Applied Logistic Regression – Assignment 2

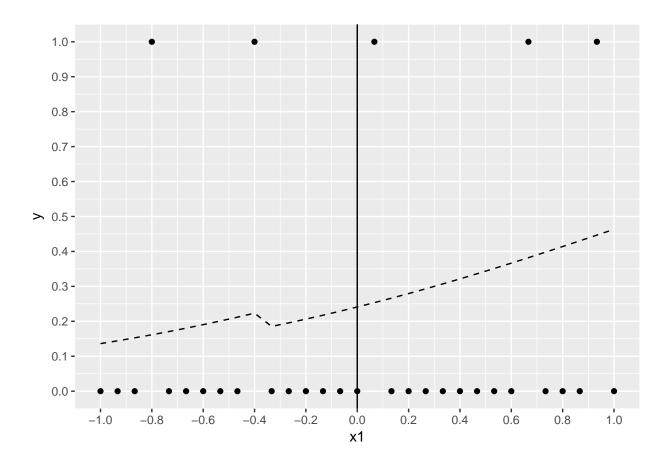
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Assignment 2

1a: Data Generation

1b: Plotting the True Probabilities

```
library(ggplot2)
the.data<-data.frame(x0,x1,x2,lp,p,y)
ggplot(data=the.data, aes(x=x1, y=y)) +
    scale_y_continuous(limits=c(0,1), breaks=seq(from=0,to=1,by=0.1)) +
    scale_x_continuous(breaks=seq(from=-1,to=1,by=0.2)) +
    geom_point(aes(y=y,x=x1)) +
    geom_line(aes(y=p,x=x1),linetype="dashed") +
    geom_vline(xintercept = 0)</pre>
```



2: Newton's Method

```
#I think the idea is to start with the grand mean (aka intercept) of the true probability:
p1<-sum(y)/n

#Since this is logistic regression, we use the log-odds (aka logit) of the probability

#rather than the probability itself. The other predictor coefficients are

#apparently given an initial value of 0:
B0<-c(log(p1/(1-p1)), 0, 0)

names(B0) <- c("x0","x1","x2")
```

So here is our starting point:

```
## x0 x1 x2
## -1.648659 0.000000 0.000000
```

Fisher Scoring Iteration 1:

```
lp0<-X%*%B0
p0<-exp(lp0)/(1+exp(lp0))
d1<-t(X)%*%(y-p0)</pre>
```

```
w<-as.vector(p0*(1-p0))
d21<--t(X)%*%diag(w)%*%X
B1<-B0-solve(d21,d1)
t(B1)

## x0 x1 x2
## [1,] -0.1788684 1.690909 -2.16969</pre>
```

The values changed, so we'll do an Iteration 2:

```
B0<-B1
lp0<-X%*%B0
p0<-exp(lp0)/(1+exp(lp0))
d1<-t(X)%*%(y-p0)
w<-as.vector(p0*(1-p0))
d21<--t(X)%*%diag(w)%*%X
B1<-B0-solve(d21,d1)
t(B1)

## x0 x1 x2
## [1,] -0.06288461 1.947025 -2.575052
```

The values changed again, so we'll do an Iteration 3:

```
B0<-B1
lp0<-X%*%B0
p0<-exp(lp0)/(1+exp(lp0))
d1<-t(X)%*%(y-p0)
w<-as.vector(p0*(1-p0))
d21<--t(X)%*%diag(w)%*%X
B1<-B0-solve(d21,d1)
t(B1)

## x0 x1 x2
## [1,] -0.03635197 1.990107 -2.638811
```

The values changed again, so we'll do an Iteration 4:

[1,] -0.03595209 1.990749 -2.639743

```
B0<-B1
lp0<-X%*%B0
p0<-exp(lp0)/(1+exp(lp0))
d1<-t(X)%*%(y-p0)
w<-as.vector(p0*(1-p0))
d21<--t(X)%*%diag(w)%*%X
B1<-B0-solve(d21,d1)
t(B1)

## x0 x1 x2
```

The values changed very little now. But I guess we'll do an Iteration 5:

```
B0<-B1

lp0<-X%*%B0

p0<-exp(lp0)/(1+exp(lp0))

d1<-t(X)%*%(y-p0)

w<-as.vector(p0*(1-p0))

d21<--t(X)%*%diag(w)%*%X

B1<-B0-solve(d21,d1)

t(B1)

## x0 x1 x2

## [1,] -0.03595201 1.990749 -2.639743
```

Only x0 changed. I suppose we'll keep going until nothing changes. Iteration 6:

```
B0<-B1
lp0<-X%*%B0
p0<-exp(lp0)/(1+exp(lp0))
d1<-t(X)%*%(y-p0)
w<-as.vector(p0*(1-p0))
d21<--t(X)%*%diag(w)%*%X
B1<-B0-solve(d21,d1)
t(B1)

## x0 x1 x2
## [1,] -0.03595201 1.990749 -2.639743
```

Nothing changed. This might be what stats people call "convergence". Thus, a total of 5 Fisher Scoring Iterations were needed.

3: R's glm() Function

Now we'll have R's glm() function carry out the same procedure, and hope for the same results:

```
Y<-cbind(y,1-y)
logreg<-glm(Y~x1+x2,family=binomial(link="logit"))
summary(logreg)
```

```
##
## Call:
## glm(formula = Y ~ x1 + x2, family = binomial(link = "logit"))
##
## Deviance Residuals:
##
       Min
                 1Q
                      Median
                                   3Q
                                           Max
## -0.9036 -0.6575 -0.5010 -0.3108
                                        2.2885
##
## Coefficients:
##
               Estimate Std. Error z value Pr(>|z|)
## (Intercept) -0.03595
                           1.35349 -0.027
                                              0.979
## x1
                1.99075
                           1.66216
                                     1.198
                                              0.231
## x2
               -2.63974
                           2.24718 -1.175
                                              0.240
```

```
##
## (Dispersion parameter for binomial family taken to be 1)
##
## Null deviance: 27.392 on 30 degrees of freedom
## Residual deviance: 25.588 on 28 degrees of freedom
## AIC: 31.588
##
## Number of Fisher Scoring iterations: 5
```

Indeed, the coefficients are the same (though the model summary rounds them). The number of Fisher Scoring iterations performed is also identical. This suggests major mistakes have not been made.

4: Variance-Covariance Matrices

```
#Here's the variance-covariance matrix of the manually computed model:
(V<--solve(d21))
##
             x0
                       x1
## x0 1.831951 1.815212 -2.821587
## x1 1.815212 2.762775 -3.321450
## x2 -2.821587 -3.321450 5.049817
#Here's the variance-covariance matrix of the model computed by the glm() function:
vcov(logreg)
               (Intercept)
                                   x1
                  1.831948 1.815206 -2.821579
## (Intercept)
## x1
                  1.815206 2.762766 -3.321437
## x2
                 -2.821579 -3.321437 5.049798
Everything is identical again, notwithstanding rounding.
```

5: Standard Errors of Model Coefficients:

```
#Predictor coefficient standard errors in the manually computed model:
sqrt(diag(V))

## x0 x1 x2
## 1.353496 1.662160 2.247180

#Predictor coefficient standard errors in glm()-computed model:
sqrt(diag(vcov(logreg)))

## (Intercept) x1 x2
## 1.353495 1.662157 2.247176
```

6: x1 as the Only Predictor

```
logreg.x1<-glm(Y~x1, family=binomial(link="logit"))
summary(logreg.x1)
##
## Call:</pre>
```

```
## glm(formula = Y ~ x1, family = binomial(link = "logit"))
##
## Deviance Residuals:
##
      Min
              1Q
                     Median
                                  3Q
                                          Max
## -0.6803 -0.6218 -0.5730 -0.5250
                                       2.0250
##
## Coefficients:
##
              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -1.6606
                           0.4934 -3.365 0.000765 ***
## x1
                0.3149
                           0.8264 0.381 0.703191
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## (Dispersion parameter for binomial family taken to be 1)
##
##
      Null deviance: 27.392 on 30 degrees of freedom
## Residual deviance: 27.245 on 29 degrees of freedom
## AIC: 31.245
## Number of Fisher Scoring iterations: 4
```

7: x2 as the Only Predictor

```
logreg.x2<-glm(Y~x2, family=binomial(link="logit"))</pre>
summary(logreg.x2)
##
## Call:
## glm(formula = Y ~ x2, family = binomial(link = "logit"))
##
## Deviance Residuals:
##
      Min
                 1Q
                     Median
                                   ЗQ
                                           Max
## -0.6681 -0.6117 -0.5553 -0.5553
                                        1.9728
##
## Coefficients:
               Estimate Std. Error z value Pr(>|z|)
## (Intercept) -1.3863
                            0.7906 -1.754
                                             0.0795 .
               -0.4055
                            1.0069 -0.403
                                             0.6872
## x2
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## (Dispersion parameter for binomial family taken to be 1)
##
      Null deviance: 27.392 on 30 degrees of freedom
## Residual deviance: 27.233 on 29 degrees of freedom
## AIC: 31.233
## Number of Fisher Scoring iterations: 4
```

8: Interpreting the Results

So, the full model (with both predictors included) reports a positive effect on the outcome for $\mathbf{x1}$ and a negative one for $\mathbf{x2}$, both of them statistically non-significant. A very similar result is obtained when we only include $\mathbf{x1}$ or $\mathbf{x2}$ in the model – the former has a positive coefficient, the latter a negative one. My interpretation is that increasing magnitude of the continuous property $\mathbf{x1}$ favors the occurrence of the outcome, while membership in group $\mathbf{x2}$ (represented by value 1 of the dichotomous predictor) disfavors it. This is well illustrated by our graph in exercise 1 – the slope is consistently ascending as a function of $\mathbf{x1}$. The single downward blip in the graph (at about $\mathbf{x1} = 0.3$) occurs when Group Identity ($\mathbf{x2}$) changes from 0 to 1.

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