



A Statistical Theory of Overfitting for Imbalanced Classification

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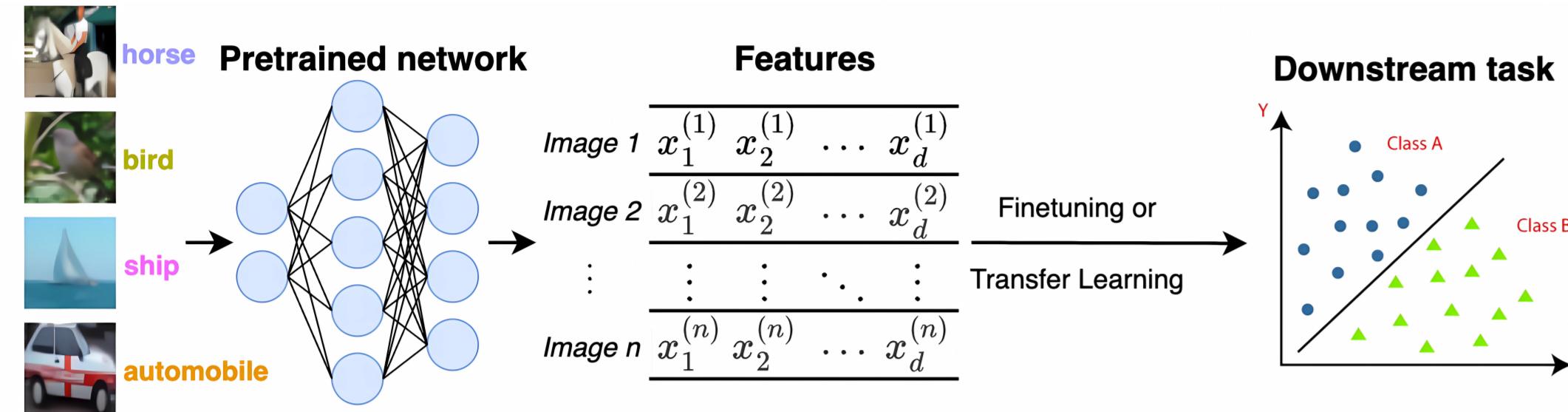
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Challenge in Imbalanced Classification

Training data $\{(\mathbf{x}_i, y_i)\}_{i=1}^n \stackrel{\text{i.i.d.}}{\sim} P_{\mathbf{x}, y}$. Features $\mathbf{x}_i \in \mathbb{R}^d$. For binary labels $y_i \in \{\pm 1\}$.



Class imbalance. $\mathbb{P}(y_i = +1) < \mathbb{P}(y_i = -1)$. (WLOG, assume “+1” is minority)

Challenges of high dimensions: a brief summary of high dimensional statistical theory:

	Low dimensions	High dimensions
Parameter estimation	$\langle \frac{\hat{\beta}}{\ \hat{\beta}\ }, \frac{\beta}{\ \beta\ } \rangle \approx 1$	$\langle \frac{\hat{\beta}}{\ \hat{\beta}\ }, \frac{\beta}{\ \beta\ } \rangle < 1$
Generalization	Training error \approx Test error	Training error $<$ Test error
Distribution of logits	1D projection of P_x	Skewed/distorted 1D projection of P_x

Challenges of data imbalance: minority classes have poor training/test errors, classical theory and finite-sample correction fail in high dimensions, the practice is heuristic-driven and ad hoc...

Key Questions

- [Q1]. Mathematically characterize overfitting in high-dim imbalanced classification?
- [Q2]. What are the adverse effects of overfitting, particularly on the **minority class**?
- [Q3]. What are the consequences for uncertainty quantification, such as calibration?

Setup of Theory

Theoretical tools: consider a **two-component Gaussian mixture model (2-GMM)**

$$\begin{aligned} \textcircled{1} \quad \text{Minority: } \mathbb{P}(y_i = +1) &= \pi, \\ \textcircled{2} \quad \text{Majority: } \mathbb{P}(y_i = -1) &= 1 - \pi, \quad \textcircled{2} \quad \mathbf{x}_i | y_i \sim \mathcal{N}(y_i \boldsymbol{\mu}, \mathbf{I}_d). \end{aligned} \quad (1)$$

Focus on linear classifier $\hat{y}(\mathbf{x}) = 2\mathbb{I}\{\hat{f}(\mathbf{x}) > 0\} - 1$ with $\hat{f}(\mathbf{x}) = \langle \mathbf{x}, \hat{\beta} \rangle + \hat{\beta}_0$, where $\hat{\beta}$, $\hat{\beta}_0$ are estimated by two standard approaches: (generalized) logistic regression and support vector machines (SVMs). ($\hat{f}(\mathbf{x})$ is strictly convex decreasing, including $\log(1 + e^{-x})$.)

$$\text{logistic regression: } \underset{\beta \in \mathbb{R}^d, \beta_0 \in \mathbb{R}}{\text{minimize}} \quad \frac{1}{n} \sum_{i=1}^n \ell(y_i(\langle \mathbf{x}_i, \beta \rangle + \beta_0)), \quad (2a)$$

$$\text{SVM: } \underset{\beta \in \mathbb{R}^d, \beta_0 \in \mathbb{R}}{\text{maximize}} \quad \kappa, \quad \text{(max-margin classifier)} \quad \text{subject to } y_i(\langle \mathbf{x}_i, \beta \rangle + \beta_0) \geq \kappa, \quad \forall i \in [n], \quad \|\beta\|_2 \leq 1. \quad (2b)$$

These two classifiers are closely related by **inductive bias** on separable data. Our theory can also be extended to multiple classes and non-isotropic covariance.

Characterizing Overfitting via Empirical Logit Distribution

Empirical logit distribution (ELD), or **training logit distribution** is defined as the empirical distribution of label-logit pairs in the training set. **Testing logit distribution (TLD)** is defined as the distribution of the label-logit pair for a test data point.

$$\text{ELD: } \hat{\nu}_n = \frac{1}{n} \sum_{i=1}^n \delta_{(y_i, \hat{f}(\mathbf{x}_i))}, \quad \text{TLD: } \hat{\nu}_n^{\text{test}} = \text{Law}(y_{\text{test}}, \hat{f}(\mathbf{x}_{\text{test}})), \quad (3)$$

Let δ_a be delta measure supported at a , and Law be the distribution of random variables/vectors.

The **discrepancy** between train/test accuracies is known as **overfitting**, which can be analyzed via ELD and TLD. For separable data, see simulation in Figure 2 (2-GMM) and real-data examples in Figure 3 (pretrained neural network).

Takeaway: overfitting = “truncation”. (for non-separable data: “shrinking/skewing” effect)

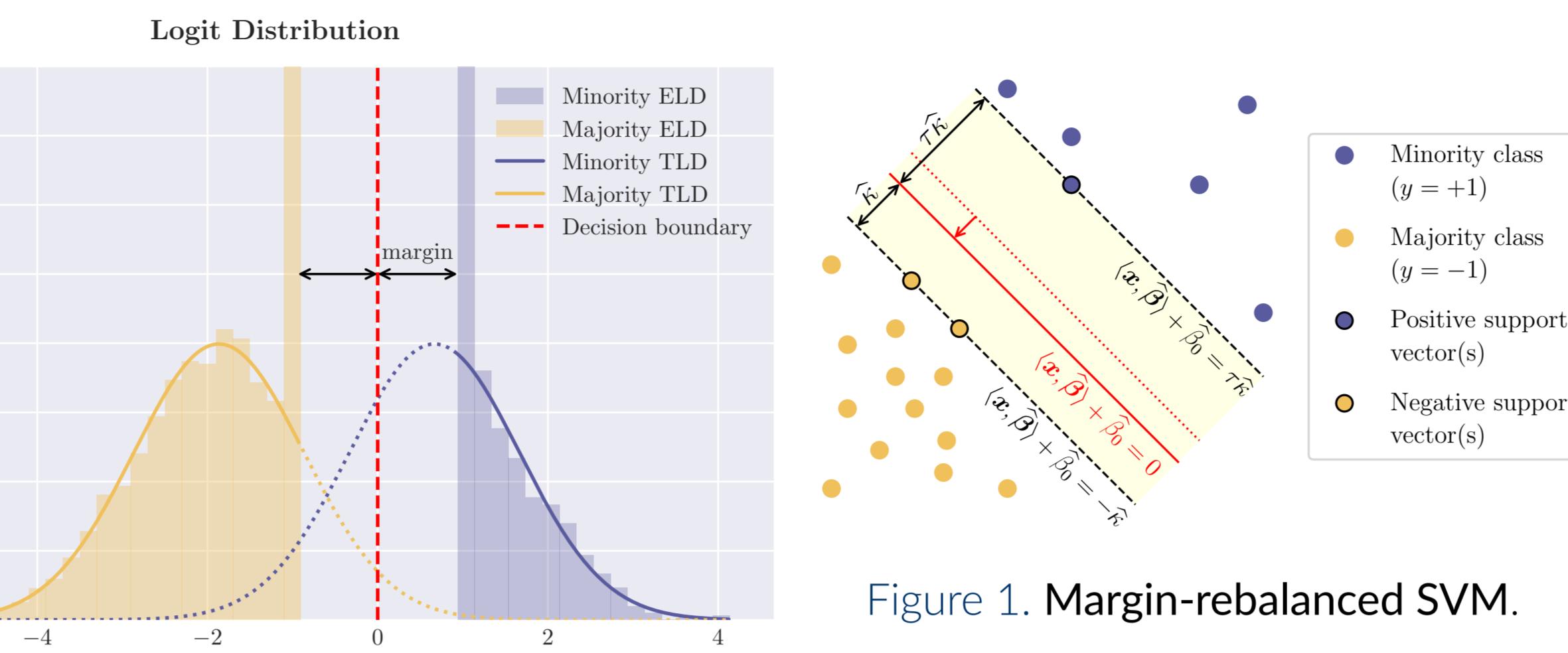


Figure 1. Margin-rebalanced SVM.

Figure 2. **Empirical logit distribution (ELD) and testing logit distribution (TLD)**. We train a max-margin classifier \hat{f} on synthetic data from a 2-component Gaussian mixture model. Colors indicate labels y_i and x -axis indicates logits $\hat{f}(\mathbf{x}_i)$. **ELD for both classes:** the rectified Gaussian distribution (histogram). **TLD for both classes:** Gaussian distribution (curve). **Overfitting effect:** The density areas below the dotted curves are overlapping in TLD \Rightarrow test errors > 0 ; but they are “pushed” to respective margin boundaries in ELD \Rightarrow separability and training errors $= 0$.

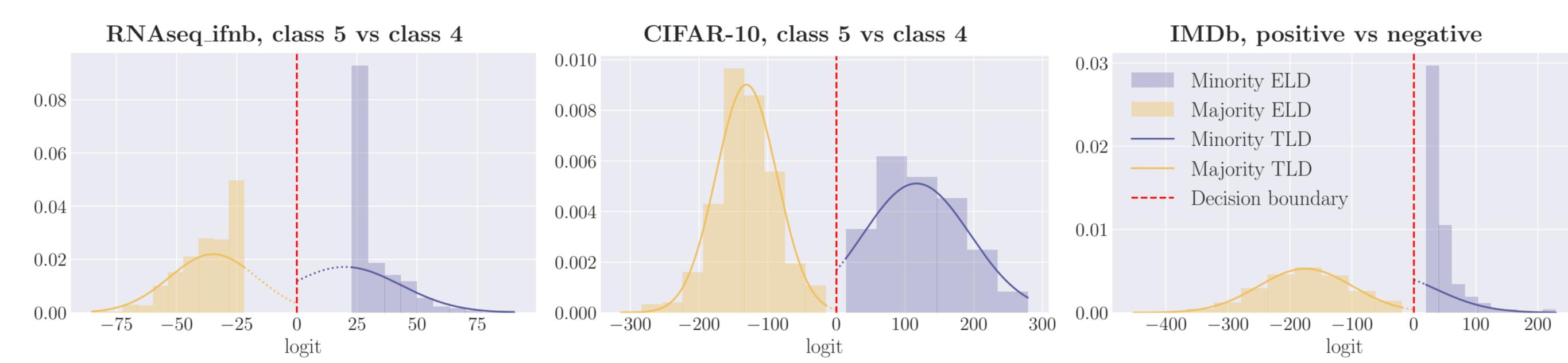


Figure 3. **ELD & TLD for real data.** Left: IFNB single-cell RNA-seq dataset (tabular data). Middle: CIFAR-10 preprocessed by the pretrained ResNet-18 for feature extraction (image data). Right: IMDb movie review preprocessed by BERT base model (110M) for feature extraction (text data).

Theoretical results: variational characterization of ELD vs. TLD

Let $(\hat{\beta}, \hat{\beta}_0, \hat{\kappa})$ be trained from (2b), where $\hat{\kappa}$ is the **margin**. Denote $\hat{\rho} := \langle \frac{\hat{\beta}}{\|\hat{\beta}\|}, \frac{\mu}{\|\mu\|} \rangle$. On a test point $(\mathbf{x}_{\text{test}}, y_{\text{test}}) \sim P_{\mathbf{x}, y}$, we consider the minority and majority errors

$$\text{Err}_+ := \mathbb{P}(\hat{f}(\mathbf{x}_{\text{test}}) \leq 0 \mid y_{\text{test}} = +1), \quad \text{Err}_- := \mathbb{P}(\hat{f}(\mathbf{x}_{\text{test}}) > 0 \mid y_{\text{test}} = -1). \quad (4)$$

Theorem (Separable data). Consider 2-GMM with $n/d \rightarrow \delta \in (0, \infty)$ as $n, d \rightarrow \infty$. There is a critical threshold $\delta_c = \delta_c(\pi, \|\mu\|)$, such that if $\delta < \delta_c$, as $n, d \rightarrow \infty$:

- Phase transition.** $\mathbb{P}\{\text{training set is linearly separable}\} \rightarrow 1$.
- Parameter convergence.** $(\hat{\rho}, \hat{\beta}_0, \hat{\kappa}) \xrightarrow{p} (\rho^*, \beta_0^*, \kappa^*)$, where $(\rho^*, \beta_0^*, \kappa^*)$ is unique optimal solution to the following **variational problem**: (we have $\rho^* > 0, \beta_0^* < 0$)

$$\underset{\rho \in [-1, 1], \beta_0 \in \mathbb{R}, \kappa > 0, \xi \in \mathcal{L}^2}{\text{maximize}} \quad \kappa, \quad \text{s.t. } \rho \|\mu\| + G + Y\beta_0 + \sqrt{1 - \rho^2}\xi \geq \kappa, \quad \mathbb{E}[\xi^2] \leq 1/\delta. \quad (5)$$

(Let \mathcal{L}^2 denote all square integrable random variables in $(\Omega, \mathcal{F}, \mathbb{P})$, and $(Y, G) \sim P_y \times \mathcal{N}(0, 1)$ where $P_y = \text{Law}(y_i)$. Note that ξ is an unknown random variable (function) to be optimized.)
- Asymptotic errors.** $\text{Err}_{\pm} \rightarrow \Phi(-\rho^* \|\mu\| \mp \beta_0^*)$. $(\Phi(t) = \mathbb{P}(\mathcal{N}(0, 1) \leq t))$
- ELD convergence.** The empirical (training) logit distribution $\hat{\nu}_n$ has limit ν_* :

$$W_2(\hat{\nu}_n, \nu_*) \xrightarrow{p} 0, \quad \text{where } \nu_* := \text{Law}(Y, Y \max\{\kappa^*, \rho^* \|\mu\| + G + Y\beta_0^*\}).$$

TLD convergence. The testing logit distribution $\hat{\nu}_n^{\text{test}}$ has limit ν_*^{test} :

$$\hat{\nu}_n^{\text{test}} \xrightarrow{w} \nu_*^{\text{test}}, \quad \text{where } \nu_*^{\text{test}} := \text{Law}(Y, Y(\rho^* \|\mu\| + G + Y\beta_0^*)).$$

Rebalancing margin is crucial

Mainstay: take a **hyperparameter $\tau > 0$** and consider the **margin-rebalanced SVM**

$$\underset{\beta \in \mathbb{R}^d, \beta_0, \kappa \in \mathbb{R}}{\text{maximize}} \quad \kappa, \quad \text{subject to } \tilde{y}_i(\langle \mathbf{x}_i, \beta \rangle + \beta_0) \geq \kappa, \quad \forall i \in [n], \quad \|\beta\| \leq 1, \quad (6)$$

where $\tilde{y}_i = \tau^{-1}$ if $y_i = +1$, otherwise $\tilde{y}_i = -1$. This **shifts the decision boundary** as shown in Figure 1. For imbalanced classification, it is common to consider the **balanced error** $\text{Err}_b := (\text{Err}_+ + \text{Err}_-)/2$. We conduct analysis under two regimes:

(i) proportional regime, where $n/d \rightarrow \delta \in (0, \infty)$ as $n, d \rightarrow \infty$. Denote $\text{Err}_+^*, \text{Err}_-^*$, Err_b^* as the limits of Err_+ , Err_- , Err_b as $n \rightarrow \infty$, respectively.

Proposition (Optimal τ in proportional regime). Define the optimal margin ratio which minimizes the asymptotic balanced error as $\tau^{\text{opt}} := \arg \min_{\tau} \text{Err}_b^*$. When $\tau = \tau^{\text{opt}} > 0$, we have $\beta_0^* = 0$, $\text{Err}_+^* = \text{Err}_-^* = \text{Err}_b^*$. (roughly speaking, $\tau^{\text{opt}} \asymp \sqrt{1/\pi}$)

A critical observation is that, changing τ only has an effect on $\hat{\beta}_0$ but not $\hat{\beta}$. When $\tau = \tau^{\text{opt}}$, we have **monotone trends** of the errors, see summary in Table 1.

(ii) high imbalance regime, in the sense $\pi \propto d^{-a}$, $\|\mu\|^2 \propto d^b$, $n \propto d^{c+1}$ as $d \rightarrow \infty$.

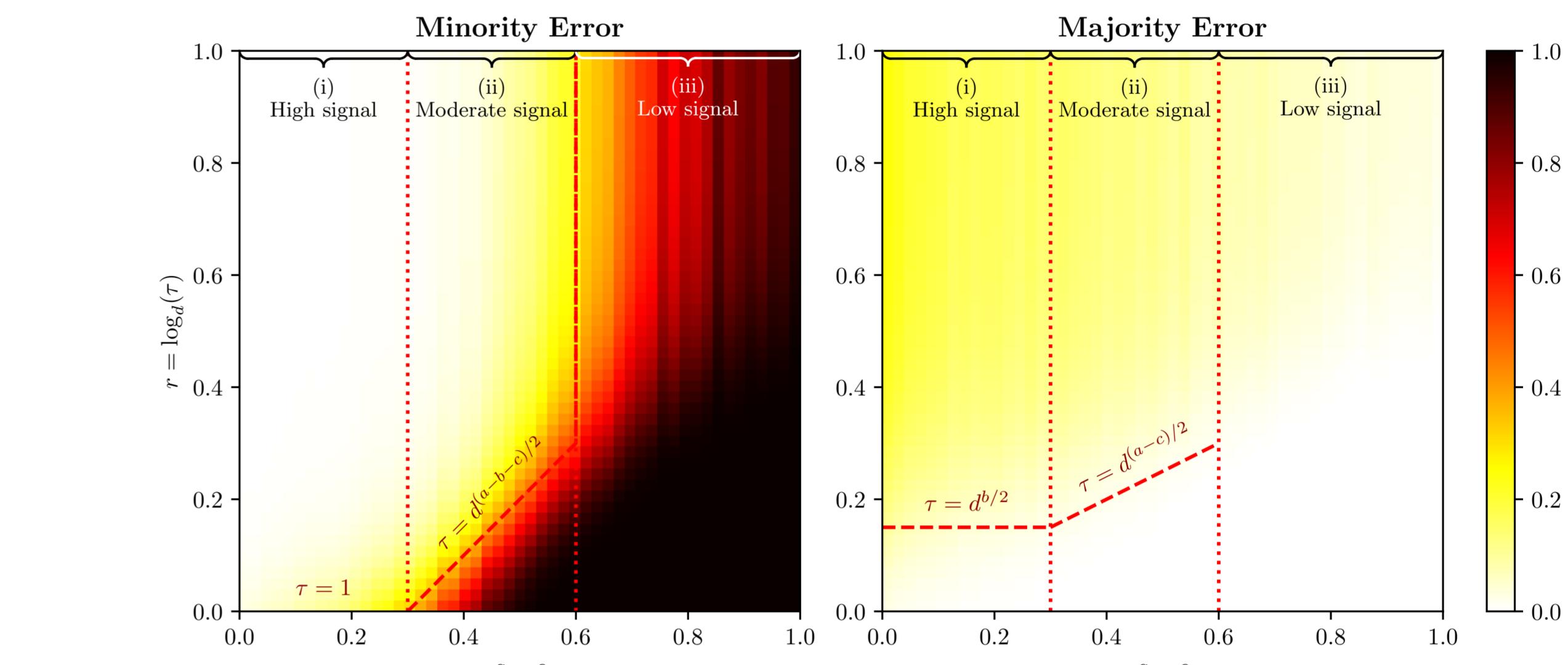


Figure 4. **Phase transition in high imbalance regime.** 2-GMM simulation under different settings of parameters (a, c) and $\tau = d^r$ ($b = 0.3$). Left: minority accuracy is (i) high for any τ under high signal, (ii) high for $\tau \gg d^{(a-b-c)/2}$ under moderate signal, but (iii) low for any τ under low signal. Right: majority accuracy is close to 1 under high/moderate signal as long as τ is not too large.

Theorem (High imbalance). Consider 2-GMM as $d \rightarrow \infty$. Suppose $a - c < 1$.

- High signal:** $a - c < b$. If take $1 \leq \tau_d \ll d^{b/2}$, then $\text{Err}_+ = o(1)$ and $\text{Err}_- = o(1)$.
- Moderate signal:** $b < a - c < 2b$. If $d^{a-b-c} \ll \tau_d \ll d^{(a-c)/2}$, then $\text{Err}_+ = o(1)$ and $\text{Err}_- = o(1)$. If naively take $\tau_d \asymp 1$, then $\text{Err}_+ = 1 - o(1)$ and $\text{Err}_- = o(1)$.
- Low signal:** $a - c > 2b$. For any τ_d , we have $\text{Err}_b \geq \frac{1}{2} - o(1)$.

Consequences for confidence estimation and calibration

Confidence: prediction probability, i.e., $\hat{p}(\mathbf{x}) := \sigma(\hat{f}(\mathbf{x}))$ where $\sigma(t) = (1 + e^{-t})^{-1}$.

Calibration: quantity uncertainty, measure faithfulness of prediction probabilities.

$\hat{p}_0(\mathbf{x}) := \mathbb{P}(y = 1 \mid \hat{p}(\mathbf{x}))$, $p^*(\mathbf{x}) := \mathbb{P}(y = 1 \mid \mathbf{x})$. Some popular miscalibration metrics: **calibration error** $\text{CalErr}(\hat{p}) := \mathbb{E}[(\hat{p}(\mathbf{x}) - \hat{p}_0(\mathbf{x}))^2]$, **mean squared error** $\text{MSE}(\hat{p}) := \mathbb{E}[(\mathbb{1}_{y=1} - \hat{p}(\mathbf{x}))^2]$, and **confidence estimation error** $\text{ConfErr}(\hat{p}) := \mathbb{E}[(\hat{p}(\mathbf{x}) - p^*(\mathbf{x}))^2]$.

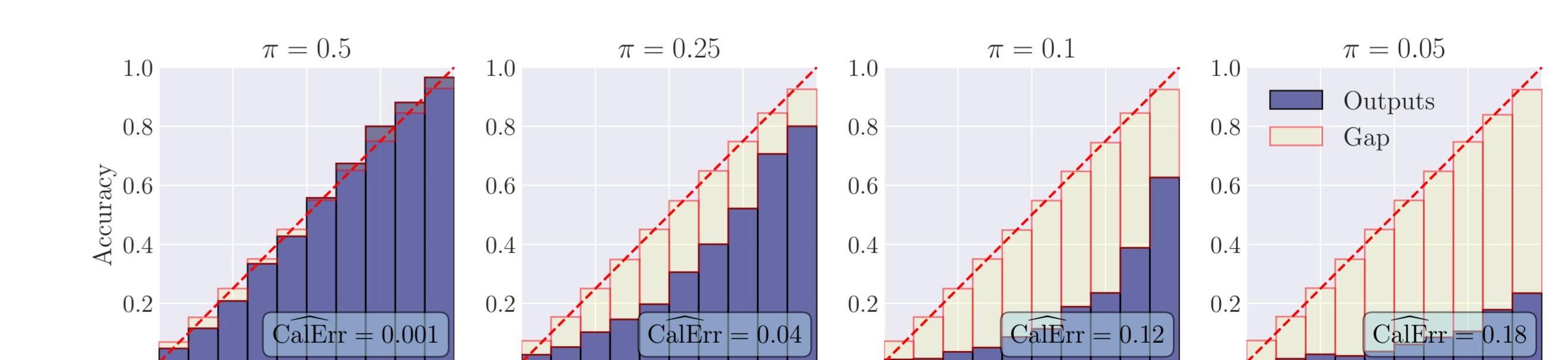


Figure 5. **Reliability diagrams: imbalance worsens calibration.** In 2-GMM simulations, we train SVMs and obtain confidence $\hat{p}(\mathbf{x})$. For each p (x -axis), we calculate $\mathbb{P}(y = 1 \mid \hat{p}(\mathbf{x}) = p)$ (y -axis) based on a test set. As imbalance increases (smaller π), the classifier becomes more miscalibrated.

Theoretical results:	$\text{Err}_+^*, \text{Err}_-^*, \text{Err}_b^*$	$\text{CalErr}^*, \text{MSE}^*, \text{ConfErr}^*$
imbalance ratio $\pi \uparrow$	↓	↓ ↓
signal strength $\ \mu\ _2 \uparrow$	↓	↓ ↓
aspect ratio $n/d \rightarrow \delta \uparrow$	↓	↓ ↓ ↓

Table 1. Monotonicity of test errors and miscalibration metrics on model parameters.