



UNIVERSITY OF
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Statistical Sciences

DoSS Summer Bootcamp Probability Module 8

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July 22, 2025

Recap

$$\begin{array}{c} L^q \supseteq L^p \\ \text{in } L^q \Rightarrow \text{in } L^p \\ \Downarrow \\ \text{a.s.} \Rightarrow \text{in 'probability'} \Rightarrow \text{in distribution} \\ \text{CDF} \end{array}$$

Learnt in last module:

- Stochastic convergence
 - ▷ Convergence in distribution
 - ▷ Convergence in probability
 - ▷ Convergence almost surely
 - ▷ Convergence in L^p
 - ▷ Relationship between convergences

Outline

- Convergence of functions of random variables
 - ▷ Slutsky's theorem
 - ▷ Continuous mapping theorem
- Laws of large numbers
 - ▷ WLLN
 - ▷ SLLN
 - ▷ Glivenko-Cantelli theorem
- Central limit theorem

Convergence of functions of random variables

Recall: Stochastic convergence If $X_n \rightarrow X$, $Y_n \rightarrow Y$ in some sense, how is the limiting property of $f(X_n, Y_n)$?

e.g. $X_n + Y_n \rightarrow ?$

$$X_n \cdot Y_n \rightarrow ?$$

$$X_n / Y_n \rightarrow ?$$

$$S_n = X_1 + \dots + X_n$$

$$\frac{S_n}{n^c} \rightarrow ?$$

could be
in prob.
in distribution
a.s.
 L^p

Convergence of functions of random variables

Recall: Stochastic convergence If $X_n \rightarrow X$, $Y_n \rightarrow Y$ in some sense, how is the limiting property of $f(X_n, Y_n)$?

Convergence of functions of random variables (a.s.)

Suppose the probability space is complete, if $X_n \xrightarrow{\text{a.s.}} X$, $Y_n \xrightarrow{\text{a.s.}} Y$, then for any real numbers a, b ,

- $aX_n + bY_n \xrightarrow{\text{a.s.}} aX + bY$;
- $X_n Y_n \xrightarrow{\text{a.s.}} XY$.

Remark:

- Still require all the random variables to be defined on the same probability space

Recall $X_n \rightarrow X$ a.s. if $\mathbb{P}(\lim X_n = X) = 1$

(P.f.) Since $X_n \rightarrow X$ a.s., there exists $N_x \subset \Omega$ s.t.

$$\underbrace{X_n \rightarrow X}_{\text{pointwise}} \text{ on } N_x \quad \text{and} \quad \underbrace{P(N_x) = 1.}$$

Since $Y_n \rightarrow Y$ a.s., there exists $N_y \subset \Omega$ s.t.

$$Y_n \rightarrow Y \text{ on } N_y \quad \text{and} \quad \underbrace{P(N_y) = 1.}$$

On $N_x \cap N_y$, we know $X_n \rightarrow X$, $Y_n \rightarrow Y$ pointwise.

Thus, we have on $N_x \cap N_y$,

$$aX_n + bY_n \rightarrow aX + bY \quad \text{pointwise}$$

$$X_n \cdot Y_n \rightarrow X \cdot Y \quad \text{pointwise.}$$

$$\begin{aligned} P(N_x \cap N_y) &= 1 - P((N_x \cap N_y)^c) \\ &= 1 - \underbrace{P(N_x^c \cup N_y^c)}_{\text{apply union bound}} \end{aligned}$$

$$\geq 1 - \left(\underbrace{P(N_x^c)}_{=0} + \underbrace{P(N_y^c)}_{=0} \right) = 1 - 0 = 1$$

$$\therefore P(N_x \cap N_y) = 1.$$

Thus we have confirmed pointwise convergence of

$$aX_n + bY_n \rightarrow X \quad \text{and} \quad X_n \cdot Y_n \rightarrow X \cdot Y$$

hold with probability 1.

Convergence of functions of random variables

Convergence of functions of random variables (probability)

Suppose the probability space is complete, if $X_n \xrightarrow{P} X, Y_n \xrightarrow{P} Y$, then for any real numbers a, b ,

- $aX_n + bY_n \xrightarrow{P} aX + bY$;
- $X_n Y_n \xrightarrow{P} XY$.

Remark:

- Still require all the random variables to be defined on the same probability space

Recall $X_n \xrightarrow{P} X$ if, $\forall \varepsilon > 0, \mathbb{P}(|X_n - X| > \varepsilon) \rightarrow 0$ as $n \rightarrow \infty$.

$X_n + \tilde{Y}_n \xrightarrow{P} X + Y$ if $X_n \xrightarrow{P} X$, $\tilde{Y}_n \xrightarrow{P} Y$
(p.f.) Let $\forall \varepsilon > 0$.

$$P\left(\underbrace{|X_n + \tilde{Y}_n - (X + Y)|}_{> \varepsilon}\right).$$

↳ by triangle inequality.

$$\{|X_n + \tilde{Y}_n - (X + Y)| > \varepsilon\} \subset \left\{|X_n - X| > \frac{\varepsilon}{2}\right\} \cup \left\{|\tilde{Y}_n - Y| > \frac{\varepsilon}{2}\right\}.$$

$$\Rightarrow \varepsilon < |X_n + \tilde{Y}_n - (X + Y)|$$

$$= |(X_n - X) + (\tilde{Y}_n - Y)|$$

$$\leq \underbrace{|X_n - X| + |\tilde{Y}_n - Y|}_{\text{it's impossible to have both of them } \leq \frac{\varepsilon}{2}}$$

it's impossible to have
both of them $\leq \frac{\varepsilon}{2}$

$$\leq P\left(\left\{|X_n - X| > \frac{\varepsilon}{2}\right\} \cup \left\{|\tilde{Y}_n - Y| > \frac{\varepsilon}{2}\right\}\right)$$

by union bound

$$\left(\leq\right) \underbrace{P\left(|X_n - X| > \frac{\varepsilon}{2}\right)}_{\rightarrow 0 \text{ since } X_n \xrightarrow{P} X} + \underbrace{P\left(|\tilde{Y}_n - Y| > \frac{\varepsilon}{2}\right)}_{\rightarrow 0 \text{ since } \tilde{Y}_n \xrightarrow{P} Y}$$

$\rightarrow 0$ as $n \rightarrow \infty$

Therefore, $X_n + \tilde{Y}_n \xrightarrow{P} X + Y$.

Convergence of functions of random variables

Convergence of functions of random variables (L^p)

Suppose the probability space is complete, if $X_n \xrightarrow{L^p} X, Y_n \xrightarrow{L^p} Y$, then for any real numbers a, b ,

- $aX_n + bY_n \xrightarrow{L^p} aX + bY$;

Remark:

- Still require all the random variables to be defined on the same probability space

$$x_n + \gamma_n \xrightarrow{L^p} x + \gamma \quad \text{if} \quad x_n \xrightarrow{L^p} x, \quad \gamma_n \xrightarrow{L^p} \gamma$$

(pf) Recall that $\|x\|_{L^p} = (\mathbb{E}|x|^p)^{1/p}$ is a norm if $p \geq 1$.

Therefore, we have triangle inequality, i.e.

$$\|x + \gamma\|_{L^p} \leq \|x\|_{L^p} + \|\gamma\|_{L^p}$$

↓ apply

$$\|(x_n + \gamma_n) - (x + \gamma)\|_{L^p} \leq \underbrace{\|x_n - x\|_{L^p}}_{\rightarrow 0} + \underbrace{\|\gamma_n - \gamma\|_{L^p}}_{\rightarrow 0} \rightarrow 0.$$

since $x_n \xrightarrow{L^p} x$
since $\gamma_n \xrightarrow{L^p} \gamma$

Convergence of functions of random variables

Remark: Convergence in distribution is different.

Even $X_n + Y_n \xrightarrow{d} X + Y$ fails
in general

Slutsky's theorem

If $X_n \xrightarrow{d} X$ and $Y_n \xrightarrow{P} c$ (c is a constant), then

- $X_n + Y_n \xrightarrow{d} X + c$;
- $X_n Y_n \xrightarrow{d} cX$;
- $X_n / Y_n \xrightarrow{d} X/c$, where $c \neq 0$.

Convergence of functions of random variables

Remark: Convergence in distribution is different.

Slutsky's theorem

If $X_n \xrightarrow{d} X$ and $Y_n \xrightarrow{P} c$ (c is a constant), then

- $X_n + Y_n \xrightarrow{d} X + c$;
- $X_n Y_n \xrightarrow{d} cX$;
- $X_n / Y_n \xrightarrow{d} X / c$, where $c \neq 0$.

Remark:

- The theorem remains valid if we replace all the convergence in distribution with convergence in probability.

Convergence of functions of random variables

Remark: The requirement that $Y_n \xrightarrow{P} c$ (c is a constant) is necessary.

Convergence of functions of random variables

Remark: The requirement that $Y_n \xrightarrow{P} c$ (c is a constant) is necessary.

Examples:

$X_n \sim \mathcal{N}(0, 1)$, $Y_n = -X_n$, then $Y_n \sim \mathcal{N}(0, 1)$ as well.

- $X_n \xrightarrow{d} Z \sim \mathcal{N}(0, 1)$, $Y_n \xrightarrow{d} Z \sim \mathcal{N}(0, 1)$;
- $X_n + Y_n \xrightarrow{d} 0$; $\nexists Z$
- $X_n Y_n = -X_n^2 \xrightarrow{d} -\chi^2(1)$; $\nexists Z^2 \sim \chi^2(1)$
- $X_n / Y_n = -1$. $\nexists 1 \in \mathbb{R}^2$

Convergence of functions of random variables

Continuous mapping theorem

Let X_n, X be random variables, if $g(\cdot) : \mathbb{R} \rightarrow \mathbb{R}$ satisfies $\mathbb{P}(X \in D_g) = 0$, then

- $X_n \xrightarrow{a.s.} X \Rightarrow g(X_n) \xrightarrow{a.s.} g(X)$;
- $X_n \xrightarrow{P} X \Rightarrow g(X_n) \xrightarrow{P} g(X)$;
- $X_n \xrightarrow{d} X \Rightarrow g(X_n) \xrightarrow{d} g(X)$;

where D_g is the set of discontinuity points of $g(\cdot)$.

g is essentially continuous w.r.t X

L^p convergence fail in general.

Let $X_n = X$ when $X \in L^p$ but $\notin L^{2p}$.

Let $g(x) = x^2$.

Then $g(X_n) \notin L^p$, so L^p convergence doesn't make sense.

Convergence of functions of random variables

Continuous mapping theorem

Let X_n, X be random variables, if $g(\cdot) : \mathbb{R} \rightarrow \mathbb{R}$ satisfies $\mathbb{P}(X \in D_g) = 0$, then

- $X_n \xrightarrow{a.s.} X \Rightarrow g(X_n) \xrightarrow{a.s.} g(X) ;$
- $X_n \xrightarrow{P} X \Rightarrow g(X_n) \xrightarrow{P} g(X) ;$
- $X_n \xrightarrow{d} X \Rightarrow g(X_n) \xrightarrow{d} g(X) ;$

where D_g is the set of discontinuity points of $g(\cdot)$.

Remark:

- If $g(\cdot)$ is continuous, then ...
- If X is a continuous random variable, and D_g only include countably many points, then ...

Law of large numbers

Weak Law of Large Numbers (WLLN)

If X_1, X_2, \dots, X_n are i.i.d. random variables, ~~$\mu = \mathbb{E}(|X_i|)$~~ $\mu = \mathbb{E} X$ $\mu < \infty$, then

$$\bar{X} = \frac{\sum_{i=1}^n X_i}{n} \xrightarrow{P} \mu.$$

Remark:

A more easy-to-prove version is the L^2 weak law, where an additional assumption $\text{Var}(X_i) < \infty$ is required.

Sketch of the proof:

$$\begin{aligned} \mathbb{E} |\bar{X} - \mu|^2 &= \text{Var}(\bar{X}) \\ &= \text{Var}\left(\frac{\sum_{i=1}^n X_i}{n}\right) \end{aligned}$$

Handwritten notes: $\bar{X} \xrightarrow{L^2} \mu$

$$= \frac{\sum_{i=1}^n \text{Var}(X_i)}{n^2}. \quad \text{since } X_i\text{'s are independent.}$$

$$= \frac{n \text{Var}(X_i)}{n^2}$$

$$= \frac{\text{Var}(X_1)}{n} \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

Therefore. $\bar{X} \rightarrow \mu$ in L^2 .

Law of large numbers

A generalization of the theorem: triangular array

Triangular array

A triangular array of random variables is a collection $\{X_{n,k}\}_{1 \leq k \leq n}$.

$$\begin{aligned} n=1 &\rightarrow X_{1,1} \xRightarrow{\text{sum}} S_1 \\ n=2 &\rightarrow X_{2,1}, X_{2,2} \xRightarrow{\text{sum}} S_2 \\ n=3 &\rightarrow X_{3,1}, X_{3,2}, X_{3,3} \xRightarrow{\text{sum}} S_3 \\ &\vdots \\ n &\rightarrow X_{n,1}, X_{n,2}, \dots, X_{n,n} \xRightarrow{\text{sum}} S_n = \sum_{k=1}^n X_{n,k} \end{aligned}$$

Remark: We can consider the limiting property of the row sum S_n .

Law of Large Numbers

L^2 weak law for triangular array

Suppose $\{X_{n,k}\}$ is a triangular array, $n = 1, 2, \dots$, $k = 1, 2, \dots, n$. Let $S_n = \sum_{k=1}^n X_{n,k}$, $\mu_n = \mathbb{E}(S_n)$, if $\sigma_n^2/b_n^2 \rightarrow 0$, where $\sigma_n^2 = \text{Var}(S_n)$ and b_n is a sequence of positive real numbers, then

$$\frac{S_n - \mu_n}{b_n} \xrightarrow{P} 0.$$

Remark:

The L^2 weak law for i.i.d. random variables is a special case of that for triangular array.

Law of large numbers

Proof:

$$\mathbb{E} \left| \frac{\bar{X}_n - \mu_n}{b_n} \right|^2 = \frac{\sigma_n^2}{b_n^2} \rightarrow 0$$

$$\text{So, } \frac{\bar{X}_n - \mu_n}{b_n} \rightarrow 0 \text{ in } L^2,$$

and hence $\rightarrow 0$ in probability. //

Law of large numbers

Proof:

Remark:

A more generalized version incorporates truncation, then the second-moment constraint is relieved.

Law of large numbers

Strong Law of Large Numbers (SLLN)

Let X_1, X_2, \dots be an i.i.d. sequence satisfying $\mathbb{E}(X_i) = \mu$ and $\mathbb{E}(|X_i|) < \infty$, then

$$\bar{X} = \frac{\sum_{i=1}^n X_i}{n} \xrightarrow{\text{a.s.}} \mu.$$

Remark: The proof needs Borel-Cantelli lemma.

Law of large numbers

Strong Law of Large Numbers (SLLN)

Let X_1, X_2, \dots be an i.i.d. sequence satisfying $\mathbb{E}(X_i) = \mu$ and $\mathbb{E}(|X_i|) < \infty$, then

$$\bar{X} = \frac{\sum_{i=1}^n X_i}{n} \xrightarrow{\text{a.s.}} \mu.$$

Remark: The proof needs Borel-Cantelli lemma.

Glivenko-Cantelli theorem

Let $X_i, i = 1, \dots, n$ i.i.d. with distribution function $F(\cdot)$, and let

$F_n(x) = \frac{1}{n} \sum_{i=1}^n I(X_i \leq x)$, then as $n \rightarrow \infty$,

$$\sup_{x \in \mathbb{R}} |F(x) - F_n(x)| \rightarrow 0, \quad \text{a.s.}$$

of X_i 's $\leq x$
 n

Law of large numbers

Proof:

Weaken version: $|F_n(x) - F(x)| \rightarrow 0$ a.s. on \mathbb{P} .

Note that $0 \leq I(X_i \leq x) \leq 1$

So, $0 \leq \underbrace{\mathbb{E} I(X_i \leq x) = \mathbb{P}(X_i \leq x) = F(x)}_{\text{finite}} \leq 1$.

By SLLN,

applied to $Y_i = I(X_i \leq x)$

$$F_n(x) = \frac{1}{n} \sum_{i=1}^n I(X_i \leq x)$$

a.s. $\rightarrow \mathbb{E} I(X_i \leq x) = F(x)$

Limit Theorems and Counterexamples

Recall: For the law of large numbers to hold, the assumption $E|X| < \infty$ is crucial.

Law of Large Numbers fail for infinite mean i.i.d. random variables

If X_1, X_2, \dots are i.i.d. to X with $E|X_i| = \infty$, then for $S_n = X_1 + \dots + X_n$,
 $P(\lim_{n \rightarrow \infty} S_n/n \in (-\infty, \infty)) = 0$.

Proof: Omitted

$\frac{S_n}{n} \rightarrow 0$ fails with probability 1.

Central Limit Theorem

$$\text{Var}(\bar{X}) = \frac{\sigma^2}{n}$$

$$\text{std}(\bar{X}) = \frac{\sigma}{\sqrt{n}} \Leftrightarrow \text{std}(\sqrt{n}\bar{X}) = \sigma$$

What is the limiting distribution of the sample mean?

Classic CLT

Suppose X_1, \dots, X_n is a sequence of i.i.d. random variables with $\mathbb{E}(X_i) = \mu$, $\text{Var}(X_i) = \sigma^2 < \infty$, then

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{d} \mathcal{N}(0, 1).$$

Remark:

- The proof involves characteristic function.
- A more generalized CLT is referred to as "Lindeberg CLT".

$$\frac{\sqrt{n}}{\sigma} \cdot \frac{1}{n} \sum_{i=1}^n (X_i - \mu) = \frac{\sum_{i=1}^n (X_i - \mu)}{\sigma \sqrt{n}}$$

Central Limit Theorem

Example:

Suppose $X_i \sim \text{Bernoulli}(p)$, i.i.d., consider $Z_n = \frac{\sum_{i=1}^n X_i - np}{\sqrt{np(1-p)}}$, then by CLT, $Z_n \sim \mathcal{N}(0, 1)$ asymptotically.


$$\text{Var}(X_i) = p(1-p)$$

Monotone Convergence Theorem

Monotone Convergence Theorem

If $X_n \geq c$ and $X_n \nearrow X$, then $EX_n \nearrow EX$

Usage:

Dominated Convergence Theorem

Dominated Convergence Theorem

If $X_n \rightarrow X$ a.s. and $|X_n| \leq Y$ a.s. for all n and Y is integrable, then $EX_n \rightarrow EX$

Usage:

Delta Method

More about CLT: Delta method

Suppose X_n are i.i.d. random variables with $EX_n = 0$, $VAR(X_n) = \sigma^2 > 0$. Let g be a measurable function that is differentiable at 0 with $g'(0) \neq 0$. Then

$$\sqrt{n} \left(g \left(\frac{\sum_{k=1}^n X_k}{n} \right) - g(0) \right) \rightarrow N(0, \sigma^2 g'(0)^2) \quad \text{weakly.}$$

Proof under stronger assumption: Here, we suppose g is continuously differentiable on \mathbb{R} . If you are interested in a general proof refer to Robert Keener's *Theoretical Statistics*.

Problem Set

Problem 1: Prove that on a complete probability space, if $X_n \xrightarrow{a.s.} X$, $Y_n \xrightarrow{a.s.} Y$, then $X_n + Y_n \xrightarrow{a.s.} X + Y$.

Problem 2: Prove that on a complete probability space, if $X_n \xrightarrow{P} X$, $Y_n \xrightarrow{P} Y$, then $X_n + Y_n \xrightarrow{P} X + Y$.

Problem 3: A bank teller serves customers standing in the queue one by one. Suppose that the service time X_i for customer i has mean $\mathbb{E}(X_i) = 2$ (minutes) and $\text{Var}(X_i) = 1$. We assume that service times for different bank customers are independent. Let Y be the total time the bank teller spends serving 50 customers. Find $\mathbb{P}(90 < Y < 110)$.