# Joint stochastic theory of bedload transport and bed elevations: derivation of heavy-tailed resting times

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# **Key Points:**

- We model fluvial bedload activity and local bed elevation as a two-species stochastic birth-death process.
- Computations show heavy-tailed power-law distributions of resting times for sediment undergoing burial, with universal tail parameters  $\alpha \approx 1.15$ .
- We discuss implications for bedload diffusion and suggest a new theoretical framework to approach the problem.

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#### Abstract

A consensus has formed that fluvial bedload resting times lie on heavy-tailed statistical distributions which may result from sediment burial. However, due to observational difficulties, only a handful of experiments have resolved these distributions, and there have been few theoretical attempts to build understanding, leaving their generating mechanism and specific characteristics uncertain. With reference to these issues, we present a new theory describing bedload transport and bed elevation changes as a joint stochastic process, deriving resting time distributions for sediment undergoing burial from the joint dynamics. Our theory predicts heavy-tailed power-law distributions of resting times with universal tail behavior completely characterized by the mean erosion rate and its scaling with bed elevation changes.

## 1 Introduction

The majority of classic studies into fluvial sediment transport have attempted to relate the bulk downstream flux of bedload to characteristics of the hydraulic forcing (e.g. Yalin, 1972), yet the relevance of this approach to environmental problems is limited, as many contemporary issues require knowledge of differences between motions of individual grains, and not just average characteristics. For example, the export of contaminants from channels (e.g. Malmon, Reneau, Dunne, Katzman, & Drakos, 2005) and the morphological response of channels to ecological restoration efforts (e.g. Gaeuman, Stewart, Schmandt, & Pryor, 2017) or changes in hydrology or sediment supply (e.g. Hassan & Bradley, 2017) are not described by bulk fluxes, highlighting individual sediment motions as an important topic for geophysics research.

A significant complication is that individual grains transport within a noisy environment, which noise sources ranging across spatial and temporal scales from fluid turbulence (Celik, Diplas, & Dancey, 2014) and the variable arrangement of bed surface grains (Gordon, Carmichael, & Isackson, 1972), to channel morphology changes (Hassan & Bradley, 2017) and unsteady flows (Phillips, Martin, & Jerolmack, 2013). As a result, the transport characteristics of individual grains are not deterministic (e.g. Einstein, 1937), even in the most controlled laboratory experiments (e.g. Böhm, Ancey, Frey, Reboud, & Ducottet, 2004; Charru, Mouilleron, & Eiff, 2004; Fathel, Furbish, & Schmeeckle, 2015; Heyman, Bohorquez, & Ancey, 2016).

This led researchers to create probabilistic theories of individual motions based on random walk concepts, whereby bedload motions are approximated as alternating sequences of steps and rests, with step lengths and resting times treated as random variables associated with statistical distributions (N. D. Bradley & Tucker, 2012; Einstein, 1937; Hassan, Church, & Schick, 1991; Nakagawa & Tsujimoto, 1976; Yano, 1969). In these theories, differences between the random motions of one grain and the next imply bedload diffusion, or a spreading apart of grains as they transport. The diffusion characteristics predicted by these models critically differ depending on whether the step length and resting time distributions have light or heavy tails (e.g. D. N. Bradley, 2017).

Heavy-tailed distributions have exceedance functions  $P(X>x)\sim x^{-\alpha}$  with tail parameters  $\alpha<2$ , meaning large values of x are relatively common, while light-tailed distributions have  $\alpha\geq 2$ , meaning large values of x are relatively rare. If both resting time and step distance distributions have light tails, the diffusion is said to be normal or Fickian, with a variance of particle positions  $\sigma_x^2$  scaling with time t as  $\sigma_x^2 \propto t$ . However, if either distribution has a heavy-tail, the diffusion is called anomalous, with a variance of particle position scaling as  $\sigma_x^2 \propto t^{\gamma}$ , where  $\gamma \neq 1$ . In this expression,  $\gamma < 1$  is called sub-diffusion and  $\gamma > 1$  is super-diffusion. In strongly asymmetric random walks such as bedload transport, heavy-tailed step lengths imply super-diffusion, while heavy-tailed resting times imply either super or sub-diffusion, depending on  $\alpha$  (Weeks & Swinney, 1998; Weeks, Urbach, & Swinney, 1996).

Tracer experiments in gravel bed rivers show anomalous bedload diffusion (D. N. Bradley, 2017; Phillips et al., 2013), light-tailed step lengths (N. D. Bradley & Tucker, 2012; Hassan, Voepel, Schumer, Parker, & Fraccarollo, 2013), and heavy-tailed resting times (D. N. Bradley, 2017; Olinde & Johnson, 2015; Pretzlav, 2016; Voepel, Schumer, & Hassan, 2013), forming a coherent experimental picture of super-diffusive bedload transport, at least at long observation timescales (e.g. Martin, Jerolmack, & Schumer, 2012; Nikora, 2002). However, field studies have not resolved the mechanism generating these heavy-tailed resting times (e.g. D. N. Bradley, 2017), and empirical distributions display clear differences in their form and characteristics, showing different tail parameters (e.g. Olinde & Johnson, 2015) and sometimes truncation (e.g. D. N. Bradley, 2017) or tempering to light tails at large resting times (e.g. Voepel et al., 2013). The mechanism generating heavy tails and these differences deserve attention.

A predominant hypothesis is that heavy-tailed resting times and anomalous diffusion originate from sediment burial (Martin, Purohit, & Jerolmack, 2014; Voepel et al., 2013; Wu et al., 2019). Conceptually, when grains rest on the bed surface, material transported from upstream can deposit on top of them, preventing entrainment until it's removed, driving up resting times and imparting a heavy tail to the distribution. Martin et al. (2014) have provided the only direct support for this hypothesis. They traced grains in a narrow flume with clear sidewalls, directly resolving burial as the generator of heavy-tailed resting times, and they described their results with a theoretical model, formally similar to an earlier effort by Voepel et al. (2013).

The models of Voepel et al. (2013) and Martin et al. (2014) treat bed elevations as a random walk and interpret resting times as return periods from above in the bed elevation time-series (e.g. Redner, 2007). Both models are successful in describing different experimental resting time distributions. However, the assumptions and results of these models are inconsistent with one another, and their treatment of bed elevations as a process independent of sediment transport is questionable, since erosion and deposition are the source of bed elevation changes (e.g. Wong, Parker, DeVries, Brown, & Burges, 2007), suggesting further study is necessary.

In this work, we approach the problem from a different angle, making an extension of the stochastic bedload transport theory of Ancey, Davison, Böhm, Jodeau, and Frey (2008) to link bed elevation changes to the erosion and deposition events of individual grains, and we derive resting times as a consequence of this theory. The key assumptions of our model are: (1) bedload erosion and deposition can be characterized by probabilities per unit time, or rates (e.g. Ancey et al., 2008; Einstein, 1950); and (2) these rates are contingent on the local bed elevation, encoding the property that erosion of sediment is emphasized from regions of exposure, while deposition is emphasized in regions of shelter (e.g. Sawai, 1987; Wong et al., 2007). Our theory generates heavy-tailed distributions with no tempering and a universal tail parameter  $\alpha \approx 1.1$  for a particular non-dimensionalization of the resting time, showing close correspondence to the results of Martin et al. (2014) and suggesting a correction to some imperfections in their results. We conclude the paper by framing our work in relation to earlier ideas and discussing the implications of this work on questions of individual bedload motions and anomalous diffusion.

# 2 Stochastic theory

We define a volume of downstream length L which contains some number n of moving particles in the water flow and some number m of stationary particles composing the bed at some time t. For simplicity, we consider all particles as approximately spherical with the same diameter 2a, so their mobility and packing characteristics are similar. Following Ancey et al. (2008), we prescribe four events which can occur at any instant to modify the populations n and m, and we characterize these events using probabilities per unit time, or rates. These are: (1) migration of a moving particle into the volume from upstream

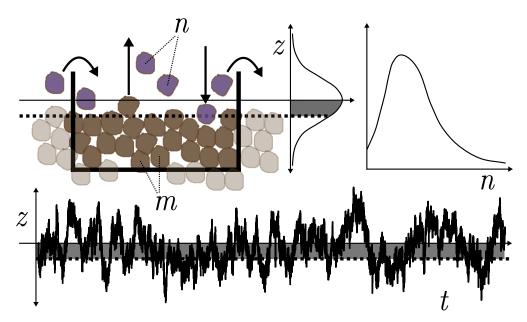


Figure 1. Definition sketch of a control volume containing n moving grains and m resting grains. Migration, entrainment, and deposition processes are represented by curved arrows, and the bed elevation at some instant is depicted by dotted line. The bed is presented in a degraded state, where  $m < m_0$ . The distributions of n and m are indicated in the upper right panel, while the bottom panel is a time-series of bed elevations (1).

 $(n \to n+1)$ ; (2) the entrainment of a stationary particle into motion within the volume  $(m \to m-1)$  and  $n \to n+1$ ; (3) the deposition of a moving particle to rest within the volume  $(m \to m+1)$  and  $n \to n-1$ ; and (4) the migration of a moving particle out of the volume to downstream  $(n \to n-1)$ . As the events occur at random intervals, they set up a joint stochastic evolution of the populations n and m characterized by a joint probability mass function (pmf) P(n,m,t) having marginal pmfs  $P(n,t) = \sum_{m} P(n,m,t)$  and  $P(m,t) = \sum_{n} P(n,m,t)$  for the number of particles in motion and rest in the volume at t. These concepts are depicted in figure 1.

The populations n and m provide the bulk bedload flux  $q_s$  and the local bed elevation z. The mean bedload transport rate is given by  $q_s \propto u_s \langle n \rangle$ , where  $u_s$  is the characteristic velocity of moving bedload and  $\langle n \rangle = \sum_{n,m} n P(n,m)$  is the mean number of grains in motion (e.g. Ancey et al., 2008; Charru et al., 2004; Furbish, Haff, Roseberry, & Schmeeckle, 2012). The bed elevation is related to m though the packing geometry of the bed. To derive this, we prescribe a mean number of grains at rest  $m_0$  and introduce a packing fraction  $\phi$  of grains in the bed. Then considering a two-dimensional bed (e.g. Einstein, 1950; Paintal, 1971), the deviation from the mean bed elevation is

$$z(m) = \frac{\pi a^2}{\phi L}(m - m_0) = z_1(m - m_0). \tag{1}$$

The constant  $z_1 = \pi a^2/(\phi L)$  is an important scale of the problem.  $z_1$  is the magnitude of bed elevation change (in an average sense across the control volume) associated with the addition or removal of a single grain. We write the rates of the four possible transitions as (e.g. Ancey et al., 2008):

$$R_{MI}(n+1,m|n,m) = \nu$$
 migration in, (2)

$$R_E(n+1, m-1|n, m) = \lambda(m) + \mu(m)n$$
 entrainment, (3)

$$R_D(n-1, m+1|n, m) = \sigma(m)n$$
 deposition, (4)

$$R_{MO}(n-1, m|n, m) = \gamma n$$
 migration out. (5)

These rates are independent of the past history of the populations and depend only on the current populations (n, m). As a result, the system is Markovian (e.g. Cox & Miller, 1965; van Kampen, 1992), meaning time intervals between subsequent transitions are exponentially distributed (e.g. Gillespie, 2007).

In (2-5),  $\nu$  and  $\gamma$  are constants characterizing migration rates of individual grains into and out of the volume. They lack any dependence on the populations n and m. In contrast,  $\lambda(m)$ ,  $\mu(m)$ , and  $\sigma(m)$ , characterizing the entrainment, collective entrainment (e.g. Ancey et al., 2008; Heyman, Ma, Mettra, & Ancey, 2014; Heyman, Mettra, Ma, & Ancey, 2013), and deposition rates of individual grains are considered to depend on m. As is well-known, bed elevation changes modify the likelihood of entrainment and deposition in a negative feedback (Sawai, 1987; Wong et al., 2007); that is, aggradation increases the likelihood of entrainment, while degradation increases the likelihood of deposition. Wong et al. (2007) concluded that bed elevation changes induce an exponential variation in entrainment and deposition probabilities, while Sawai (1987) concluded that the variation is linear. For simplicity, we incorporate the scaling of Sawai (1987) and note its equivalence to the Wong et al. (2007) scaling when bed elevation changes are small. Because experimental distributions of bed elevations are usually symmetrical, (Martin et al., 2014; Singh, Fienberg, Jerolmack, Marr, & Foufoula-Georgiou, 2009; Wong et al., 2007), we expect the erosion and deposition feedbacks to be anti-symmetrical. That is, as bed elevation changes drive up (down) erosion rates, so they drive down (up) deposition rates to the same degree.

Summarizing these ideas, the entrainment and deposition rates can be written  $\chi(m) = \chi_0(1 \pm z_1 z(m)/(2l)^2)$ , where  $\chi = \lambda, \mu, \sigma$ , and the entrainment parameters take the plus sign, while deposition takes the minus, and we have introduced a length scale l. As we'll see, the variance of bed elevation turns out to be given by  $\text{var}(z) = (lz_1)^2$ . Accordingly, l characterizes the range of bed elevation variations, which could be interpreted as the active layer depth (e.g. Church, 2017). Another perspective is that l is the distance of bed elevation change at which the entrainment and deposition rates are significantly affected. With these substitutions, the local bed elevation-dependent entrainment and deposition rates (3-4) can be written:

$$R_E(n+1, m-1|n, m) = [\lambda_0 + \mu_0 n] \left[ 1 + \frac{z_1 z(m)}{(2l)^2} \right],$$
 entrainment, (6)

$$R_D(n-1, m+1|n, m) = \sigma_0 \left[ 1 - \frac{z_1 z(m)}{(2l)^2} \right] n,$$
 deposition. (7)

At z(m) = 0, the rates reduce to those of the Ancey et al. (2008) theory. Away from this elevation, entrainment and deposition are alternatively suppressed and accentuated depending on the sign of z(m).

In terms of the transition rates (2-7), we can obtain the Master equation for the probability flow using the forward Kolmogorov equation  $\partial P(n,m;t)/\partial t = \sum_{n',m'} R(n,m)P(n',m';t)$  (e.g. Ancey et al., 2008; Cox & Miller, 1965; Gillespie, 1992) as

$$\frac{\partial P}{\partial t}(n,m;t) = \nu P(n-1,m;t) + \{\lambda(m+1) + [n-1]\mu(m+1)\}P(n-1,m+1;t) + [n+1]\sigma(m-1)P(n+1,m-1;t) + [n+1]\gamma P(n+1,m;t) - \{\nu + \lambda(m) + n\mu(m) + n\sigma(m) + n\gamma\}P(n,m;t).$$
(8)

The joint probability distribution P(n,m;t) solving this equation will fully characterize the statistics of n and m. We anticipate that solutions will adjust from the initial conditions to a steady-state distribution  $P_s(n,m)$ , independent of time, if the constant factors in the transition rates are representative of equilibrium bedload transport conditions. This Master equation describes a two-species stochastic birth-death model (e.g. Cox & Miller, 1965) of a type well-known in the population ecology literature (e.g. Pielou, 1977; Swift, 2002). In our context, the two species are the moving and stationary grains in the volume.

#### 3 Numerical simulations

Unfortunately, (8) does not appear to admit an analytical solution (but see Swift (2002) for a standard method which fails in this case). The difficulty stems from the product terms between n and m. In response, we resort to numerical methods, simulating (8) using the Gillespie algorithm (Gillespie, 1977, 1992, 2007). The Gillespie algorithm leverages the defining property of a Markov process: when transition rates do not depend on the past, time intervals between transitions are exponentially distributed (e.g. Cox & Miller, 1965).

**Table 1.** Parameters from Ancey et al. (2008) experiments describing the rates of migration in, entrainment, deposition, and migration out when z(m) = 0. All units are  $s^{-1}$  (probability/time). Elevation changes modulate these rates in accord with (??-??)

Flow	ν	$\lambda_0$	$\mu_0$	$\sigma_0$	$\gamma$
(a)	5.45	6.59	3.74	4.67	0.77
(g)	7.74	8.42	4.34	4.95	0.56
(i)	15.56	22.07	3.56	4.52	0.68
(1)	15.52	14.64	4.32	4.77	0.48
(n)	15.45	24.49	3.64	4.21	0.36

Therefore, to step the Markov process through a single transition, it's enough to draw a random value from the exponential distribution of transition intervals to determine the time of the next transition. Then, drawing another random value to choose the type of transition which occurs using the relative probabilities (2-7), the transition can be enacted by stepping t,, n and m by the appropriate shifts. This procedure can be iterated to form an exact realization of the stochastic process (e.g. Gillespie, 2007).

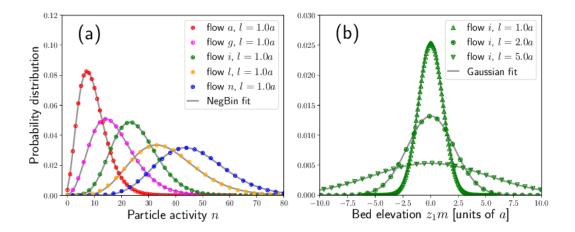
Using this algorithm, we simulated 5 flow conditions with 10 different values of l taken across a range from l=a (1 grain radius) to l=10a (10 grain radii). These values lie in the range exhibited in the majority of available experimental data (Martin et al., 2014; Singh et al., 2009; Wong et al., 2007). For the migration, entrainment, and deposition parameters at each flow condition  $(\nu, \lambda_0, \mu_0, \sigma_0, \gamma)$ , we used values mea-

sured by Ancey et al. (2008) in a series of flume experiments. These are summarized in table 1. Flow conditions are labeled (a), (g), and so on, roughly in order of increasing bedload flux (see Ancey et al. (2008) for more details). In all simulations, we take the packing fraction  $\phi=0.6$ , a typical value for a pile of spheres (e.g. Bennett, 1972), and set L=22.5cm and a=0.3cm, in accord with the Ancey et al. (2008) experiments. Each simulation was run for 1500hrs of virtual time, a period we selected by trial and error to ensure sufficient convergence of the statistics.

## 4 Results

Our simulations show intensely noisy time-series of bedload activities n, and temporally correlated bed elevation time-series (which can be seen in the bottom panel of figure 1). From our chosen initial conditions, all simulations show a rapid attainment of steady state conditions followed by a stable stochastic dynamics of n and m which support a stable time-dependent probability  $P_s(n,m)$ . We compute this joint distribution by counting occurrences of the states (n,m) in the simulated time series. From this joint distribution we compute marginals P(n) and P(m) as explained in section 2.

Some of these marginal distributions are displayed in figure 2. Neglecting changes in bed elevation, Ancey et al. (2008) analytically derived negative binomial (NegBin) distributions for the bedload activity n, and this functional form appears to be preserved through the coupling of their stochastic theory to bed elevation changes, with all of our computed distributions accepting clean NegBin fits. For m, our computations obviously show Gaussian



**Figure 2.** Marginal distributions for n and m for a small subset of simulations. Some points have been omitted for clarity.

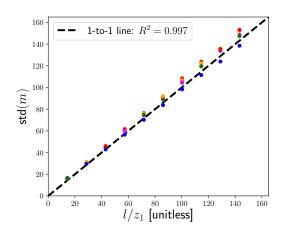
distributions, consistent with our assumptions about the scaling of erosion and deposition rates with bed elevation changes (e.g. Wong et al., 2007).

From the marginal distributions, we calculate means and variances of bedload activity and elevation (m). The mean bed elevation is just the initial condition  $m_0$ . m fluctuates around this value because it sets the equilibrium position of the elevation-related feedbacks in (1). The variance of m appears given by  $z_1^2 \text{var}(m) = l^2$ , as indicated in figure 3. l is a measure of bed elevation fluctuations. The moments of n are more difficult to understand. Without dwelling on the issue, the moments of n shift with the ratio  $l/z_1$ , somehow resulting from feedback between bed elevation changes and bedload transport.

Now we describe the analysis of bedload resting times from time-series of m. Following Voepel et al. (2013) and Martin et al. (2014), we concentrate on a particular bed elevation  $m_*$ , and find all time intervals separating deposition events at  $m = m_*$ from erosion events at  $m = m_* + 1$ . Binning these return times to  $m_*$  and counting the occurrences in each bin (using logarithmically spaced bins for presentation purposes), we obtain a non-exceedance distribution of return times  $t_r$  held conditional to the elevation  $m_*$ :  $P(T > t_r | m_*)$ . Using the marginal probability distribution of bed elevations, we derive the unconditional nonexceedance distribution of resting times as a sum over all elevations (Martin et al., 2014; Nakagawa & Tsujimoto, 1980; Voepel et al., 2013; Yang & Sayre, 1971):

$$P(T > t_r) = \sum_{m'} P(m')P(T > t_r|m'). \quad (9)$$

Some of these results are displayed in figure 4.. Comparing panels (a) and (c) shows that the resting time distributions depend



**Figure 3.** Data from all simulations is plotted to show that l controls the standard deviation of bed elevations:  $l^2 = z_1^2 \text{var}(m)$ . We might interpret l as the depth of the active layer (e.g. Church, 2017).

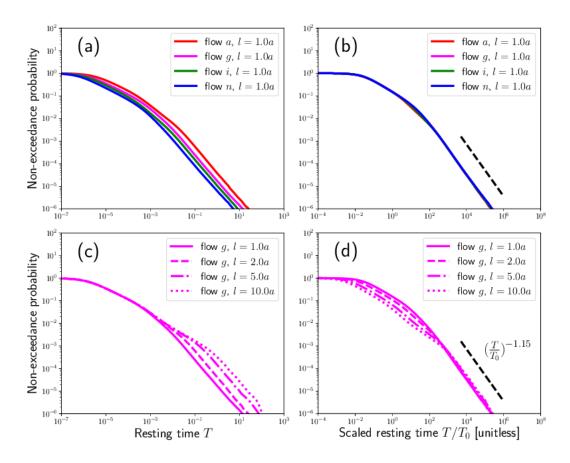


Figure 4. Resting time statistics vary in different ways with flow conditions and the variance of bed elevations. Panel (a) represents differing flow conditions at a fixed l value, while panel (c) is fixed flow conditions at a variable l value. When scaled by  $T_0$  (10), both of types of difference collapse away in the tails of the distributions, as shown in panels (b) and (d). Finally, in panel (d), the black dotted line indicates a power law decay with tail parameter  $\alpha = 1.15$ .

on flow condition and the standard deviation of bed elevation (l) in different ways. However, as shown in panels (b) and (d), a characteristic timescale  $T_0$  can collapse the tails of the distributions regardless.

Heuristically, we can obtain  $T_0$  by thinking of the bed as having a characteristic drift velocity obtained from the mean erosion rate and a certain length scale. Formally, the mean erosion rate is  $E = \sum_{n,m} R_E(n,m)P(n,m)$ . This is the number of grains leaving the bed per unit time. If bedload transport is in equilibrium, the deposition rate D satisfies E = D. Since the removal of a single grain changes the bed elevation by  $z_1$ , the bed elevations change with a characteristic velocity  $z_1E$ . Since the characteristic deviation of elevation is l, the time required for the bed to shift through a characteristic deviation is

$$T_0 = \frac{l}{z_1 E}. (10)$$

This is the timescale collapsing the distributions in figure 4. Apparently, for return times roughly satisfying  $T/T_0 > 10^3$ , all resting time non-exceedance distributions decay as a heavy-tailed power law with parameter  $\alpha \approx 1.15$ .

#### 5 Discussion

Now we summarize our key findings and discuss their scope and implications. First, we have generalized the stochastic bedload theory of Ancey et al. (2008) to provide a joint description of bed elevations and transport.

- 1. You generalized the stochastic bedload theory of ancey to provide joint description of bedload transport and bed elevation changes, including fluctuations and feedbacks. Ancey et al. (2008) was a single cell model of bedload transport which was generalized in Ancey and Heyman (2014) to an N cell model. This facilitated the study of spatial correlations in transport in Heyman et al. (2014). Could the single cell model developed here be a first step toward a stochastic framework from which to study morphology? Maybe, but closure would be a nightmare (e.g. Heyman et al., 2016).
- 2. Now we draw connections to earlier work. The master equation (??) has limits of  $l \to 0$  and  $l \to \infty$  which generate the theories of Martin et al. (2014) and Ancey et al. (2008) as limit cases.
- 3. We predict the expected form of the resting time distribution which would be expected in narrowly graded sediment under equilibrium sediment transport. Of course, it's unclear to what extent this corresponds to a natural stream. At the very least, it's apparently true that flood cycles can sometimes be treated as equilibrium even though they're conceptually not (e.g. Phillips et al., 2013).
- 4. Our resting time distributions collapse to an identical tail for  $T/T_0 > 10^3$  where  $T_0 = l/(z_1 E)$  is a characteristic timescale of bed elevation change. This implies anomalous superdiffusion of bedload tracers. Can describe what this entails.. infinite slowdown of tracers and so on.
- 5. Finally, we corroborate the study of Martin et al. (2014). He gets similar results. However, his theory did not address sediment transport, and his scaling was not complete. Specifically, because the depth of the active layer shifts with discharge, and Martin was only scaling with the factor E, his experiments show a systematic drift in the tail of the resting time distributions that would not collapse away. That's because he didn't involve l.
- 6. Deriving a bed elevation theory from the basis in sediment transport implies equations which do not reproduce or relate to Voepel et al. (2013). The concepts in this theory are useful, but the treatment of bed elevations as a bounded random walk without persistence are incorrect. The OU process of Martin et al. (2014) is more appropriate, especially if it's ammended to include l.
- 7. four implications: (1) joint stochastic theory, (2) heavy-tailed resting times, prototypical,  $\alpha=1.15$ , (3) anomalous superdiffusion of bedload tracers is expected if burial occurs, (4) however, there is still no proof that burial is the mechanism of burial.. discuss D. N. Bradley (2017) suggestion of multiple return processes occurring simultaneously in a channel, highlight the possibility that some of them are even heavier-tailed than burial, which means when they act they obscure the burial signature (i.e. the 1.15 tail of me and Martin.)

Of course, a great deal of effort has been devoted to modeling the coupling between sediment transport and morphology (e.g. Pelosi & Parker, 2014). However, to our knowledge, the simple population dynamics theory of bed elevation changes and transport we've presented here might be the first example of a morphodynamics theory

ay be the first example of what might be called a Lagrangian theory of morphodynamics. That is, a means to predict morphodynamics – in this case just simple bed evolution, from modeling the motion characteristics of individual grains. In closure to this first point, we note the Ancey et al. (2008) model of bedload transport based on a single cell was subsequently extended to an array of adjacent cells (Ancey & Heyman, 2014), generated

a framework which Heyman et al. (2014) used to study spatial correlations in bedload transport.

### 6 Conclusion

Conclude by wrapping discussion with introduction and implications especially. 1 paragraph

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