# Slice sampling

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# Slice sampling

Suppose the target distribution is  $p(\theta|y)$  with scalar  $\theta$ . Then,

$$p(\theta|y) = \int_0^{p(\theta|y)} 1 \, du$$

Thus,  $p(\theta|y)$  can be thought of as the marginal distribution of

$$(\theta, U) \sim \mathsf{Unif}\{(\theta, u) : 0 < u < p(\theta|y)\}\$$

where u is an auxiliary variable.

Slice sampling performs the following Gibbs sampler:

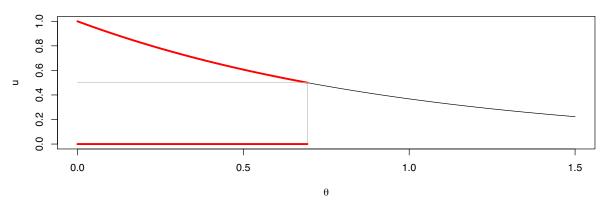
- 1.  $u^t | \theta^{t-1}, y \sim \text{Unif}\{u : 0 < u < p(\theta^{t-1}|y)\}$  and
- 2.  $\theta^t | u^t, y \sim \text{Unif}\{\theta : u^t < p(\theta|y)\}.$

# Slice sampler for exponential distribution

Consider the target  $\theta|y\sim Exp(1)$ , then

$$\{\theta : u < p(\theta|y)\} = (0, -\log(u)).$$

#### **Target disribution**



## Slice sampling in R

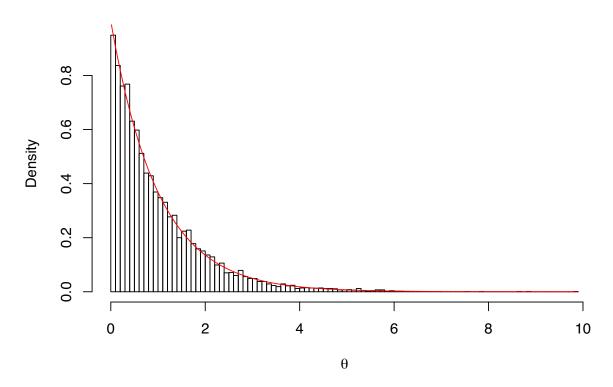
```
slice = function(n,init_theta,target,A) {
  u = theta = rep(NA,n)
  theta[1] = init_theta
  u[1] = runif(1,0,target(theta[1])) # This never actually gets used

for (i in 2:n) {
  u[i] = runif(1,0,target(theta[i-1]))
  endpoints = A(u[i],theta[i-1]) # The second argument is used in the second example
  theta[i] = runif(1, endpoints[1],endpoints[2])
  }
  return(list(theta=theta,u=u))
}
```

```
set.seed(6)
A = function(u,theta=NA) c(0,-log(u))
res = slice(10, 0.1, dexp, A)
```

# Histogram of draws

#### Slice sampling approximation to Exp(1) distribution



### Normal model with unknown mean

Let

$$Y_i \overset{ind}{\sim} N(\theta, 1)$$
 and  $\theta \sim La(0, 1)$ 

then

$$p(\theta|y) \propto \left[\prod_{i=1}^{n} N(y_i; \theta, 1)\right] La(\theta; 0, 1)$$

```
n = 5
y = rnorm(n,.2)
f = Vectorize(function(theta, y.=y) exp(sum(dnorm(y., theta, log=TRUE)) + dexp(abs(theta), log=TRUE)))

# Function to numerically find endpoints
A = function(u, xx, f.=f) {
  left_endpoint = uniroot(function(x) f.(x) - u, c(-10^10, xx))
  right_endpoint = uniroot(function(x) f.(x) - u, c( 10^10, xx))
  c(left_endpoint$root, right_endpoint$root)
}
```

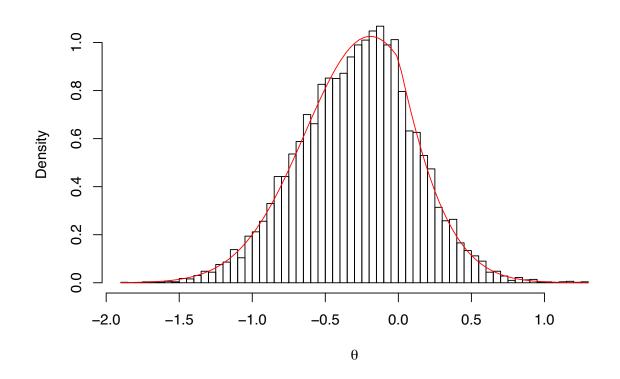
```
res = slice(20, mean(y), f, A)
```

# Slice sampling using numerically calculated endpoints

# Histogram of draws

#### Slice sampling approximation to posterior

Posterior



### An alternative augmentation

Suppose

$$Y_i \stackrel{ind}{\sim} N(\theta, 1)$$
 and  $\theta \sim La(0, 1)$ 

but now, we will use the augmentation

$$p(u, \theta) \propto p(\theta) I(0 < u < p(y|\theta))$$

The full conditional distributions are now

- 1.  $u|\theta, y \sim Unif(0, p(y|\theta))$  and
- 2.  $\theta | u, y \sim p(\theta) I(u < p(y|\theta))$ .

# Sampling $\theta|u,y$

Now we need to sample from

$$p(\theta)I(u < p(y|\theta)).$$

If  $p(\theta)$  is unimodal, then this is equivalent to

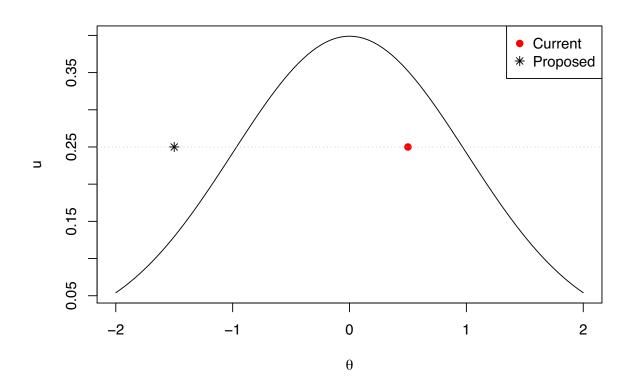
$$p(\theta)I(\theta_L(u) < \theta < \theta_U(u))$$

for some bounds  $\theta_L(u)$  and  $\theta_U(u)$  which depend on u.

One way to learn these is to sample from  $p(\theta)$  and update the bounds, e.g. if  $\theta^{(i-1)}$  is our current value in the chain, we know  $u < p(y|\theta^{(i-1)})$  or, equivalently,  $\theta_L(u) < \theta^{(i-1)} < \theta_U(u)$ . Letting  $u^{(i)}$  be the current value for the auxiliary variable and setting  $\theta_L(u^{(i)})$  [ $\theta_U(u^{(i)})$ ] to the lower [upper] bound of the support for  $\theta$ , we can

- 1. Sample  $\theta^* \sim p(\theta) I(\theta_L(u^{(i)}) < \theta < \theta_U(u^{(i)})$ .
- 2. Set  $\theta^{(i)} = \theta^*$  if  $u^{(i)} < p(y|\theta^*)$ , otherwise
  - a. set  $\theta_L(u^{(i)}) = \theta^*$  if  $\theta^* < \theta^{(i-1)}$  or
  - b. set  $\theta_U(u^{(i)}) = \theta^*$  if  $\theta^* > \theta^{(i-1)}$  and

# Learning the endpoints

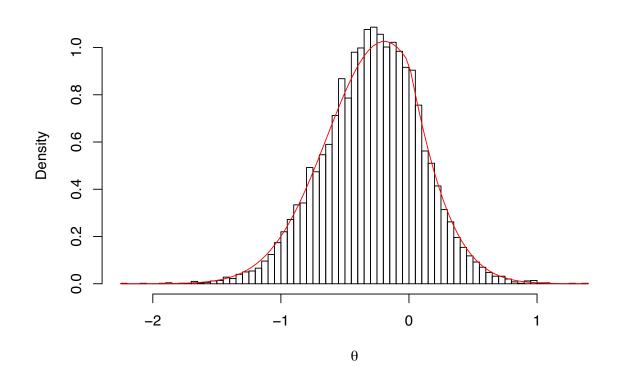


### R code

```
slice2 = function(n, init_theta, like, qprior) {
   u = theta = rep(NA, n)
   theta[1] = init_theta
   u[1] = runif(1, 0, like(theta[1]))
   for (i in 2:n) {
        u[i] = runif(1, 0, like(theta[i - 1]))
        success = FALSE
        endpoints = 0:1
        while (!success) {
            # Inverse CDF
            up = runif(1, endpoints[1], endpoints[2])
            theta[i] = qprior(up)
            if (u[i] < like(theta[i])) {</pre>
                success = TRUE
            } else {
                # Updated endpoints when proposed value is rejected
                if (theta[i] > theta[i - 1])
                  endpoints[2] = up
                if (theta[i] < theta[i - 1])</pre>
                  endpoints[1] = up
   return(list(theta = theta, u = u))
```

# Histogram

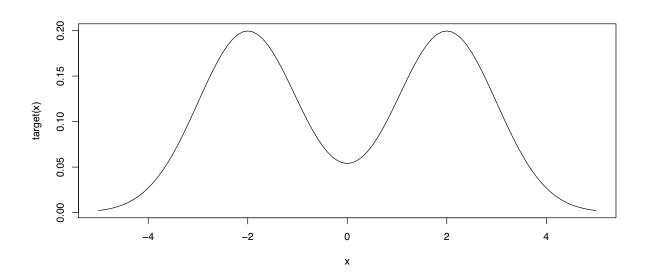
#### Slice sampling approximation to posterior



### Bimodal target distributions

### Consider the posterior

$$p(\theta|y) = \frac{1}{2}N(\theta; -2, 1) + \frac{1}{2}N(\theta; 2, 1)$$



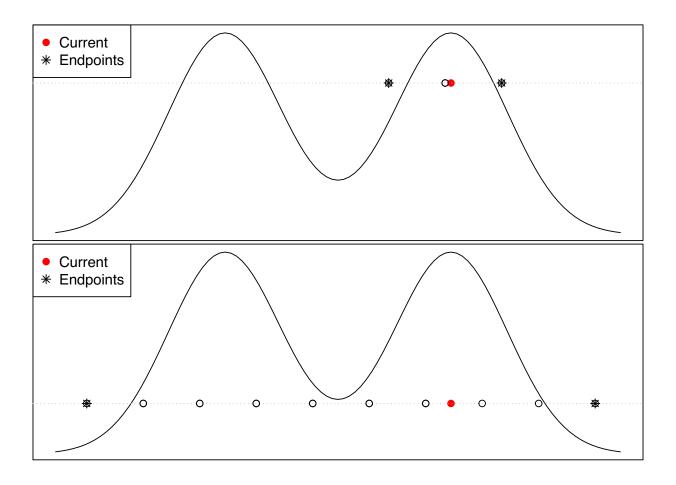
# Stepping-out slice sampler

### To sample from $\theta|u,y$ , let

- $\bullet$   $\theta^{(i-1)}$  be the current draw for  $\theta$
- ullet  $u^{(1)}$  be the current draw for the auxiliary variable u
- ullet w be a tuning parameter that you choose

### Perform the following

- 1. Randomly place an interval  $(\theta_L(u^{(i)}), \theta_U(u^{(i)}))$  of length w around the current value  $\theta^{(i-1)}$ .
- 2. Step the endpoints of this interval out in increments of w until  $u^{(i)} > p(\theta_L(u^{(i)})|y)$  and  $u^{(i)} > p(\theta_B(u^{(i)})|y)$ .
- 3. Sample  $\theta^* \sim Unif(\theta_L(u^{(i)}), \theta_L(u^{(i)}))$ .
- 4. If  $u^{(i)} < p(\theta^*|y)$ , then set  $\theta^{(i)} = \theta^*$ , otherwise
  - a. set  $\theta_L(u^{(i)}) = \theta^*$  if  $\theta^* < \theta^{(i-1)}$  or
  - b. set  $\theta_U(u^{(i)}) = \theta^*$  if  $\theta^* > \theta^{(i-1)}$  and
  - c. return to Step 3.

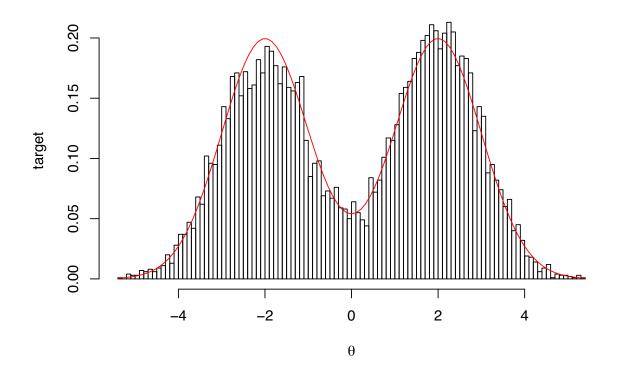


```
create_interval = function(theta, u, target, w, max_steps) {
 L = theta - runif(1,0,w)
 R = I + W
 # Step out
 J = floor(max_steps * runif(1))
 K = (max\_steps - 1) - J
 while ((u < target(L)) & J > 0) {
   L = L - w
    J = J - 1
 while ((u < target(R)) & K > 0) {
   R = R + w
   K = K - 1
 return(list(L=L,R=R))
shrink_and_sample = function(theta, u, target, int) {
 L = int L; R = int R
 repeat {
   theta_prop = runif(1, L, R)
   if (u < target(theta_prop))</pre>
     return(theta_prop)
    # shrink
   if (theta_prop > theta)
      R = theta_prop
   if (theta_prop < theta)</pre>
     L = theta_prop
```

# Sampling from mixture of normals

```
res = slice(n = 1e4, init_theta=0, target=target, w=1, max_steps=10)
```

#### Stepping out slice sampler for bimodal target



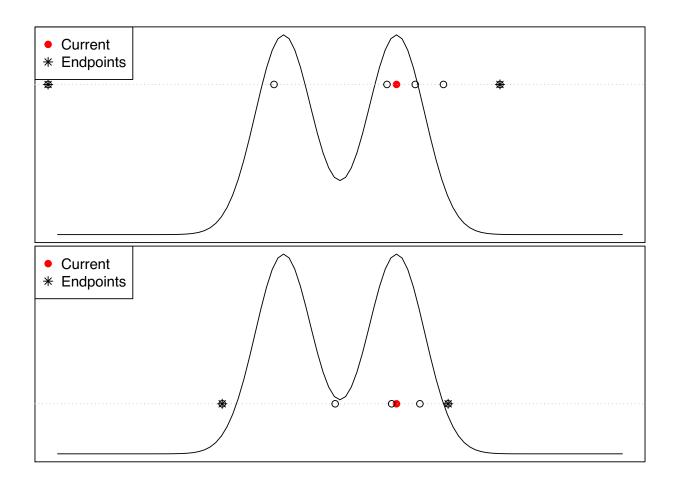
## Doubling slice sampler

To sample from  $\theta|u,y$ , let

- ullet  $\theta^{(i-1)}$  be the current draw for  $\theta$
- $\bullet$   $u^{(1)}$  be the current draw for the auxiliary variable u
- ullet w be a tuning parameter that you choose

### Perform the following

- 1. Randomly place an interval  $(\theta_L(u^{(i)}), \theta_U(u^{(i)}))$  of length w around the current value  $\theta^{(i-1)}$ .
- 2. Randomly double the size of the interval to either the left or right until  $u^{(i)} > p(\theta_L(u^{(i)})|y)$  and  $u^{(i)} > p(\theta_B(u^{(i)})|y)$ .
- 3. Sample  $\theta^* \sim Unif(\theta_L(u^{(i)}), \theta_L(u^{(i)}))$ .
- 4. If  $u^{(i)} < p(\theta^*|y)$  and a reversibility criterion is satisfied, then set  $\theta^{(i)} = \theta^*$ , otherwise
  - a. set  $\theta_L(u^{(i)}) = \theta^*$  if  $\theta^* < \theta^{(i-1)}$  or
  - b. set  $\theta_U(u^{(i)}) = \theta^*$  if  $\theta^* > \theta^{(i-1)}$  and
  - c. return to Step 3.



## Reversibility criterion

This procedure works backward through the intervals that the doubling procedure would pass through to arrive at [the doubled interval] when starting from the new point, checking that none of [the intermediate intervals] has both ends outside the slice, which would lead to earlier termination of the doubling procedure.

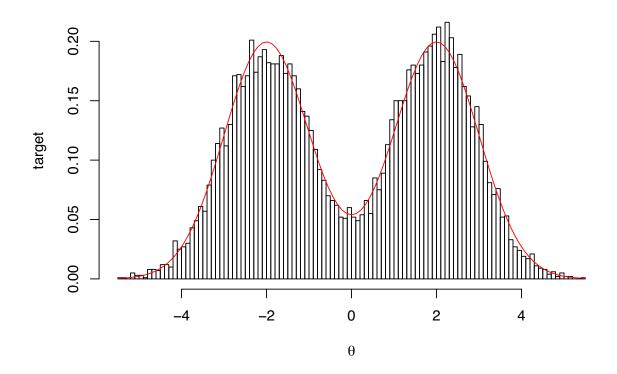
```
accept = function(theta0, theta1, L, R, u, w) {
    D = FALSE
    while (R - L > 1.1 * w) {
        M = (L + R)/2
        if ((theta0 < M & theta1 >= M) | (theta0 >= M & theta1 < M))
            D = TRUE
        if (theta1 < M) {
            R = M
        } else {
            L = M
        }
        if (D & u >= target(L) & u >= target(R)) {
            return(FALSE)
        }
    }
    return(TRUE)
}
```

```
slice = function(n, init_theta, target, w, max_doubling) {
    u = theta = rep(NA, n)
   theta[1] = init_theta
   for (i in 2:n) {
        u[i] = runif(1, 0, target(theta[i - 1]))
       L = theta[i - 1] - runif(1, 0, w)
       R = L + w
        # Step out
       K = max_doubling
        while ((u[i] < target(L) | u[i] < target(R)) & K > 0) {
            if (runif(1) < 0.5) {
               L = L - (R - L)
            } else {
                R = R + (R - L)
            K = K - 1
        # Sample and shrink
        repeat {
            theta[i] = runif(1, L, R)
            if (u[i] < target(theta[i]) & accept(theta[i - 1], theta[i], L, R, u[i], w))
                break
            # shrink
            if (theta[i] > theta[i - 1])
               R = theta[i]
            if (theta[i] < theta[i - 1])</pre>
               L = theta[i]
   return(list(theta = theta, u = u))
```

# Doubling slice sampler for bimodal target

```
res = slice(n=1e4, init_theta=0, target=target, w=1, max_doubling=10)
```

#### Stepping out slice sampler for bimodal target



# Multivariate slice sampling

Suppose, we are interested in sampling from

$$p(\theta_1, \theta_2|y) = \int_0^{p(\theta_1, \theta_2|y)} 1 \, du$$

- Treat each variable independently, i.e.
  - 1.  $u|\theta_1, \theta_2, y \sim Unif(0, p(\theta_1, \theta_2|y))$
  - 2.  $\theta_1 | u, \theta_2, y \sim Unif(u < p(\theta_1, \theta_2 | y))$
  - 3.  $\theta_2|u,\theta_1,y \sim Unif(u < p(\theta_1,\theta_2|y))$
  - Use overrelaxation to avoid random walk behavior
- Hyperrectangle slice sampling
  - $\bullet$  Replace interval constructed from w with a hyperrectangle W placed randomly over the slice
  - Shrink as points are rejected
- Reflective slice sampling
  - Candidate samples are kept within the bounds by reflecting the direction of sampling when the boundary is hit