Solution 9: Parallel Computing and Simulations

July 23, 2024

Simulations on Cauchy

Suppose $\mathbf{X} = (X_1, \dots, X_n)$ is an i.i.d. sample from the shifted Cauchy distribution with density

$$f(x \mid \theta) = \frac{1}{\pi (1 + (x - \theta)^2)}, \quad x \in \mathbb{R}$$

Our goal is to compare the following 4 estimators of the parameter θ .

• Sample mean

$$\hat{\theta}_n^{(1)} = \bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$$

• Sample median

$$\hat{\theta}_{n}^{(2)} = M_{n} = \frac{1}{2} \left(X_{(k)} + X_{(k+1)} \right)$$

• Modified sample mean

$$\hat{\theta}_n^{(3)} = M_n + \frac{2}{n} \cdot \frac{\partial \ell}{\partial \theta} \bigg|_{\theta = M_n}$$

where ℓ is the log-likelihood function.

• Maximum likelihood estimator (MLE) $\hat{\theta}_n^{(4)}$ defined by

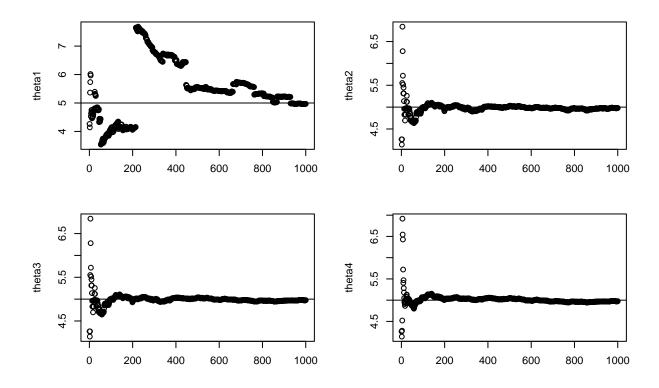
$$\ell\left(\hat{\theta}_{n}^{(4)} \mid X\right) = \max_{\theta \in \mathbb{R}} \ell(\theta \mid \boldsymbol{X})$$

where ℓ is the log-likelihood function.

- 1. Derive the likelihood function and log-likelihood function.
- 2. Simulate data from Cauchy distribution with location 5, and scale 1.
- 3. Choose your number of simulations.
- 4. Verify consistency of the estimators.
- 5. Calculate the mean square error of the estimators.
- 6. Calculate the coverage probability of the estimators. Calculate $\mathbf{P}_{\theta}\left(\left|\hat{\theta}_{n}-\theta\right|\leq\varepsilon\right)$, for $\varepsilon=0.1$, and $\theta=5$.

Solution

```
# Simulate Cauchy.
n <- 1000
set.seed(123456)
cauchy_samples <- reauchy(n, location = 5, scale = 1)</pre>
# Function to calculate the derivative of the log-likelihood.
dloglik <- function(x, theta) {</pre>
  d1 \leftarrow x - theta
  d2 < -1 + (x - theta)^2
  return(2*sum(d1/d2))
# Function to calculate the log-likelihood.
loglik <- function(theta) {</pre>
   x <- cauchy_samples
   n <- length(x)
   11 <- -n*log(pi)
   12 \leftarrow -sum(\log(1 + (x - theta)^2))
   return(11 + 12)
}
nloglik<- function(theta) {</pre>
  return(-loglik(theta))
# Consistency.
theta1 <- vector(length = n)</pre>
theta2 <- vector(length = n)</pre>
theta3 <- vector(length = n)</pre>
theta4 <- vector(length = n)</pre>
full_samples <- cauchy_samples</pre>
for (k in 1:n) {
  cauchy_samples <- NULL</pre>
  cauchy_samples <- full_samples[1:k]</pre>
  #print(cauchy_samples)
  theta1[k] <- mean(cauchy_samples)</pre>
  #print(theta1[i])
  theta2[k] <- median(cauchy_samples)</pre>
  theta3[k] <- theta2[k] + 2/n * dloglik(cauchy_samples, theta2[k])
  theta4[k] \leftarrow optimize(nloglik, interval = c(-10,10))$minimum
}
par(mfrow = c(4, 2), mar = c(1, 4.1, 4.1, 2.1))
plot(theta1)
abline(h = 5)
plot(theta2)
abline(h = 5)
plot(theta3)
abline(h = 5)
plot(theta4)
abline(h = 5)
```



```
nsim <- 1000
theta1 <- vector(length = nsim)
theta2 <- vector(length = nsim)
theta3 <- vector(length = nsim)
theta4 <- vector(length = nsim)</pre>
```

```
for (i in 1:nsim) {
  set.seed(1234 + i)
  n <- 1000
  cauchy_samples <- reauchy(n, location = 5, scale = 1)</pre>
  theta1[i] <- mean(cauchy_samples)</pre>
  theta2[i] <- median(cauchy samples)</pre>
  theta3[i] <- theta2[i] + 2/n * dloglik(cauchy_samples, theta2[i])</pre>
  theta4[i] <- optimize(nloglik, interval = c(-10,10))$minimum
mse <- function(hat_theta, theta) {</pre>
 return(mean(hat_theta - theta)^2)
}
cat("mse1 =", mse(theta1, 5), "\n")
## mse1 = 1.530928
cat("mse2 =", mse(theta2, 5), "\n")
## mse2 = 6.172925e-06
cat("mse3 =", mse(theta3, 5), "\n")
## mse3 = 1.968817e-06
cat("mse4 =", mse(theta4, 5), "\n")
## mse4 = 1.835778e-06
# Coverage probability
cov_prob <- function(hat_theta, theta, eps = 0.1) {</pre>
  return(mean(abs(hat_theta - theta) <= eps))</pre>
cat("cov_prob1 =", cov_prob(theta1, 5), "\n")
## cov_prob1 = 0.064
cat("cov_prob2 =", cov_prob(theta2, 5), "\n")
## cov_prob2 = 0.957
cat("cov_prob3 =", cov_prob(theta3, 5), "\n")
## cov_prob3 = 0.972
```

```
cat("cov_prob4 =", cov_prob(theta4, 5), "\n")
```

cov_prob4 = 0.973