

# Proof of Markov and Chebyshev's inequalities

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In statistical inference, we often wish to compute an upper bound on the probability that a random variable, which may be a parameter of interest, deviates from some positive value  $k$ .

Two inequalities are often used for these questions: Markov's inequality and Chebyshev's inequality. We will derive Markov's inequality and show how we can use it to obtain Chebyshev's inequality.

Markov's inequality states that for some positive value  $k$  when  $\theta$  is non-negative, if the expected value  $E(\theta)$  is defined, then

$$P(\theta > k) \leq E(\theta)/k$$

To show why this is true we may write the  $E(\theta)$  as

$$E(\theta) = \int_0^{\infty} \theta f(\theta) d\theta$$

In the above, since  $\theta$  is non-negative, the lower bound on the integral is zero.

We can rewrite the integral on the LHS to specify some positive value  $k$  we are interested in as

$$E(\theta) = \int_0^k \theta f(\theta) d\theta + \int_k^{\infty} \theta f(\theta) d\theta$$

Since every term in the above is non-negative, the sum of the integrals must be larger than either individual integral, which implies

$$E(\theta) = \int_0^k \theta f(\theta) d\theta + \int_k^{\infty} \theta f(\theta) d\theta \geq \int_k^{\infty} \theta f(\theta) d\theta$$

And since  $k$  is positive and  $\theta$  is non-negative, integrating the product  $\theta f(\theta)$  will yield a value that is at least as large, or larger, than the integral  $\int_k^{\infty} k f(\theta) d\theta$ . Therefore we have

$$E(\theta) \geq \int_k^{\infty} \theta f(\theta) d\theta \geq k \int_k^{\infty} f(\theta) d\theta$$

And since  $\int_k^\infty f(\theta)d\theta$  is just the probability that  $\theta$  is larger than  $k$ , we can write

$$\begin{aligned} E(\theta) &\geq kP(\theta > k) \\ \implies P(\theta > k) &\leq E(\theta)/k \end{aligned}$$

which is the expression of Markov's inequality. We may now use this to derive Chebyshev's inequality, which is

$$P(|\theta - E(\theta)| > k) \leq Var(\theta)/k^2$$

To show this, we may treat  $|\theta - E(\theta)|$  as a non-negative random variable and substitute it into Markov's inequality

$$\implies P(|\theta - E(\theta)| > k) \leq E(|\theta - E(\theta)|)/k$$

We may square both sides on the inequality inside the probability, as well as the term of the RHS  $|\theta - E(\theta)|/k$ , to obtain

$$\implies P([\theta - E(\theta)]^2 > k^2) \leq E[\theta - E(\theta)]^2/k^2$$

And since the definition of variance is

$$E([\theta - E(\theta)]^2) = Var(\theta)$$

We may write

$$\implies P(|\theta - E(\theta)| > k) \leq Var(\theta)/k^2$$

which is the desired result.