

[과제 #4]

20213064_김종민

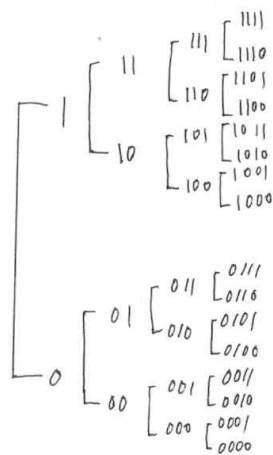
Assignment #4

problems 2.35

A random hexadecimal character in the form of four binary digits is read from a storage device.

(a) Draw the tree diagram and the sample space for this experiment.

↳



sample space =

{0000, 0001, 0010, 0011, 0100, 0101, 0110, 0111, 1000, 1001, 1010, 1011, 1100, 1101, 1110, 1111}

(b) Given that the first bit is zero, what is the probability of more zeros than ones?

↳ 첫번째 비트가 0인 경우 = {0111, 0110, 0101, 0100, 0011, 0010, 0001, 0000} ... 8개

위의 집합에서 1보다 0이 더 많은 경우 = {0100, 0010, 0001, 0000} ... 4개

$$\therefore \text{probability is } \frac{4}{8} = \frac{1}{2}$$

(c) Given that the first two bits are 10, what is the probability of more zeros than ones?

↳ 시작 비트가 10인 경우 = {1011, 1010, 1001, 1000} ... 4개

위의 집합에서 1보다 0이 더 많은 경우 = {1000} ... 1개

$$\therefore \text{probability is } \frac{1}{4}$$

(d) Given that there are more zeros than ones, what is the probability that the first bit is a zero?

↳ 1보다 0이 더 많은 경우 = {1000, 0100, 0010, 0001, 0000} ... 5개

위의 집합에서 시작 비트가 0인 경우 = {0100, 0010, 0001, 0000} ... 4개

$$\therefore \text{probability is } \frac{4}{5}$$

problems 7.10 For each of the following random variables =

(i) geometric, $p = 1/8$

(ii) uniform, $m=0, n=1$

(iii) Poisson, $\alpha = 2$

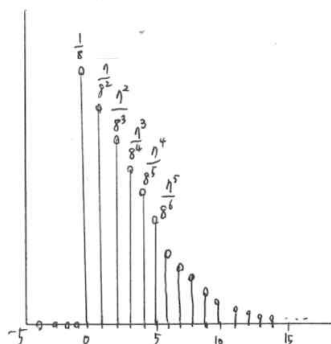
(a) sketch the PMF

↳ (i) geometric PMF

$$f_k(k) = p(1-p)^k \quad 0 \leq k < \infty$$

$$p = \frac{1}{8}$$

$$f_k(k) = \frac{1}{8} \left(\frac{7}{8}\right)^k$$



$(1-p)^6 \leq 0.5$ 를 만족하는 p 를 찾으려 한다.

$(1-p)^6 = 0.5$ 라 할때, 이를 만족하는 p 는 0.1091이다.

$$\therefore p = 0.1091$$

(ii) Uniform

$$Pr(K > 5) = 1 - (f_K(0) + f_K(1) + f_K(2) + f_K(3) + f_K(4) + f_K(5))$$

uniform random variable은 parameter가 m, n 두 개입니다.

처음 문제에서 주어진 바와 같이 $m=0$ 이라고 하겠습니다.

$$Pr(K > 5) = 1 - \left(\frac{1}{n+1} + \frac{1}{n+1} + \frac{1}{n+1} + \frac{1}{n+1} + \frac{1}{n+1} + \frac{1}{n+1} \right) = 1 - \frac{6}{n+1} = \frac{1}{2}$$

$$\frac{(n+1)-6}{n+1} = \frac{1}{2}, \quad \frac{n-5}{n+1} = \frac{1}{2}, \quad n=11$$

$$\therefore m=0, n=11$$

(iii) poisson

poisson의 PMF는 앞서 구한바와 같이 $f_K(k) = \frac{\alpha^k}{k!} e^{-\alpha}$ 입니다.

$$Pr(K > 5) = 1 - (f_K(0) + f_K(1) + f_K(2) + f_K(3) + f_K(4) + f_K(5))$$

$$1 - \left(\frac{\alpha^0}{0!} e^{-\alpha} + \frac{\alpha^1}{1!} e^{-\alpha} + \frac{\alpha^2}{2!} e^{-\alpha} + \frac{\alpha^3}{3!} e^{-\alpha} + \frac{\alpha^4}{4!} e^{-\alpha} + \frac{\alpha^5}{5!} e^{-\alpha} \right) \leq \frac{1}{2}$$

$$e^{-\alpha} \left(1 + \alpha + \frac{\alpha^2}{2} + \frac{\alpha^3}{6} + \frac{\alpha^4}{24} + \frac{\alpha^5}{120} \right) = \frac{1}{2} \text{을 만족하는}$$

α 를 구하면

$$\alpha = 5.69$$

$$\therefore \alpha = 5.69$$

해를 구하는 과정의 추가적인 설명이 필요한 것 같아 위의 geometric random variable과 poisson random variable의 parameter를 구하는 파이썬 코드를 추가로 첨부합니다.

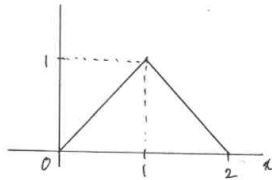
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In [4]: #geometric random variable의 해를 구하기
import numpy as np
from scipy.optimize import fsolve
# 방정식을 정의합니다
def equation(p):
    return (1 - p)**6 - 1/2
# 초기 추정치를 설정합니다
initial_guess = 0.5
# fsolve를 사용하여 방정식을 풉니다
solution = fsolve(equation, initial_guess)
# 결과를 출력합니다
p_value = solution[0]
print(f"방정식을 만족하는 p 값은: {p_value}")

방정식을 만족하는 p 값은: 0.10910128185966073
```

```
In [2]: #poisson random variable의 해를 구하기
import numpy as np
from scipy.optimize import fsolve
# 방정식 정의
def equation(x):
    return np.exp(-x) * (1 + x + x**2/2 + x**3/6 + x**4/24 + x**5/120) - 0.5
# 초기 추정값 설정
initial_guess = 0.5
# fsolve를 사용하여 방정식 풀이
solution = fsolve(equation, initial_guess)
print(f"미지수 x의 값은: {solution[0]}")

미지수 x의 값은: 5.670161188712067
```

problems 4.41 The probability density function of a random variable X is shown below:



(a) Write a set of algebraic expressions that describe the density function.

↳

$$f_X(x) = \begin{cases} x & (0 \leq x < 1) \\ 2-x & (1 \leq x < 2) \\ 0 & \text{otherwise} \end{cases}$$

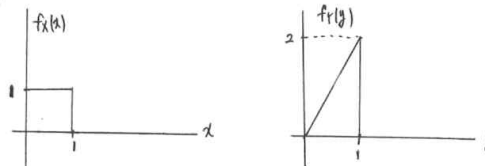
(b) Compute the mean and variance of this random variable.

$$\hookrightarrow E(X) = \int_0^2 x \cdot f_X(x) dx = \int_0^1 x \cdot x dx + \int_1^2 x(2-x) dx = \frac{1}{3} + \frac{2}{3} = 1$$

$$E(X^2) = \int_0^2 x^2 \cdot f_X(x) dx = \int_0^1 x^2 \cdot x dx + \int_1^2 x^2(2-x) dx = \frac{1}{4} + \frac{11}{12} = \frac{7}{6}$$

$$V(X) = E(X^2) - (E(X))^2 = \frac{7}{6} - 1 = \frac{1}{6} \quad \therefore E(X)=1, \quad V(X)=\frac{1}{6}$$

problems 5.29 Random variables X and Y are independent and described by the density function shown below:



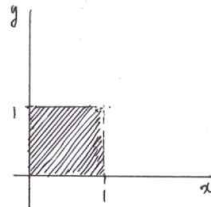
(a) Give an algebraic expression for the joint density $f_{X,Y}(x,y)$ (including limits) and sketch the region where it is nonzero below.

$$\hookrightarrow f_X(x) = \begin{cases} 1 & (0 \leq x \leq 1) \\ 0 & \text{otherwise} \end{cases} \quad f_Y(y) = \begin{cases} 2y & (0 \leq y \leq 1) \\ 0 & \text{otherwise} \end{cases}$$

확률변수 X 와 Y 가 independent 하므로 $f_{X,Y}(x,y) = f_X(x) \cdot f_Y(y)$ 가 성립한다.

$$f_{X,Y}(x,y) = 1 \cdot 2y = 2y$$

$$f_{X,Y}(x,y) = \begin{cases} 2y & (0 \leq x \leq 1) \text{ and } (0 \leq y \leq 1) \\ 0 & \text{otherwise} \end{cases}$$



nonzero region은 $(0,0), (0,1), (1,0), (1,1)$ 을 꼭짓점으로 하는 정사각형.

(b) What is $\Pr[X \leq Y]$?

$$\hookrightarrow 0 \leq X \leq Y \leq 1$$

$$\int_0^1 \int_0^y f_{X,Y}(x,y) dx dy = \int_0^1 \int_0^y 2y dx dy = 2 \int_0^1 y \int_0^y 1 \cdot dx dy = 2 \int_0^1 y^2 dy = \frac{2}{3}$$

$$\therefore \Pr[X \leq Y] = \frac{2}{3}$$

(c) What is the variance of Y ?

$$\hookrightarrow E(Y) = \int_0^1 y \cdot f_Y(y) dy = \int_0^1 2y^2 dy = \frac{2}{3}$$

$$E(Y^2) = \int_0^1 y^2 \cdot f_Y(y) dy = \int_0^1 2y^3 dy = \frac{1}{2}$$

$$V(Y) = E(Y^2) - (E(Y))^2 = \frac{1}{2} - \left(\frac{2}{3}\right)^2 = \frac{1}{18} \quad \therefore V(Y) = \frac{1}{18}$$

(d) What is the correlation coefficient ρ_{XY} ?

$$\hookrightarrow \text{Correlation coefficient } \rho_{XY} = \frac{\text{COV}(X,Y)}{\sigma_X \sigma_Y} \text{ 을 정의합니다.}$$

$$\text{COV}(X,Y) = E(XY) - m_X m_Y$$

X 와 Y 가 independent 하기 때문에 $\text{COV}(X,Y) = E(X)E(Y) - m_X m_Y$ 가 성립하고 이는 0입니다.

$$\therefore \rho_{XY} = \frac{0}{\sigma_X \sigma_Y} = 0$$

$$\therefore \text{correlation coefficient } \rho_{XY} = 0$$

problems 6.11

In a communication receiver, IID signal samples (random variables) X_i collected at a 25-element antenna are linearly combined with equal weights, i.e., $Y = \sum_{i=1}^{25} X_i$. The signal sample X_i collected at the i th antenna element is uniformly distributed over the interval $[-1, 1]$. Consider that the central Limit Theorem applies. Calculate the probability $\Pr[Y \geq 0]$.

\hookrightarrow 균등 확률변수는 구간 (a,b) 에서 평균이 $\frac{a+b}{2}$, 분산이 $\frac{(b-a)^2}{12}$ 입니다.

각각의 X_i 는 $[-1, 1]$ 구간에서 균등하게 분포하므로 X_i 의 평균 $\mu_{X_i} = \frac{a+b}{2} = \frac{-1+1}{2} = 0$ 이고

$$X_i \text{의 분산 } \sigma_{X_i}^2 = \frac{(b-a)^2}{12} = \frac{(1-(-1))^2}{12} = \frac{1}{3} \text{ 입니다.}$$

$Y = \sum_{i=1}^{25} X_i$ 이고 중심극한 정리에 의해, 많은 수의 IID 확률변수들의 합이 n 과 $n\sigma$ 를 갖는 정규분포로 근사할 수 있습니다.

$$\mu_Y = 25\mu \text{ (}\mu \text{는 } X_i \text{들의 평균)} = 25 \times 0 = 0$$

$$\sigma_Y^2 = 25\sigma^2 \text{ (}\sigma^2 \text{은 } X_i \text{들의 분산)} = 25 \times \frac{1}{3} = \frac{25}{3}$$

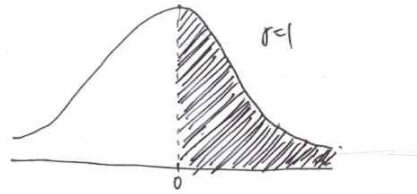
Y 는 $N(0, \frac{25}{3})$ 인 정규분포를 따릅니다.

이를 표준 정규분포로 정규화해주면

$$Z = \frac{Y - \mu_Y}{\sigma_Y} = \frac{Y - 0}{\frac{5}{\sqrt{3}}} = \frac{Y\sqrt{3}}{5}$$

$$\Pr[Y \geq 0] = \Pr\left[\frac{Y\sqrt{3}}{5} \geq 0\right] = \Pr[Z \geq 0]$$

표준 정규 분포 $Z \sim \mathcal{N}(0, 1)$ 을 따르는 PDF는
 오른쪽과 같습니다.



위와 같은 확률은 빗금친 부분이고 이는 전체 확률의 절반, $\frac{1}{2}$ 입니다.

$$\therefore \Pr[Y \geq 0] = \frac{1}{2}$$

problems 7-1 The joint PDF of two random variables X_1 and X_2 is given by

$$f_{X_1, X_2}(x_1, x_2) = \begin{cases} 1/12 & -1 \leq x_1 \leq 1, -3 \leq x_2 \leq 3 \\ 0 & \text{otherwise} \end{cases}$$

(a) Find the first moment of the random variables in vector notation.

$$\hookrightarrow E(X_1) = \int_{-1}^1 \int_{-3}^3 x_1 \cdot f_{X_1, X_2}(x_1, x_2) dx_2 dx_1 = \int_{-1}^1 \int_{-3}^3 \frac{x_1}{12} dx_2 dx_1 = \int_{-1}^1 \frac{x_1}{2} dx_1 = 0$$

$$E(X_2) = \int_{-1}^1 \int_{-3}^3 x_2 \cdot f_{X_1, X_2}(x_1, x_2) dx_2 dx_1 = \int_{-1}^1 \int_{-3}^3 \frac{x_2}{12} dx_2 dx_1 = \int_{-1}^1 0 dx_1 = 0$$

$$\therefore \underline{\underline{m_X = E(X) = \begin{pmatrix} E(X_1) \\ E(X_2) \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \end{pmatrix}}}$$

(b) Determine the correlation matrix.

$$\hookrightarrow \text{correlation matrix} \hat{=} R_X = E\{XX^T\} = \begin{bmatrix} E(X_1^2) & E(X_1 X_2) & \dots & E(X_1 X_K) \\ E(X_2 X_1) & E(X_2^2) & \dots & E(X_2 X_K) \\ \vdots & \vdots & \ddots & \vdots \\ E(X_K X_1) & E(X_K X_2) & \dots & E(X_K^2) \end{bmatrix} \text{ 입니다.}$$

$$E(X_1^2) = \int_{-1}^1 \int_{-3}^3 x_1^2 \cdot f_{X_1, X_2}(x_1, x_2) dx_2 dx_1 = \frac{1}{12} \int_{-1}^1 x_1^2 \int_{-3}^3 1 dx_2 dx_1 = \frac{1}{2} \int_{-1}^1 x_1^2 dx_1 = \frac{1}{3}$$

$$E(X_2^2) = \int_{-1}^1 \int_{-3}^3 x_2^2 \cdot f_{X_1, X_2}(x_1, x_2) dx_2 dx_1 = \frac{1}{12} \int_{-1}^1 \int_{-3}^3 x_2^2 dx_2 dx_1 = \frac{3}{2} \int_{-1}^1 1 dx_1 = 3$$

$$E(X_1 X_2) = \int_{-1}^1 \int_{-3}^3 x_1 x_2 \cdot f_{X_1, X_2}(x_1, x_2) dx_2 dx_1 = \int_{-1}^1 x_1 \int_{-3}^3 \frac{x_2}{12} dx_2 dx_1 = \frac{1}{12} \int_{-1}^1 x_1 \cdot 0 dx_1 = 0$$

$$\therefore \underline{\underline{R_X = \begin{bmatrix} \frac{1}{3} & 0 \\ 0 & 3 \end{bmatrix}}}$$

(c) Determine the covariance matrix.

$$\hookrightarrow C_X = R_X - m_X m_X^T$$

$$R_X = \begin{bmatrix} \frac{1}{3} & 0 \\ 0 & 3 \end{bmatrix} \quad m_X m_X^T = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \begin{bmatrix} 0 & 0 \end{bmatrix} = \begin{bmatrix} 0 & 0 \\ 0 & 0 \end{bmatrix}$$

$$C_X = \begin{bmatrix} \frac{1}{3} & 0 \\ 0 & 3 \end{bmatrix} - \begin{bmatrix} 0 & 0 \\ 0 & 0 \end{bmatrix} = \begin{bmatrix} \frac{1}{3} & 0 \\ 0 & 3 \end{bmatrix} \quad \therefore C_X = \begin{bmatrix} \frac{1}{3} & 0 \\ 0 & 3 \end{bmatrix}$$

problems 1.4 Using (1.19) through (1.21), show that the estimate for the covariance matrix can be put in the form (1.22)

$$\hookrightarrow (1.19) \quad M_n = \frac{1}{n} \sum_{i=1}^n X_i \quad (1.20) \quad R_n = \frac{1}{n} \sum_{i=1}^n X_i X_i^T \quad (1.21) \quad C_n = \frac{1}{n-1} \sum_{i=1}^n (X_i - M_n) (X_i - M_n)^T$$

$$(1.22) \quad C_n = \frac{n}{n-1} (R_n - M_n M_n^T)$$

지금부터 (1.19), (1.20), (1.21)의 식을 이용하여 (1.22)의 식을 유도해보겠습니다.

$$\text{우선 (1.21) 식을 정리하겠습니다. } C_n = \frac{1}{n-1} \sum_{i=1}^n (X_i - M_n) (X_i - M_n)^T = \frac{1}{n-1} \sum_{i=1}^n (X_i X_i^T - X_i M_n^T - M_n X_i^T + M_n M_n^T)$$

$$= \frac{1}{n-1} \left(\sum_{i=1}^n X_i X_i^T - \sum_{i=1}^n X_i M_n^T - M_n \sum_{i=1}^n X_i^T + \sum_{i=1}^n M_n M_n^T \right)$$

$$\sum_{i=1}^n X_i X_i^T \text{는 (1.20)을 이용하여 } n \cdot R_n \text{으로 표현할 수 있고 } \sum_{i=1}^n X_i M_n^T \text{는 (1.19)를 이용하여 } \left(\sum_{i=1}^n X_i \right) M_n^T = n \cdot \left(\frac{1}{n} \sum_{i=1}^n X_i \right) M_n^T$$

$$= n M_n M_n^T, \quad M_n \sum_{i=1}^n X_i^T \text{도 마찬가지로 } M_n \sum_{i=1}^n X_i^T = M_n \cdot n \left(\frac{1}{n} \sum_{i=1}^n X_i^T \right) = n M_n M_n^T, \text{ 마지막으로 } \sum_{i=1}^n M_n M_n^T = n M_n M_n^T \text{로}$$

나타낼 수 있습니다. 이는 위의 식에 적용하면 아래와 같습니다.

$$\frac{1}{n-1} \left(\sum_{i=1}^n X_i X_i^T - \sum_{i=1}^n X_i M_n^T - M_n \sum_{i=1}^n X_i^T + \sum_{i=1}^n M_n M_n^T \right) = \frac{1}{n-1} \left(n R_n - n M_n M_n^T - n M_n M_n^T + n M_n M_n^T \right) = \frac{1}{n-1} (n R_n - n M_n M_n^T)$$

$$= \frac{n}{n-1} (R_n - M_n M_n^T) \quad (1.22)$$

\therefore (1.19), (1.20), (1.21)을 통해 (1.22)를 유도했습니다.

problems 1-10 A pair of random variables X_1 and X_2 is defined by $X_1 = 3U - 4V$
 $X_2 = 2U + V$
 where U and V are independent random variables with mean 0 and variance 1.

(a) Find a matrix A such that $\begin{bmatrix} X_1 \\ X_2 \end{bmatrix} = A \begin{bmatrix} U \\ V \end{bmatrix}$

$$\hookrightarrow \begin{bmatrix} X_1 \\ X_2 \end{bmatrix} = \begin{bmatrix} 3U - 4V \\ 2U + V \end{bmatrix} = A \begin{bmatrix} U \\ V \end{bmatrix}, \quad A = \begin{bmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{bmatrix} \text{ if if.}$$

$$\begin{bmatrix} 3U - 4V \\ 2U + V \end{bmatrix} = \begin{bmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{bmatrix} \begin{bmatrix} U \\ V \end{bmatrix} = \begin{bmatrix} a_{11}U + a_{12}V \\ a_{21}U + a_{22}V \end{bmatrix} \quad \begin{matrix} a_{11} = 3, a_{12} = -4 \\ a_{21} = 2, a_{22} = 1 \end{matrix}$$

$$\therefore A = \begin{bmatrix} 3 & -4 \\ 2 & 1 \end{bmatrix}$$

(b) What are R_X and C_X of the random vector X ?

$$\hookrightarrow X = \begin{bmatrix} X_1 \\ X_2 \end{bmatrix}, K = \begin{bmatrix} U \\ V \end{bmatrix}, X = AK$$

$$E(U) = 0$$

$$E(V) = 0$$

$$E(X_1) = E(3U - 4V) = 3E(U) - 4E(V) = 0$$

$$V(U) = 1$$

$$V(V) = 1$$

$$E(X_2) = E(2U + V) = 2E(U) + E(V) = 0$$

$$V(U) = E(U^2) - (E(U))^2$$

$$V(V) = E(V^2) - (E(V))^2$$

$$E(U^2) = 1$$

$$E(V^2) = 1$$

$$R_X = E(XX^T) = E(AK(AK)^T) = E(AK \cdot K^T A^T) = A E(KK^T) A^T = A \cdot R_K A^T$$

$$R_K = \begin{bmatrix} E(U^2) & E(UV) \\ E(VU) & E(V^2) \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$R_X = A R_K A^T = \begin{bmatrix} 3 & -4 \\ 2 & 1 \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} 3 & 2 \\ -4 & 1 \end{bmatrix} = \begin{bmatrix} 25 & 2 \\ 2 & 5 \end{bmatrix}$$

$$m_X = \begin{bmatrix} E(X_1) \\ E(X_2) \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \quad m_X^T = \begin{bmatrix} 0 & 0 \end{bmatrix}$$

$$C_X = R_X - m_X m_X^T = \begin{bmatrix} 25 & 2 \\ 2 & 5 \end{bmatrix} - \begin{bmatrix} 0 \\ 0 \end{bmatrix} \begin{bmatrix} 0 & 0 \end{bmatrix} = \begin{bmatrix} 25 & 2 \\ 2 & 5 \end{bmatrix}$$

$$\therefore R_X = \begin{bmatrix} 25 & 2 \\ 2 & 5 \end{bmatrix}, \quad C_X = \begin{bmatrix} 25 & 2 \\ 2 & 5 \end{bmatrix}$$

(c) What are the means and variance of X_1 and X_2 ?

$$\hookrightarrow E(X_1) = E(3U - 4V) = 3E(U) - 4E(V) = 0.$$

$$V(X_1) = V(3U - 4V) = 3^2 V(U) + 4^2 V(V) = 25.$$

$$E(X_2) = E(2U + V) = 2E(U) + E(V) = 0.$$

$$V(X_2) = V(2U + V) = 2^2 V(U) + 1^2 V(V) = 5.$$

$$\therefore E(X_1) = 0, E(X_2) = 0, V(X_1) = 25, V(X_2) = 5.$$

(d) What is the correlation $E(X_1 X_2)$?

$$\hookrightarrow E(X_1 X_2) = E((3U - 4V)(2U + V)) = E(6U^2 - 5UV - 4V^2) = 6E(U^2) - 5E(UV) - 4E(V^2)$$

(b)에서 이미 $E(U^2) = 1$, $E(V^2) = \frac{1}{5}$ 이고 U, V 는 독립이기 때문에 $E(UV) = E(U) \cdot E(V)$ 입니다.

$$6E(U^2) - 5E(UV) - 4E(V^2) = 6 \cdot 1 - 5 \cdot E(U) \cdot E(V) - 4E(V^2) = 6 - 0 - 4 = 2$$

$$\therefore E(X_1 X_2) = 2.$$