# Bedload diffusion theory

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### 1 The two-state random walk

The formalism of two-state random walks [e.g. Weiss, 1976, 1994; Masoliver, 2016; Masoliver and Lindenberg, 2017] ultimately composes all existing models of bedload diffusion [e.g. Nikora et al., 2001; Nikora, 2002; Zhang et al., 2012; Fan et al., 2016] and provides a framework to build new models from. Assuming residence time distributions  $\psi_i(t)$  in the states with associated survival probabilities  $\Psi_i(t) = \int_t^{\infty} dt' \psi_i(t')$ , probabilities of moving a distance x up to time t within a sojourn  $f_i(x,t)$ , denoting  $g_i(x,t) = f_i(x,t)\psi_i(t)$  and  $G_i(x,t) = f_i(x,t)\Psi_i(x,t)$ , and letting  $\omega_i(x,t)$  be the probability that a sojourn in state i ends at x,t provides

$$p_i(x,t) = \theta_i G_i(x,t) + \int_0^\infty dx' \int_0^\infty dt' \omega_{\bar{i}}(x',t') G_i(x-x',t-t')$$
(1)

and

$$\omega_i(x,t) = \theta_i g_i(x,t) + \int_0^\infty dx' \int_0^\infty dt' \omega_{\bar{i}}(x',t') g_i(x-x',t-t'). \tag{2}$$

Here  $\theta_i$  are the initial probabilities of being in each state with  $\theta_1 + \theta_2 = 1$ .  $\bar{i}$  is the opposite of i and i = 1, 2. The probability of the two-state random walker being at position x at time t is

$$p(x,t) = p_1(x,t) + p_2(x,t). (3)$$

Denoting the laplace transform with respect to the variable q as  $\mathcal{L}_q$  and associating variables  $\eta$  and s with  $\mathcal{L}_x$  and  $\mathcal{L}_t$ , taking  $\mathcal{L}_x\mathcal{L}_t$  of (2) provides a much simpler algebraic problem for the probability p:

$$\tilde{\omega}_i = \frac{[\theta_i + \theta_{\bar{i}}\tilde{g}_{\bar{i}}]\tilde{g}_i}{1 - \tilde{g}_1\tilde{g}_2} \tag{4}$$

and (c.f. *Masoliver* [2016] eq. 20)

$$\tilde{p}_{\pm} = \left(\theta_i + \tilde{\omega}_{\bar{i}}\right) \tilde{G}_i = \frac{\theta_i + \theta_{\bar{i}} \tilde{g}_{\bar{i}}}{1 - \tilde{g}_1 \tilde{g}_2} \tilde{G}_i. \tag{5}$$

Therefore the double transform of the joint PDF reads [c.f. Weiss, 1994, eq. 6.33 pg. 243]

$$\tilde{p}(\eta, s) = \frac{\theta_1[\tilde{G}_1 + \tilde{g}_1\tilde{G}_2] + \theta_2[\tilde{G}_2 + \tilde{g}_2\tilde{G}_1]}{1 - \tilde{g}_1\tilde{g}_2}.$$
(6)

This is a direct generalization of the famous Montroll-Weiss formula for a single state continuous-time random walk.

For the evaluation of this formula, a useful fact to take account of is

$$\mathcal{L}_t\{\Psi_{\pm}(t); s\} = \int_0^\infty dt e^{-st} \int_t^\infty dt' \psi_{\pm}(t') = \frac{1 - \psi(s)}{s}.$$
 (7)

Finally, owing to the definition of the double Laplace transform of p(x,t):

$$\tilde{p}(\eta, s) = \int_0^\infty dt e^{-st} \int_0^\infty dx e^{-\eta x} p(x, t) \tag{8}$$

we see the (double) inverse transform p(x,t) is not necessary to study the moments  $\langle x(t)^k \rangle = \int_0^\infty x^k p(x,t)$  of an ensemble of tracers since the moments follow from

$$\mathcal{L}_t\{\langle x(t)^k \rangle; s\} = (-)^k \partial_{\eta}^k \tilde{p}(\eta, s) \Big|_{\eta=0}. \tag{9}$$

### 1.1 Moments of a two-state random walk

Weeks and Swinney [1998] starts in motion, for future reference.

### 2 The Einstein theory

Taking  $g_1(x,t) = \delta(x)k_1e^{-k_1t}$  (rest) and  $g_2(x,t) = k_2e^{-k_2x}\delta(t)$  (step) reproduces the *Einstein* [1937] diffusion theory. In this case the double transforms are:

$$\tilde{g}_1(\eta, s) = \frac{k_1}{k_1 + s} \tag{10}$$

$$\tilde{g}_2(\eta, s) = \frac{k_2}{k_2 + \eta},$$
(11)

and the survival functions are

$$\Psi_1(t) = \int_t^\infty dt' k_1 e^{-k_1 t'} = e^{-k_1 t} \tag{12}$$

and

$$\Psi_2(t) = \int_t^\infty dt' \delta(t') = 0, \tag{13}$$

meaning  $G_1(x,t) = \delta(x)e^{-k_1t}$  and  $G_2(x,t) = 0$ , providing

$$\tilde{G}_1(\eta, s) = \frac{1}{k_1 + s}. (14)$$

Taking  $\theta_1 = 1$  and  $\theta_2 = 0$ , so the dynamics start at rest, the MW generalization (6) is

$$\tilde{p}(\eta, s) = \frac{\tilde{G}_1}{1 - \tilde{g}_1 \tilde{g}_2} = \frac{1}{s + \frac{k_1 \eta}{k_2 + n}}.$$
(15)

The Laplace transform of the mean is

$$\langle \tilde{x} \rangle = \frac{k_1}{s^2 k_2} \tag{16}$$

so in real space it's  $\langle x \rangle = k_- t/k_+$  as expected [e.g. Einstein, 1937; Nakagawa and Tsujimoto, 1976]. The Laplace transform of the second moment is

$$\langle \tilde{x}^2 \rangle = 2 \left( \frac{k_1}{k_2} \right)^2 \left[ \frac{1}{k_1} \frac{1}{s^2} + \frac{1}{s^3} \right],$$
 (17)

implying a second moment  $\langle x^2 \rangle = (k_1/k_2)^2 [2t/k_1 + t^2]$  and a variance exemplifying the normal diffusion of bedload:

$$\sigma_x^2 = \frac{2k_1}{k_2^2}t. (18)$$

This is depicted in figure 1. Of course, for the Einstein theory a closed form solution of the pdf p(x,t) is possible to obtain [e.g. Einstein, 1937; Hubbell and Sayre, 1964; Daly and Porporato, 2006; Daly, 2019]. The first transform in (15) inverts easily for

$$\tilde{p}(\eta, t) = \exp\left\{-\frac{k_1 \eta}{k_2 + \eta} t\right\}. \tag{19}$$

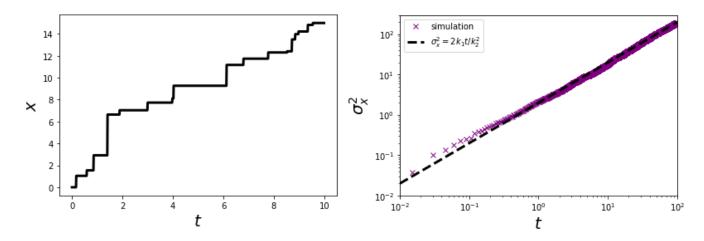


Figure 1: Left panel shows the Einstein walk in x-t space while the right panel shows the linear variance. There is a single range of normal diffusion when steps are instantaneous.

Incidentally this single Laplace transform of p provides the cumulant generating function  $c(\eta,t) = \log \tilde{p}(-\eta,t)$  from which the variance follows from a more simple computation:  $\sigma_x^2(t) = \partial_\eta^2 c(\eta,t)\big|_{\eta=0}$ . The second inversion gives [e.g. Daly, 2019]

$$p(x,t) = e^{-k_1 t - k_2 x} \left\{ \sqrt{\frac{k_1 k_2 t}{x}} I_1 \left( 2\sqrt{k_1 k_2 x t} \right) + \delta(x) \right\}.$$
 (20)

This exact solution has been the benchmark theory of bedload diffusion for over 100 years. I have only derived it within a more general framework of multi-state random walks [e.g. Weiss, 1994].

A final note – taking the expression for  $\tilde{p}(\eta, t)$  and inverting as in Weiss [1994] pg. 247 gives the equation

$$[k_1\partial_x + k_2\partial_t + \partial_x\partial_t]p = 0, (21)$$

after some jangling. I'm surprised this is not the normal diffusion equation. I'm curious if it's correct and if the solution involving  $I_1$  solves it. Decomposing  $p = e^{-k_1 t - k_2 x} \pi$  gives

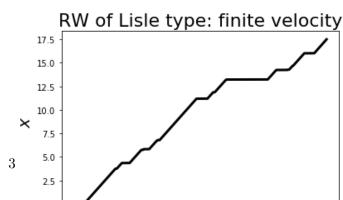
$$\partial_x \partial_t \pi = k_1 k_2 \pi. \tag{22}$$

Does this provide  $\pi \propto \sqrt{\frac{k_1 k_2 t}{x}} I_1(2\sqrt{k_1 k_2 x t})$ ? Probably not.

# 3 The Lisle Theory

Apart from formulations of *Einstein* [1937] using different step length and resting time distributions than exponential [e.g. *Sayre and Hubbell*, 1965], the first significant advancement from *Einstein* [1937] was due to *Lisle et al.* [1998]. This type of random walk is depicted in figure 2.

I need to carefully investigate whether Gordon et al. [1972] did it too. They imparted a finite duration to bedload motions instead of considering them instantaneous like Einstein. In this way they derived two stages of bedload diffusion. This approach is closely related to the so-called persistent diffusion model [Balakrishnan and Chaturvedi, 1988; Van Den Broeck, 1990], the diffusion of a particle driven



by dichotomous Markov noise [e.g. Horsthemke and Lefever, 1984; Risken, 1989; Bena, 2006]. The mathematics were essentially developed by Takacs (1957).

It is obtained by the choice  $g_1(x,t) = \delta(x)k_1e^{-k_1t}$  (rest) and  $g_2(x,t) = \delta(x-vt)k_2e^{-k_2t}$  (motion). Hence motions occur with velocity v for a duration characterized by an exponential distribution with mean  $1/k_2$ , while rests occur for a duration characterized by an exponential distribution with mean  $1/k_1$ . We consider each of the extreme initial conditions in turn.

#### 3.1 Rest initial state

If the process starts from rest, this means  $\theta_1 = 1$  and  $\theta_2 = 0$ . In this case the Laplace transforms are

$$\tilde{g}_1(\eta, s) = \frac{k_1}{k_1 + s},$$
(23)

$$\tilde{g}_2(\eta, s) = \frac{k_2}{k_2 + \eta v + s} \tag{24}$$

and

$$\tilde{G}_i(\eta, s) = \frac{1}{k_i} \tilde{g}_i(\eta, s). \tag{25}$$

Plugging these into (6) gives

$$\tilde{p}(\eta, s) = \frac{k + s + \eta v}{(k_1 + s)\eta v + (k + s)s},\tag{26}$$

where  $k = k_1 + k_2$ . This inverts to

$$\tilde{p}(x,s) = \frac{k(k+s)}{v(k_1+s)^2} \exp\left[-\frac{s(k+s)}{v(k_1+s)}x\right] + \frac{1}{k_1+s}\delta(x). \tag{27}$$

Using the property  $\mathcal{L}\{f(x+a);s\} = e^{as}\mathcal{L}\{f(x);s\}$  along with the transform of a modified Bessel function gives

$$p(x,t) = \delta(x)e^{-k_{1}t} + \frac{k}{v}\exp\left[-\frac{k_{2}x}{v} - k_{1}\left(t - \frac{x}{v}\right)\right]\Theta(t - x/v) \times \left\{I_{0}\left(2\sqrt{\frac{k_{1}k_{2}x}{v}\left(t - \frac{x}{v}\right)}\right)\right\}$$

$$\sqrt{\frac{k_{2}v(t - x/v)}{k_{1}x}}I_{1}\left(2\sqrt{\frac{k_{1}k_{2}x}{v}\left(t - \frac{x}{v}\right)}\right)\right\}$$
 (28)

for the distribution of x at time t. The non-dimensionalization proposed by Lisle is  $\xi = k_+ x/v$  and  $\tau = k_-(t-x/v)$ . In this notation the result appears as

$$p(\xi,\tau) = \frac{k_2}{v}\delta(\xi)e^{-\tau - k_1\xi/k_2} + \frac{k}{v}e^{-\xi - \tau}\Theta(\tau)\Theta(\xi)\left\{I_0\left(2\sqrt{\xi\tau}\right) + \frac{k_2}{k_1}\sqrt{\frac{\tau}{\xi}}I_1\left(2\sqrt{\xi\tau}\right)\right\}$$
(29)

#### 3.1.1 Analytical solution of moments from rest

The first derivative gives

$$\langle \tilde{x} \rangle = v k_1 \frac{1}{s^2 (k+s)},\tag{30}$$

while the second gives

$$\langle \tilde{x}^2 \rangle = 2v^2 k_1 \frac{k_1 + s}{s^3 (k+s)^2}.$$
 (31)

Therefore the mean follows from 82:

$$\frac{k^2}{vk_1}\langle x\rangle = e^{-kt} + kt - 1,\tag{32}$$

and the second moment follows from 83 and 84:

$$\frac{k^4}{2v^2k_1}\langle x^2\rangle = k_1 \left[ \frac{(kt)^2}{2} - kt + 1 - e^{-kt} \right] + k_2 \left[ kt - 2 + (kt+2)e^{-kt} \right]$$
(33)

Manipulating the mean provides

$$\frac{k^4}{2v^2k_1}\langle x\rangle^2 = k_1\left(\frac{1}{2}e^{-2kt} + \frac{(kt)^2}{2} + \frac{1}{2} + (kt-1)e^{-kt} - kt\right)$$
(34)

so the variance is

$$\frac{k^4}{2v^2k_1}\sigma_x^2 = k_1 \left[ \frac{1}{2} - \frac{1}{2}e^{-2kt} - kte^{-kt} \right] + k_2 \left[ kt - 2 + (kt+2)e^{-kt} \right]. \tag{35}$$

Taylor expanding shows asymptotic behavior

$$\sigma_x^2 \sim \begin{cases} \frac{1}{3} k_1 v^2 t^3, & t \to 0\\ 2 \frac{k_1 k_2 v^2}{k^3} t, & t \to \infty. \end{cases}$$
 (36)

There is a cross-over from ballistic to normal diffusion.

#### 3.1.2 Asymptotic solution of moments from rest

The  $t \to \infty$  behavior comes from expanding 30 and 31 at  $s \to 0$  and transforming. The expansions are

$$\langle \tilde{x} \rangle \sim v k_1 \left( \frac{1}{ks^2} - \frac{1}{k^2 s} \right)$$
 (37)

$$\langle \tilde{x}^2 \rangle \sim 2v^2 k_1 \left( \frac{k_1}{k^2 s^3} + \frac{k_2 - k_1}{k^3 s^2} \right)$$
 (38)

giving

$$\sigma_x^2 \sim \frac{2v^2k_1^2t^2}{2k^2} + \frac{2v^2k_1(k_2 - k_1)t}{k^3} - \frac{v^2k_1^2t^2}{k^2} + \frac{2v^2k_1^2t}{k^3}$$
(39)

$$\sigma_x^2 \sim 2 \frac{k_1 k_2 v^2}{k^3} t,$$
 (40)

in agreement with 61. This was tricky to figure out. You have to keep the constant term in the mean. Similarly the  $t \to 0$  behavior comes from expanding 30 and 31 at  $1/s \to 0$  and transforming. The expansions are

$$\langle \tilde{x} \rangle \sim \frac{vk_1}{s^3} \left( 1 - \frac{k}{s} \right)$$
 (41)

$$\langle \tilde{x}^2 \rangle \sim \frac{2v^2 k_1}{s^4} \left( 1 + [k_1 - 2k] \frac{1}{s} \right)$$
 (42)

giving asymptotic variance (at  $t \to 0$ )

$$\sigma_x^2 \sim \frac{1}{3}k_1 v^2 t^3 \tag{43}$$

in agreement with 61.

#### 3.2 Motion initial state

Now I'll try using the opposite initial condition (the one chosen by Lisle et al. [1998]):  $\theta_2 = 1$  and  $\theta_1 = 0$ . In this case

$$\tilde{p}(\eta, s) = \frac{k+s}{v(k_1+s)\eta + (k+s)s},\tag{44}$$

and a first inverse transform gives

$$p(x,s) = \frac{k+s}{v(k_1+s)} \exp\left[-\frac{(k+s)s}{v(k_1+s)}x\right],\tag{45}$$

which is the same as the other case but without a delta function term at x = 0. This can be manipulated to

$$p(x,t) = \mathcal{L}^{-1}\left\{\frac{k+s}{v(k_1+s)}\exp\left[-\frac{(k+s)s}{v(k_1+s)}x\right];t\right\}$$
(46)

$$= e^{-k_1 t} \mathcal{L}^{-1} \left\{ \frac{k_2 + s}{vs} \exp \left[ -\frac{(k_2 + s)(s - k_1)}{vs} x \right]; t \right\}$$
 (47)

$$= e^{-k_1 t - (k_2 - k_1)x/v} \mathcal{L}^{-1} \left\{ \frac{k_2 + s}{vs} \exp\left[\frac{k_1 k_2}{vs} x - \frac{xs}{v}\right]; t \right\}$$
(48)

$$= e^{-k_1 t - (k_2 - k_1)x/v} \mathcal{L}^{-1} \left\{ \frac{k_2 + s}{vs} \exp\left[\frac{k_1 k_2}{vs} x\right]; t - x/v \right\}$$
(49)

$$= e^{-k_1 t - (k_2 - k_1)x/v} \left[ \mathcal{L}^{-1} \left\{ \frac{k_2}{vs} \exp\left[ \frac{k_1 k_2}{vs} x \right]; t - x/v \right\} + \mathcal{L}^{-1} \left\{ \frac{1}{v} \exp\left[ \frac{k_1 k_2}{vs} x \right]; t - x/v \right\} \right]$$
 (50)

$$= e^{-k_1 t - (k_2 - k_1)x/v} \left[ \frac{k_2}{v} I_0 \left( 2\sqrt{\frac{k_1 k_2 x}{v} \left( t - \frac{x}{v} \right)} \right) + \frac{1}{v} \mathcal{L}^{-1} \left\{ \exp \left[ \frac{k_1 k_2}{v s} x \right] - 1; t - x/v \right\} + \frac{1}{v} \delta(t - x/v) \right]$$
(51)

$$=e^{-k_1t-(k_2-k_1)x/v}\left[\frac{k_2}{v}I_0\left(2\sqrt{\frac{k_1k_2x}{v}(t-\frac{x}{v})}\right)+\frac{1}{v}\sqrt{\frac{k_1k_2x}{v(t-x/v)}}I_1\left(2\sqrt{\frac{k_1k_2x}{v}(t-\frac{x}{v})}\right)+\frac{1}{v}\delta(t-x/v)\right].$$
(52)

A key property here was  $e^{-ax}f(x) = \mathcal{L}^{-1}\{f(s+a);x\}$ , and those in the appendix. This type of math is not easy for me. In the non-dimensional variables this becomes

$$p(\xi,\tau) = e^{-\tau - \xi} \left[ \frac{k_2}{v} I_0(2\sqrt{\xi\tau}) + \frac{k_1}{v} \sqrt{\frac{\xi}{\tau}} I_1(2\sqrt{\xi\tau}) + \frac{k_1}{v} \delta(\tau) \right].$$
 (53)

This appears totally aligned with *Lisle et al.* [1998].

#### 3.2.1 moments having started in motion

Taking derivatives (it's easier this time) gives

$$\langle \tilde{x} \rangle = v \frac{k_1 + s}{s^2(s+k)} \tag{54}$$

$$\langle \tilde{x}^2 \rangle = 2v^2 \frac{(k_1 + s)^2}{s^3 (k + s)^2}.$$
 (55)

Using 82 and 86 gives

$$\frac{k^2}{v}\langle x \rangle = k_1 k t + k_2 (1 - e^{-kt}). \tag{56}$$

Using 83, 84, and 85 gives

$$\frac{k^4}{2v^2}\langle x^2 \rangle = k_1^2 \frac{(kt)^2}{2} + k_1 k_2 \left[ 2kt - 2 + 2e^{-kt} \right] + k_2^2 \left[ 1 - (1+kt)e^{-kt} \right]. \tag{57}$$

Manipulating the mean to

$$\frac{k^4}{2v^2}\langle x\rangle^2 = k_1^2 \frac{(kt)^2}{2} + k_1 k_2 \left[\frac{1}{2} + \frac{1}{2}e^{-2kt} - e^{-kt}\right] + k_2^2 \left[kt - kte^{-kt}\right]$$
(58)

provides a variance

$$\frac{k^4}{2v^2k_2}\sigma_x^2 = k_1\left[kt + (2+kt)e^{-kt} - 2\right] + k_2\left[\frac{1}{2} - \frac{1}{2}e^{-2kt} - kte^{-kt}\right].$$
 (59)

Expanding this reveals asymptotic behavior

$$\sigma_x^2 \sim \begin{cases} \frac{1}{3} k_2 v^2 t^3, & t \to 0\\ 2 \frac{k_1 k_2 v^2}{k^3} t, & t \to \infty. \end{cases}$$
 (60)

The conclusion is initial condition does not affect the asymptotic scaling. This is still super-diffusion  $\sigma_x^2 \propto t^3$  crossing to normal diffusion  $\sigma_x^2 \propto t$ . It only affects the coefficient of this scaling.

#### 3.2.2 asymptotic approach to the moments starting in motion

Same story exactly. Not very interesting

$$\sigma_x^2 \sim \begin{cases} \frac{1}{3} k_2 v^2 t^3, & t \to 0\\ 2\frac{k_1 k_2 v^2}{k^3} t, & t \to \infty. \end{cases}$$
 (61)

with no hiccups.

#### 3.3 Mixed initial state

With an arbitrary mixed state the double transformed density is

$$\tilde{p}(\eta, s) = \theta_1 \frac{k + s + \eta v}{(k_1 + s)\eta v + (k + s)s} + \theta_2 \frac{k + s}{v(k_1 + s)\eta + (k + s)s}$$
(62)

$$=\frac{k+s+\theta_1\eta v}{(k_1+s)\eta v+(k+s)s}\tag{63}$$

with the identity  $\theta_1 + \theta_2 = 1$ .

#### 3.3.1 analytical approach to the moments in an arbitrary mixed state

Taking one derivative gives

$$\partial_{\eta}\tilde{p}(\eta,s) = -v \frac{(\theta_2 s + k_1)(k+s)}{[(k_1 + s)\eta v + (k+s)s]^2},\tag{64}$$

and a second gives

$$\partial_{\eta}^{2} \tilde{p}(\eta, s) = 2v^{2} \frac{(\theta_{2}s + k_{1})(k+s)(k_{1}+s)}{[(k_{1}+s)\eta v + (k+s)s]^{3}},\tag{65}$$

so the transformed first and second moments are

$$\langle \tilde{x} \rangle = v \frac{k_1 + \theta_2 s}{s^2 (k+s)} \tag{66}$$

$$\langle \tilde{x}^2 \rangle = 2v^2 \frac{(k_1 + \theta_2 s)(k_1 + s)}{s^3 (k_1 + s)^2}.$$
 (67)

These expressions seem correct since  $\theta_2 = 0$  and  $\theta_2 = 1$  cases provide the earlier expressions. Inverting the mean obtains

$$\frac{k^2}{v}\langle x\rangle = k_1 \left[\theta_1 e^{-kt} + kt - \theta_1\right] + k_2 \theta_2 \left[1 - e^{-kt}\right]$$

$$\tag{68}$$

$$= k_1 \left[ (1 - \theta_2)e^{-kt} + kt + (\theta_2 - 1) \right] + k_2 \theta_2 \left[ 1 - e^{-kt} \right], \tag{69}$$

which still reduces to both earlier results. After more work the second moment becomes

$$\frac{k^4}{2v^2} \langle x^2 \rangle = k_1^2 \left[ \frac{(kt)^2}{2} - kt + 1 - e^{-kt} + \theta_2 \left\{ kt - 1 + e^{-kt} \right\} \right] 
+ k_1 k_2 \left[ kt - 2 + (kt + 2)e^{-kt} + \theta_2 \left\{ kt - kte^{-kt} \right\} \right] 
+ k_2^2 \theta_2 \left[ 1 - (1 + kt)e^{-kt} \right]$$
(70)

which still reduces to earlier results. Proceeding from here gets very difficult and messy. It's possible.

#### 3.3.2 asymptotic approach to the moments in an arbitrary mixed state

Expanding the earlier expressions for  $s \to \infty$  gives

$$\langle \tilde{x} \rangle = v \left[ \frac{\theta_2}{s^2} + \frac{k_1 - \theta_2 k}{s^3} - \dots \right]$$
 (71)

$$\langle \tilde{x}^2 \rangle = 2v^2 \left[ \frac{\theta_2}{s^3} + \frac{k_1 \theta_2 + k_1 - 2k \theta_2}{s^4} + \dots \right]$$
 (72)

transforming to

$$\langle x \rangle \sim v \left[ \theta_2 t + (k_1 - \theta_2 k) \frac{t^2}{2} \right]$$
 (73)

$$\langle x^2 \rangle \sim 2v^2 \left[ \frac{\theta_2 t^2}{2} + \frac{t^3}{3!} (k_1 \theta_2 + k_1 - 2k \theta_2) \right].$$
 (74)

The square of the mean is  $\langle x \rangle^2 \sim v^2 \theta_2^2 t^2 + v^2 \theta_2 (k_1 - \theta_2 k) t^3$ , so the variance is (as  $t \to 0$ )

$$\sigma_x^2 \sim v^2 \theta_1 \theta_2 t^2 + \frac{1}{3} (\theta_1 k_1 + \theta_2 k_2) v^2 t^3.$$
 (75)

This reproduces both earlier results and explains the link between *Lisle et al.* [1998] and my other investigations with the dichotomous Markov noise [e.g. *Horsthemke and Lefever*, 1984; *Bena*, 2006].

### 3.4 Summary of Lisle process

This process supports two stages of diffusion: short-time super-diffusion and long time normal-diffusion. The results are all correct and verified by simulations in figure 3. The two stages of diffusion and their

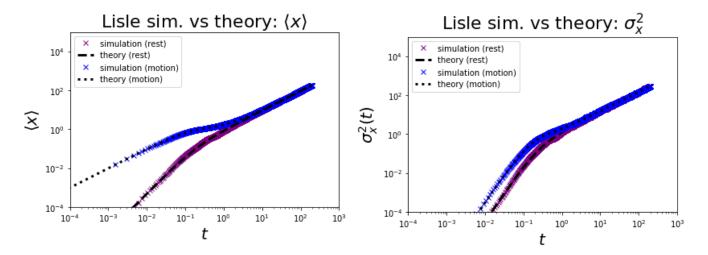


Figure 3: Variance and mean scaling compared to numerical simulations

asymptotic behavior are indicated in figure 4. The initial ballistic diffusion is at least  $\sigma_x^2 \propto t^2$  and at most  $\sigma_x^2 \propto t^3$ , and the crossover to normal diffusion occurs around max $\{1/k_1, 1/k_2\}$ .

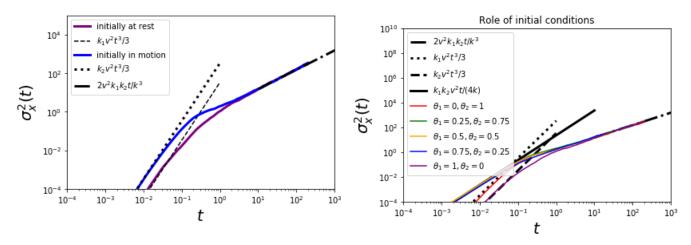


Figure 4: Scaling behavior of variance. Left panel shows pure initial conditions while the action-packed right panel shows mixed initial conditions. Mixed initial express a  $\sigma_x^2 \propto t^2$  term.

## 4 Heavy-tailed resting time in a two-state walk

Set  $\psi_1(t) = A_{\alpha}t^{-\alpha}$  for  $t \ge t_{\alpha}$ .  $A_{\alpha} = (\alpha - 1)t_{\alpha}^{\alpha - 1}$  [e.g. Weeks et al., 1996, eq. 12]. Consider everything else the same as the Lisle et al. [1998] process. The laplace transform of  $\psi_1(t)$  is

$$\tilde{\psi}(s) = A_{\alpha} s^{\alpha - 1} \Gamma(1 - \alpha, st_{\alpha}). \tag{76}$$

The incomplete gamma function is defined by

$$\Gamma(q,x) = \int_{x}^{\infty} t^{q-1}e^{-t}dt. \tag{77}$$

There's a need to expand this at small arguments. Crucially, Weeks et al. [1996] did not use the standard Pareto distribution notation.

### 4.1 $t \to \infty$ expected behavior

 $\alpha < 1/2$  implies subdiffusion.  $1/2 < \alpha < 2$  implies superdiffusion,  $\alpha > 2$  implies normal diffusion.

## 5 A new generalization: Randomly stopped Lisle process

Consider the choice  $\psi_1(t) = k_1 e^{-k_1 t}$ ,  $f_1(x,t) = \delta(x)$ ,  $\psi_2(t) = k_2 e^{-k_2 t}$ , and

$$f_2(x,t|T) = \delta(x - vt)\Theta(T - t) + \delta(x)\Theta(t - T). \tag{78}$$

This describes a randomly stopped variant of the Lisle process. When t > T, the motion state becomes a second rest state: the motion is turned off. In this case,  $g_1(\eta, s) = \frac{k_1}{k_1 + s}$  and  $G_1(\eta, s) = g_1(\eta, s)/k_1$  as before, while

$$g_2(\eta, s|T) = \frac{k_2}{k_2 + s} e^{-(k_2 + s)T} + \frac{k_2}{k_2 + \eta v + s} \left(1 - e^{-(k_2 + \eta v + s)T}\right)$$
(79)

and  $G_2(\eta, s|T) = g_2(\eta, s|T)/k_2$ . You can maybe solve this conditional to T to obtain p(x, t|T). Then inputting a distribution for the trapping time T, the over-all distribution will be  $p(x, t) = \int dT p(x, t|T) f(T)$ .

Assuming the walker starts at rest, the double transformed probability is

$$\tilde{p}(\eta, s) = \tag{80}$$

# Appendix: Laplace transforms

This is just a reference of useful Laplace transforms for these types of studies.

$$\mathcal{L}\left\{\frac{1}{s^{k+1}};s\right\} = \frac{t^k}{k!};\tag{81}$$

$$\mathcal{L}^{-1}\left\{p^{-2}(p+a)^{-1};x\right\} = \frac{1}{a^2}(e^{-ax} + ax - 1),\tag{82}$$

[Prudnikov et al., 1986, 2.1.2.33];

$$\mathcal{L}^{-1}\left\{p^{-2}(p+a)^{-2};x\right\} = \frac{1}{a^3}[ax - 2 + (ax+2)e^{-ax}],\tag{83}$$

[Prudnikov et al., 1986, 2.1.2.49];

$$\mathcal{L}^{-1}\left\{p^{-3}(p+a)^{-2};x\right\} = \frac{1}{a^4} \left[\frac{(ax)^2}{2} - 2ax - (ax+3)e^{-ax} + 3\right],\tag{84}$$

derived from the previous result using the differentiation property;

$$\mathcal{L}^{-1}\left\{p^{-1}(p+a)^{-2};x\right\} = \frac{1}{a^2}\left(1 - (1+ax)e^{-ax}\right),\tag{85}$$

[Prudnikov et al., 1986, 2.1.2.47];

$$\mathcal{L}^{-1}\left\{p^{-1}(p+a)^{-1};x\right\} = \frac{1}{a}(1-e^{-ax}),\tag{86}$$

[Prudnikov et al., 1986, 2.1.2.31];

$$\mathcal{L}^{-1}\{e^{as}\tilde{f}(s);x\} = \mathcal{L}^{-1}\{\tilde{f}(s);x+a\},\tag{87}$$

the shifting property;

$$\mathcal{L}^{-1}\left\{\frac{1}{s^{\nu}}\exp(a/s);t\right\} = \left(\frac{t}{a}\right)^{(\nu-1)/2} I_{\nu-1}\left(2\sqrt{at}\right),\tag{88}$$

valid for  $\nu > 1$  from [Prudnikov et al., 1986, 2.2.2.1];

$$\mathcal{L}^{-1}\{e^{a/p} - 1; x\} = \sqrt{\frac{a}{x}} I_1(2\sqrt{ax}), \tag{89}$$

and [Prudnikov et al., 1986, 2.2.2.8]. All of these Laplace transforms are verified from Wolfram Alpha so I expect no typos. Finally,

$$\mathcal{L}\left\{\frac{a}{bs+c};s\right\} = \frac{a}{b}e^{-cx/b} \tag{90}$$

and

$$\mathcal{L}\left\{\frac{as}{bs+c};s\right\} = \frac{a}{b}\left[\delta(x) - \frac{c}{b}e^{-cx/b}\right]. \tag{91}$$

The Laplace transform of a first derivative is

$$\mathcal{L}\lbrace f'(t)\rbrace = s\tilde{f}(s) - f(0), \tag{92}$$

while a second derivative is

$$\mathcal{L}\{f''(t)\} = s^2 \tilde{f}(s) - sf(0) - f'(0). \tag{93}$$

## Appendix: Asymptotics of stable laws

When a resting time distribution  $\psi(t)$  has divergent moments, the behavior of its Laplace transform  $\tilde{\psi}(s)$  for  $s \to 0$  cannot be analyzed by simply expanding the exponential in its definition for small s as in

$$\tilde{\psi}(s) \sim \int_0^\infty \{1 - st + (st)^2 / 2 + \dots\} \psi(t),$$
(94)

since the moments of t involved in this expression are divergent. Instead, there is a different way (outlined around Weiss [1994] eq 2.95) leveraging the asymptotic form of  $\psi(t)$  for  $t \to \infty$ :

$$\psi(t) \sim At^{-\alpha - 1}. (95)$$

Writing (using the normalization property of  $\psi(t)$ )

$$\tilde{\psi}(s) = 1 - (1 - \tilde{\psi}(s)) = 1 - \int_0^\infty dt (1 - e^{-st}) \psi(t)$$
 (96)

and setting  $\tau = st$  gives

$$\tilde{\psi}(s) = 1 - \frac{1}{s} \int_0^\infty d\tau (1 - e^{-\tau}) \psi\left(\frac{\tau}{s}\right). \tag{97}$$

Clearly this integral is dominated by the asymptotic behavior of  $\psi(t)$ :

$$\tilde{\psi}(s) \sim 1 - Bs^{\alpha} \tag{98}$$

This is an important result for analyzing asymptotics of random walks involving power-law pausing time densities. The coefficient B is obtained (if necessary) by integrating the asymptotic power law using its coefficients and small-time cutoff. Such an argument is essential to the paper of Weeks and Swinney [1998]. They give B in terms of an incomplete Gamma function, which I could do if I wanted.

# Appendix: Modified Bessel functions of the first kind

A cool property is

$$I_n(x) = T_n \left(\frac{d}{dx}\right) I_0(x) \tag{99}$$

where  $T_n$  is a Chebyshev polynomial of the first kind:

$$T_0(x) = 1 \tag{100}$$

$$T_1(x) = x \tag{101}$$

$$T_2(x) = 2x^2 - 1 (102)$$

$$\vdots \tag{103}$$

This is available at http://mathworld.wolfram.com/ModifiedBesselFunctionoftheFirstKind.html. In particular,

$$I_1(x) = I_0'(x). (104)$$

The large x expansion  $(x \gg |\nu^2 - 1/4|)$  of the modified Bessel function of the first kind is

$$I_{\nu}(x) \approx \frac{e^x}{\sqrt{2\pi x}} \left( 1 + \frac{(1-2\nu)(1+2\nu)}{8x} + \dots \right)$$
 (105)

while the small x expansion  $(0 \le x \le \sqrt{\nu+1})$  is

$$I_{\nu}(x) = \frac{1}{\Gamma(\nu+1)} \left(\frac{x}{2}\right)^{\nu}.$$
 (106)

These are from https://pdfs.semanticscholar.org/48dc/ca3cffb78de80ab37b84a992379c2f30bdda.pdf. The recurrence relations are

$$I_{\nu-1}(x) - I_{\nu+1}(x) = \frac{2\nu}{x} I_{\nu}(x)$$
(107)

$$I_{\nu-1}(x) + I_{\nu+1}(x) = 2\frac{d}{dx}I_{\nu}(x). \tag{108}$$

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