

Solution 5: Statistical inference (II)

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EM and Newton-Raphson implementation

The ABO-gene or ABO-locus is on chromosome 9. It has 3 alleles (antigens) (A, B, O) and it determines 4 blood type (A, B, AB, O).

We have a large random sample obtained from Berlin (Bernstein 1925, Sham's book page 44):

- $n_A = 9123$ blood type A
- $n_B = 2987$ blood type B
- $n_{AB} = 1269$ blood type AB
- $n_O = 7725$ blood type O

For instance, $n_A = 9123 = n_{AA} + n_{AO}$: Among 9123 blood type A individuals, some have genotype AA and the others have genotype AO .

Our interest is to estimate the allele frequencies of alleles A , B , and O . i.e. $p = \text{freq}(\text{allele } A)$, $q = \text{freq}(\text{allele } B)$, $1 - p - q = \text{freq}(\text{allele } O)$.

1. Write out the log-likelihood $L(p, q)$.
2. Is there a closed-form solution of this log-likelihood function?
3. Formulate the problem as a missing data problem and use the Newton-Raphson algorithm to find the MLEs, \hat{p} and \hat{q} , that maximize the log-likelihood, $\ln L(p, q)$.
4. (Advanced) Use the EM algorithm to find the Maximum Likelihood Estimates (MLEs) of parameters, \hat{p} and \hat{q} .

Hint: Lei Sun's STA2080 Modern genetic statistics notes ([link](#)).

Solution

1. Let $X = (n_A, n_B, n_{AB}, n_O)$, X follows the multinomial distribution,

$$L(p, q) = \binom{n}{n_A, n_B, n_{AB}, n_O} (p^2 + 2p(1 - p - q))^{n_A} (q^2 + 2q(1 - p - q))^{n_B} (2pq)^{n_{AB}} ((1 - p - q)^2)^{n_O}$$

The general approach to estimate allele frequency is maximum likelihood estimation. To find MLE, take the derivatives of the log-likelihood function, and find the pair of (p, q) that set the derivatives to 0.

The log transformation is performed since finding a maximizer of $L(p, q)$ is equivalent to finding a maximizer of $\ln L(p, q)$. The log-likelihood is,

$$\ln L(p, q) \sim n_A \ln(p^2 + 2p(1 - p - q)) + n_B \ln(q^2 + 2q(1 - p - q)) + n_{AB} \ln(2pq) + n_O \ln((1 - p - q)^2) \quad (1)$$

Take the partial derivatives of the log-likelihood and set them to 0,

$$\frac{\partial \ln L(p, q)}{\partial p} = \frac{2(1 - p - q)}{p(2 - p - 2q)} n_A + \frac{2}{2p + q - 2} n_B + \frac{1}{p} n_{AB} - \frac{2}{1 - p - q} n_O = 0 \quad (2)$$

$$\frac{\partial \ln L(p, q)}{\partial q} = \frac{2}{p + 2q - 2} n_A + \frac{2(1 - p - q)}{q(2 - 2p - q)} n_B + \frac{1}{q} n_{AB} - \frac{2}{1 - p - q} n_O = 0 \quad (3)$$

2. It is hard to find the explicit form of the (p, q) from equation (2) and (3) **directly**. We consider solve the problem **iteratively**.
3. Newton-Raphson algorithm is another method that can be applied to estimate the ABO allele frequency. The algorithm is initially designed to find the roots (or zeros) of a real-valued function iteratively.

In the ABO blood type settings, maximizing the likelihood is equivalent to finding the roots of the derivative of log-likelihood function. Since it is hard to find the roots directly, one can approximate it iteratively. Here denote the parameters to be estimated as $\vec{\theta} = (p, q)$, denote the log-like hood shown in equation (1) as $f(\vec{\theta}) = \ln L(\theta)$, then the partial derivatives of the log-likelihood (score function) is $f'(\vec{\theta}) = f'(p, q)$, the second derivatives (Hessian matrix, or observed information $-I(\theta)$) as $f''(\vec{\theta}) = f''(p, q)$. The explicit forms of the score function and observed information are shown in the appendix.

4. Expectation-Maximum (EM) algorithm is a method for obtaining the Maximum likelihood estimates (MLE) of parameters iteratively. It usually contains two parts: **E-step (expectation)** computes an expected value of the log-likelihood using current estimate for the parameters; **M-step (Maximization)** calculates MLE based on the log likelihood in the E-step, then updates the estimates of the parameters.

The EM algorithms are often used when the model contains unobserved data. In the ABO settings, one could observe the phenotype counts (n_A, n_B, n_{AB}, n_O) , while the genotype counts n_{AA} or n_{AO} in blood A group and n_{BB} or n_{BO} in blood B group are missing. This leads to a problem when applying direct counting (Sham, 1998) to estimate p, q , the allele frequency of A and B respectively. The following steps will show how this missing data problem could be solved by the EM algorithm.

In each iteration k , firstly, the **E-step** computes the expected value of the log-likelihood $h(p, q)$ using the observed data $n_{\text{obs}} = (n_A, n_B, n_{AB}, n_O)$ and current parameter value $p^{(k)}, q^{(k)}$:

$$Q_k(p, q) = E[h_k(p, q) | n_{\text{obs}}, p^{(k)}, q^{(k)}] \quad (4)$$

and $h_k(p, q) \sim 2n_{AA} \log p^{(k)} + n_{AO} \log(2p^{(k)}(1 - p^{(k)} - q^{(k)})) + 2n_{BB} \log q^{(k)} + n_{BO} \log(2p^{(k)}(1 - p^{(k)} - q^{(k)})) + n_{AB} \log(2p^{(k)}q^{(k)}) + 2n_O \log(1 - p^{(k)} - q^{(k)})$

Since the log-likelihood is linear w.r.t. the missing data, therefore, when taking the expectation for each component, the missing data can be imputed in this case. For example, since $E[2n_{AA} \log p^{(k)} | n_{\text{obs}}, p^{(k)}, q^{(k)}] = 2 \log p^{(k)} E[n_{AA} | n_{\text{obs}}, p^{(k)}, q^{(k)}]$, and under the HWE assumption,

$$E[n_{AA} | n_{\text{obs}}, p^{(k)}, q^{(k)}] = \frac{\text{freq}(AA)}{\text{freq}(AA) + \text{freq}(AO)} n_A = \frac{p^{(k)} p^{(k)}}{p^{(k)} p^{(k)} + 2p^{(k)} (1 - p^{(k)} - q^{(k)})} n_A := n_{AA}^{(k)}$$

By the similar calculation, we can get the imputed data $n_{AA}^{(k)}, n_{AO}^{(k)}, n_{BB}^{(k)}, n_{BO}^{(k)}$ for each iteration. Then the **M-step** computes the MLE based on the likelihood in equation (4). To be specific,

$$\frac{\partial Q_k(p, q)}{\partial p} = \frac{2n_{AA}^{(k)} + n_{AO}^{(k)} + n_{AB}}{p} - \frac{n_{AO}^{(k)} + n_{BO}^{(k)} + n_O}{1 - p - q} = 0$$

$$\frac{\partial Q_k(p, q)}{\partial q} = \frac{2n_{BB}^{(k)} + n_{BO}^{(k)} + n_{AB}}{q} - \frac{n_{AO}^{(k)} + n_{BO}^{(k)} + n_O}{1 - p - q} = 0$$

Thus, the updated values of parameters are

$$p^{(k+1)} = \frac{2n_{AA}^{(k)} + n_{AO}^{(k)} + n_{AB}}{2n} \quad q^{(k+1)} = \frac{2n_{BB}^{(k)} + n_{BO}^{(k)} + n_{AB}}{2n}$$

```
max <- 10000
epsilon <- 1e-5
iter <- 0
p <- 0.3333333
q <- 0.3333333
diff1 <- 1
diff2 <- 1

ep <- NULL
eq <- NULL
elike <- NULL
ep[1] <- p
eq[1] <- q
elike[1] <- log_like(nA, nB, nAB, nO, p, q)

while (diff1 > epsilon & diff2 > epsilon & iter < max) {

  # E-step
  nAA <- nA * (p*p) / (p*p + 2*p*(1 - p - q))
  nAO <- nA * 2*p*(1 - p - q) / (p*p + 2*p*(1 - p - q))
  nBB <- nB * (q*q) / (q*q + 2*q*(1 - p - q))
  nBO <- nB * 2*q*(1 - p - q) / (q*q + 2*q*(1 - p - q))

  # M-step
  p.new <- (2*nAA + nAO + nAB) / (2*n)
  q.new <- (2*nBB + nBO + nAB) / (2*n)

  diff1 <- abs(p.new - p)
  diff2 <- abs(q.new - q)

  p <- p.new
  q <- q.new

  log_lik <- log_like(nA, nB, nAB, nO, p, q)

  iter <- iter + 1
  ep[iter+1] <- p
  eq[iter+1] <- q
}
```

```

  elike[iter+1] <- log_lik
}

knitr::kable(
  data.frame(c(0:(length(ep)-1)), ep, eq, elike),
  col.names = c("iteration $k$", "$p^{\{(k)\}}$", "$q^{\{(k)\}}$", "log likelihood"), booktabs = TRUE,
  align = "cccr",
  caption = 'Results for EM algorithm'
)

```

```

## Warning in attr(x, "align"): 'xfun::attr()' is deprecated.
## Use 'xfun::attr2()' instead.
## See help("Deprecated")

## Warning in attr(x, "format"): 'xfun::attr()' is deprecated.
## Use 'xfun::attr2()' instead.
## See help("Deprecated")

```

Table 1: Results for EM algorithm

| iteration k | $p^{(k)}$ | $q^{(k)}$ | log likelihood |
|---------------|-----------|-----------|----------------|
| 0 | 0.3333333 | 0.3333333 | -32186.43 |
| 1 | 0.3182572 | 0.1244235 | -24998.46 |
| 2 | 0.2942165 | 0.1079404 | -24827.44 |
| 3 | 0.2888920 | 0.1066936 | -24822.90 |
| 4 | 0.2879007 | 0.1065736 | -24822.76 |
| 5 | 0.2877236 | 0.1065579 | -24822.76 |
| 6 | 0.2876923 | 0.1065555 | -24822.76 |

```

df <- function(p, q) {
  dfp <- 2*(1 - p - q)*nA / (p*(2 - p - 2*q)) + 2*nB / (2*p + q - 2) + nAB/p - 2*n0 / (1 - p - q)
  dfq <- 2*nA / (p + 2*q - 2) + 2*(1 - p - q)*nB / (q*(2 - 2*p - q)) + nAB/q - 2*n0 / (1 - p - q)
  c(dfp, dfq)
}

d2f <- function(p, q) {
  pp <- -4*(1 - p - q)^2*nA / (p^2*(2 - p - 2*q)^2) - 2*nA / (p*(2 - p - 2*q)) -
    4*nB / ((2 - 2*p - q)^2) - nAB / (p^2) - 2*n0 / ((1 - p - q)^2)
  pq <- 4*(1 - p - q)*nA / (p*(2 - p - 2*q)^2) - 2*nA / (p*(2 - p - 2*q)) -
    2*nB / ((2 - 2*p - q)^2) - 2*n0 / ((1 - p - q)^2)
  qp <- -2*nA / ((2 - p - 2*q)^2) + 4*(1 - p - q)*nB / (q*(2 - 2*p - q)^2) -
    2*nB / (q*(2 - 2*p - q)) - 2*n0 / ((1 - p - q)^2)
  qq <- -4*nA / ((2 - p - 2*q)^2) - 4*(1 - p - q)^2*nB / ((q^2*(2 - 2*p - q)^2)) -
    2*nB / (q*(2 - 2*p - q)) - nAB / (q^2) - 2*n0 / ((1 - p - q)^2)
  matrix(c(pp, pq, qp, qq), nrow = 2, ncol = 2, byrow = T)
}

max <- 10000
epsilon <- 1e-5
iter <- 0
p <- 0.3333333
q <- 0.3333333
diff1 <- 1

```

```

diff2 <- 1
nA <- 9123
nB <- 2987
nAB <- 1269
n0 <- 7725

theta <- c(p, q)
rp <- NULL
rq <- NULL
rlike <- NULL
rp[1] <- p
rq[1] <- q
rlike[1] <- log_like(nA, nB, nAB, n0, p, q)

while (diff1 > epsilon & diff2 > epsilon & iter < max) {
  p <- theta[1]
  q <- theta[2]

  theta.new <- theta - solve(d2f(p,q)) %*% df(p, q)

  diff1 <- abs(theta[1] - theta.new[1])
  diff2 <- abs(theta[2] - theta.new[2])

  theta <- theta.new

  log_lik <- log_like(nA, nB, nAB, n0, theta[1], theta[2])

  iter <- iter + 1

  rp[iter+1] <- theta[1]
  rq[iter+1] <- theta[2]
  rlike[iter+1] <- log_lik
}

knitr::kable(
  data.frame(c(0:(length(rp)-1)), rp, rq, rlike),
  col.names = c("iteration $k$", "$p^{(k)}$", "$q^{(k)}$", "log likelihood"), booktabs = TRUE,
  align = "cccr",
  caption = 'Results for Newton-Raphson algorithm'
)

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## See help("Deprecated")

```

Table 2: Results for Newton-Raphson algorithm

| iteration k | $p^{(k)}$ | $q^{(k)}$ | log likelihood |
|---------------|-----------|-----------|----------------|
| 0 | 0.3333333 | 0.3333333 | -32186.43 |
| 1 | 0.4175144 | 0.0171487 | -29913.64 |
| 2 | 0.2981320 | 0.0318471 | -27106.11 |
| 3 | 0.3034715 | 0.0546993 | -25646.39 |
| 4 | 0.2950282 | 0.0820775 | -24968.81 |
| 5 | 0.2892635 | 0.1013210 | -24828.66 |
| 6 | 0.2877549 | 0.1063252 | -24822.77 |
| 7 | 0.2876857 | 0.1065546 | -24822.76 |
| 8 | 0.2876856 | 0.1065550 | -24822.76 |