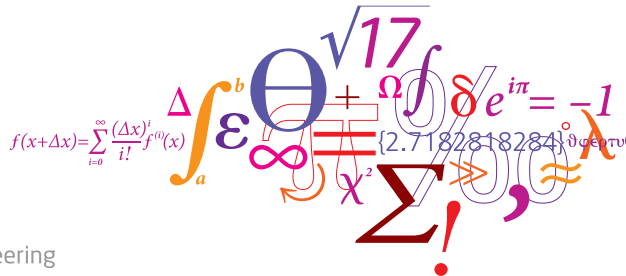


Probability and Statistics review

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Outline

- Introduction
- Random variable, atom, and event
- Joint distribution
- Conditional probability
- Bayes theorem
- Independence
- Expectation
- Continuous random variables

(Based on David MacKay, David Blei,
<https://www.cs.princeton.edu/courses/archive/spring12/cos424/pdf/lecture02.pdf>)

Teaser

Consider the “Monty Hall problem”

- There are three doors:
 - One has a car (picture)
 - Two have a goat (picture)
- ① Participant chooses one door
- ② Host (*Monty Hall*) opens another door
- ③ Host's opened door is always a goat
- Should the participant change his/her choice?



Random variable, atom, and event

- In Algebra a variable, x , is an unknown value
 - E.g. $2x = 4$
 - It can take at most one value at a time
- A **random variable** represents simultaneously a set of values
- Necessary in contexts where we *cannot* determine a unique value
 - Of course, theoretically, it also corresponds to one value...
 - But we can only determine its distribution
 - E.g. $p(5 < X < 10) = 0.5$
- It can be a single value, a vector, a matrix...

Random variable, atom, and event

- Random variables take on values in a sample space
- They can be discrete or continuous
- For example:
 - Coin flip: $\{H, T\}$
 - Height: Positive values $(0, \infty)$
 - Temperature: real values $(-\infty, \infty)$
 - Number of words in a document: Positive integers $\{1, 2, \dots, \infty\}$
- We call the values of random variables *atoms*

Random variable, atom, and event

- A *discrete probability distribution* assigns probability to every atom in the sample space
- For example, if X is an (unfair) coin, then
 - $p(X = H) = 0.7$
 - $p(X = T) = 0.3$
- The sum of probabilities of *any* distribution is 1

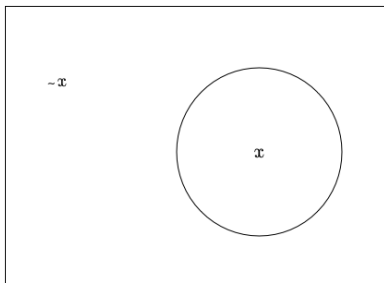
$$\sum_x p(X = x) = 1$$

- And all probabilities have to be greater or equal to 0
- Probabilities of disjunctions are sums over part of the space.
E.g., the probability that a die is bigger than 3:

$$p(X > 3) = p(X = 4) + p(X = 5) + p(X = 6)$$

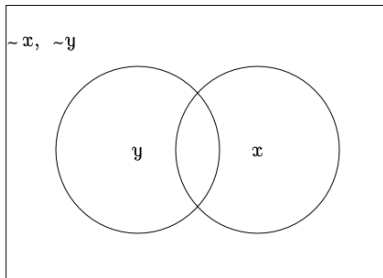
Random variable, atom, and event

- The figure below is helpful to understand these concepts well



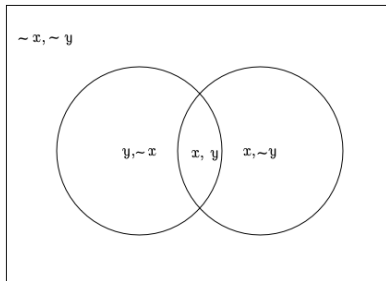
- An *atom* is a point in the box. All atoms together form the *sample space*
- An *event* is a subset of atoms. Two events in the picture are x and $\sim x$
- The probability of an event is the sum of the probabilities of its atoms

- In practice, we often combine many variables/events at the same time



- The **joint distribution** is a distribution over the configuration of all the random variables in the ensemble
 - For the figure, the function $p(X, Y)$ gives the probability of all possible combinations of X and Y
 - Notice that $X \in \{x, \sim x\}$ and $Y \in \{y, \sim y\}$
 - Therefore $X, Y \in \{(x, y), (x, \sim y), (\sim x, y), (\sim x, \sim y)\}$

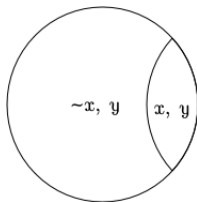
Joint distribution



- Some useful properties:
 - Union: $p(X \cup Y) = p(X) + p(Y) - p(X, Y)$
 - **Marginalization:** $p(X) = \sum_Y p(X, Y)$
This property is referred to as the **sum rule of probability!**

Conditional probability

- What about when we have observed one event, but want to know the probability of another one?
- The **conditional probability** of X given Y is the probability of event X when event Y is known



- So, we only concentrate on the subset of events where the specific value of Y occurs
- In the above figure, we focus on when $Y = y$

$$p(X|Y = y) = \frac{p(X, Y = y)}{p(Y = y)}$$

The chain rule (or product rule)

- Consider the conditional probability rule

$$p(X|Y) = \frac{p(X, Y)}{p(Y)}$$

- It allows us to derive the chain rule, which defines the joint distribution as a product of conditionals:

$$\begin{aligned} p(X, Y) &= p(X, Y) \frac{p(Y)}{p(Y)} \\ &= p(X|Y) p(Y) \end{aligned}$$

- In general, for any set of variables

$$p(X_1, X_2, \dots, X_N) = \prod_{n=1}^N p(X_n | X_1, X_2, \dots, X_{n-1})$$

- For example:

$$p(X, Y, Z) = p(X) p(Y|X) p(Z|Y, X)$$

Bayes theorem

- Using the chain rule, we can trivially say:

$$p(X|Y)p(Y) = p(Y|X)p(X)$$

which means that [**Bayes theorem**]:

$$p(X|Y) = \frac{p(Y|X)p(X)}{p(Y)}$$

- The Bayes theorem is an important foundation for Bayesian statistics, and particularly for Probabilistic Graphical Models!

Playtime!

- Open “1.Probability_Review.ipynb” in Jupyter
- Do Part 1, estimated duration 20 min

Independence

- Random variables are *independent* if knowing about X tells us nothing about Y

$$p(Y|X) = p(Y)$$

- This means that their joint distribution is

$$p(X, Y) = p(X) p(Y)$$

- A few examples:
 - Two lottery numbers that two (unacquainted) people chose. Are these two numbers independent?
 - Two persons, A, and B, start their trip in different parts of town. The transport mode for A is X and for B, it is Y . Are these two choices independent?
 - It's a rainy day. Two accidents happen on different roads of the city. Are these two, independent events?
 - The speeds in adjacent road sections x

Independence

- Example: two coins, C_1, C_2 with $p(H|C_1) = 0.6, p(H|C_2) = 0.2$
 - ① Suppose that I randomly choose a number $Z \in \{1, 2\}$ (with equal probability), and take coin C_Z
 - ② I flip it twice, with results (X_1, X_2)

Are X_1 and X_2 independent? What about if I know Z ?

Conditional independence

- X and Y are *conditionally independent* given Z

$$p(X|Y, Z) = p(X|Z)$$

- So, we can say that

$$X \perp\!\!\!\perp Y|Z \implies p(X, Y|Z) = p(X|Z) p(Y|Z)$$

- *If we know Z , then knowing about Y tells us nothing about X*

Teaser - Monty Hall

- We can now solve the Monty Hall problem

X = true location of the car

Y = door that host opened

Z = choice of participant

We want to know $p(X|Y, Z)$, from Bayes rule we have:

$$p(X|Y, Z) = \frac{p(Y|X, Z)p(X)}{p(Y|Z)} = \frac{p(Y|X, Z)p(X)}{\sum_X p(Y, X|Z)} = \frac{p(Y|X, Z)p(X)}{\sum_X p(Y|X, Z)p(X)}$$

we can also say:

$$p(X|Y, Z) \propto p(Y|X, Z)p(X)$$

We know that, $p(X)$, the prior probability of the location is:

$$p(X) = \frac{1}{3}$$

But what about $p(Y|X, Z)$?

Teaser - Monty Hall

But what about $p(Y|X, Z)$?

It is the *likelihood* of host choosing location Y , *given that* he knows X and Z .

	$Y = 1$	$Y = 2$	$Y = 3$
$X = 1, Z = 1$	0	0.5	0.5
$X = 1, Z = 2$	0	0	1
$X = 1, Z = 3$	0	1	0
$X = 2, Z = 1$	0	0	1
$X = 2, Z = 2$	0.5	0	0.5
$X = 2, Z = 3$	1	0	0
$X = 3, Z = 1$	0	1	0
$X = 3, Z = 2$	1	0	0
$X = 3, Z = 3$	0.5	0.5	0

Table: $p(Y|X, Z)$

Ok, let's try a scenario. Let's assume that the participant chose door 3 and the host opened door 2. Then our table becomes:

	$Y = 2$
$X = 1, Z = 3$	1
$X = 2, Z = 3$	0
$X = 3, Z = 3$	0.5

Table: $p(Y = 2|X, Z = 3)$

Teaser - Monty Hall

In this scenario, we want to calculate

$$p(X|Y = 2, Z = 3) \propto p(Y = 2|X, Z = 3)P(X)$$

and we have

	$Y = 2$
$X = 1, Z = 3$	1
$X = 2, Z = 3$	0
$X = 3, Z = 3$	0.5

$$P(X) = \frac{1}{3}$$

Table: $p(Y = 2|X, Z = 3)$

- Let's just calculate for the two possible cases (X is either in door 1 or 3!):

$$p(Y = 2|X = 1, Z = 3) \times \frac{1}{3} = \frac{1}{3}$$

$$p(Y = 2|X = 3, Z = 3) \times \frac{1}{3} = \frac{1}{2} * \frac{1}{3} = \frac{1}{6}$$

- So we can get our normalizing quantity: $\sum_X p(Y|X, Z)p(X) = \frac{1}{3} + \frac{1}{6} = \frac{1}{2}$

Teaser - Monty Hall

Remember: to calculate the distribution of X (i.e. “where the car probably is”), we need to calculate

$$p(X|Y, Z) = \frac{p(Y|X, Z)p(X)}{\sum_X p(Y|X, Z)p(X)}$$

- If we follow our example, we get

$$p(X|Y = 2, Z = 3) = \frac{p(Y = 2|X, Z = 3) \times \frac{1}{3}}{\frac{1}{2}}$$

- using the calculations from the previous slide, we have

$$p(X = 1|Y = 2, Z = 3) = \frac{\frac{1}{3}}{\frac{1}{2}} = \frac{2}{3}$$
$$p(X = 3|Y = 2, Z = 3) = \frac{\frac{1}{6}}{\frac{1}{2}} = \frac{1}{3}$$

- By this reasoning, we **always** have $\frac{2}{3}$ chances when we change doors, and keep $\frac{1}{3}$ if we keep it!

Playtime!

- Open “1.Probability_Review.ipynb” in Jupyter
- Do Part 2, estimated duration 30 min

Expectation

- The *expected value* of a random variable is the probability-weighted average of all possible values
- In other words, it is the *mean* of the distribution of this random variable

$$\mathbb{E}[X] = \sum_x x p(X = x)$$

- More generically (remember the $f(x)$ can be itself a random variable)

$$\mathbb{E}[f(X)] = \sum_x f(x) p(X = x)$$

Playtime!

- Open “1.Probability_Review.ipynb” in Jupyter
- Do Part 3, estimated duration 10 min

Continuous random variables

- We've only used discrete random variables so far (e.g., dice, cards)
- Random variables can be continuous
- We need a density function $p(x)$, which integrates to one.

$$\int_{-\infty}^{\infty} p(x) dx = 1$$

- Probabilities are integrals over $p(x)$
- An *event* is thus defined by an interval of possible values of the random variable

$$P(a \leq X \leq b) = \int_a^b p(x) dx$$

- Notice that we use X , x , P , and p !...

Some distributions - Gaussian

- By far, the most common one...
- Two parameters:
 - Mean, μ
 - Standard deviation, σ (or, variance, σ^2)
- $p(x)$ is defined as

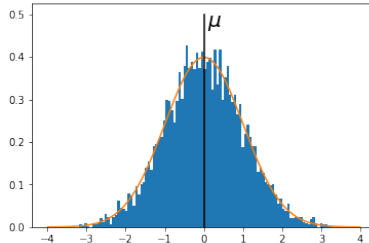
$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)}$$

- Often represented as:

$$p(x) \sim \mathcal{N}(\mu, \sigma^2)$$

Some distributions - Gaussian

- Support is $] - \infty, \infty[$
- Symmetrical



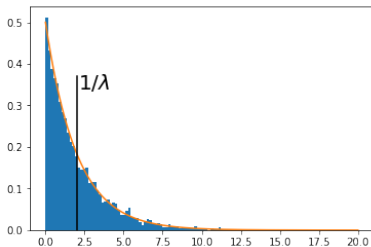
- The Central limit theorem (CLT) establishes that *the distribution of the sampling means approaches a normal distribution as the sample size gets larger, no matter what the shape of the population distribution.*

Some distributions - Exponential

- Exponential distribution, with *rate* λ

$$p(x) = \lambda e^{-\lambda x}$$

- Support is $[0, \infty[$

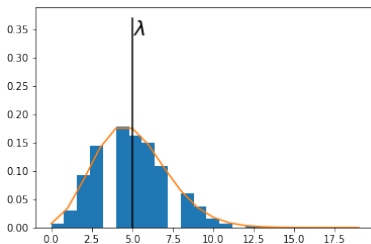


Some distributions - Poisson

- Poisson distribution, with *rate* λ

$$p(k) = \frac{\lambda^k e^{-\lambda}}{k!}$$

- for $k = 0, 1, 2, \dots$
- Pretty common in transportation (e.g. arrival rates)



1

¹ In fact, this distribution relates to a discrete random variable, so we include it to emphasize that not only continuous variables can be parameterized as a probability distribution.

Independent and identically distributed random variables (iid)

- Independent
- Identically distributed

If we repeatedly flip the same coin N times and record the outcome, then X_1, \dots, X_N are **iid**

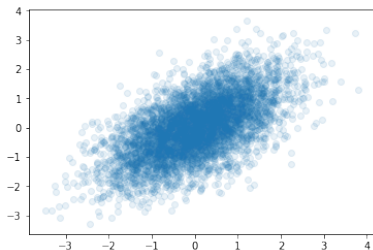
- The iid assumption can be extremely useful in data analysis

Multivariate distributions

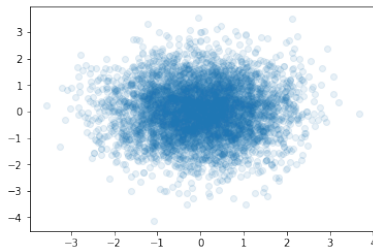
- So far, we've been working with single variable distributions
- Multivariate means it's the same as above, but with more variables at the same time!
- In practice, joint distribution of variables that share a common structure
- In some cases (e.g. Poisson), it is not a trivial problem
- In others (e.g. Gaussian), it is well studied, and extensively applied

$$p(\mathbf{x}) = \frac{1}{\sqrt{2\pi}|\Sigma|} e^{-\frac{1}{2}(\mathbf{x}-\mu)^T \Sigma^{-1}(\mathbf{x}-\mu)}$$

- Bivariate Gaussian



$$\Sigma = \begin{bmatrix} 1 & 0.6 \\ 0.6 & 1 \end{bmatrix}$$



$$\Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

Playtime!

- Open “2. Probability_Review.ipynb”
- Do part 1. Est. time is 15 min

A note on notation

- So far we have been using a rather standard statistics notation
 - X is a random variable and x is atom/event
 - We write e.g. $p(X = x)$
- In the machine learning literature, this notation is typically simplified
 - Lowercase letters, such as x , represent random variables
 - We simply write $p(x)$. Everything else should be clear from the context!
- This allows us to have
 - Bold letters denote vectors (e.g. \mathbf{x} , where the i^{th} element is referred as x_i)
 - Matrices are represented by bold uppercase letters such as \mathbf{X}
 - Roman letters, such as N , denote constants
- This is the notation that we will adopt from now on!

The likelihood function

- Imagine you have the data. For example:
 - N readings of traffic counts at a certain time, each one called x_i , $i = 1 \dots N$
- You assume it follows some parametric distribution (e.g. Gaussian)
- How do you determine its parameters, Θ ?
- The likelihood function, $L(\Theta)$, should be:

$$L(\Theta) = \prod_i^N p(x_i | \Theta)$$

- Notice that this is the joint distribution of all **independent** data points!
- In the case of the Gaussian, we should have $\Theta = \{\mu, \sigma\}$
- The likelihood function, $L(\Theta)$, would be

$$L(\Theta) = \prod_i^N \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}}$$

The likelihood function

- In the case of the Gaussian, we should have $\Theta = \{\mu, \sigma\}$
- The likelihood function, $L(\Theta)$, would be

$$L(\Theta) = \prod_i^N \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}}$$

- If you actually *had* the true parameters, the likelihood function would have the maximum value, right?
- So, this becomes an optimization problem:
 - Find the values of Θ that maximize the function $L(\Theta)$

The log-likelihood function

- For practical reasons, we apply a logarithmic transformation to the likelihood function
 - Less prone to numeric error (numerical stability)
 - Computationally faster
- In the case of the Gaussian distribution, the log likelihood becomes:

$$-\frac{N}{2}(\log(2\pi) + \log(\sigma^2)) - \frac{1}{2\sigma^2} \sum_i (x_i - \mu)^2$$

Maximum likelihood estimate (MLE)

- The maximum likelihood estimate is the value of the parameter that maximizes the log likelihood (equivalently, the likelihood)
- In the case of the Gaussian, the MLE corresponds to:

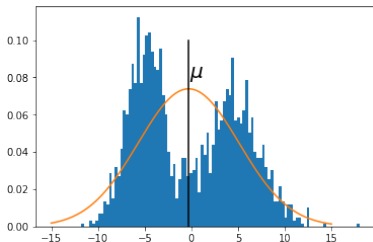
$$\hat{\mu} = \frac{\sum_{i=1}^N x_i}{N}, \quad \text{i.e. the *sample mean*}$$

$$\hat{\sigma}^2 = \frac{\sum_{i=1}^N (x_i - \hat{\mu})^2}{N}, \quad \text{i.e. the *sample variance*}$$

Maximum likelihood estimate (MLE)

DISCLAIMER:

- The fact that you get a MLE doesn't mean you found a good model!



- You need to know your data...

Playtime!

- Open “2. Probability_Review.ipynb”
- Do part 2. Est. time is 30 min