James K. Pringle 550.621 Probability Dr. Jim Fill Assignment 4 April 2, 2014

# **Assignment 4**

Chung Exercise 7.2.4

## Chung Exercise 7.2.4

Prove the sufficiency part of Theorem 7.2.1 without using Theorem 7.1.2, but by elaborating the proof of the latter.

### Chung Theorem 7.2.1

Assume  $\sigma_{nj}^2 < \infty$  for each n and j and the reduction hypotheses

$$\sum_{j=1}^{k_n} \sigma^2(X_{nj}) = 1 \tag{1}$$

$$E(X_{nj}) = 0 (2)$$

of Sec 7.1. In order that as  $n \to \infty$  the two conclusions below both hold

- (I)  $S_n$  converges in dist. to  $\Phi$  (standard normal)
- (II) the double array (2) of Sec. 7.1 is holospoudic,

it is necessary and sufficient that for each  $\eta > 0$ , we have

$$\sum_{j=1}^{k_n} \int_{|x| > \eta} x^2 dF_{nj}(x) \to 0 \tag{3}$$

### Assumptions

By way of notation, let  $F_{nj}(x)$  be the d.f. of  $X_{nj}$ . Assume the hypotheses of **Chung Theorem 7.2.1** and assume (3). Note by (1)

$$1 = \sum_{j=1}^{k_n} \sigma^2(X_{nj}) = \sum_{j=1}^{k_n} \int x^2 dF_{nj}(x) = \sum_{j=1}^{k_n} \left( \int_{|x| \le \eta} + \int_{|x| > \eta} \right) x^2 dF_{nj}(x)$$

so that

$$0 = \lim_{n \to \infty} \sum_{j=1}^{k_n} \int_{|x| > \eta} x^2 dF_{nj}(x) = 1 - \lim_{n \to \infty} \sum_{j=1}^{k_n} \int_{|x| \le \eta} x^2 dF_{nj}(x)$$

Thus

$$\lim_{n \to \infty} \sum_{j=1}^{k_n} \int_{|x| \le \eta} x^2 dF_{nj}(x) = 1 \tag{4}$$

### Discussion

The proof examines the convergence of ch.f.'s of the rows of the double array. By hypothesis,  $S_n$  is the sum of the  $k_n$  independent random variables in the n-th row. By independence, the ch.f. of  $S_n$  is the product of the ch.f.'s of the random variables of that row. Define  $f_n$  to be the ch.f. of  $S_n$ . Define  $f_{nj}$  to be the ch.f. of  $X_{nj}$ 

$$f_n(t) = f_{S_n}(t) = \prod_{j=1}^{k_n} f_{X_{nj}}(t) = \prod_{j=1}^{k_n} f_{nj}(t)$$

By Chung Theorem 6.3.2 it is sufficient to show for all  $t \in \mathbb{R}$  that  $f_n(t)$  converges to the ch.f. of standard normal,  $e^{-\frac{t^2}{2}}$  in order that  $S_n$  converges in distribution to  $\Phi$ .

Next follow some lemmas that are required for the proof.

#### Lemma 1

Let  $\{\theta_{nj}, 1 \leq j \leq k_n, 1 \leq n\}$  be a double array of complex numbers satisfying the following conditions as  $n \to \infty$ 

- (i)  $\max_{1 \le j \le k_n} |\theta_{nj}| \to 0$
- (ii)  $\sum_{j=1}^{k_n} |\theta_{nj}| \leq M < \infty$ , where M does not depend on n
- (iii)  $\sum_{j=1}^{k_n} \theta_{nj} \to \theta$ , where  $\theta$  is a (finite) complex number.

Then we have

$$\prod_{i=1}^{k_n} (1 + \theta_{nj}) \to e^{\theta}$$

#### Lemma 2

According to Billingsley 26.4<sub>1</sub>

$$|e^{ix} - (1+ix)| \le \min\{\frac{1}{2}x^2, 2|x|\}$$

### Lemma 3

According to Billingsley 26.4<sub>2</sub>

$$|e^{ix} - (1 + ix - \frac{1}{2}x^2)| \le \min\{\frac{1}{6}|x|^3, x^2\}$$

# Proof of the Main Result

Fix  $t \in \mathbb{R}$ . Notice

$$f_n(t) = \prod_{j=1}^{k_n} f_{nj}(t) = \prod_{j=1}^{k_n} 1 + (f_{nj}(t) - 1)$$

To apply **Lemma 1**, let  $\theta_{nj} = f_{nj}(t) - 1$ .

#### Condition (i)

Let  $\eta > 0$ . Calculating,

$$\max_{1 \le j \le k_n} |\theta_{nj}| = \max_{1 \le j \le k_n} |f_{nj}(t) - 1| 
= \max_{1 \le j \le k_n} |f_{nj}(t) - 1 - itE(X_{nj})| by (2) 
= \max_{1 \le j \le k_n} |\int (e^{itx} - 1 - itx)dF_{nj}(x)| 
\le \max_{1 \le j \le k_n} \int |e^{itx} - 1 - itx|dF_{nj}(x) by modulus inequality 
\le \max_{1 \le j \le k_n} \int \frac{1}{2} (tx)^2 dF_{nj}(x) by Lemma 2 
= \max_{1 \le j \le k_n} \frac{1}{2} t^2 \left( \int_{|x| \le \eta} x^2 dF_{nj}(x) + \int_{|x| > \eta} x^2 dF_{nj}(x) \right) 
\le \frac{1}{2} t^2 \eta^2 + \frac{1}{2} t^2 \max_{1 \le j \le k_n} \int_{|x| > \eta} x^2 dF_{nj}(x) 
\le \frac{1}{2} t^2 \eta^2 + \frac{1}{2} t^2 \sum_{j=1}^{k_n} \int_{|x| > \eta} x^2 dF_{nj}(x)$$

since the max of a set of positive numbers is less than or equal to the sum of all of them. Thus, by (3),

$$0 \le \lim_{n \to \infty} \left( \max_{1 \le j \le k_n} |\theta_{nj}| \right) \le \frac{1}{2} t^2 \eta^2 + \frac{1}{2} t^2 \lim_{n \to \infty} \left( \sum_{j=1}^{k_n} \int_{|x| > \eta} x^2 dF_{nj}(x) \right) = \frac{1}{2} t^2 \eta^2$$

Since there is no dependence on  $\eta$  in the second term above, and since  $\eta$  is arbitrary and positive,

$$0 \le \lim_{n \to \infty} \left( \max_{1 \le j \le k_n} |\theta_{nj}| \right) \le \lim_{\eta \to 0} \left( \frac{1}{2} t^2 \eta^2 \right) = 0$$

Hence Condition (i) holds.

#### Condition (ii)

Calculating,

$$\sum_{j=1}^{k_n} |\theta_{nj}| = \sum_{j=1}^{k_n} |f_{nj}(t) - 1|$$

$$\leq \sum_{j=1}^{k_n} \int \frac{1}{2} (tx)^2 dF_{nj}(x) \quad \text{by same calculations as for Condition (i)}$$

$$= \frac{1}{2} t^2 \sum_{j=1}^{k_n} \int x^2 dF_{nj}(x)$$

$$= \frac{1}{2} t^2 \sum_{j=1}^{k_n} \sigma^2(X_{nj})$$

$$= \frac{1}{2} t^2 \quad \text{by (1)}.$$

Since this bound does not depend on n and it is finite, Condition (ii) holds.

#### Condition (iii)

Let  $\eta > 0$ . It is claimed that  $\theta = -\frac{t^2}{2}$ , so that

$$\sum_{j=1}^{k_n} \theta_{nj} \to \theta.$$

Calculating,

$$\left| \left( \sum_{j=1}^{k_n} \theta_{nj} \right) - \theta \right| = \left| \sum_{j=1}^{k_n} \left( f_{nj}(t) - 1 \right) + \frac{t^2}{2} \right|$$

$$= \left| \sum_{j=1}^{k_n} \left( f_{nj}(t) - 1 \right) - \frac{(it)^2}{2} \left( \sum_{j=1}^{k_n} \sigma^2(X_{nj}) \right) \right| \quad \text{by (1)}$$

$$= \left| \sum_{j=1}^{k_n} \left( f_{nj}(t) - 1 - itE(X_{nj}) \right) - \frac{(it)^2}{2} \left( \sum_{j=1}^{k_n} E(X_{nj}^2) \right) \right| \quad \text{by (2)}$$

$$= \left| \sum_{j=1}^{k_n} \left( f_{nj}(t) - 1 - itE(X_{nj}) - \frac{(it)^2}{2} E(X_{nj}^2) \right) \right|$$

$$= \left| \sum_{j=1}^{k_n} \left( \int e^{itx} - 1 - itx + \frac{(tx)^2}{2} dF_{nj}(x) \right) \right|$$

$$\leq \sum_{j=1}^{k_n} \int \left| e^{itx} - 1 - itx + \frac{(tx)^2}{2} dF_{nj}(x) \right| \text{ by triangle and modulus inequalities}$$

$$= \sum_{j=1}^{k_n} \left( \int_{|x| \le \eta} \left| e^{itx} - 1 - itx + \frac{(tx)^2}{2} dF_{nj}(x) + \int_{|x| > \eta} \left| e^{itx} - 1 - itx + \frac{(tx)^2}{2} dF_{nj}(x) \right| \right)$$

$$\leq \sum_{j=1}^{k_n} \left( \int_{|x| \le \eta} \frac{|x|^3}{6} dF_{nj}(x) + \int_{|x| > \eta} x^2 dF_{nj}(x) \right) \quad \text{by Lemma 3}$$

$$\leq \frac{\eta}{6} \sum_{i=1}^{k_n} \int_{|x| < \eta} x^2 dF_{nj}(x) + \sum_{i=1}^{k_n} \int_{|x| > \eta} x^2 dF_{nj}(x)$$

Taking the limit of this nonegative quantity as n tends to infinity,

$$0 \leq \lim_{n \to \infty} \left| \left( \sum_{j=1}^{k_n} \theta_{nj} \right) - \theta \right|$$

$$\leq \lim_{n \to \infty} \left( \frac{\eta}{6} \sum_{j=1}^{k_n} \int_{|x| \leq \eta} x^2 dF_{nj}(x) + \sum_{j=1}^{k_n} \int_{|x| > \eta} x^2 dF_{nj}(x) \right)$$

$$= \frac{\eta}{6} \lim_{n \to \infty} \left( \sum_{j=1}^{k_n} \int_{|x| \leq \eta} x^2 dF_{nj}(x) \right) + \lim_{n \to \infty} \left( \sum_{j=1}^{k_n} \int_{|x| > \eta} x^2 dF_{nj}(x) \right)$$

$$= \frac{\eta}{6} \quad \text{by (4) and (3)}$$

Since  $\eta$  is arbitrary and positive, by squeezing,

$$0 \le \lim_{n \to \infty} \left| \left( \sum_{j=1}^{k_n} \theta_{nj} \right) - \theta \right| \le \lim_{\eta \to 0} \frac{\eta}{6} = 0$$

and it follows that

$$\lim_{n \to \infty} \left| \left( \sum_{j=1}^{k_n} \theta_{nj} \right) - \theta \right| = 0$$

Hence,  $\sum_{j=1}^{k_n} \theta_{nj}$  converges to  $\theta = -\frac{t^2}{2}$  as n tends to infinity. Condition (iii) holds. Therefore by Lemma 1,

$$\lim_{n \to \infty} f_{S_n}(t) = \lim_{n \to \infty} f_n(t) = \prod_{j=1}^{k_n} 1 + (f_{nj}(t) - 1) = \prod_{j=1}^{k_n} 1 + \theta_{nj} = e^{\theta} = e^{-\frac{t^2}{2}}$$

Note  $t \in \mathbb{R}$  is arbitrary. Thus for all  $t \in \mathbb{R}$  it is the case that  $f_{S_n}(t)$  converges to the ch.f. of standard normal,  $e^{-\frac{t^2}{2}}$ , and by **Chung Theorem 6.3.2**  $S_n$  converges in distribution to  $\Phi$ . Thus (I).

Condition (II), holospoudicity, follows from **Chung Theorem 7.1.1** since **Condition (i)** holds. Alternatively, the proof of holospoudicity copies over from Chung since it does not rely on **Chung Theorem 7.1.2**. Therefore, (3) is sufficient for (I) and (II).

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