate meta-analysis based on generalized least squares

# Usage

```
mvmeta(b, V)
```

# **Arguments**

b the parameter estimates.

V the triangular variance-covariance matrix.

### **Details**

This function accepts a data matrix of parameter estimates and their variance-covariance matrix from individual studies and obtain a generalized least squares (GLS) estimate and heterogeneity statistic.

For instance, this would be appropriate for combining linear correlation coefficients of single nucleotide polymorphisms (SNPs) for a given region.

# Value

The returned value is a list containing:

- d the compact parameter estimates.
- Psi the compact covariance-covariance matrix.
- X the design matrix.
- beta the pooled parameter estimates.
- cov.beta the pooled variance-covariance matrix.
- X2 the Chi-squared statistic for heterogeneity.
- df the degrees(s) of freedom.
- p the p value.

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#### Author(s)

Jing Hua Zhao

# References

Hartung J, Knapp G, Sinha BK (2008). Statistical Meta-analysis with Applications. Wiley.

### See Also

metareg

```
## Not run:
# example 11.3 from Hartung et al.
b <- matrix(c(</pre>
0.808, 1.308, 1.379, NA, NA,
NA, 1.266, 1.828, 1.962, NA,
NA, 1.835, NA, 2.568, NA,
NA, 1.272, NA, NA, 2.038,
1.171, 2.024, 2.423, 3.159, NA,
0.681, NA, NA, NA, NA), ncol=5, byrow=TRUE)
psi1 <- psi2 <- psi3 <- psi4 <- psi5 <- psi6 <- matrix(0,5,5)
psi1[1,1] <- 0.0985
psi1[1,2] <- 0.0611
psi1[1,3] <- 0.0623
psi1[2,2] <- 0.1142
psi1[2,3] <- 0.0761
psi1[3,3] <- 0.1215
psi2[2,2] <- 0.0713
psi2[2,3] <- 0.0539
psi2[2,4] <- 0.0561
psi2[3,3] <- 0.0938
psi2[3,4] <- 0.0698
psi2[4,4] <- 0.0981
psi3[2,2] <- 0.1228
psi3[2,4] <- 0.1119
psi3[4,4] <- 0.1790
psi4[2,2] <- 0.0562
psi4[2,5] <- 0.0459
psi4[5,5] <- 0.0815
psi5[1,1] <- 0.0895
psi5[1,2] <- 0.0729
psi5[1,3] <- 0.0806
psi5[1,4] <- 0.0950
psi5[2,2] <- 0.1350
psi5[2,3] <- 0.1151
psi5[2,4] <- 0.1394
psi5[3,3] <- 0.1669
```

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```
psi5[3,4] <- 0.1609
psi5[4,4] <- 0.2381

psi6[1,1] <- 0.0223

V <- rbind(psi1[upper.tri(psi1,diag=TRUE)],psi2[upper.tri(psi2,diag=TRUE)],
psi3[upper.tri(psi3,diag=TRUE)],psi4[upper.tri(psi4,diag=TRUE)],
psi5[upper.tri(psi5,diag=TRUE)],psi6[upper.tri(psi6,diag=TRUE)])

mvmeta(b,V)

## End(Not run)</pre>
```

pbsize

Power for population-based association design

# **Description**

Power for population-based association design

### Usage

```
pbsize(kp, gamma = 4.5, p = 0.15, alpha = 5e-08, beta = 0.2)
```

### **Arguments**

kp population disease prevalence.

gamma genotype relative risk assuming multiplicative model.

p frequency of disease allele.

alpha type I error rate. beta type II error rate.

### **Details**

This function implements Long et al. (1997) statistics for population-based association design. This is based on a contingency table test and accurate level of significance can be obtained by Fisher's exact test.

### Value

The returned value is scaler containing the required sample size.

#### Author(s)

Jing Hua Zhao extracted from rm.c.

### References

Scott WK, Pericak-Vance MA, et al. (1997). Genetic analysis of complex diseases. Science 275: 1327.

Long AD, Grote MN, Langley CH (1997). Genetic analysis of complex traits. Science 275: 1328.

Rosner B (2000). Fundamentals of Biostatistics, 5th Edition, Duxbury.

Armitage P, Colton T (2005). Encyclopedia of Biostatistics, 2nd Edition, Wiley.

pbsize2

#### See Also

fbsize

```
kp <- c(0.01, 0.05, 0.10, 0.2)
models <- matrix(c(</pre>
    4.0, 0.01,
    4.0, 0.10,
    4.0, 0.50,
    4.0, 0.80,
    2.0, 0.01,
    2.0, 0.10,
    2.0, 0.50,
    2.0, 0.80,
    1.5, 0.01,
    1.5, 0.10,
    1.5, 0.50,
    1.5, 0.80), ncol=2, byrow=TRUE)
outfile <- "pbsize.txt"</pre>
cat("gamma","p","p1","p5","p10","p20\n",sep="\t",file=outfile)
for(i in 1:dim(models)[1])
{
  g \leftarrow models[i,1]
  p <- models[i,2]</pre>
  n <- vector()</pre>
  for(k in kp) n <- c(n,ceiling(pbsize(k,g,p)))</pre>
  \verb|cat(models[i,1:2],n,sep="\t",file=outfile,append=TRUE|)|
  cat("\n",file=outfile,append=TRUE)
}
table5 <- read.table(outfile,header=TRUE,sep="\t")</pre>
unlink(outfile)
# Alzheimer's disease
g < -4.5
p <- 0.15
alpha <- 5e-8
beta <- 0.2
z1alpha <- qnorm(1-alpha/2)
                                # 5.45
z1beta <- qnorm(1-beta)</pre>
q <- 1-p
pi <- 0.065
                                 # 0.07 and zbeta generate 163
k < -pi*(g*p+q)^2
s \leftarrow (1-pi*g^2)*p^2+(1-pi*g)*2*p*q+(1-pi)*q^2
# LGL formula
lambda <- pi*(g^2*p+q-(g*p+q)^2)/(1-pi*(g*p+q)^2)
# mine
lambda <- pi*p*q*(g-1)^2/(1-k)
n <- (z1alpha+z1beta)^2/lambda</pre>
cat("\nPopulation-based result: Kp = ",k, "Kq = ",s, "n = ",ceiling(n),"\n")
```

118 pbsize2

#### **Description**

Power for case-control association design

#### Usage

```
pbsize2(
    N,
    fc = 0.5,
    alpha = 0.05,
    gamma = 4.5,
    p = 0.15,
    kp = 0.1,
    model = "additive"
)
```

# **Arguments**

N The sample size.

fc The proportion of cases in the sample.

alpha Type I error rate.

gamma The genotype relative risk (GRR).

p Frequency of the risk allele.

kp The prevalence of disease in the population.

model Disease model, i.e., "multiplicative", "additive", "dominant", "recessive", "overdominant".

### **Details**

This extends pbsize from a multiplicative model for a case-control design under a range of disease models. Essentially, for given sample sizes(s), a proportion of which (fc) being cases, the function calculates power estimate for a given type I error (alpha), genotype relative risk (gamma), frequency of the risk allele (p), the prevalence of disease in the population (kp) and optionally a disease model (model). A major difference would be the consideration of case/control ascertainment in pbsize.

Internally, the function obtains a baseline risk to make the disease model consistent with Kp as in tscc and should produce accurate power estimate. It provides power estimates for given sample size(s) only.

# Value

The returned value is the power for the specified design.

### See Also

The design follows that of pbsize.

```
## Not run:
# single calculation
m <- c("multiplicative","recessive","dominant","additive","overdominant")
for(i in 1:5) print(pbsize2(N=50,alpha=5e-2,gamma=1.1,p=0.1,kp=0.1, model=m[i]))
# a range of sample sizes</pre>
```

pedtodot 119

pedtodot

Converting pedigree(s) to dot file(s)

# **Description**

Converting pedigree(s) to dot file(s)

# Usage

```
pedtodot(
  pedfile,
  makeped = FALSE,
  sink = TRUE,
  page = "B5",
  url = "https://jinghuazhao.github.io/",
  height = 0.5,
  width = 0.75,
  rotate = 0,
  dir = "none"
)
```

# **Arguments**

pedfile	a pedigree file in GAS or LINKAGE format, note if individual's ID is character then it is necessary to specify as.is=T in the read.table command.
makeped	a logical variable indicating if the pedigree file is post-makeped.
sink	a logical variable indicating if .dot file(s) are created.
page	a string indicating the page size, e.g, A4, A5, B5, Legal, Letter, Executive, " $x$ , $y$ ", where $x$ , $y$ is the customized page size.
url	Unified Resource Locator (URL) associated with the diagram(s).
height	the height of node(s).
width	the width of node(s).
rotate	if set to 90, the diagram is in landscape.
dir	direction of edges, i.e., "none", "forward", "back", "both". This will be useful if the diagram is viewed by lneato.

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#### **Details**

This function converts GAS or LINKAGE formatted pedigree(s) into .dot file for each pedigree to be used by dot in graphviz, which is a flexible package for graphics freely available.

Note that a single PostScript (PDF) file can be obtained by dot, fdp, or neato.

```
dot -Tps -o
or
fdp -Tps -o
or
neato -Tps -o
```

See relevant documentations for other formats.

To preserve the original order of pedigree(s) in the data, you can examine the examples at the end of this document.

Under Cygwin/Linux/Unix, the PostScript file can be converted to Portable Document Format (PDF) default to Acrobat.

```
ps2pdf
```

Use ps2pdf12, ps2pdf13, or ps2pdf14 for appropriate versions of Acrobat according to information given on the headline of .

Under Linux, you can also visualize the .dot file directly via command,

dotty &

We can extract the code below (or within pedtodot.Rd) to pedtodot and then use command:

sh pedtodot

#### Value

For each pedigree, the function generates a .dot file to be used by dot. The collection of all pedigrees (\*.dot) can also be put together.

# Note

This is based on the gawk script program pedtodot by David Duffy with minor changes.

# Author(s)

```
David Duffy, Jing Hua Zhao
```

#### See Also

package sem in CRAN and Rgraphviz in BioConductor https://www.bioconductor.org/.

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```
3 2 1 7/7 7/10
 6
        3 1
              1
                 7/7 7/10
                 7/7 7/10
 7
        3 2
              1
8
              1 -/- -/-
        0 1
        4 1 1 7/9 3/10
9
    8
10
        0 2 1 -/- -/-
   0
   2 10 2 1 7/7 7/7
11
12 2 10 2 2 6/7 7/7
13 0
       0 1 1 -/- -/-
14 13 11 1 1 7/8 7/8
15 0 0 1 1 -/- -/-
16 15 12 2 1 6/6 7/7
p2 <- as.data.frame(p1)</pre>
names(p2) <-c("id","fid","mid","sex","aff","GABRB1","D4S1645")</pre>
p3 <- data.frame(pid=10081,p2)</pre>
attach(p3)
pedtodot(p3)
# Three examples of pedigree-drawing
# assuming pre-MakePed LINKAGE file in which IDs are characters
pre<-read.table("pheno.pre",as.is=TRUE)[,1:6]</pre>
pedtodot(pre)
dir()
# for post-MakePed LINKAGE file in which IDs are integers
ped <-read.table("pheno.ped")[,1:10]</pre>
pedtodot(ped,makeped=TRUE)
dir()
# for a single file with a list of pedigrees ordered data
sink("gaw14.dot")
pedtodot(ped,sink=FALSE)
sink()
file.show("gaw14.dot")
# more details
pedtodot(ped, sink=FALSE, page="B5", url="https://jinghuazhao.github.io/")
# An example from Richard Mott and in the demo
filespec <- system.file("tests/ped.1.3.pre")</pre>
pre <- read.table(filespec,as.is=TRUE)</pre>
pre
pedtodot(pre,dir="forward")
## End(Not run)
```

pedtodot\_verbatim

Pedigree-drawing with graphviz

### **Description**

Pedigree-drawing with graphviz

### Usage

```
pedtodot_verbatim(f, run = FALSE, toDOT = FALSE, ...)
```

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### **Arguments**

f A data frame containing pedigrees, each with pedigree id, individual id, father

id, mother id, sex and affection status.

run A flag to run dot/neato on the generated .dot file(s).

toDOT A flag to generate script for DOT::dot().

... Other flag(s) for DOT::dot().

### **Details**

Read a GAS or LINKAGE format pedigree, return a digraph in the dot language and optionally call dot/neato to make pedigree drawing.

This is a verbatim translation of the original pedtodot implemneted in Bash/awk in contrast to pedtodot which was largely a mirror. To check independently, try xsel  $-i < (cat pedtodot_verbatim.R)$  or cat pedtodot\_verbatim.R | xsel -i and paste into an R session.

### Value

No value is returned but outputs in .dot, .pdf, and .svg.

### Note

Adapted from Bash/awk script by David Duffy

# **Examples**

```
## Not run:
# pedigree p3 in pedtodot
   pedtodot_verbatim(p3,run=TRUE,toDOT=TRUE,return="verbatim")
## End(Not run)
```

pfc

Probability of familial clustering of disease

# Description

Probability of familial clustering of disease

### Usage

```
pfc(famdata, enum = 0)
```

#### **Arguments**

famdata collective information of sib size, number of affected sibs and their frequencies.

enum a switch taking value 1 if all possible tables are to be enumerated.

# Details

To calculate exact probability of familial clustering of disease

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#### Value

The returned value is a list containing (tailp,sump,nenum are only available if enum=1:

- p the probabitly of familial clustering.
- stat the deviances, chi-squares based on binomial and hypergeometric distributions, the degrees of freedom should take into account the number of marginals used.
- tailp the exact statistical significance.
- sump sum of the probabilities used for error checking.
- nenum the total number of tables enumerated.

### Note

Adapted from family.for by Dani Zelterman, 25/7/03

#### Author(s)

Dani Zelterman, Jing Hua Zhao

#### References

```
Yu C, Zelterman D (2001). "Exact inference for family disease clusters." Communications in Statistics - Theory and Methods, 30(11), 2293-2305. ISSN 0361-0926, doi:10.1081/STA100107686.

Yu C, Zelterman D (2002). "Statistical inference for familial disease clusters." Biometrics, 58(3),
```

481-91. ISSN 0006-341X (Print) 0006-341x, doi:10.1111/j.0006341x.2002.00481.x.

# See Also

kin.morgan

```
## Not run:
# IPF among 203 siblings of 100 COPD patients from Liang KY, SL Zeger,
# Qaquish B. Multivariate regression analyses for categorical data
# (with discussion). J Roy Stat Soc B 1992, 54:3-40
# the degrees of freedom is 15
famtest<-c(</pre>
1, 0, 36,
1, 1, 12,
2, 0, 15,
2, 1,
      7,
2, 2,
      1,
3, 0,
3, 1,
       7,
3, 2,
       3,
3, 3,
       2,
4, 0,
       3,
4, 1,
      3,
4, 2, 1,
6, 0, 1,
6, 2, 1,
6, 3, 1,
6, 4, 1,
```

124 pfc.sim

```
6, 6, 1)
test<-t(matrix(famtest,nrow=3))
famp<-pfc(test)
## End(Not run)</pre>
```

pfc.sim

Probability of familial clustering of disease

# **Description**

Probability of familial clustering of disease

# Usage

```
pfc.sim(famdata, n.sim = 1e+06, n.loop = 1)
```

# **Arguments**

famdata collective information of sib size, number of affected sibs and their frequencies.

n. sim number of simulations in a single Monte Carlo run.

n. loop total number of Monte Carlo runs.

# **Details**

To calculate probability of familial clustering of disease using Monte Carlo simulation.

#### Value

The returned value is a list containing:

- n.sim a copy of the number of simulations in a single Monte Carlo run.
- n.loop the total number of Monte Carlo runs.
- p the observed p value.
- tailpl accumulated probabilities at the lower tails.
- tailpu simulated p values.

#### Note

Adapted from runi.for from Change Yu, 5/6/4

### Author(s)

Chang Yu, Dani Zelterman

# References

Yu C, Zelterman D (2001). "Exact inference for family disease clusters." *Communications in Statistics - Theory and Methods*, **30**(11), 2293-2305. ISSN 0361-0926, doi:10.1081/STA100107686.

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### See Also

pfc

# **Examples**

```
## Not run:
# Li FP, Fraumeni JF Jr, Mulvihill JJ, Blattner WA, Dreyfus MG, Tucker MA,
# Miller RW. A cancer family syndrome in twenty-four kindreds.
# Cancer Res 1988, 48(18):5358-62.
# family_size #_of_affected frequency
famtest<-c(</pre>
1, 0, 2,
1, 1, 0,
2, 0, 1,
2, 1, 4,
2, 2, 3,
3, 0, 0,
3, 1, 2,
3, 2, 1,
3, 3, 1,
4, 0, 0,
4, 1, 2,
5, 0, 0,
5, 1, 1,
6, 0, 0,
6, 1, 1,
7, 0, 0,
7, 1, 1,
8, 0, 0,
8, 1, 1,
8, 2, 1,
8, 3, 1,
9, 3, 1)
test<-matrix(famtest,byrow=T,ncol=3)</pre>
famp<-pfc.sim(test)</pre>
## End(Not run)
```

pgc

Preparing weight for GENECOUNTING

# Description

Preparing weight for GENECOUNTING

# Usage

```
pgc(data, handle.miss = 1, is.genotype = 0, with.id = 0)
```

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#### **Arguments**

data the multilocus genotype data for a set of individuals. handle.miss a flag to indicate if missing data is kept, 0 = no, 1 = yes.

is genotype a flag to indicate if the data is already in the form of genotype identifiers. with id a flag to indicate if the unique multilocus genotype identifier is generated.

#### **Details**

This function is a R port of the GENECOUNTING/PREPARE program which takes an array of genotype data and collapses individuals with the same multilocus genotype. This function can also be used to prepare for the genotype table in testing Hardy-Weinberg equilibrium.

### Value

The returned value is a list containing:

- cdata the collapsed genotype data.
- wt the frequency weight.
- obscom the observed number of combinations or genotypes.
- idsave optional, available only if with.id = 1.

#### Note

Built on pgc.c.

### Author(s)

Jing Hua Zhao

# References

Zhao JH, Sham PC (2003). "Generic number systems and haplotype analysis." *Comput Methods Programs Biomed*, **70**(1), 1-9. ISSN 0169-2607 (Print) 0169-2607, doi:10.1016/s01692607(01)00193-6.

# See Also

```
genecounting, hwe.hardy
```

```
## Not run:
require(gap.datasets)
data(hla)
x <- hla[,3:8]

# do not handle missing data
y<-pgc(x,handle.miss=0,with.id=1)
hla.gc<-genecounting(y$cdata,y$wt)

# handle missing but with multilocus genotype identifier
pgc(x,handle.miss=1,with.id=1)</pre>
```

plot.hap.score 127

```
# handle missing data with no identifier
pgc(x,handle.miss=1,with.id=0)
## End(Not run)
```

plot.hap.score

Plot haplotype frequencies versus haplotype score statistics

# **Description**

Method function to plot a class of type hap.score

# Usage

```
## S3 method for class 'hap.score' plot(x, ...)
```

# **Arguments**

x The object returned from hap.score (which has class hap.score).

... Optional arguments.

### Value

Nothing is returned.

This is a plot method function used to plot haplotype frequencies on the x-axis and haplotype-specific scores on the y-axis. Because hap.score is a class, the generic plot function can be used, which in turn calls this plot.hap.score function.

### References

Schaid DJ, Rowland CM, Tines DE, Jacobson RM, Poland GA (2002) Score tests for association of traits with haplotypes when linkage phase is ambiguous. Amer J Hum Genet 70:425-34

### See Also

```
hap.score
```

```
## Not run:
save <- hap.score(y, geno, trait.type = "gaussian")

# Example illustrating generic plot function:
plot(save)

# Example illustrating specific method plot function:
plot.hap.score(save)

## End(Not run)</pre>
```

128 print.hap.score

print.hap.score

Print a hap.score object

### **Description**

Method function to print a class of type hap.score

# Usage

```
## S3 method for class 'hap.score'
print(x, ...)
```

### **Arguments**

x The object returned from hap.score (which has class hap.score).

... Optional argunents.

#### Value

Nothing is returned.

This is a print method function used to print information from hap.score class, with haplotype-specific information given in a table. Because hap.score is a class, the generic print function can be used, which in turn calls this print.hap.score function.

### References

Schaid DJ, Rowland CM, Tines DE, Jacobson RM, Poland GA (2002) Score tests for association of traits with haplotypes when linkage phase is ambiguous. Amer J Hum Genet 70:425-34

### See Also

```
hap.score
```

```
## Not run:
save <- hap.score(y, geno, trait.type = "gaussian")

# Example illustrating generic print function:
print(save)

# Example illustrating specific method print function:
print.hap.score(save)

## End(Not run)</pre>
```

pvalue 129

pvalue

P value for a normal deviate

# **Description**

P value for a normal deviate

### Usage

```
pvalue(z, decimals = 2)
```

# **Arguments**

z normal deviate.

decimals number of decimal places.

# Value

P value as a string variable.

# **Examples**

```
pvalue(-1.96)
```

qqfun

Quantile-comparison plots

# **Description**

Quantile-comparison plots

# Usage

```
qqfun(
  Х,
  distribution = "norm",
  ylab = deparse(substitute(x)),
  xlab = paste(distribution, "quantiles"),
  main = NULL,
  las = par("las"),
  envelope = 0.95,
  labels = FALSE,
  col = palette()[4],
  lcol = palette()[2],
  xlim = NULL,
  ylim = NULL,
  lwd = 1,
  pch = 1,
  bg = palette()[4],
  cex = 0.4,
```

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```
line = c("quartiles", "robust", "none"),
...
)
```

### **Arguments**

X	vector of numeric values.
distribution	root name of comparison distribution – e.g., norm for the normal distribution; $\boldsymbol{t}$ for the t-distribution.
ylab	label for vertical (empirical quantiles) axis.
xlab	label for horizontal (comparison quantiles) axis.
main	label for plot.
las	if 0, ticks labels are drawn parallel to the axis; set to 1 for horizontal labels (see graphics::par).
envelope	confidence level for point-wise confidence envelope, or FALSE for no envelope.
labels	vector of point labels for interactive point identification, or FALSE for no labels.
col	color for points; the default is the <i>fourth</i> entry in the current color palette (see grDevices::palette and graphics::par).
lcol	color for lines; the default is the second entry as above.
xlim	the x limits $(x1, x2)$ of the plot. Note that $x1 > x2$ is allowed and leads to a reversed axis.
ylim	the y limits of the plot.
lwd	line width; default is 1 (see graphics::par). Confidence envelopes are drawn at half this line width.
pch	plotting character for points; default is 1 (a circle, see graphics::par).
bg	background color of points.
cex	factor for expanding the size of plotted symbols; the default is .4.
line	"quartiles" to pass a line through the quartile-pairs, or "robust" for a robust-regression line; the latter uses the rlm function in the MASS package. Specifying line = "none" suppresses the line.
	arguments such as df to be passed to the appropriate quantile function.

# **Details**

Plots empirical quantiles of a variable against theoretical quantiles of a comparison distribution.

Draws theoretical quantile-comparison plots for variables and for studentized residuals from a linear model. A comparison line is drawn on the plot either through the quartiles of the two distributions, or by robust regression.

Any distribution for which quantile and density functions exist in R (with prefixes q and d, respectively) may be used. Studentized residuals are plotted against the appropriate t-distribution.

This is adapted from qq.plot of package car with different values for points and lines, more options, more transparent code and examples in the current setting. Another similar but sophisticated function is qqmath of package lattice.

### Value

NULL. These functions are used only for their side effect (to make a graph).

qqunif 131

# Author(s)

John Fox, Jing Hua Zhao

### References

Davison, A. C. (2003) Statistical Models. Cambridge University Press.

Leemis, L. M., J. T. Mcqueston (2008) *Univariate distribution relationships*. The American Statistician 62:45-53

### See Also

```
stats::qqnorm, qqunif, gcontrol2
```

# **Examples**

```
## Not run:
p <- runif(100)
alpha <- 1/log(10)
qqfun(p,dist="unif")
qqfun(-log10(p),dist="exp",rate=alpha,pch=21)
#library(car)
#qq.plot(p,dist="unif")
#qq.plot(-log10(p),dist="exp",rate=alpha)
#library(lattice)
#qqmath(~ -log10(p), distribution = function(p) qexp(p,rate=alpha))
## End(Not run)</pre>
```

qqunif

Q-Q plot for uniformly distributed random variable

# **Description**

Q-Q plot for uniformly distributed random variable

# Usage

```
qqunif(
    u,
    type = "unif",
    logscale = TRUE,
    base = 10,
    col = palette()[4],
    lcol = palette()[2],
    ci = FALSE,
    alpha = 0.05,
    ...
)
```

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### **Arguments**

u a vector of uniformly distributed random variables.

type string option to specify distribution: "unif"=uniform, "exp"=exponential.

logscale to use logscale.

base the base of the log function.

col color for points.

lcol color for the diagonal line.

ci logical option to show confidence interval.

alpha 1-confidence level, e.g., 0.05.

... other options as appropriae for the qqplot function.

### **Details**

This function produces Q-Q plot for a random variable following uniform distribution with or without using log-scale. Note that the log-scale is by default for type "exp", which is a plot based on exponential order statistics. This appears to be more appropriate than the commonly used procedure whereby the expected value of uniform order statistics is directly log-transformed.

#### Value

The returned value is a list with components of a qqplot:

- x expected value for uniform order statistics or its -log(,base) counterpart.
- y observed value or its -log(,base) counterpart.

### Author(s)

Jing Hua Zhao

### References

Balakrishnan N, Nevzorov VB. A Primer on Statistical Distributions. Wiley 2003.

Casella G, Berger RL. Statistical Inference, Second Edition. Duxbury 2002.

Davison AC. Statistical Models. Cambridge University Press 2003.

### See Also

qqfun

```
## Not run:
# Q-Q Plot for 1000 U(0,1) r.v., marking those <= 1e-5
u_obs <- runif(1000)
r <- qqunif(u_obs,pch=21,bg="blue",bty="n")
u_exp <- r$y
hits <- u_exp >= 2.30103
points(r$x[hits],u_exp[hits],pch=21,bg="green")
legend("topleft",sprintf("GC.lambda=\"
## End(Not run)
```

qtl2dplot 133

qtl2dplot

2D QTL plot

# Description

```
2D QTL plot
```

# Usage

```
qtl2dplot(
  d,
  chrlen = gap::hg19,
  snp_name = "SNP",
  snp_chr = "Chr",
  snp_pos = "bp",
  gene_chr = "p.chr",
  gene_start = "p.start",
  gene_end = "p.end",
  trait = "p.target.short",
  gene = "p.gene",
  TSS = FALSE,
  cis = "cis",
  value = "log10p",
  plot = TRUE,
  cex.labels = 0.6,
  cex.points = 0.6,
  xlab = "QTL position",
  ylab = "Gene position"
)
```

# Arguments

cex.labels

d	Data to be used.
chrlen	lengths of chromosomes for specific build: hg18, hg19, hg38.
snp_name	variant name.
snp_chr	variant chromosome.
snp_pos	variant position.
gene_chr	gene chromosome.
gene_start	gene start position.
gene_end	gene end position.
trait	trait name.
gene	gene name.
TSS	to use TSS when TRUE.
cis	cis variant when TRUE.
value	A specific value to show.
plot	to plot when TRUE.

Axis label extension factor.

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```
cex.points Data point extension factor.
xlab X-axis title.
ylab Y-axis title.
```

# **Details**

This function is both used as its own for a 2d plot and/or generate data for a plotly counterpart.

#### Value

positional information.

# **Examples**

```
## Not run:
INF <- Sys.getenv("INF")
d <- read.csv(file.path(INF,"work","INF1.merge.cis.vs.trans"),as.is=TRUE)
r <- qtl2dplot(d)
## End(Not run)</pre>
```

qt12dplotly

2D QTL plotly

# **Description**

```
2D QTL plotly
```

# Usage

```
qtl2dplotly(
   d,
   chrlen = gap::hg19,
   qtl.id = "SNPid:",
   qtl.prefix = "QTL:",
   qtl.gene = "Gene:",
   target.type = "Protein",
   TSS = FALSE,
   xlab = "QTL position",
   ylab = "Gene position",
   ...
)
```

### **Arguments**

```
d Data in qtl2dplot() format.

chrlen Lengths of chromosomes for specific build: hg18, hg19, hg38.

qtl.id QTL id.

qtl.prefix QTL prefix.

qtl.gene QTL gene.
```

qtl3dplotly 135

Additional arguments, e.g., target, log10p, to qtl2dplot.

```
target.type Type of target, e.g., protein.

TSS to use TSS when TRUE.

xlab X-axis title.

ylab Y-axis title.
```

### Value

. . .

A plotly figure.

# **Examples**

```
## Not run:
INF <- Sys.getenv("INF")
d <- read.csv(file.path(INF,"work","INF1.merge.cis.vs.trans"),as.is=TRUE)
r <- qtl2dplotly(d)
htmlwidgets::saveWidget(r,file=file.path(INF,"INF1.qtl2dplotly.html"))
r
## End(Not run)</pre>
```

qtl3dplotly

3D QTL plot

# **Description**

```
3D QTL plot
```

# Usage

```
qtl3dplotly(
   d,
   chrlen = gap::hg19,
   zmax = 300,
   qtl.id = "SNPid:",
   qtl.prefix = "QTL:",
   qtl.gene = "Gene:",
   target.type = "Protein",
   TSS = FALSE,
   xlab = "QTL position",
   ylab = "Gene position",
   ...
)
```

# Arguments

d Data in qtl2d() format.

chrlen Lengths of chromosomes for specific build: hg18, hg19, hg38.

zmax Maximum value (e.g., -log10p) to truncate, above which they would be set to

this value.

136 qtlClassifier

```
qtl.id QTL id.
qtl.prefix QTL prefix.
qtl.gene QTL target gene.
target.type Type of target, e.g., protein.
TSS to use TSS when TRUE.
xlab X-axis title.
ylab Y-axis title.
... Additional arguments, e.g., to qtl2dplot().
```

### Value

A plotly figure.

### **Examples**

qtlClassifier

A QTL cis/trans classifier

# **Description**

A QTL cis/trans classifier

# Usage

```
qtlClassifier(geneSNP, SNPPos, genePos, radius)
```

### **Arguments**

geneSNP data.frame with columns on gene, SNP and biomarker (e.g., expression, pro-

tein).

SNPPos data.frame containing SNP, chromosome and position.

genePos data.frame containing gene, chromosome, start and end positions.

radius flanking distance.

# **Details**

The function obtains QTL (simply called SNP here) cis/trans classification based on gene positions.

read.ms.output 137

#### Value

It returns a geneSNP-prefixed data.frame with the following columns:

- geneChrom gene chromosome.
- geneStart gene start.
- geneEnd gene end.
- SNPChrom pQTL chromosome.
- SNPPos pQTL position.
- Type cis/trans labels.

#### Note

This is adapted from iBMQ/eqtlClassifier as an xQTL (x=e, p, me, ...) classifier.

#### See Also

```
cis.vs.trans.classification
```

### **Examples**

```
## Not run:
  merged <- read.delim("INF1.merge",as.is=TRUE)</pre>
  hits <- merge(merged[c("CHR","POS","MarkerName","prot","log10p")],</pre>
                 inf1[c("prot", "uniprot")],by="prot")
  names(hits) <- c("prot","Chr","bp","SNP","log10p","uniprot")</pre>
  options(width=200)
  geneSNP <- merge(hits[c("prot","SNP","log10p")],</pre>
                    inf1[c("prot", "gene")], by="prot")[c("gene", "SNP", "prot", "log10p")]
  SNPPos <- hits[c("SNP","Chr","bp")]</pre>
  genePos <- inf1[c("gene","chr","start","end")]</pre>
  cvt <- qtlClassifier(geneSNP,SNPPos,genePos,1e6)</pre>
  cvt
  cistrans <- cis.vs.trans.classification(hits,inf1,"uniprot")</pre>
 cis.vs.trans <- with(cistrans,data)</pre>
 cistrans.check <- merge(cvt[c("gene", "SNP", "Type")],cis.vs.trans[c("p.gene", "SNP", "cis.trans")],</pre>
                            by.x=c("gene","SNP"),by.y=c("p.gene","SNP"))
  with(cistrans.check,table(Type,cis.trans))
## End(Not run)
```

 $\verb"read.ms.output"$ 

A utility function to read ms output

# Description

A utility function to read ms output

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#### Usage

```
read.ms.output(
  msout,
  is.file = TRUE,
  xpose = TRUE,
  verbose = TRUE,
  outfile = NULL,
  outfileonly = FALSE
)
```

# **Arguments**

```
msout an ms output.

is.file a flag indicating ms output as a system file or an R object.

xpose a flag to obtain the tranposed format as it is (when TRUE).

verbose when TRUE, display on screen every 1000 for large nsam.

outfile to save the haplotypes in a tab-delimited ASCII file.

outfileonly to reset gametes to NA when nsam/nreps is very large and is useful with outfile.
```

### **Details**

This function reads in the output of the program ms, a program to generate samples under a variety of neutral models.

The argument indicates either a file name or a vector of character strings, one string for each line of the output of ms. As with the second case, it is appropriate with system(,intern=TRUE), see example below.

### Value

The returned value is a list storing the results:

- call system call to ms.
- seed random number seed to ms.
- nsam number of copies of the locus in each sample.
- nreps the number of independent samples to generate.
- segsites a vector of the numbers of segregating sites.
- times vectors of time to most recent ancester (TMRCA) and total tree lengths.
- positions positions of polymorphic sites on a scale of (0,1).
- gametes a list of haplotype arrays.
- probs the probability of the specified number of segregating sites given the genealogical history of the sample and the value to -t option.

# Author(s)

D Davison, RR Hudson, JH Zhao

ReadGRM 139

#### References

Hudson RR (2002). "Generating samples under a Wright-Fisher neutral model of genetic variation." *Bioinformatics*, **18**(2), 337-8. ISSN 1367-4803 (Print) 1367-4803, doi:10.1093/bioinformatics/18.2.337.

Press WH, SA Teukolsky, WT Vetterling, BP Flannery (1992). Numerical Recipes in C. Cambridge University Press, Cambridge.

# **Examples**

```
## Not run:
# Assuming ms is on the path

system("ms 5 4 -s 5 > ms.out")
msout1 <- read.ms.output("ms.out")

system("ms 50 4 -s 5 > ms.out")
msout2 <- read.ms.output("ms.out",outfile="out",outfileonly=TRUE)

msout <- system("ms 5 4 -s 5 -L", intern=TRUE)
msout3 <- read.ms.output(msout,FALSE)

## End(Not run)</pre>
```

ReadGRM

A function to read GRM file

# **Description**

A function to read GRM file

# Usage

```
ReadGRM(prefix = 51)
```

# **Arguments**

prefix

file root.

ReadGRMBin

A function to read GRM binary files

# **Description**

A function to read GRM binary files

# Usage

```
ReadGRMBin(prefix, AllN = FALSE, size = 4)
```

140 revStrand

# Arguments

prefix file root.

AllN a logical variable.

size size.

### **Details**

Modified from GCTA documentation

revStrand

Allele on the reverse strand

# Description

Allele on the reverse strand

# Usage

```
revStrand(allele)
```

# **Arguments**

allele

Allele to reverse.

# **Details**

The function obtains allele on the reverse strand.

# Value

Allele on the reverse strand.

```
## Not run:
   alleles <- c("a","c","G","t")
   reverse_strand(alleles)
## End(Not run)</pre>
```

runshinygap 141

runshinygap

Start shinygap

# Description

Start shinygap

### Usage

```
runshinygap(...)
```

# **Arguments**

... Additional arguments passed to the 'runApp' function from the 'shiny' package.

# **Details**

This function starts the interactive 'shinygap' shiny web application that allows for flexible model specification.

The 'shiny' based web application allows for flexible model specification for the implemented study designs.

### Value

These are design specific.

s2k

Statistics for 2 by K table

# Description

Statistics for 2 by K table

# Usage

```
s2k(y1, y2)
```

# **Arguments**

y1 a vector containing the first row of a 2 by K contingency table.

y2 a vector containing the second row of a 2 by K contingency table.

# **Details**

This function calculates one-to-others and maximum accumulated chi-squared statistics for a 2 by K contingency table.

sentinels

#### Value

The returned value is a list containing:

- x2a the one-to-other chisquare.
- x2b the maximum accumulated chisquare.
- col1 the column index for x2a.
- col2 the column index for x2b.
- p the corresponding p value.

# Note

The lengths of y1 and y2 should be the same.

### Author(s)

Chihiro Hirotsu, Jing Hua Zhao

### References

Hirotsu C, Aoki S, Inada T, Kitao Y (2001). "An exact test for the association between the disease and alleles at highly polymorphic loci with particular interest in the haplotype analysis." *Biometrics*, **57**, 769-778.

### **Examples**

```
## Not run:
# an example from Mike Neale
# termed 'ugly' contingency table by Patrick Sullivan
y1 <- c(2,15,16,35,132,30,25,7,12,24,10,10,0)
y2 <- c(0, 6,31,49,120,27,15,8,14,25, 3, 9,3)
result <- s2k(y1,y2)
## End(Not run)</pre>
```

sentinels

Sentinel identification from GWAS summary statistics

# Description

Sentinel identification from GWAS summary statistics

# Usage

```
sentinels(
  p,
  pid,
  st,
  debug = FALSE,
  flanking = 1e+06,
  chr = "Chrom",
```

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```
pos = "End",
b = "Effect",
se = "StdErr",
log_p = NULL,
snp = "MarkerName",
sep = ","
)
```

### **Arguments**

р an object containing GWAS summary statistics. a phenotype (e.g., protein) name in pGWAS. pid st row number as in p. debug a flag to show the actual data. the width of flanking region. flanking chr Chromosome name. pos Position. Effect size. b Standard error. se log\_p log(P). Marker name. snp field delimitor. sep

### Details

This function accepts an object containing GWAS summary statistics for signal identification as defined by flanking regions. As the associate P value could be extremely small, the effect size and its standard error are used.

A distance-based approach was consequently used and reframed as an algorithm here. It takes as input signals multiple correlated variants in particular region(s) which reach genomewide significance and output three types of sentinels in a region-based manner. For a given protein and a chromosome, the algorithm proceeds as follows:

# Algorithm sentinels

Step 1. for a particular collection of genomewide significant variants on a chromosome, the width of the region is calculated according to the start and end chromosomal positions and if it is smaller than the flanking distance, the variant with the smallest P value is taken as sentinel (I) otherwise goes to step 2.

Step 2. The variant at step 1 is only a candidate and a flanking region is generated. If such a region contains no variant the candidate is recorded as sentinel (II) and a new iteration starts from the variant next to the flanking region.

Step 3. When the flanking is possible at step 2 but the P value is still larger than the candidate at step 2, the candidate is again recorded as sentinel (III) but next iteration starts from the variant just after the variant at the end position; otherwise the variant is updated as a new candidate where the next iteration starts.

Note Type I signals are often seen from variants in strong LD at a cis region, type II results seen when a chromosome contains two trans signals, type III results seen if there are multiple trans signals.

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Typically, input to the function are variants reaching certain level of significance and the function identifies minimum p value at the flanking interval; in the case of another variant in the flanking window has smaller p value it will be used instead.

For now key variables in p are "MarkerName", "End", "Effect", "StdErr", "P.value", where "End" is as in a bed file indicating marker position, and the function is set up such that row names are numbered as 1:nrow(p); see example below. When log\_p is specified, log(P) is used instead, which is appropriate with output from METAL with LOGPVALUE ON. In this case, the column named log(P) in the output is actually log10(P).

#### Value

The function give screen output.

```
## Not run:
 ## OPG as a positive control in our pGWAS
 require(gap.datasets)
data(OPG)
 p <- reshape::rename(OPGtbl, c(Chromosome="Chrom", Position="End"))</pre>
 chrs <- with(p, unique(Chrom))</pre>
 for(chr in chrs)
 {
   ps <- subset(p[c("Chrom","End","MarkerName","Effect","StdErr")], Chrom==chr)</pre>
   row.names(ps) <- 1:nrow(ps)</pre>
   sentinels(ps, "OPG", 1)
 subset(OPGrsid,MarkerName=="chr8:120081031_C_T")
 subset(OPGrsid,MarkerName=="chr17:26694861_A_G")
 ## log(P)
 p <- within(p, {logp <- log(P.value)})</pre>
 for(chr in chrs)
   ps <- subset(p[c("Chrom","End","MarkerName","logp")], Chrom==chr)</pre>
   row.names(ps) <- 1:nrow(ps)</pre>
   sentinels(ps, "OPG", 1, log_p="logp")
 ## to obtain variance explained
 tbl <- within(OPGtbl, chi2n <- (Effect/StdErr)^2/N)
 s <- with(tbl, aggregate(chi2n,list(prot),sum))</pre>
 names(s) \leftarrow c("prot", "h2")
 sd <- with(tbl, aggregate(chi2n,list(prot),sd))</pre>
 names(sd) <- c("p1","sd")
 m <- with(tbl, aggregate(chi2n,list(prot),length))</pre>
 names(m) <- c("p2","m")
h2 <- cbind(s,sd,m)
 ord <- with(h2, order(h2))
 sink("h2.dat")
 print(h2[ord, c("prot","h2","sd","m")], row.names=FALSE)
 sink()
png("h2.png", res=300, units="in", width=12, height=8)
np <- nrow(h2)
 with(h2[ord,], {
     plot(h2, cex=0.4, pch=16, xaxt="n", xlab="protein", ylab=expression(h^2))
     xtick \leftarrow seq(1, np, by=1)
     axis(side=1, at=xtick, labels = FALSE)
```

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```
text(x=xtick, par("usr")[3],labels = prot, srt = 75, pos = 1, xpd = TRUE, cex=0.5)
})
dev.off()
write.csv(tbl,file="INF1.csv",quote=FALSE,row.names=FALSE)
## End(Not run)
```

snpHWE

Functions for single nucleotide polymorphisms

# **Description**

These are a set of functions specifically for single nucleotide polymorphisms (SNPs), which are biallelic markers. This is particularly relevant to the genomewide association studies (GWAS) using GeneChips and in line with the classic generalised single-locus model. snpHWE is from Abecasis's website and yet to be adapted for chromosome X.

### Usage

```
snpHWE(g)
PARn(p, RRlist)
snpPVE(beta, se, N)
snpPAR(RR, MAF, unit = 2)
```

### **Arguments**

g	Observed genotype vector.
p	genotype frequencies.
RRlist	A list of RRs.
beta	Regression coefficient.
se	Standard error for beta.
N	Sample size.
RR	Relative risk.
MAF	Minar allele frequency.
unit	Unit to exponentiate for homozygote.

### **Details**

snpHWE gives an exact Hardy-Weinberg Equilibrium (HWE) test and it return -1 in the case of misspecification of genotype counts.

snpPAR calculates the the population attributable risk (PAR) for a particular SNP. Internally, it calls for an internal function PARn, given a set of frequencies and associate relative risks (RR). Other 2x2 table statistics familiar to epidemiologists can be added when necessary.

snpPVE provides proportion of variance explained (PVE) estimate based on the linear regression coefficient and standard error. For logistic regression, we can have similar idea for log(OR) and log(SE(OR)).

snptest\_sample

### Author(s)

Jing Hua Zhao, Shengxu Li

snptest\_sample

A utility to generate SNPTEST sample file

# **Description**

A utility to generate SNPTEST sample file

# Usage

```
snptest_sample(
  data,
  sample_file = "snptest.sample",
  ID_1 = "ID_1",
  ID_2 = "ID_2",
  missing = "missing",
  C = NULL,
  D = NULL,
  P = NULL
)
```

# Arguments

```
Data to be used.
data
sample_file
                  Output filename.
ID_1
                  ID_1 as in the sample file.
ID_2
                  ID_2 as in the sample file.
                  Missing data column.
missing
С
                  Continuous variables.
D
                  Discrete variables.
Р
                  Phenotypic variables.
```

# Value

Output file in SNPTEST's sample format.

```
## Not run:
d <- data.frame(ID_1=1,ID_2=1,missing=0,PC1=1,PC2=2,D1=1,P1=10)
snptest_sample(d,C=paste0("PC",1:2),D=paste0("D",1:1),P=paste0("P",1:1))
## End(Not run)</pre>
```

tscc 147

tscc

Power calculation for two-stage case-control design

# Description

Power calculation for two-stage case-control design

# Usage

```
tscc(model, GRR, p1, n1, n2, M, alpha.genome, pi.samples, pi.markers, K)
```

# **Arguments**

model	any in c("multiplicative", "additive", "dominant", "recessive").
GRR	genotype relative risk.
p1	the estimated risk allele frequency in cases.
n1	total number of cases.
n2	total number of controls.
М	total number of markers.
alpha.genome	false positive rate at genome level.
pi.samples	sample% to be genotyped at stage 1.
pi.markers	markers% to be selected (also used as the false positive rate at stage 1).
K	the population prevalence.

# **Details**

This function gives power estimates for two-stage case-control design for genetic association.

The false positive rates are calculated as follows,

$$P(|z1| > C1)P(|z2| > C2, sign(z1) = sign(z2))$$

and

for replication-based and joint analyses, respectively; where C1, C2, and Cj are threshoulds at stages 1, 2 replication and joint analysis,

$$z1 = z(p1, p2, n1, n2, pi.samples)$$
 
$$z2 = z(p1, p2, n1, n2, 1 - pi.samples)$$
 
$$zj = sqrt(pi.samples) * z1 + sqrt(1 - pi.samples) * z2$$

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#### Value

The returned value is a list containing a copy of the input plus output as follows,

- model any in c("multiplicative", "additive", "dominant", "recessive").
- GRR genotype relative risk.
- p1 the estimated risk allele frequency in cases.
- pprime expected risk allele frequency in cases.
- p expected risk allele frequency in controls.
- n1 total number of cases.
- n2 total number of controls.
- M total number of markers.
- alpha.genome false positive rate at genome level.
- pi.samples sample% to be genotyped at stage 1.
- pi.markers markers% to be selected (also used as the false positive rate at stage 1).
- K the population prevalence.
- C threshoulds for no stage, stage 1, stage 2, joint analysis.
- power power corresponding to C.

#### Note

solve. skol is adapted from CaTS.

### Author(s)

Jing Hua Zhao

### References

Skol AD, Scott LJ, Abecasis GR, Boehnke M (2006). "Joint analysis is more efficient than replication-based analysis for two-stage genome-wide association studies." *Nat Genet*, **38**(2), 209-13. ISSN 1061-4036 (Print) 1061-4036, doi:10.1038/ng1706.

```
## Not run:
K <- 0.1
p1 <- 0.4
n1 <- 1000
n2 <- 1000
M <- 300000
alpha.genome <- 0.05
GRR <- 1.4
p1 <- 0.4
pi.samples <- 0.2
pi.markers <- 0.1

options(echo=FALSE)
cat("sample%,marker%,GRR,(thresholds x 4)(power estimates x 4)","\n")
for(GRR in c(1.3,1.35,1.40))
{
    cat("\n")</pre>
```

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```
for(pi.samples in c(1.0,0.5,0.4,0.3,0.2))
     if(pi.samples==1.0) s <- 1.0</pre>
     else s <- c(0.1,0.05,0.01)
     for(pi.markers in s)
       x <- tscc("multiplicative",GRR,p1,n1,n2,M,alpha.genome,</pre>
                pi.samples,pi.markers,K)
       1 <- c(pi.samples,pi.markers,GRR,x$C,x$power)</pre>
       1[1],1[2],1[3],1[4],1[5],1[6],1[7],1[8],1[9],1[10],1[11])
       cat(1,"\n")
     }
     cat("\n")
  }
}
options(echo=TRUE)
## End(Not run)
```

whscore

Whittemore-Halpern scores for allele-sharing

# Description

Whittemore-Halpern scores for allele-sharing

# Usage

```
whscore(allele, type)
```

# **Arguments**

allele a matrix of alleles of affected pedigree members.

type 0 = pairs, 1 = all.

# **Details**

Allele sharing score statistics.

### Value

The returned value is the value of score statistic.

### Note

adapted from GENEHUNTER.

# Author(s)

Leonid Kruglyak, Jing Hua Zhao

WriteGRMBin

### References

Kruglyak L, Daly MJ, Reeve-Daly MP, Lander ES (1996). "Parametric and nonparametric linkage analysis: a unified multipoint approach." *Am J Hum Genet*, **58**(6), 1347-63. ISSN 0002-9297 (Print) 0002-9297.

Whittemore AS, Halpern J (1994). "A class of tests for linkage using affected pedigree members." *Biometrics*, **50**(1), 118-27. ISSN 0006-341X (Print) 0006-341x.

Whittemore AS, Halpern J (1994). "Probability of gene identity by descent: computation and applications." *Biometrics*, **50**(1), 109-17. ISSN 0006-341X (Print) 0006-341x.

# **Examples**

```
## Not run:
c<-matrix(c(1,1,1,2,2,2),ncol=2)
whscore(c,type=1)
whscore(c,type=2)
## End(Not run)</pre>
```

WriteGRM

A function to write GRM file

# Description

A function to write GRM file

# Usage

```
WriteGRM(prefix = 51, id, N, GRM)
```

# **Arguments**

prefix file root.
id id.

 ${\sf N}$  sample size.  ${\sf GRM}$  a  ${\sf GRM}$ .

WriteGRMBin

A function to write GRM binary file

# **Description**

A function to write GRM binary file

### Usage

```
WriteGRMBin(prefix, grm, N, id, size = 4)
```

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# Arguments

prefix file root.
grm a GRM.
N Sample size.
id id.
size size.

ху

Conversion of chromosome names to strings

# Description

Conversion of chromosome names to strings

# Usage

xy(x)

# Arguments

Х

(alpha)numeric value indicating chromosome.

# **Details**

This function converts x=1:24 to 1:22, X, Y

# Value

As indicated.

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