

# Least Absolute Shrinkage and Selection Operator (LASSO)

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# Outline

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- LASSO vs. Ridge

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# Motivation and Intuition

- Variable Selection

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# Motivation and Intuition

- Variable Selection
- Too Many Covariates in the Model
- Shrinkage Method to Push  $\beta$ 's towards zero
- But Why?

- The Lasso coefficients  $\hat{\beta}_\lambda^L$  minimizes the following equation:

$$\sum_{i=1}^n (y_i - \beta_0 - \sum_{j=1}^P \beta_j x_{ij})^2 + \lambda \sum_{j=1}^P |\beta_j|$$

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- $L_1$  Norm Penalty

# LASSO Estimation

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- $L_1$  Norm Penalty
- Ridge Regression is  $L_2$  Norm  $\rightarrow \beta^2$

# LASSO vs. Ridge

## LASSO

$$\sum_{i=1}^n (y_i - \beta_0 - \sum_{j=1}^P \beta_j x_{ij})^2 + \lambda \sum_{j=1}^P |\beta_j|$$

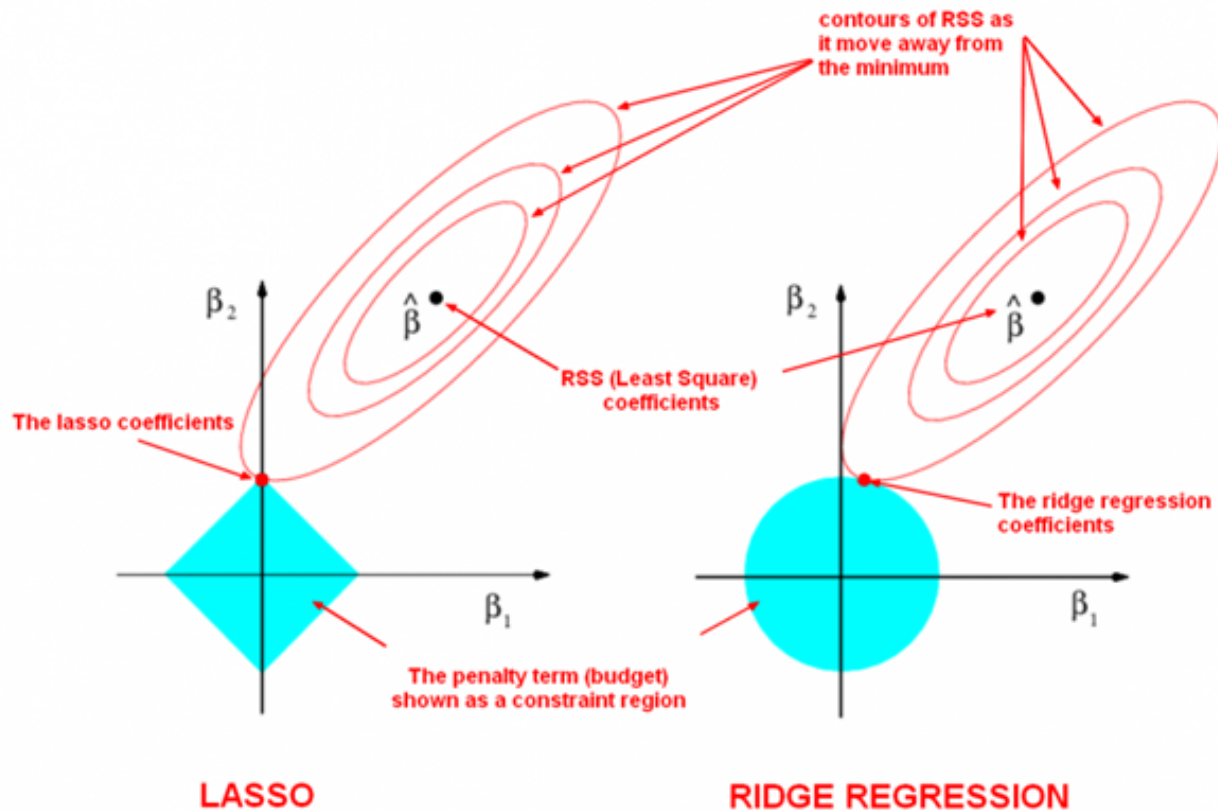
$$RSS + \lambda \sum_{j=1}^P |\beta_j|$$

## Ridge Regression

$$\sum_{i=1}^n (y_i - \beta_0 - \sum_{j=1}^P \beta_j x_{ij})^2 + \lambda \sum_{j=1}^P \beta_j^2$$

$$RSS + \lambda \sum_{j=1}^P \beta_j^2$$

# LASSO vs. Ridge

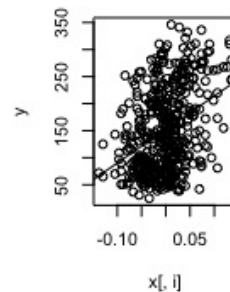
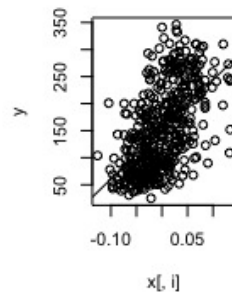
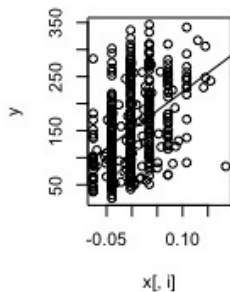
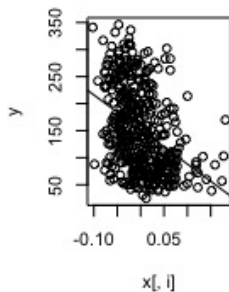
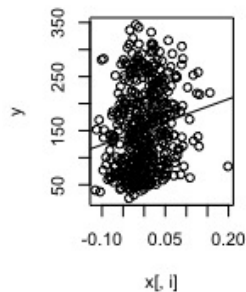
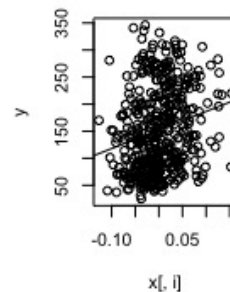
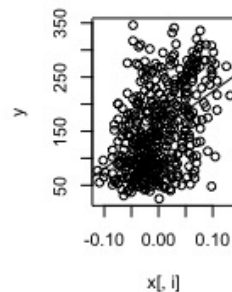
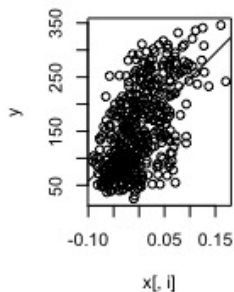
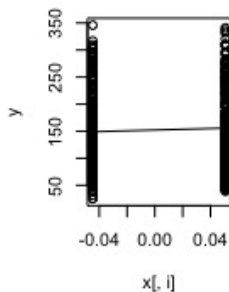
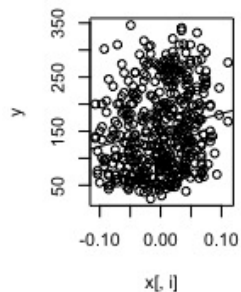


# Code and Example

```
library(lars)  
library(glmnet)  
data(diabetes)
```

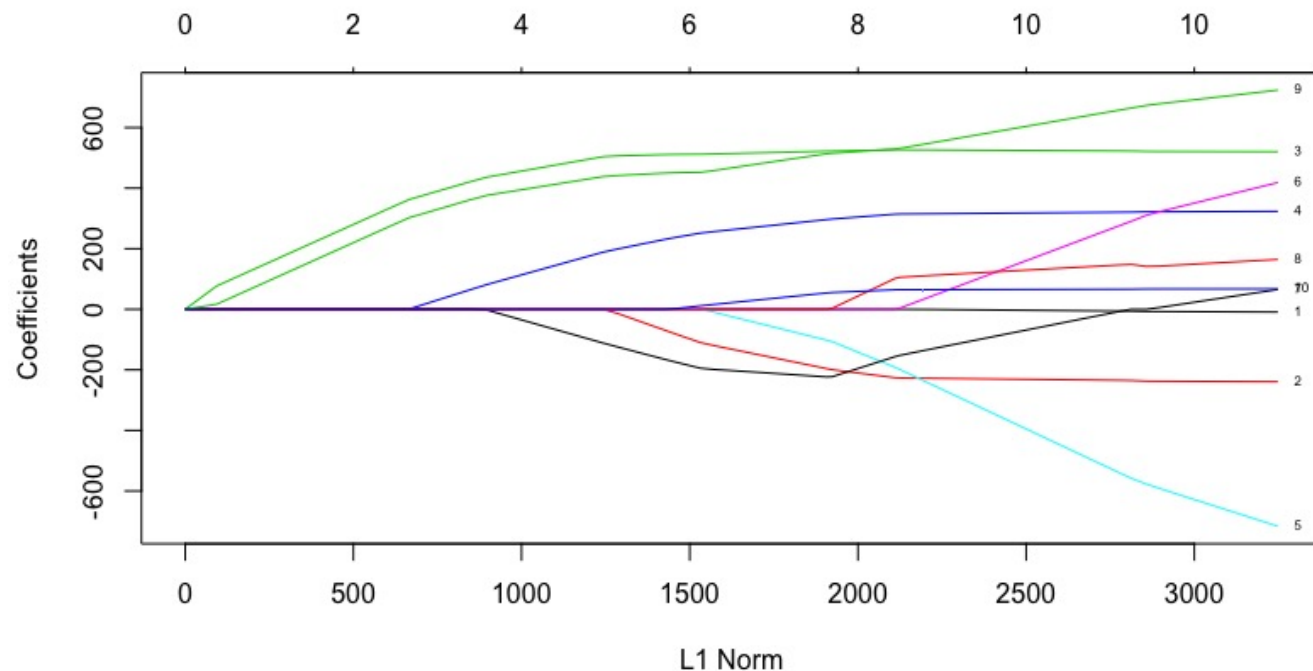
- age
- sex
- bmi
- map
- tc
- ldl
- hdl
- tch
- ltg
- glu

```
for(i in 1:10){
  plot(x[,i], y)
  abline(lm(y~x[,i]))
}
```



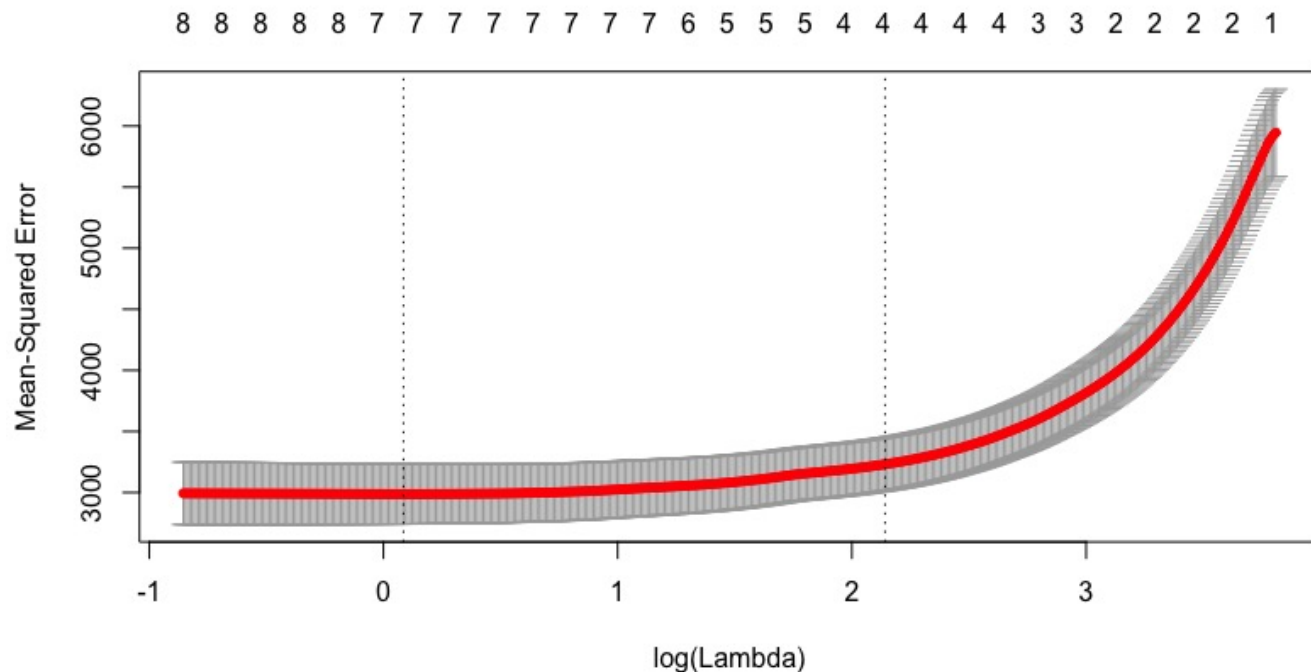
# Code and Example

```
model_lasso <- glmnet(x, y)
plot.glmnet(model_lasso, xvar = "norm", label = TRUE)
```



# Code and Example

```
cv_fit <- cv.glmnet(x=x, y=y, alpha = 1, nlambda = 1000)
plot.cv.glmnet(cv_fit)
cv_fit$lambda.min # 1.069371
```





# Code and Example (OLS)

```
summary(model_ols)
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )	
(Intercept)	152.133	2.576	59.061	< 2e-16	***
xage	-10.012	59.749	-0.168	0.867000	
xsex	-239.819	61.222	-3.917	0.000104	***
xbmi	519.840	66.534	7.813	4.30e-14	***
xmap	324.390	65.422	4.958	1.02e-06	***
xtc	-792.184	416.684	-1.901	0.057947	.
xldl	476.746	339.035	1.406	0.160389	
xhdl	101.045	212.533	0.475	0.634721	
xtch	177.064	161.476	1.097	0.273456	
xltg	751.279	171.902	4.370	1.56e-05	***
xglu	67.625	65.984	1.025	0.305998	

---

Residual standard error: 54.15 on 431 degrees of freedom

Multiple R-squared: 0.5177, Adjusted R-squared: 0.5066

F-statistic: 46.27 on 10 and 431 DF, p-value: < 2.2e-16

# Code and Example (LASSO)

```
fit <- glmnet(x=x, y=y, alpha = 1, lambda=cv_fit$lambda.min)
fit$beta
10 x 1 sparse Matrix of class "dgCMatrix"
      s0
age      .
sex -193.51972
bmi  521.56400
map  294.92639
tc   -98.53276
ldl    .
hdl -222.65684
tch    .
ltg  511.53755
glu   52.60584
```

# Why the LASSO is Uniquely Important: Causal Estimates

## Lasso adjustments of treatment effect estimates in randomized experiments

Adam Bloniarz<sup>a,1</sup>, Hanzhong Liu<sup>a,1</sup>, Cun-Hui Zhang<sup>b</sup>, Jasjeet S. Sekhon<sup>a,c</sup>, and Bin Yu<sup>a,d,2</sup>

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$$\hat{\sigma}_{e^{(a)}}^2 = \frac{1}{n_A - df^{(a)}} \sum_{i \in A} \left( a_i - \bar{a}_A - (\mathbf{x}_i - \bar{\mathbf{x}}_A)^T \hat{\boldsymbol{\beta}}_{\text{Lasso}}^{(a)} \right)^2,$$

$$\hat{\sigma}_{e^{(b)}}^2 = \frac{1}{n_B - df^{(b)}} \sum_{i \in B} \left( b_i - \bar{b}_B - (\mathbf{x}_i - \bar{\mathbf{x}}_B)^T \hat{\boldsymbol{\beta}}_{\text{Lasso}}^{(b)} \right)^2,$$

where  $df^{(a)}$  and  $df^{(b)}$  are degrees of freedom defined by the following:

$$df^{(a)} = \hat{s}^{(a)} + 1 = \left\| \hat{\boldsymbol{\beta}}_{\text{Lasso}}^{(a)} \right\|_0 + 1; \quad df^{(b)} = \hat{s}^{(b)} + 1 = \left\| \hat{\boldsymbol{\beta}}_{\text{Lasso}}^{(b)} \right\|_0 + 1.$$

Define the variance estimate of  $\sqrt{n}(\widehat{ATE}_{\text{Lasso}} - ATE)$  as follows:

$$\hat{\sigma}_{\text{Lasso}}^2 = \frac{n}{n_A} \hat{\sigma}_{e^{(a)}}^2 + \frac{n}{n_B} \hat{\sigma}_{e^{(b)}}^2. \quad [20]$$

Thanks!