

# Occasionally Misspecified

JJ Forneron, Boston University

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  - good fit for most of the sample
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- Look for an estimator that is:
  1. robust to small amounts of large contamination
  2. asymptotically unbiased
  3. not too hard to compute

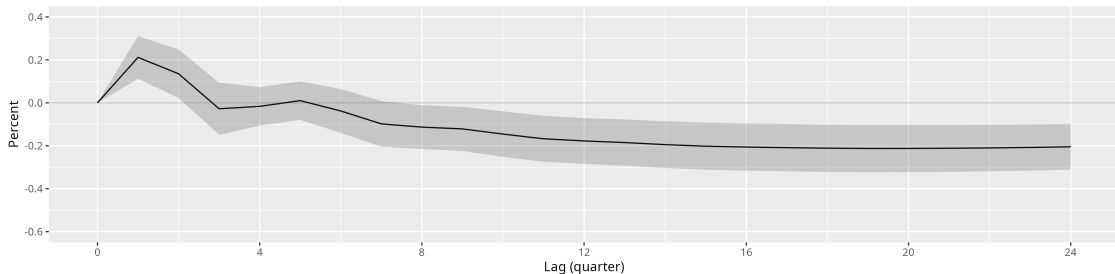
# Motivating Example: Price Puzzle

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- Estimand: effect of monetary shock on inflation
- Method: recursive VAR (OLS)
  - Variables: Interest Rates  $R_t$ , Inflation  $\pi_t$ , Unemployment:  $u_t$
  - Specification: 4 lags
  - Data: US from 1960Q1 to 2000Q4  
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  - Data: US from 1960Q1 to 2000Q4  
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- Results: when  $R_t \nearrow$ ,  $\pi_t \nearrow$



# Price Puzzle: Background

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  - structural breaks: e.g (Castelnuovo and Surico, 2010)
  - ...
  - meta-analysis: Rusnák et al. (2013) look at 1000 regressions ( $\gg T$ )

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$$\pi_t = \beta_0 + \beta_1 R_{t-1} + \beta_2 u_{t-1} + \beta_3 \pi_{t-1} + \cdots + \beta_{10} R_{t-4} + \beta_{11} u_{t-4} + \beta_{12} \pi_{t-4} + e_{\pi,t}.$$

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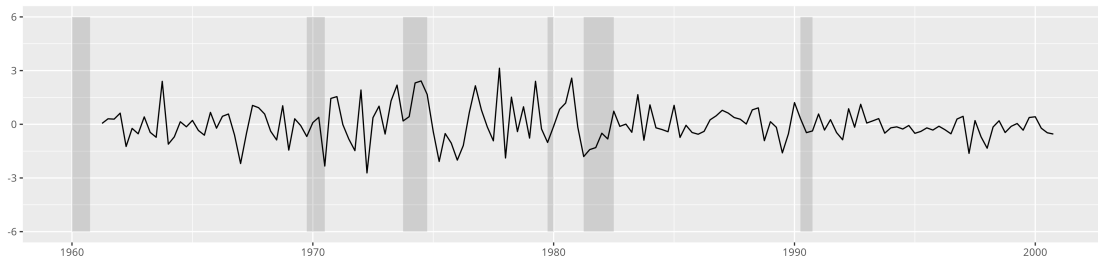
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- Let's look at some regression diagnostics

# Price Puzzle: Diagnostics

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- Residuals  $\hat{e}_{\pi,t}$  (standardized)

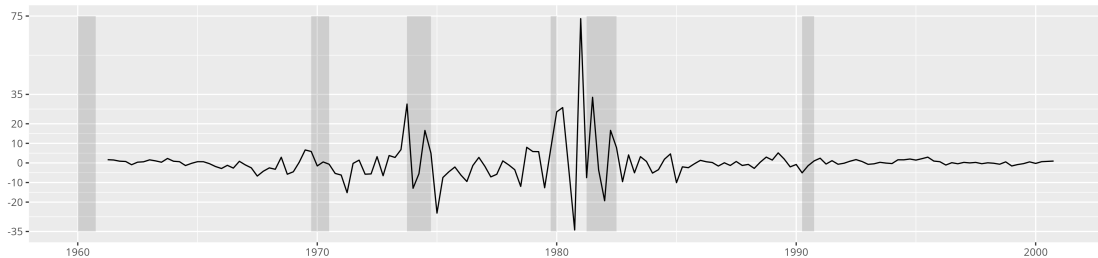


- Skewness: 0.36, Kurtosis: 3.78



# Price Puzzle: Diagnostics

- Contributions to  $\hat{\beta}_{1n}$  – based on  $(X'X/n)^{-1}x_t y_t$  (avg =  $\hat{\beta}_n$ )



- 1981Q1:  $75/n \simeq 0.47$  vs.  $\hat{\beta}_1 = 0.21$ , 3.5 standard errors
- Skewness: 3.24, Kurtosis: 27.81

# Price Puzzle: Concerns

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- 1979-1982: Federal Reserve no longer sets  $R_t$  directly
  - Non-borrowed reserves targeting (misspecification)
  - Increased volatility in  $R_t$  (leverage)

⇒ **Leveraged outliers**

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  - + Reduce the influence of outliers
  - Biased/Inconsistent for asymmetric data

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## ⇒ Leveraged outliers

- Robust estimation & inference desirable
  - + Reduce the influence of outliers
  - Biased/Inconsistent for asymmetric data
  - Robust M-estimators **not robust to leverage** (Hamilton, 1992)

# Literature

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- Robust M-estimation:

$$\min_{\theta} \sum_{t=1}^n \psi(y_t - x_t' \theta),$$

$\psi$  non-quadratic: LAD, Huber loss, trimming, etc.

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- Local asymptotics imply  $\hat{\theta}_n$  consistent

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- More recent, finite-sample: Median-of-Means
  - $K$ -subsamples,  $K$  means, return median
  - robust up to  $n_o \leq K/2 - 1$  outliers
  - cv. rate depends on  $n/K$
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- $n_o \rightarrow \infty$  requires  $K \rightarrow \infty$
- bias of order:  $\sqrt{K/n}$

# Setting: Sample

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- GMM estimation:  $\mathbb{E}_P[g(z_t; \theta)] = 0 \Leftrightarrow \theta = \theta_0$
- Data  $(z_1, \dots, z_n)$  with  $n = n_P + n_o$

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- Sample mean  $\bar{g}_n(\theta)$  can be
  - **asymptotically biased** if  $n_o n^\alpha / \sqrt{n} = O(1)$
  - **inconsistent** if  $n_o n^\alpha / n = O(1)$
  - **divergent** if  $n_o n^\alpha / n \rightarrow \infty$

# Setting: Estimator

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1. **Moments:** find  $\hat{\psi}_n(\theta; \nu) = (\hat{\mu}_n, \hat{\Sigma}_n)$  minimizing:

$$Q_n(\psi; \theta) = \frac{\nu + p}{n} \sum_{t=1}^n \log \left( 1 + \frac{\|g(z_t; \theta) - \mu\|_{\Sigma^{-1}}^2}{\nu} \right) + \log |\Sigma| + \frac{\kappa_1}{\nu} \|\mu\|_{\Sigma^{-1}}^2 + \frac{\kappa_2}{\nu} \text{trace}(\Sigma)$$

for  $0 < \nu, \kappa_1, \kappa_2 < \infty$  over

$$\Psi = \{(\mu, \Sigma), \mu \in \mathbb{R}^p, 0 < s_0 \leq \lambda_{\min}(\Sigma) \leq \lambda_{\max}(\Sigma) \leq +\infty\}$$

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3. **Estimation:** find  $\tilde{\theta}_n$  such that

$$\|\tilde{\mu}_n(\tilde{\theta}_n)\|_{W_n}^2 \leq \inf_{\theta \in \Theta} \|\tilde{\mu}_n(\theta)\|_{W_n}^2 + o_p(n^{-1})$$

# Tuning parameters, Some Properties

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- $0 < \nu < \infty$ : controls robustness,
  - $\nu = \infty$ :  $\hat{\mu}_n(\theta; \infty) = \bar{g}_n(\theta)$  sample mean,  $\hat{\Sigma}_n(\theta; \infty)$  sample covariance
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- Weighted average representation:

$$\hat{\mu}_n(\theta; \nu) = \sum_t \omega_t(\theta; \nu) g(z_t; \theta), \quad \tilde{\mu}_n(\theta; \nu) = \sum_t \tilde{\omega}_t(\theta; \nu) g(z_t; \theta)$$

where  $0 \leq \omega_t$ ,  $\sum_t \omega_t \leq 1$ ,  $\tilde{\omega}_t = 2\omega_t(\nu) - \omega_t(\nu/2)$

$\Rightarrow$  Robust-LS is weighted-LS with weights  $\tilde{w}_t$

# Some intuition

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- Simplified estimator:  $\theta_0 = \mathbb{E}_P(z_t)$

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- Finite-Sample: a)  $\sup_z |z|/(1 + |z|^2/\nu) \leq \sqrt{\nu}/2$ , b) Bernstein's inequality:

$$\mathbb{P} \left( \sup_{z_t \in \mathcal{O}_n, t > n_P} |\hat{\mu}_n(\nu) - \mu(\nu)| \geq \frac{\sqrt{\nu} n_o}{n} + \frac{n_P}{n} \frac{x}{\sqrt{n_P}} \right) \leq 2 \exp \left( -\frac{x^2}{2} \sigma_\nu^2 + 2/3 \sqrt{\nu/n_P} x \right)$$

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- Two biases:

- outlier:  $\sqrt{\nu} n_o / n$
- asymmetry:  $\mu(\nu) - \theta_0 = \mathbb{E}_P[\hat{\mu}_n(\nu)] - \theta_0$

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- Asymmetry bias is at most  $\mathbb{E}_P(|z_t|^3)/\nu$ :

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- Repeat bias correction  $\tilde{\tilde{\mu}}(\nu) = 2\tilde{\mu}(\nu) - \tilde{\mu}(\nu/2)$ :

$$\tilde{\tilde{\mu}}(\nu) = \theta_0 + \frac{2}{\nu^2} \mathbb{E}_P \left( \frac{z_t^5(1 - z_t^4/\nu)}{(1 + z_t^2/\nu)(1 + 2z_t^2/\nu)(1 + 2z_t^2/\nu)(1 + 4z_t^2/\nu)} \right)$$

numerator has 3 roots, better small sample properties (simulations)

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- Optimal choice  $\nu \asymp (n/n_o)^{2/5}$

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- Two biases:
  - Outlier bias:  $\sqrt{\nu}n_o/n$
  - Asymmetry bias:  $1/\nu^2$
- Optimal choice  $\nu \asymp (n/n_o)^{2/5}$
- Asymptotic normality requires:  $\sqrt{n}/\nu^2 = o(1), \sqrt{\nu/nn_o} = o(1) \Rightarrow n_o = o(n^{3/8})$ 
  - $n_o$  cannot increase too quickly. . .
  - $P$  symmetric, need:  $n_o = o(n^{1/2})$
  - $\nu \asymp n^{1/4} \log(n)$  implies  $n_o = o(n^{3/8}/\sqrt{\log(n)})$



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- Undersmoothing (no bias correction):  $n_o = o(n^{1/4})$

# Finite Sample Results

---

$$Q_\nu(\psi; \theta) = \mathbb{E}_P \left[ (\nu + p) \log \left( 1 + \|g(z_t; \theta) - \mu\|_{\Sigma^{-1}}^2 / \nu \right) \right] + \log |\Sigma| + \frac{\kappa_1}{\nu} \|\mu\|_{\Sigma^{-1}}^2 + \frac{\kappa_2}{\nu} \text{trace}(\Sigma).$$

- $Q_\nu$ : population analog of  $Q_n$  with  $n_o = 0$ , let  $\psi(\theta; \nu)$  minimize  $Q_\nu$

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- $Q_\nu$ : population analog of  $Q_n$  with  $n_o = 0$ , let  $\psi(\theta; \nu)$  minimize  $Q_\nu$
- If i.  $\mathbb{E}_P[\|g(z_t; \theta)\|^2] < \infty, \forall \theta$ ,  
 ii.  $\|g(z_t; \theta_1) - g(z_t; \theta_2)\| \leq G_t \|\theta_1 - \theta_2\|, \mathbb{E}_P(\|G_t\|^2) < \infty$ , then, for iid data:

$$\mathbb{P} \left( \sup_{\theta \in \Theta} \sup_{z_t \in \mathcal{O}_n, t > n_P} \left\{ Q_\nu(\hat{\psi}_n(\theta; \nu); \theta) - Q_\nu(\psi(\theta; \nu); \theta) \right\} \geq C_{\mathcal{O}} \frac{n_o(\nu + p)}{n} [1 + \log(n)] \right. \\ \left. + L \frac{n_P}{n} (\nu + p) \log(1 + \nu p) \left[ \sqrt{\frac{x}{n_P}} + \frac{x}{n_P} + \sqrt{\frac{C_n}{n_P}} + \frac{C_n}{n_P} \right] \right) \leq 4 \exp(-x),$$

where  $C_n = 1 + (k + 2p^2)[\log(p) + \log(\nu) + \log(n_P)]$

# Large Sample Results: Oracle Equivalence

---

- Further assume that  $\sup_{\theta \in \Theta} \mathbb{E}_P[\|g(z_t; \theta)\|^4] < \infty$ , and

$$n_o = o\left(\frac{n}{\nu \log(n)}\right), \quad \nu \log(\nu) = o\left(\sqrt{\frac{n}{\log(n)}}\right).$$

- Let  $M_{r,\delta} = \max(\mathbb{E}_P(\|g(z_t; \theta_0)\|^{r+\delta}), \mathbb{E}_P(\|G_t\|^{r+\delta}))$  and  $\bar{g}_{n_P}(\theta) = \frac{1}{n_P} \sum_{t=1}^{n_P} g(z_t; \theta)$

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$$\sup_{\theta \in \Theta} \left( \sup_{z_t \in \mathcal{O}_n, t > n_P} \|\hat{\mu}_n(\theta; \nu) - \bar{g}_{n_P}(\theta)\| \right) = O_p \left( \max \left[ \frac{1}{\nu}, \frac{\sqrt{\nu} n_o}{n} \right] \right)$$

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- If  $M_{5,\delta} < \infty$  for some  $\delta > 0$ :

$$\sup_{\theta \in \Theta} \left( \sup_{z_t \in \mathcal{O}_n, t > n_P} \|\tilde{\mu}_n(\theta; \nu) - \bar{g}_{n_P}(\theta)\| \right) = O_p \left( \max \left[ \frac{1}{\nu^2}, \frac{\sqrt{\nu} n_o}{n} \right] \right).$$

# Large Samples: Oracle Equivalence (estimator)

---

- Assume standard regularity conditions for  $1 \leq t \leq n_P$ , and

$$\frac{\sqrt{n}}{\nu^2} = o(1), \text{ and } \sqrt{\frac{\nu}{n}} n_o = o(1).$$

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- Let  $\hat{\theta}_{n_P}$  minimize  $\|\bar{g}_{n_P}(\theta)\|_{W_n}$ ;  $\tilde{\theta}_n$  minimize  $\|\tilde{\mu}_n(\theta; \nu)\|_{W_n}$  then:

$$\sup_{z_t \in \mathcal{O}_n, t > n_P} \|\sqrt{n_P}(\tilde{\theta}_n - \hat{\theta}_{n_P})\| = o_p(1), \quad \text{and} \quad \sqrt{n_P}(\tilde{\theta}_n - \theta_0) \xrightarrow{d} \mathcal{N}(0, V)$$

for any sequence  $(z_t)_{t > n_P} \in \mathcal{O}_n$ ,  $V = \text{avar}(\hat{\theta}_{n_P})$



# Monte Carlo

---

- Simulation design – for  $1 \leq t \leq n_P$

$$y_t = \theta_0 + \theta_1 x_{1t} + \theta_2 x_{2t} + \theta_3 x_{3t} + e_t,$$

- $n = 150$ ,  $x_{1t}, x_{2t}, x_{3t}, e_t = (\chi_5^2 - 5)/\sqrt{10}$ .  $\theta_0 = (0, 1, 1, 1)$

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- Outliers:  $y_t = x_t' \theta_{\dagger}$ ,  $\theta_{\dagger} = (0, 1/2, 1/2, 1/2)$ .  $x_{1t} = x_{2t} = x_{3t} = \sqrt{n}$
- $n_o = 0, 1, 5, 10, \dots$

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---

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- Outliers:  $y_t = x_t' \theta_{\dagger}$ ,  $\theta_{\dagger} = (0, 1/2, 1/2, 1/2)$ .  $x_{1t} = x_{2t} = x_{3t} = \sqrt{n}$
- $n_o = 0, 1, 5, 10, \dots$
- Several estimators:
  - OLS,
  - oracle estimator,
  - Robust M-estimator with Huber loss,
  - This paper's:  $\hat{\theta}_n, \tilde{\theta}_n, \tilde{\tilde{\theta}}_n$
  - Undersmoothing: no bias correction,  $\nu^2$

# Monte Carlo: $n = 150$ , $n_o = 0$

	100 × RMSE							Rejection Rate						
	$n_o = 0$													
	$\hat{\theta}_n^{ols}$	$\hat{\theta}_{np}^{ols}$	$\hat{\theta}_n^{rlm}$	$\hat{\theta}_n$	$\tilde{\theta}_n$	$\tilde{\tilde{\theta}}_n$	$\hat{\theta}_n^{un}$	$\hat{\theta}_n^{ols}$	$\hat{\theta}_{np}^{ols}$	$\hat{\theta}_n^{rlm}$	$\hat{\theta}_n$	$\tilde{\theta}_n$	$\tilde{\tilde{\theta}}_n$	$\hat{\theta}_n^{un}$
$\theta_0$	8.05	8.05	12.00	11.84	9.31	8.11	7.94	0.04	0.04	0.24	0.29	0.14	0.05	0.06
$\theta_1$	8.00	8.00	7.15	7.97	7.79	7.78	7.92	0.06	0.06	0.06	0.11	0.08	0.07	0.06
$\theta_2$	8.10	8.10	7.46	8.45	8.21	8.11	8.06	0.04	0.04	0.05	0.10	0.06	0.05	0.05
$\theta_3$	8.19	8.19	7.43	8.55	8.30	8.16	8.14	0.06	0.06	0.06	0.10	0.07	0.06	0.06

Average  $\hat{\nu}_n$ : 35.85, 16.00, 11.00, 10.71 for  $n_0 = 0, 1, 5, 10$  respectively. Each  $\hat{\nu}_n$  is selected on a grid  $[\nu_0, \dots, \nu_J]$  where  $\nu_0 = 8.77$ ,  $\nu_J = 584.69$ .

# Monte Carlo: $n = 150$ , $n_o = 1$

	100 $\times$ RMSE							Rejection Rate						
	$n_o = 1$													
	$\hat{\theta}_n^{ols}$	$\hat{\theta}_{n_P}^{ols}$	$\hat{\theta}_n^{rlm}$	$\hat{\theta}_n$	$\tilde{\theta}_n$	$\tilde{\tilde{\theta}}_n$	$\hat{\theta}_n^{un}$	$\hat{\theta}_n^{ols}$	$\hat{\theta}_{n_P}^{ols}$	$\hat{\theta}_n^{rlm}$	$\hat{\theta}_n$	$\tilde{\theta}_n$	$\tilde{\tilde{\theta}}_n$	$\hat{\theta}_n^{un}$
$\theta_0$	10.71	8.04	13.01	14.18	10.97	8.52	10.32	0.03	0.04	0.20	0.46	0.23	0.08	0.08
$\theta_1$	38.57	8.07	15.23	8.27	7.97	7.87	32.24	0.00	0.06	0.01	0.14	0.10	0.07	0.40
$\theta_2$	38.39	8.11	15.09	8.73	8.36	8.14	32.08	0.01	0.04	0.01	0.12	0.06	0.06	0.38
$\theta_3$	39.94	8.20	15.75	8.83	8.49	8.27	33.47	0.00	0.06	0.00	0.12	0.09	0.07	0.39

Average  $\hat{\nu}_n$ : 35.85, 16.00, 11.00, 10.71 for  $n_0 = 0, 1, 5, 10$  respectively. Each  $\hat{\nu}_n$  is selected on a grid  $[\nu_0, \dots, \nu_J]$  where  $\nu_0 = 8.77$ ,  $\nu_J = 584.69$ .

# Monte Carlo: $n = 150$ , $n_o = 5$

	100 × RMSE							Rejection Rate						
	$n_o = 5$													
	$\hat{\theta}_n^{ols}$	$\hat{\theta}_{np}^{ols}$	$\hat{\theta}_n^{rlm}$	$\hat{\theta}_n$	$\tilde{\theta}_n$	$\tilde{\tilde{\theta}}_n$	$\hat{\theta}_n^{un}$	$\hat{\theta}_n^{ols}$	$\hat{\theta}_{np}^{ols}$	$\hat{\theta}_n^{rlm}$	$\hat{\theta}_n$	$\tilde{\theta}_n$	$\tilde{\tilde{\theta}}_n$	$\hat{\theta}_n^{un}$
$\theta_0$	11.98	8.14	16.57	16.98	13.38	9.82	13.45	0.10	0.04	0.24	0.59	0.38	0.13	0.16
$\theta_1$	47.57	8.40	47.17	9.02	8.62	8.40	46.72	0.99	0.06	0.99	0.12	0.08	0.06	0.99
$\theta_2$	47.48	8.26	48.25	9.28	8.80	8.53	47.14	0.99	0.04	1.00	0.12	0.05	0.03	1.00
$\theta_3$	49.17	8.28	49.48	9.33	8.94	8.72	48.65	0.98	0.06	0.98	0.10	0.08	0.04	0.98

Average  $\hat{\nu}_n$ : 35.85, 16.00, 11.00, 10.71 for  $n_0 = 0, 1, 5, 10$  respectively. Each  $\hat{\nu}_n$  is selected on a grid  $[\nu_0, \dots, \nu_J]$  where  $\nu_0 = 8.77$ ,  $\nu_J = 584.69$ .

# Monte Carlo: $n = 150$ , $n_o = 10$

	100 × RMSE							Rejection Rate						
	$n_o = 10$													
	$\hat{\theta}_n^{ols}$	$\hat{\theta}_{np}^{ols}$	$\hat{\theta}_n^{rlm}$	$\hat{\theta}_n$	$\tilde{\theta}_n$	$\tilde{\tilde{\theta}}_n$	$\hat{\theta}_n^{un}$	$\hat{\theta}_n^{ols}$	$\hat{\theta}_{np}^{ols}$	$\hat{\theta}_n^{rlm}$	$\hat{\theta}_n$	$\tilde{\theta}_n$	$\tilde{\tilde{\theta}}_n$	$\hat{\theta}_n^{un}$
$\theta_0$	12.21	8.21	17.33	16.78	13.27	10.35	14.13	0.09	0.04	0.23	0.47	0.22	0.07	0.17
$\theta_1$	49.14	8.54	48.38	10.22	11.68	19.76	48.65	0.99	0.04	0.99	0.01	0.01	0.09	1.00
$\theta_2$	49.05	8.31	49.67	10.76	12.40	20.28	48.92	0.99	0.04	0.99	0.01	0.01	0.09	1.00
$\theta_3$	50.52	8.51	50.70	11.04	13.00	20.96	50.19	0.98	0.06	0.98	0.00	0.01	0.09	0.99

Average  $\hat{\nu}_n$ : 35.85, 16.00, 11.00, 10.71 for  $n_0 = 0, 1, 5, 10$  respectively. Each  $\hat{\nu}_n$  is selected on a grid  $[\nu_0, \dots, \nu_J]$  where  $\nu_0 = 8.77$ ,  $\nu_J = 584.69$ .

# Back to the Price Puzzle

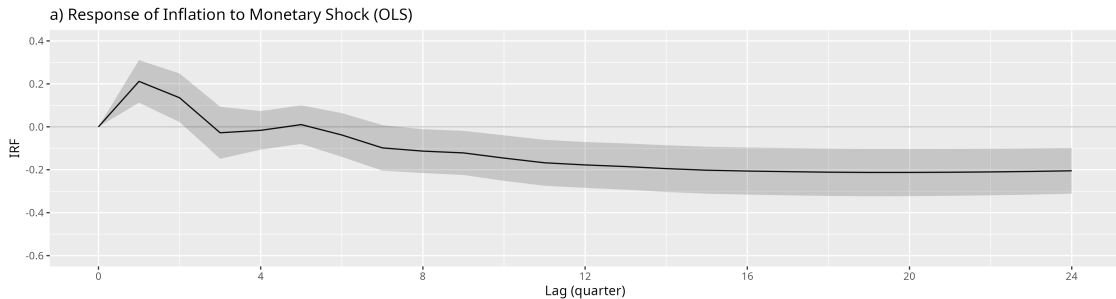
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- Re-estimate the model using OLS,  $\hat{\theta}_n$ ,  $\tilde{\theta}_n$ ,  $\tilde{\tilde{\theta}}_n$
- Same data
- Same model

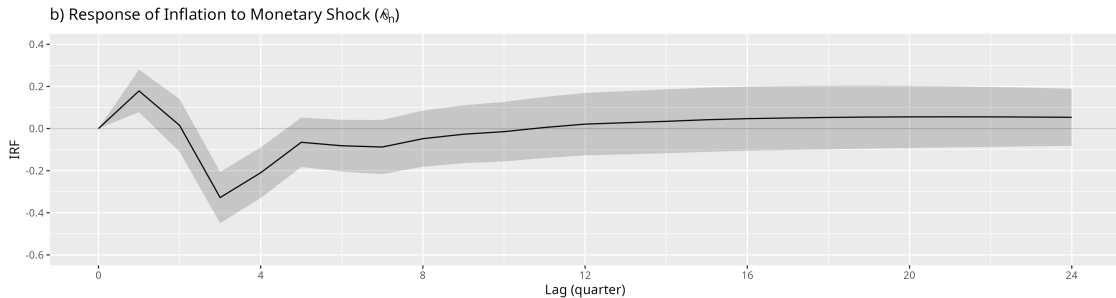


# Price Puzzle: OLS

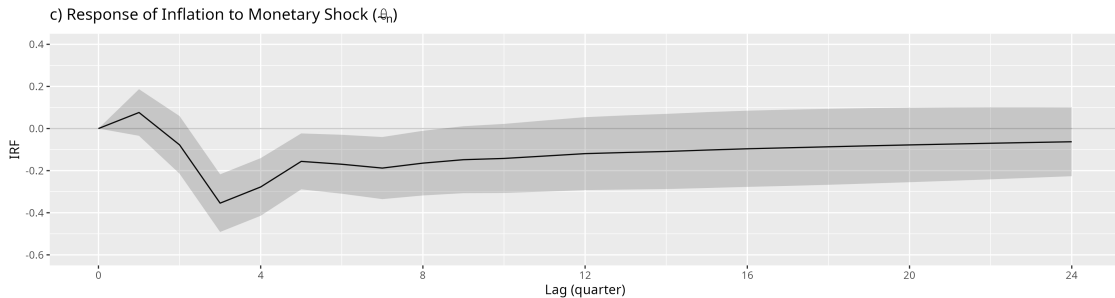
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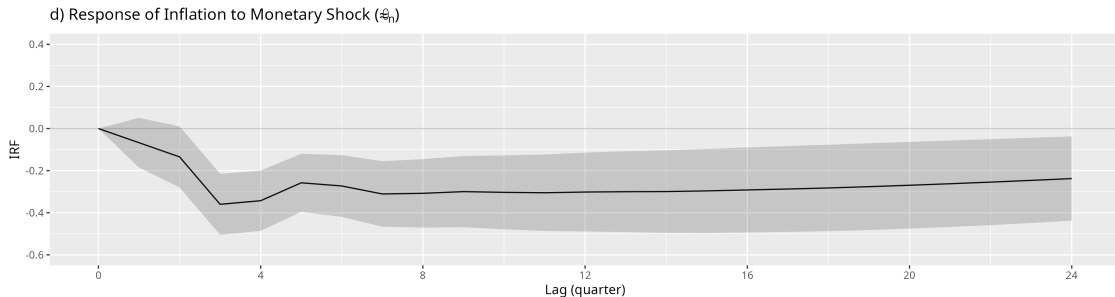
# Price Puzzle: Robust, No Bias Correction



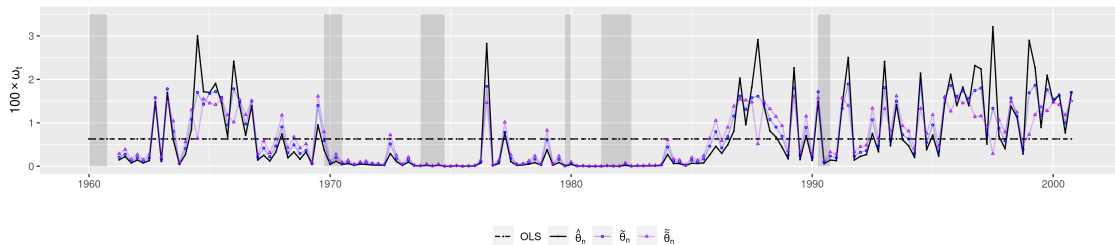
# Price Puzzle: Robust, Bias Correction



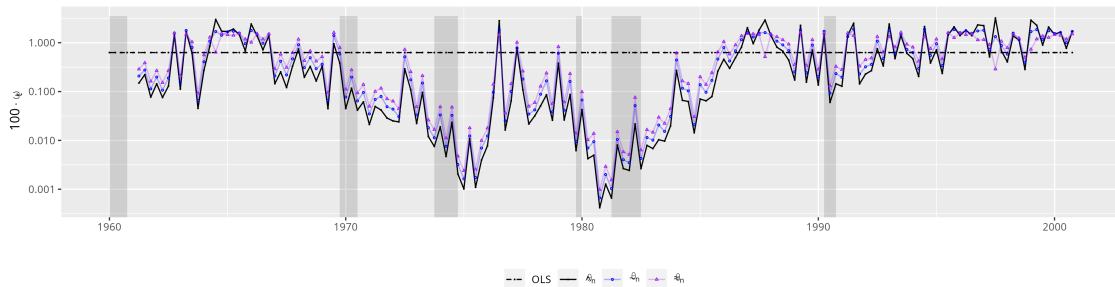
# Price Puzzle: Robust, Repeat Bias Correction



# Price Puzzle: Estimation Weights



# Price Puzzle: Estimation Weights (log scale)



# Conclusion

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- Misspecification can occur
  - don't always know which  $t$  are involved
  - diagnostics useful but not definitive
- Robust estimation:
  - more resilient
  - potentially biased/inconsistent
- This paper: simple estimates, bias correction
  - some finite and large sample results
  - robustness to leveraged outliers
  - weights: make the results transparent

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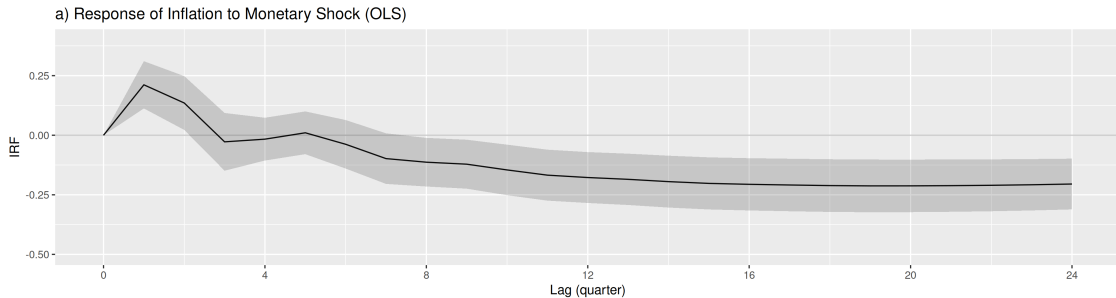
# Back to the Price Puzzle

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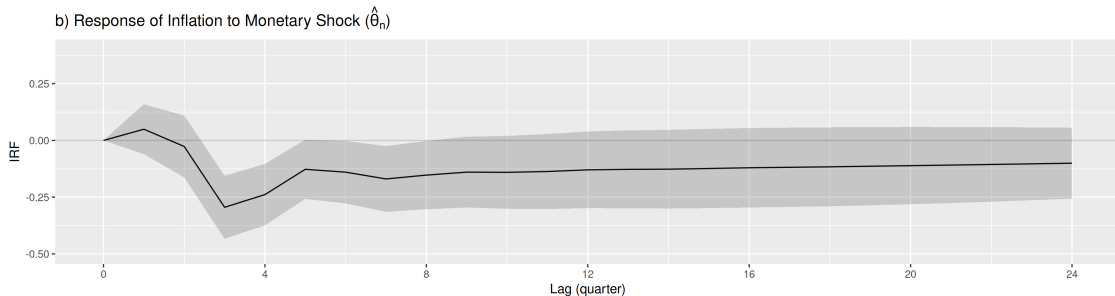
- Re-estimate with a larger  $\nu = 15$  vs.  $\nu = 8.99$  in the main results

# Price Puzzle: OLS

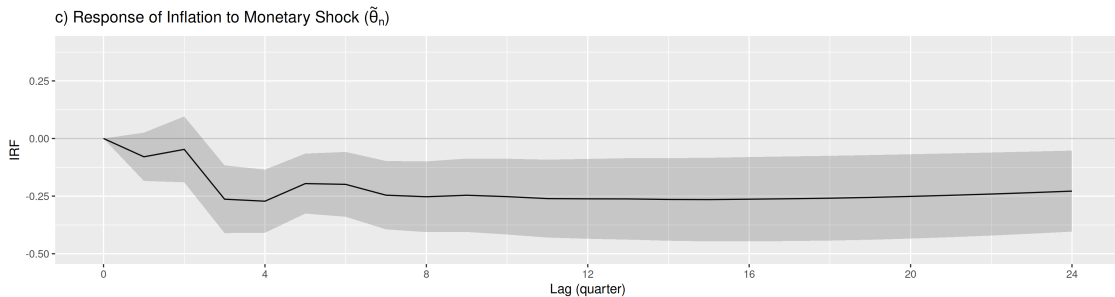
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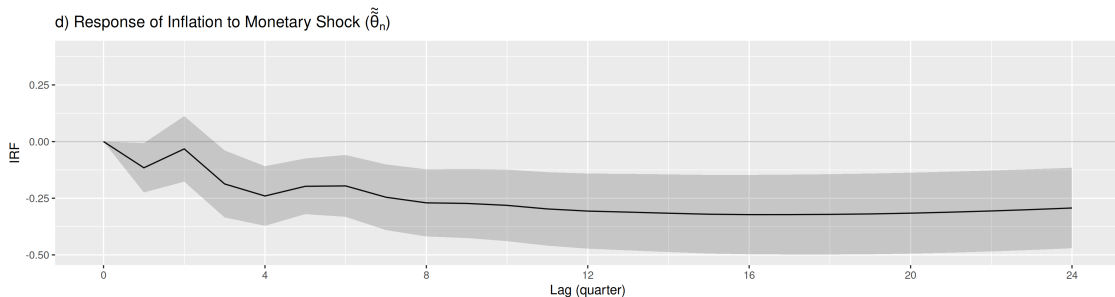
# Price Puzzle: Robust, No Bias Correction



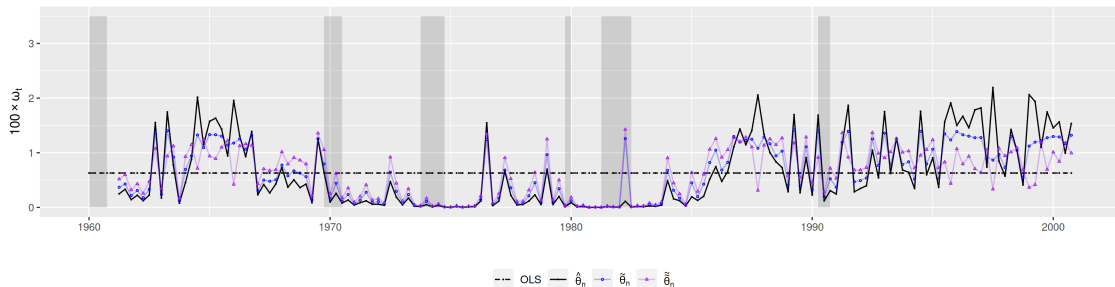
# Price Puzzle: Robust, Bias Correction



# Price Puzzle: Robust, Repeat Bias Correction



# Price Puzzle: Estimation Weights



# Price Puzzle: Estimation Weights (log scale)

