

# Marginal structural models for clustered data: the positivity and no unmeasured confounding assumptions

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1. Motivating example: dialysis and the ANZDATA Registry
2. Clustering by dialysis centre
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**Question:** which dialysis treatment modality is associated with longest survival times?

- Haemodialysis (HD)
  - **Home HD:** performed by the patient at home;
  - **Facility HD:** performed in a hospital/dialysis centre.
  - Vascular access (VA) types:
    - Arterio-venous fistula or graft: **AVF/AVG**
    - Central venous catheter: **CVC**
- **Peritoneal dialysis (PD)**

**ANZDATA:** Australia and New Zealand Dialysis and Transplant Registry

- Collects data from all dialysis patients in Australia and NZ.
- Changes between treatment modalities recorded as they occur.
- Data (including comorbidities, vascular access) collected at dialysis start and at yearly surveys.

# The ANZDATA dataset used for analysis

All patients commencing dialysis between October 1 2003 and December 31 2011, undergoing at least 90 days of dialysis.

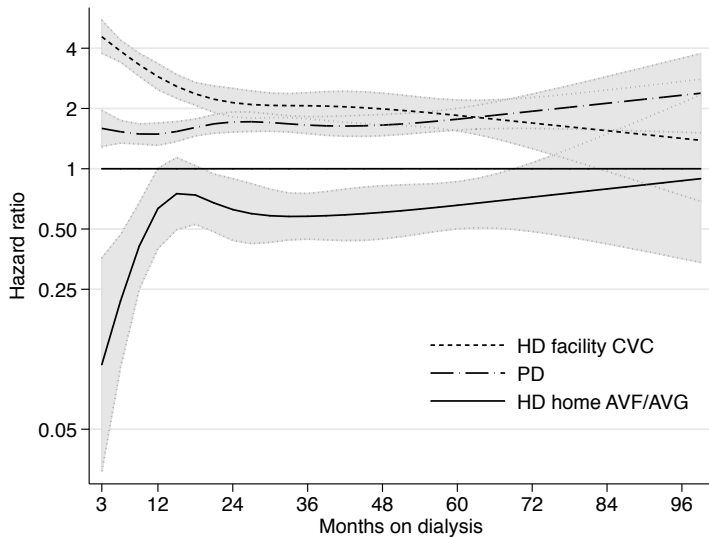
## **20,191 patients:**

- 210,741 90-day periods of follow-up
- 6,971 deaths
- 2,966 kidney transplants
- 267 recovered kidney function

Over their treatment course, **30% of all patients had changes** in dialysis modality/VA

- Modality/VA choice thought to be affected by, and affect, comorbidities (e.g. coronary artery disease).
- We use marginal structural models (MSMs) to estimate the effect of modality/VA on mortality.

# Estimated HRs, relative to facility HD AVF/AVG



# MSM assumptions

- Consistency (treatment version irrelevance);
- Positivity: all patients have a positive probability of receiving all treatments;
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## Problems:

- 1 Patients are **clustered** within dialysis centres.
- 2 ANZDATA is a registry, so set of **measured confounders is limited**. Furthermore, the impact of unmeasured confounding may differ across clusters.

How do these problems impact upon validity of causal inference assumptions?

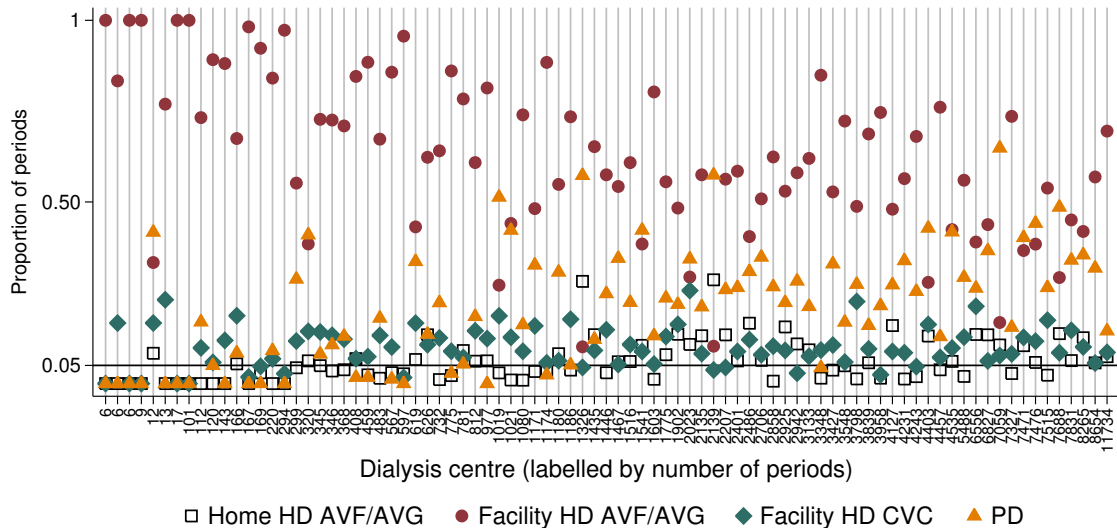


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# Clustering by dialysis centre

- All patients have a dialysis centre which is responsible for administering their treatment.
  - 85 dialysis centres are represented in our dataset.
  - There are differences in practice and survival across centres.
- An extreme difference: **not all dialysis types are available/represented in all centres** (or occur rarely within a centre).
  - In violation of the positivity assumption...

# Clustering of treatments within the 85 centres



# Clustering and the positivity assumption

**Positivity:** patient  $i$  at time  $t$ , on treatment  $A_i(t) = a \in \mathcal{A}$ , with baseline covariates  $V$  and time-varying covariates  $L$

$$\frac{P(A_i(t) = a | A_i(t-1), V_i)}{P(A_i(t) = a | A_i(t-1), L_i(t-1), V_i)} < \infty \quad \forall a \in \mathcal{A}$$

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**Cluster positivity:** treatments available at centre  $\mathcal{C}_j$  denoted by  $a_k \in \mathcal{A}_j$

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If only cluster positivity holds, HRs for particular treatments only defined for centres in which treatment option is available.

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- Include only those centres in which all treatments are possible
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**Both approaches:** to account for unexplained variation between centres, include fixed effects for centres in treatment, censoring and survival models.

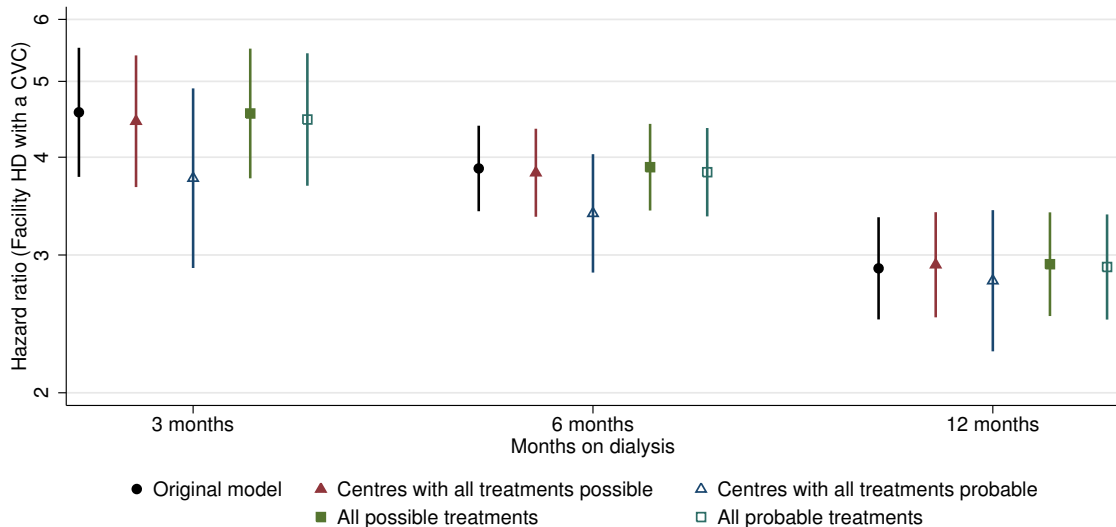
# Analyses accounting for clustering by centre $C_j$ , treatments $a \in \mathcal{A}$

Centres with only one treatment possible/probable must be excluded from all analyses.

- Exclude 11  $C_j$  with  $< 150$  periods (545 periods excluded in total)
- Leaves **74 centres,  $\approx 208000$  periods**

	Restriction	Included centres	Total no. of periods (1000s)
1	Centres w/ $P(A = a C_j) > 0, \forall a \in \mathcal{A}$	68	192
2	Centres w/ $P(A = a C_j) > 0.05, \forall a \in \mathcal{A}$	34	128
3	Treatments w/ $P(A = a C_j) > 0$	74	208
4	Treatments w/ $P(A = a C_j) > 0.05$	70	207

# HRs for Facility HD CVC, accounting for clustering by centre



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# Unmeasured confounding and clustering: confounding function

- $D(t) = 1$  if death at time  $t$
- $D_a(t)$ : **counterfactual** outcome had this patient received dialysis type  $a$ .
- The confounding function:

$$c(a) = \frac{P(D_a(t) = 1 | A(t) = a, V = v)}{\frac{1}{\sum_{a^* \in \mathcal{A} \setminus \{a\}} P(a^*)} \sum_{a^* \in \mathcal{A} \setminus \{a\}} P(a^*) P(D_a(t) = 1 | A(t) = a^*, V = v)},$$
$$P(a^*) = P(A(t) = a^* | V = v).$$

- $c(a)$ : **HR of death comparing patients on  $a$  to those not on  $a$ , had those patients been (contrary to the fact!) on  $a$ .**

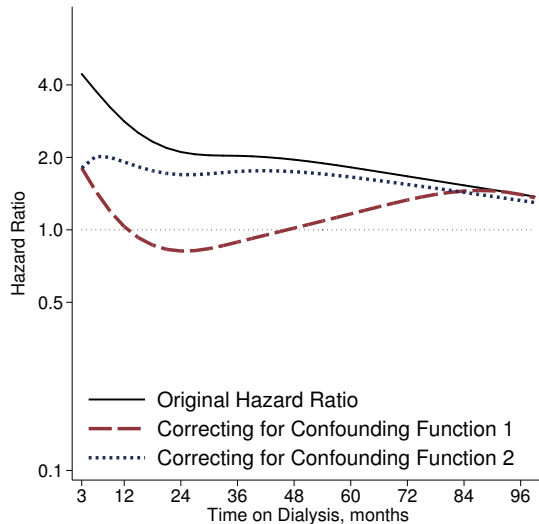
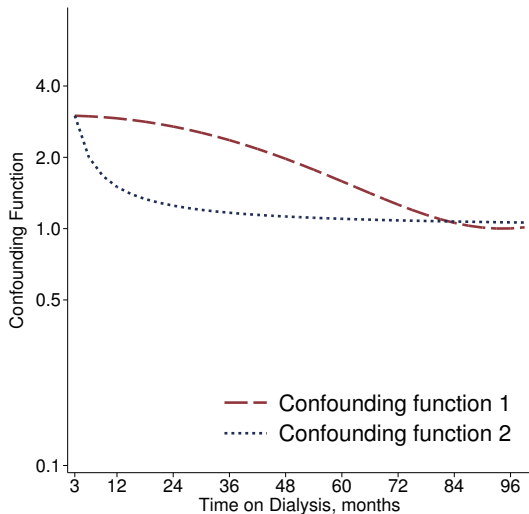
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- $c(a)$ : **HR of death comparing patients on  $a$  to those not on  $a$ , had those patients been (contrary to the fact!) on  $a$ .**
  - $c(a) = 1$ : no difference in the risk of death of patients on  $a$  and those not on  $a$ .
  - **$c(\text{Facility HD CVC}) > 1$** : Facility HD CVC patients have a greater risk of death than those patients on PD/ Home HD/ Facility HD AVF/AVG (had those patients been on Facility HD CVC).
- Possible that the impact of unmeasured confounding differs across dialysis centres.





# HRs accounting for unmeasured confounding: Facility HD + CVC



- Clustering is not often accounted for in the application of MSMs:
  - The link between clustering and the positivity assumption must be considered!
  - If treatment options are restricted (instead of centres): HRs defined only for those centres in which the treatment is available.
  - There remain questions about the best way to account for clustering.
- Sensitivity of the estimates to the impact of unmeasured confounding should be assessed:
  - Confounding functions are useful in this assessment
- Future research: quantifying the amount of unmeasured confounding accounted for by including cluster effects.



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