# COLT 2021 RL Theory Tutorial: Solutions

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## Solutions for Natural Policy Gradient Exercises

## 1 Closed form NPG update

First, observe that

$$\nabla_{\theta} \log \pi_{\theta}(a \mid s) = e_{s,a} - \sum_{a'} e_{s,a'} \pi_{\theta}(a' \mid s). \tag{1}$$

which implies that

$$\forall s : \mathbb{E}_{a \sim \pi_{\theta}(\cdot \mid s)} \left[ \nabla \log \pi_{\theta}(a \mid s) \right] = 0$$

Next, by the definition of the Moore-Penrose pseudoinverse  $(F_{\rho}^{\theta})^{\dagger}\nabla V^{\pi_{\theta}}(\rho)$  is equal to the minimum norm solution of

$$\min_{w} \|\nabla V^{\pi_{\theta}}(\rho) - F_{\rho}^{\theta} w\|_{2}^{2}$$

Let us calculate this latter matrix vector product:

$$F_{\rho}^{\theta} w = \mathbb{E}_{s \sim d_{\rho}^{\pi_{\theta}}} \mathbb{E}_{a \sim \pi_{\theta}(\cdot \mid s)} \left[ \nabla \log \pi_{\theta}(a \mid s) (w^{\top} \nabla \log \pi_{\theta}(a \mid s)) \right]$$

$$= \mathbb{E}_{s \sim d_{\rho}^{\pi_{\theta}}} \mathbb{E}_{a \sim \pi_{\theta}(\cdot \mid s)} \left[ \nabla \log \pi_{\theta}(a \mid s) (w_{s,a} - \mathbb{E}_{a' \sim \pi(\cdot \mid s)} w_{s,a'}) \right]$$

$$= \mathbb{E}_{s \sim d_{\rho}^{\pi_{\theta}}} \mathbb{E}_{a \sim \pi_{\theta}(\cdot \mid s)} \left[ w_{s,a} \cdot \nabla \log \pi_{\theta}(a \mid s) \right]$$

Next, we use the advantage version of the policy gradient theorem to write:

$$\nabla V^{\pi_{\theta}}(\rho) = \frac{1}{1 - \gamma} \mathbb{E}_{s \sim d_{\rho}^{\pi_{\theta}}} \mathbb{E}_{a \sim \pi_{\theta}(\cdot \mid s)} \left[ A^{\pi_{\theta}}(s, a) \nabla \log \pi_{\theta}(a \mid s) \right]$$

Intuitively, at this point we can see that  $w_{s,a} = A^{\pi_{\theta}}(s,a)/(1-\gamma)$  is a solution to the least squares problem, which is quite close to proving the result.

To be more formal, using (1) again, we see that the  $(s,a)^{th}$  element of both of these vectors are

$$[F_{\rho}^{\theta}w]_{s,a} = d_{\rho}^{\pi_{\theta}}(s)\pi_{\theta}(a \mid s) \left(w_{s,a} - \sum_{a'} w_{s,a'}\pi_{\theta}(a' \mid s)\right),$$
$$[\nabla V^{\pi_{\theta}}(\rho)]_{s,a} = d_{\rho}^{\pi_{\theta}}(s)\pi_{\theta}(a \mid s) \left(A^{\pi_{\theta}}(s,a)\right).$$

Here we are using that  $\mathbb{E}_{a \sim \pi_{\theta}(\cdot|s)} A^{\pi_{\theta}}(s, a) = 0$ . Now we can more clearly see that all solutions with 0 square loss are of the form  $w_{s,a} = A^{\pi_{\theta}}(s,a) + v_s$  where the second term depends only on the state (it is constant across actions for that state). This proves the claim regarding the update for  $\theta$  and the update for  $\pi_{\theta}$  follows immediately, since the state-dependent offset can be absorbed into the normalization term.

#### 2 Performance difference lemma

The proof is based on an un-rolling argument. Observe that

$$\begin{split} V^{\pi_1}(\rho) - V^{\pi_2}(\rho) &= \mathbb{E}_{s \sim \rho} \left[ \mathbb{E}_{a \sim \pi_1(\cdot \mid s)} Q^{\pi_1}(s, a) - V^{\pi_2}(s) \right] \\ &= \mathbb{E}_{s \sim \rho} \left[ \mathbb{E}_{a \sim \pi_1(\cdot \mid s)} (Q^{\pi_1}(s, a) - Q^{\pi_2}(s, a)) \right] + \mathbb{E}_{s \sim \rho} \left[ \mathbb{E}_{a \sim \pi_1(\cdot \mid s)} Q^{\pi_2}(s, a) - V^{\pi_2}(s) \right] \\ &= \mathbb{E}_{s \sim \rho} \left[ \mathbb{E}_{a \sim \pi_1(\cdot \mid s)} (Q^{\pi_1}(s, a) - Q^{\pi_2}(s, a)) \right] + \mathbb{E}_{s, a \sim \rho \circ \pi_1} \left[ A^{\pi_2}(s, a) \right] \end{split}$$

Now the first term gives the difference value starting from the second state visited by  $\pi_1$ :

$$\mathbb{E}_{s \sim \rho} \left[ \mathbb{E}_{a \sim \pi_1(\cdot | s)} (Q^{\pi_1}(s, a) - Q^{\pi_2}(s, a)) \right] = \gamma \mathbb{E}_{s, a, s' \sim \rho \circ \pi_1} V^{\pi_1}(s') - V^{\pi_2}(s')$$

Applying the same argument as above, we can express this in terms of the advantage function and value starting from the third state. To express this more concisely, let  $P_{\rho}^{\pi_1}$  be the distribution over infinitely long trajectories  $\tau = (s_0, a_0, s_1, a_1, \ldots)$  sampled by starting from  $\rho$  and taking actions according to  $\pi_1$ . With this notation we have

$$V^{\pi_1}(\rho) - V^{\pi_2}(\rho) = \mathbb{E}_{\tau \sim P_{\rho}^{\pi_1}} \left[ \sum_{t=0}^{\infty} \gamma^t A^{\pi_2}(s_t, a_t) \right] = \frac{1}{1 - \gamma} \mathbb{E}_{s, a \sim d_{\rho}^{\pi_1}} \left[ A^{\pi_2}(s, a) \right]$$

## 3 NPG regret analysis

We use a potential function argument, where our potential is the KL divergence between the comparator  $\tilde{\pi}$  and our iterate  $\tilde{\pi}^{(t)}$  on the distribution induced by  $\tilde{\pi}$ . Using smoothness, we have

$$\begin{split} &\mathbb{E}_{s \sim d_{\rho}^{\tilde{\pi}}} \left[ \text{KL}(\tilde{\pi}(\cdot \mid s) || \pi^{(t)}(\cdot \mid s)) - \text{KL}(\tilde{\pi}(\cdot \mid s) || \pi^{(t+1)}(\cdot \mid s)) \right] \\ &= \mathbb{E}_{s,a \sim d_{\rho}^{\tilde{\pi}}} \left[ \log \left( \frac{\tilde{\pi}(a \mid s)}{\pi^{(t)}(a \mid s)} \right) - \log \left( \frac{\tilde{\pi}(a \mid s)}{\pi^{(t+1)}(a \mid s)} \right) \right] \\ &= \mathbb{E}_{s,a \sim d_{\rho}^{\tilde{\pi}}} \left[ \log \pi^{(t+1)}(a \mid s) - \log \pi^{(t)}(a \mid s) \right] \\ &\geq \mathbb{E}_{s,a \sim d_{\rho}^{\tilde{\pi}}} \left[ \left\langle \nabla \log \pi^{(t)}(a \mid s), \theta^{(t+1)} - \theta^{(t)} \right\rangle - \frac{\beta}{2} \|\theta^{(t+1)} - \theta^{(t)}\|_{2}^{2} \right] \\ &= \mathbb{E}_{s,a \sim d_{\rho}^{\tilde{\pi}}} \left[ \eta \left\langle \nabla \log \pi^{(t)}(a \mid s), w^{(t)} \right\rangle - \frac{\eta^{2}\beta}{2} \|w^{(t)}\|_{2}^{2} \right] \\ &= \eta \mathbb{E}_{s,a \sim d_{\rho}^{\tilde{\pi}}} \left[ A^{(t)}(s,a) \right] - \frac{\eta^{2}\beta}{2} \|w^{(t)}\|_{2}^{2} - \eta \cdot \text{err}_{t} \\ &= (1 - \gamma) \eta \left( V^{\tilde{\pi}}(\rho) - V^{(t)}(\rho) \right) - \frac{\eta^{2}\beta}{2} \|w^{(t)}\|_{2}^{2} - \eta \cdot \text{err}_{t}. \end{split}$$

The only inequality here is the lower bound implied by our smoothness assumption on  $\pi_{\theta}(a \mid s)$ . The last equality is the performance difference lemma.

Now we can obtain a telescoping sum involving the KL divergences:

$$\begin{split} & \min_{0 \leq t < T} \left( V^{\tilde{\pi}}(\rho) - V^{(t)}(\rho) \right) \leq \frac{1}{T} \sum_{t=0}^{T-1} \left( V^{\tilde{\pi}}(\rho) - V^{(t)}(\rho) \right) \\ & \leq \frac{1}{1-\gamma} \frac{1}{T} \sum_{t=0}^{T-1} \left[ \frac{1}{\eta} \mathbb{E}_{s \sim d_{\rho}^{\tilde{\pi}}} \left[ \text{KL}(\tilde{\pi}(\cdot \mid s) || \pi^{(t)}(\cdot \mid s)) - \text{KL}(\tilde{\pi}(\cdot \mid s) || \pi^{(t+1)}(\cdot \mid s)) \right] + \frac{\eta \beta}{2} || w^{(t)} ||_{2}^{2} + \text{err}_{t} \right] \\ & \leq \frac{1}{1-\gamma} \left( \frac{\log |\mathcal{A}|}{T\eta} + \frac{\eta \beta W}{2} + \frac{1}{T} \sum_{t=0}^{T-1} \text{err}_{t} \right). \end{split}$$

## 4 NPG Error Analysis (Sketch)

There are two steps remaining in the analysis of NPG with the softmax policy class in tabular settings. These steps involve controlling the err<sub>t</sub> terms above and highlight how the initial/exploratory distribution  $\rho$  must ensure sufficient exploration.

The first step is fit  $w^{(t)}$  somehow. As discussed previously, we want  $w^{(t)}(s,a) \approx Q^{(t)}(s,a)$ . Since we can sample from  $\rho$ , a natural population objective is

$$L_t(w) = \mathbb{E}_{s \sim \rho, a \sim \text{unif}(\mathcal{A})} \left[ (Q^{(t)}(s, a) - w(s, a))^2 \right]. \tag{2}$$

We will optimize this objective from samples, in the usual way. We can get unbiased estimates of  $Q^{(t)}(s, a)$  where  $(s, a) \sim \rho \circ \operatorname{unif}(\mathcal{A})$  by (1) sampling the initial state/action and subsequently executing  $\pi^{(t)}$ , (2) terminating the episode at each time t with probability  $1 - \gamma$ , and (3) reporting the undiscounted sum of rewards up to termination. Call this random variable  $\hat{R}$  and let  $t^*$  denote the time step that we terminate. Then for any (s, a) pair sampled from  $\rho \circ \operatorname{unif}(\mathcal{A})$ , linearity of expectation gives:

$$\mathbb{E}_{\pi}[\hat{R} \mid s, a] = \mathbb{E}_{\pi}\left[\sum_{T=1}^{\infty} \mathbf{1}\{t^{\star} = T\} \sum_{t=1}^{T} r_{t} \mid s, a\right] = \mathbb{E}_{\pi}\left[\sum_{t=1}^{\infty} r_{t} \mathbf{1}\{t^{\star} \geq t\} \mid s, a\right] = \mathbb{E}_{\pi}\left[\sum_{t=1}^{\infty} \gamma^{t} r_{t} \mid s, a\right] = Q^{\pi}(s, a)$$

With this procedure and since we are using the tabular representation, we can ensure that  $L_t(w^{(t)}) \lesssim 1/N$  if we collect N rollouts.

The last step is to use this guarantee to ensure that  $\operatorname{err}_t$  is small. The main conceptual point is that we have to perform a distribution shift from  $\rho \circ \operatorname{unif}$  to  $d_{\rho}^{\tilde{\pi}}$ . With an importance weighting argument, we can obtain

$$\operatorname{err}_{t} := \mathbb{E}_{(s,a) \sim d_{\rho}^{\tilde{\pi}}} \left[ A^{(t)}(s,a) - \left\langle \nabla \log \pi^{(t)}(a \mid s), w^{(t)} \right\rangle \right]$$

$$= \mathbb{E}_{(s,a) \sim d_{\rho}^{\tilde{\pi}}} \left[ Q^{(t)}(s,a) - \mathbb{E}_{a' \sim \pi^{(t)}(\cdot \mid s)} Q^{(t)}(s,a') - w^{(t)}(s,a) + \mathbb{E}_{a' \sim \pi^{(t)}(\cdot \mid s)} w^{(t)}(s,a') \right]$$

$$\leq 2 \sqrt{|\mathcal{A}| \mathbb{E}_{s \sim d_{\rho}^{\tilde{\pi}}, a \sim \operatorname{unif}(\mathcal{A})} \left[ (Q^{(t)}(s,a) - w^{(t)}(s,a))^{2} \right]}$$

$$\leq 2 \sqrt{|\mathcal{A}| \sum_{s} \frac{d_{\rho}^{\tilde{\pi}}(s)}{\rho(s)} \cdot \rho(s) \mathbb{E}_{a \sim \operatorname{unif}(\mathcal{A})} \left[ (Q^{(t)}(s,a) - w^{(t)}(s,a))^{2} \right]}$$

$$\leq 2 \sqrt{|\mathcal{A}| \cdot \left\| \frac{d_{\rho}^{\tilde{\pi}}}{\rho} \right\|_{\infty} \cdot L_{t}(w^{(t)})}.$$

The density ratio term measures how well  $\rho$  covers the states visited by the comparator policy  $\tilde{\pi}$ . This highlights the role of the reset distribution in providing good coverage for policy gradient methods.

Remark 1 (Omitted details). There are some details we are glossing over here. First, while the sample-based estimates of  $Q^{(t)}(s,a)$  are bounded with high probability, they are technically unbounded, so the least squares generalization argument needs to account for this. Second, we need to set the smoothness parameter  $\beta$  and the norm bound W appropriately. For W, we can add a constraint in the least squares problem, since  $Q^{(t)}(s,a) \in [0,\frac{1}{1-\gamma}]$ . For  $\beta$ , since we are in the tabular representation, one can verify that  $\beta = 1$  suffices.

## Solutions for UCB-VI Exercises

#### 1 Prove Simulation lemma

By definition, we have:

$$\begin{split} V_0^{\pi} - \widehat{V}_0^{\pi} &= \mathbb{E}_{s \sim \mu, a \sim \pi(\cdot \mid s_0)} \left[ Q_0^{\pi}(s, a) - \widehat{Q}_0^{\pi}(s, a) \right] \\ &= \mathbb{E}_{s \sim \mu, a \sim \pi(\cdot \mid s_0)} \left[ r_0(s, a) - \widehat{r}_0(s, a) + \mathbb{E}_{s' \sim P_0(s, a)} V_1^{\pi}(s') - \mathbb{E}_{s' \sim \widehat{P}_0(s, a)} \widehat{V}_1^{\pi}(s') \right] \\ &= \mathbb{E}_{s \sim \mu, a \sim \pi(\cdot \mid s_0)} \left[ r_0(s, a) - \widehat{r}_0(s, a) + \mathbb{E}_{s' \sim P_0(s, a)} \widehat{V}_1^{\pi}(s') - \mathbb{E}_{s' \sim \widehat{P}_0(s, a)} \widehat{V}_1^{\pi}(s') + \mathbb{E}_{s' \sim P_0(s, a)} V_1^{\pi}(s') - \mathbb{E}_{s' \sim P_0(s, a)} \widehat{V}_1^{\pi}(s') \right] \\ &= \mathbb{E}_{s \sim \mu, a \sim \pi(\cdot \mid s_0)} \left[ r_0(s, a) - \widehat{r}_0(s, a) + \mathbb{E}_{s' \sim P_0(s, a)} \widehat{V}_1^{\pi}(s') - \mathbb{E}_{s' \sim \widehat{P}_0(s, a)} \widehat{V}_1^{\pi}(s') \right] + \mathbb{E}_{s_1 \sim d_1^{\pi}} \left[ V_1^{\pi}(s') - \widehat{V}_1^{\pi}(s') \right] \end{split}$$

Now recursively apply the same argument on the second term of the RHS of the above equation H times, and we can conclude the lemma.

#### 2 Prove Optimism

We prove optimism via induction. In the base case, we consider the fictitious step H where we have  $\hat{V}_H(s) = 0$  and  $V_H^{\star}(s) = 0$  for all s. Thus, the base case holds.

The inductive hypothesis is that at time step h+1 with  $h \leq H-1$ , we have optimism, i.e.,  $\widehat{V}_{h+1}(s) \geq V_{h+1}^{\star}(s), \forall s$ . We will prove that at time step h, we have optimism as well.

Recall the update procedure for  $\widehat{V}_h$ . When  $\widehat{V}_h(s) = H$ , we have  $\widehat{V}_h(s) = H \ge V_h^*(s)$  (since  $\|V_h^*\|_{\infty} \le H$ ). Thus, we have:

$$\begin{split} \widehat{Q}_{h}(s,a) - Q_{h}^{\star}(s,a) &= r(s,a) + b_{h}(s,a) + \mathbb{E}_{s' \sim \widehat{P}_{h}(s,a)} \widehat{V}_{h+1}(s') - \left( r(s,a) + \mathbb{E}_{s' \sim P_{h}(s,a)} V_{h+1}^{\star}(s') \right) \\ &= b_{h}(s,a) + \mathbb{E}_{s' \sim \widehat{P}_{h}(s,a)} \widehat{V}_{h+1}(s') - \mathbb{E}_{s' \sim P_{h}(s,a)} V_{h+1}^{\star}(s') \\ &\geq b_{h}(s,a) + \mathbb{E}_{s' \sim \widehat{P}_{h}(s,a)} V_{h+1}^{\star}(s') - \mathbb{E}_{s' \sim P_{h}(s,a)} V_{h+1}^{\star}(s') \\ &\geq b_{h}(s,a) - \left| \mathbb{E}_{s' \sim \widehat{P}_{h}(s,a)} V_{h+1}^{\star}(s') - \mathbb{E}_{s' \sim P_{h}(s,a)} V_{h+1}^{\star}(s') \right| \geq 0, \end{split}$$

where the first inequality uses the inductive hypothesis that  $\widehat{V}_{h+1}(s) \geq V_{h+1}^{\star}(s), \forall s$  and the fact that  $\widehat{P}(s'|s,a) \geq 0$  for all s,a,s', and the last inequality uses the given assumption on the bonuses  $b_h(s,a)$ . Note that the above derivation holds for any (s,a), so we have proved that  $\widehat{Q}_h$  is optimistic. This immediately implies that  $\widehat{V}_h$  is also optimistic, since, if we let  $a^{\star}(s) = \operatorname{argmax}_a Q_h^{\star}(s,a)$  we have

$$V_h^{\star}(s) = Q_h^{\star}(s, a^{\star}(s)) \le \widehat{Q}_h(s, a^{\star}(s)) \le \max_{a} \widehat{Q}_h(s, a) \le \widehat{V}_h(s),$$

which verifies the inductive claim.

#### 3 Regret Decomposition

Since  $\widehat{V}_h(s) \geq V_h^{\star}(s)$ , we can upper bound the regret as follows:

$$V^{\star} - V^{\widehat{\pi}} \le \mathbb{E}_{s \sim \mu} \left[ \widehat{V}_0(s) - V_0^{\widehat{\pi}}(s) \right]$$

Note here that  $\widehat{V}_0(s)$  is the value of policy  $\widehat{\pi}$  in the bonus augmented MDP  $\widetilde{\mathcal{M}}$ , while  $V_0^{\widehat{\pi}}(s)$  is the value of policy  $\widehat{\pi}$  in the true MDP  $\mathcal{M}$ . Thus, we are in a position to apply an argument similar to the simulation lemma.

$$\begin{split} \mathbb{E}_{s \sim \mu} \left[ \widehat{V}_0(s) - V_0^{\widehat{\pi}}(s) \right] &= \mathbb{E}_{s \sim \mu} \left[ \widehat{Q}_0(s, \widehat{\pi}(s)) - Q_0^{\widehat{\pi}}(s, \widehat{\pi}(s)) \right] \\ &\leq \mathbb{E}_{s \sim \mu} \left[ b_0(s, \widehat{\pi}(s)) + \mathbb{E}_{s' \sim \widehat{P}_0(\cdot \mid s, \widehat{\pi}(s))} \widehat{V}_1(s') - \mathbb{E}_{s' \sim P_0(s, \widehat{\pi}(s))} V_1^{\widehat{\pi}}(s') \right] \\ &= \mathbb{E}_{s \sim \mu} \left[ b_0(s, \widehat{\pi}(s)) + \mathbb{E}_{s' \sim P_0(\cdot \mid s, \widehat{\pi}(s))} \widehat{V}_1(s') - \mathbb{E}_{s' \sim P_0(s, \widehat{\pi}(s))} V_1^{\widehat{\pi}}(s') + \mathbb{E}_{s' \sim \widehat{P}_0(\cdot \mid s, \widehat{\pi}(s))} \widehat{V}_1(s') - \mathbb{E}_{s' \sim P_0(\cdot \mid s, \widehat{\pi}(s))} \widehat{V}_1(s') \right] \\ &\leq \mathbb{E}_{s \sim \mu} \left[ b_0(s, \widehat{\pi}(s)) + H \left\| \widehat{P}_0(s, \widehat{\pi}(s)) - P_0(s, \widehat{\pi}(s)) \right\|_1 \right] + \mathbb{E}_{s \sim d_1^{\widehat{\pi}}} \left[ \widehat{V}_1(s) - V_1^{\widehat{\pi}}(s) \right] \end{split}$$

Now recursively apply the same procedure on the second term of RHS of the above inequality H times, and we can conclude the proof. Note that this is basically the simulation lemma derivation except that in the first inequality above, we use the fact that  $\widehat{Q}_h(s,a) = \min \left\{ H, b_h(s,a) + r_h(s,a) + \mathbb{E}_{s' \sim \widehat{P}_h(s,a)} \widehat{V}_{h+1}(s') \right\}$  and the inequality that  $\min\{a,b\} \leq b$ .

## 4 Proving UCB-VI has valid bonus

Starting from this section, to simplify, we will use  $\lesssim$  to suppress absolute constants. First, via standard concentration on discrete distributions (e.g., Proposition A.4 in the AJKS monograph), for a fixed (s, a, h, t) tuple, we must have that with probability at least  $1 - \delta$ ,

$$\left\|\widehat{P}_{t,h}(s,a) - P_h(s,a)\right\|_1 \lesssim \sqrt{\frac{S\ln(1/\delta)}{N_{t,h}(s,a)}}.$$

Via a union bound over all  $s \in \mathcal{S}, a \in \mathcal{A}, h \in [H], t \in [N]$ , we must have that with probability at least  $1 - \delta$ :

$$\left\|\widehat{P}_{t,h}(s,a) - P_h(s,a)\right\|_1 \lesssim \sqrt{\frac{S\ln(SAHN/\delta)}{N_{t,h}(s,a)}}.$$

Now let us consider bounding  $\widehat{P}_{t,h}(s,a)^{\top}V_{h+1}^{\star} - P_h(s,a)^{\top}V_{h+1}^{\star}$  (note that here we abuse notation a bit and treat  $P_h(s,a)$  and  $V_h^{\star}$  as vectors in  $\mathbb{R}^{|S|}$ ). We have:

$$\widehat{P}_{t,h}(s,a)^{\top}V_{h+1}^{\star} - P_h(s,a)^{\top}V_{h+1}^{\star} = \frac{1}{N_{t,h}(s,a)} \sum_{i=1}^{t-1} \mathbf{1}\{(s_{i,h},a_{i,h}) = (s,a)\}V_{h+1}^{\star}(s_{i,h+1}) - P_h(s,a)^{\top}V_{h+1}^{\star},$$

and note that for iteration i where  $s_{i,h}, a_{i,h} = s, a$ , we have  $V_{h+1}^{\star}(s_{i,h+1})$  being an unbiased estimate of  $P_h(s, a)^{\top}V_{h+1}^{\star}$ . Also note that  $\|V_{h+1}^{\star}\|_{\infty} \leq H$  which implies that our random variables are bounded in [0, H]. Thus, we can apply the standard Hoeffding's inequality. For a fixed (s, a, h, t) tuple, with probability at least  $1 - \delta$ , we have:

$$\left| \widehat{P}_{t,h}(s,a)^{\top} V_{h+1}^{\star} - P_{h}(s,a)^{\top} V_{h+1}^{\star} \right| \lesssim H \sqrt{\frac{\ln(1/\delta)}{N_{t,h}(s,a)}}$$

Again with union bound over all s, a, h, t tuples, we immediately have that with probability at least  $1 - \delta$ ,

$$\left| \widehat{P}_{t,h}(s,a)^{\top} V_{h+1}^{\star} - P_h(s,a)^{\top} V_{h+1}^{\star} \right| \lesssim H \sqrt{\frac{\ln(SAHN/\delta)}{N_{t,h}(s,a)}}$$

This concludes the proof.

## 5 Concluding the proof: bounding the confidence sum

We proceed to upper bound the confidence sum as follows:

$$\sum_{t=1}^{N} \sum_{h=0}^{H-1} \frac{1}{\sqrt{N_{t,h}(s_{t,h}, a_{t,h})}} = \sum_{h=0}^{H-1} \sum_{s,a} \sum_{i=1}^{N_{N,h}(s,a)} \frac{1}{\sqrt{i}} \le \sum_{h=0}^{H-1} \sum_{s,a} \sqrt{N_{N,h}(s,a)}$$
$$\le \sum_{h=0}^{H-1} \sqrt{SA} \sum_{s,a} N_{N,h}(s,a) = \sum_{h=0}^{H-1} \sqrt{SAN} = H\sqrt{SAN}.$$

Here in the first inequality, we use that  $\sum_{i=1}^{n} 1/\sqrt{i} \leq \sqrt{n}$ , and the second inequality is by Cauchy-Schwarz. This concludes the proof.