

# Artificial Intelligence: HW 2

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November 6, 2023

## 1 Linear Regression

### 1.a

Let  $f$  be the target function. I'll use the superscript with parenthesis to describe the n-th sample vector.

$$f(\boldsymbol{\omega}, \overline{\mathbf{x}}^{(n)}) = \sum_n \frac{1}{2} \left( t^{(n)} - \boldsymbol{\omega}^T \overline{\mathbf{x}}^{(n)} \right)^2 \quad (1)$$

To minimize  $f$ , differentiate it by  $\boldsymbol{\omega}$  and find the  $\boldsymbol{\omega}_0$  which makes the derivative zero. To make expression simple, I used the Einstein notation.

$$\begin{aligned} \frac{\partial f}{\partial \omega_j} &= -(t^{(n)} - \boldsymbol{\omega}^T \overline{\mathbf{x}}^{(n)}) \cdot \frac{\partial}{\partial \omega_j} \boldsymbol{\omega}^T \overline{\mathbf{x}}^{(n)} \\ &= -(t^{(n)} - \boldsymbol{\omega}^T \overline{\mathbf{x}}^{(n)}) x_j^{(n)} \\ &= 0 \end{aligned}$$

By enumerating the  $\frac{\partial f}{\partial \omega_j}$  horizontally, one can get  $\frac{\partial f}{\partial \boldsymbol{\omega}}$ .

$$\sum_n t^{(n)} \begin{pmatrix} x_1^{(n)} \\ \vdots \\ x_M^{(n)} \end{pmatrix}^T = \sum_n \begin{pmatrix} \boldsymbol{\omega}^T \overline{\mathbf{x}}^{(n)} x_1^{(n)} \\ \vdots \\ \boldsymbol{\omega}^T \overline{\mathbf{x}}^{(n)} x_M^{(n)} \end{pmatrix}^T \quad (2)$$

The left hand side is simply  $\sum_n t^{(n)} \overline{\mathbf{x}}^{(n)T}$ . From the linearity of vector summation rule, the right hand side is  $\left( \left( \sum_n \overline{\mathbf{x}}^{(n)} \cdot \overline{\mathbf{x}}^{(n)T} \right) \boldsymbol{\omega} \right)^T$ . By taking transpose to both sides, one can get the following equation.

$$\left[ \sum_n \overline{\mathbf{x}}^{(n)} \cdot \overline{\mathbf{x}}^{(n)T} \right] \boldsymbol{\omega} = \sum_n t^{(n)} \overline{\mathbf{x}}^{(n)} \quad (3)$$

Therefore,  $\mathbf{A} = \sum_n \overline{\mathbf{x}}^{(n)} \cdot \overline{\mathbf{x}}^{(n)T}$  and  $\mathbf{b} = \sum_n t^{(n)} \overline{\mathbf{x}}^{(n)}$

### 1.b

$\overline{\mathbf{x}}^{(1)} = (1, 0)^T, t^{(1)} = 1$ .  $\overline{\mathbf{x}}^{(2)} = (1, \epsilon)^T, t^{(2)} = 1$ .  $\mathbf{A} = \overline{\mathbf{x}}^{(1)} \cdot \overline{\mathbf{x}}^{(1)T} + \overline{\mathbf{x}}^{(1)} \cdot \overline{\mathbf{x}}^{(2)T} = \begin{pmatrix} 2 & \epsilon \\ \epsilon & \epsilon^2 \end{pmatrix}$   
 $\mathbf{b} = \overline{\mathbf{x}}^{(1)} + \overline{\mathbf{x}}^{(2)} = (2, \epsilon)^T$ . Since  $\mathbf{A}$  is invertible (determinant is nonzero.),

$$\boldsymbol{\omega} = \mathbf{A}^{-1} \mathbf{b} \quad (4)$$

$$= \frac{1}{\epsilon^2} \begin{pmatrix} \epsilon^2 & -\epsilon \\ -\epsilon & 2 \end{pmatrix} \begin{pmatrix} 2 \\ \epsilon \end{pmatrix} \quad (5)$$

$$= \begin{pmatrix} 1 \\ 0 \end{pmatrix} \quad (6)$$

### 1.c

$\mathbf{A}$  is same to the above one.  $\mathbf{b} = (1 + \epsilon) \cdot \overline{\mathbf{x}^{(1)}} + \overline{\mathbf{x}^{(2)}} = (2 + \epsilon, \epsilon)^T$

$$\boldsymbol{\omega} = \mathbf{A}^{-1}\mathbf{b} \quad (7)$$

$$= \frac{1}{\epsilon^2} \begin{pmatrix} \epsilon^2 & -\epsilon \\ -\epsilon & 2 \end{pmatrix} \begin{pmatrix} 2 + \epsilon \\ \epsilon \end{pmatrix} \quad (8)$$

$$= \begin{pmatrix} 1 + \epsilon \\ -1 \end{pmatrix} \quad (9)$$

### 1.d

$\boldsymbol{\omega}_b = (1, 0)^T, \boldsymbol{\omega}_c = (1.1, -1)^T$ . The difference of  $\Delta\boldsymbol{\omega} = \boldsymbol{\omega}_c - \boldsymbol{\omega}_b = (\epsilon, -1)^T = (0.1, -1)^T$

## 2 Linear Regression with Regularization

### 2.a

**Claim 1** :  $\mathbf{A}$  is positive semi-definite.

**proof**

$\mathbf{A}$  is trivially symmetry matrix.  $\forall \mathbf{v} \in \mathbb{R}^n, \mathbf{v}^T \mathbf{A} \mathbf{v} = \sum_n \mathbf{v}^T \overline{\mathbf{x}^{(n)}} \cdot \overline{\mathbf{x}^{(n)}}^T \mathbf{v} = \sum_n \|\mathbf{v}^T \overline{\mathbf{x}^{(n)}}\|^2 \geq 0. \blacksquare$

**Claim 2** :  $\mathbf{A}\mathbf{x} = \lambda\mathbf{x} \iff \mathbf{A}^{-1}\mathbf{x} = \lambda^{-1}\mathbf{x}$  where  $\lambda \neq 0$  and  $\mathbf{A}$  is invertible.

**proof**

$$\mathbf{A}\mathbf{x} = \lambda\mathbf{x} \iff \mathbf{A}^{-1}\mathbf{A}\mathbf{x} = \lambda\mathbf{A}^{-1}\mathbf{x} \iff \lambda^{-1}\mathbf{x} = \mathbf{A}^{-1}\mathbf{x}. \blacksquare$$

Let  $S(\mathbf{A})$  be the set of all eigenvalues of  $\mathbf{A}$ .

$$S(\mathbf{A}) \equiv \{\lambda_i | \text{for some } \mathbf{x} \in \mathbb{R}, \mathbf{A}\mathbf{x} = \lambda_i\mathbf{x}\} \quad (10)$$

$\forall \tilde{\lambda} \in S(\mathbf{A} + \lambda\mathbf{I})$  s.t.  $(\mathbf{A} + \lambda\mathbf{I})\mathbf{x} = \tilde{\lambda}\mathbf{x}$ .

By multiplying  $\mathbf{x}^T, \mathbf{x}^T(\mathbf{A} + \lambda\mathbf{I})\mathbf{x} = \mathbf{x}^T \mathbf{A} \mathbf{x} + \lambda = \tilde{\lambda} \geq \lambda$ . ( $\because \mathbf{A}$  is positive semi-definite.)

This implies that  $\min(S(\mathbf{A} + \lambda\mathbf{I})) \geq \lambda$ . Equivalently, due to the **Claim 2**, this also means that  $\max(S((\mathbf{A} + \lambda\mathbf{I})^{-1})) \leq \lambda^{-1}$ . By noticing that  $\max(S((\mathbf{A} + \lambda\mathbf{I})^{-1})) = \rho((\mathbf{A} + \lambda\mathbf{I})^{-1})$ , the proof is done. Note that for the equality,  $\mathbf{A}\mathbf{x} = \mathbf{0}$  must have nontrivial solution.

### 2.b

For both problems, the  $\mathbf{A} + \lambda\mathbf{I}$  is following.

$$\mathbf{A} + \lambda\mathbf{I} = \begin{pmatrix} 2 + \lambda & \epsilon \\ \epsilon & \epsilon^2 + \lambda \end{pmatrix} \quad (11)$$

Since (11) is invertible, one can get  $\boldsymbol{\omega}_b, \boldsymbol{\omega}_c$ .

$$\begin{aligned} \boldsymbol{\omega}_b &= \frac{1}{(1 + \lambda)\epsilon^2 + \lambda(\lambda + 2)} \cdot \begin{pmatrix} \epsilon^2 + \lambda & -\epsilon \\ -\epsilon & 2 + \lambda \end{pmatrix} \cdot \begin{pmatrix} 2 \\ \epsilon \end{pmatrix} \\ &= \frac{1}{(1 + \lambda)\epsilon^2 + \lambda(\lambda + 2)} \cdot \begin{pmatrix} \epsilon^2 + 2\lambda \\ \epsilon\lambda \end{pmatrix} \\ &= \begin{pmatrix} 0.973 \\ 0.044 \end{pmatrix} \end{aligned}$$

$$\begin{aligned}
\omega_c &= \frac{1}{(1+\lambda)\epsilon^2 + \lambda(\lambda+2)} \cdot \begin{pmatrix} \epsilon^2 + \lambda & -\epsilon \\ -\epsilon & 2 + \lambda \end{pmatrix} \cdot \begin{pmatrix} 2 + \epsilon \\ \epsilon \end{pmatrix} \\
&= \frac{1}{(1+\lambda)\epsilon^2 + \lambda(\lambda+2)} \cdot \begin{pmatrix} \epsilon^3 + \epsilon^2 + \lambda\epsilon + 2\lambda \\ -\epsilon^2 + \lambda\epsilon \end{pmatrix} \\
&= \begin{pmatrix} 1.026 \\ -0.044 \end{pmatrix}
\end{aligned}$$

Furthemore,  $\Delta\omega = \omega_c - \omega_b$  can be obtained.

$$\Delta\omega = \frac{1}{(1+\lambda)\epsilon^2 + \lambda(\lambda+2)} \begin{pmatrix} \epsilon^3 + \lambda\epsilon \\ -\epsilon^2 \end{pmatrix} = \begin{pmatrix} 0.0531 \\ -0.088 \end{pmatrix} \quad (12)$$

## 2.c

One can notice that  $\Delta\omega$  with regularization is much smaller than  $\Delta\omega$  without regularization. This implies that regularization makes the parameters less variable with respect to small noise in input data. This can be verified by the following figure.

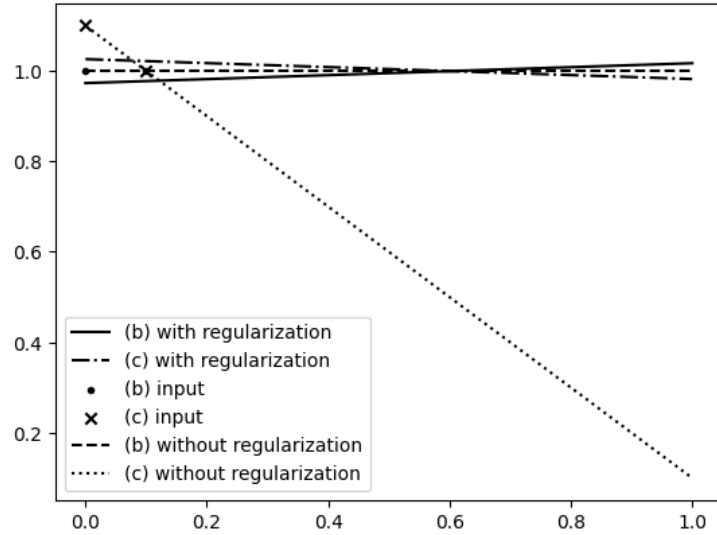


Figure 1: The linear regression result of (b) and (c)

## 3 LR with Regularization: A Probabilistic Perspective

$$\begin{aligned}
\Pr(\omega) &= \mathcal{N}(\mathbf{0}, \frac{1}{\lambda} \mathbf{I}) \\
&= \frac{\lambda^{N/2}}{2\pi^{N/2}} \exp\left(-\frac{1}{2} \omega^T \lambda \omega\right)
\end{aligned}$$

$$\begin{aligned}
t^{(i)} &= \omega^T \mathbf{x}^{(i)} + \epsilon^{(i)} \quad \& \quad \Pr(\epsilon^{(i)}) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\epsilon^{(i)2}/2\sigma^2\right) \\
\Rightarrow \Pr(t^{(i)} \mid \omega, \mathbf{x}^{(i)}) &= \frac{1}{\sqrt{(2\pi)}\sigma} \exp\left(-(t^{(i)} - \omega^T \mathbf{x}^{(i)})^2/2\sigma^2\right)
\end{aligned}$$

$$\begin{aligned}
\Pr(\mathbf{t} \mid \boldsymbol{\omega}^T \mathbf{x}^{(i)}) &= \prod_{i=1}^N \Pr(t^{(i)} \mid \boldsymbol{\omega}^T, \mathbf{x}) \\
&= \frac{1}{2\pi^{N/2}\sigma^N} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^N (t^{(i)} - \boldsymbol{\omega}^T \mathbf{x}^{(i)})^2\right)
\end{aligned}$$

From the discussion above, the posterior probability is following.

$$\Pr(\boldsymbol{\omega} \mid \mathbf{x}, \mathbf{t}) \propto \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^N (t^{(i)} - \boldsymbol{\omega}^T \mathbf{x}^{(i)})^2 - \frac{\lambda}{2} \boldsymbol{\omega}^T \boldsymbol{\omega}\right) \quad (13)$$

Minimizing (13) is equivalent to minimizing its logarithm.

$$\log(\Pr(\boldsymbol{\omega} \mid \mathbf{x}, \mathbf{t})) \quad (14)$$

## 4 Logistic Regression