

# HW4

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## Homework 4

### 3.5.1

1

$f(x;t) = 1/(pi(1 + (x - t)^2)) \rightarrow L(t) = \text{product}(f(x;t)) = pi^{-n} - \text{Product}_i^n(1 + (t - x)^2)$ . Thus,  $\log(L(t)) = l(t) = -n\ln(pi) - \text{Sum}(\ln(1 + (t - x)^2))$ . To obtain the first derivative, the  $-n\ln(pi)$  is removed as it equals 0, and  $\ln(1 + (t-x)^2) \rightarrow (t-x)/(1 + (t-x)^2)$  via the power and chain rules. This can be done again to obtain the second derivative. The information can then be found by evaluating  $I(t) = -E(l''(t)|t)$ .

2

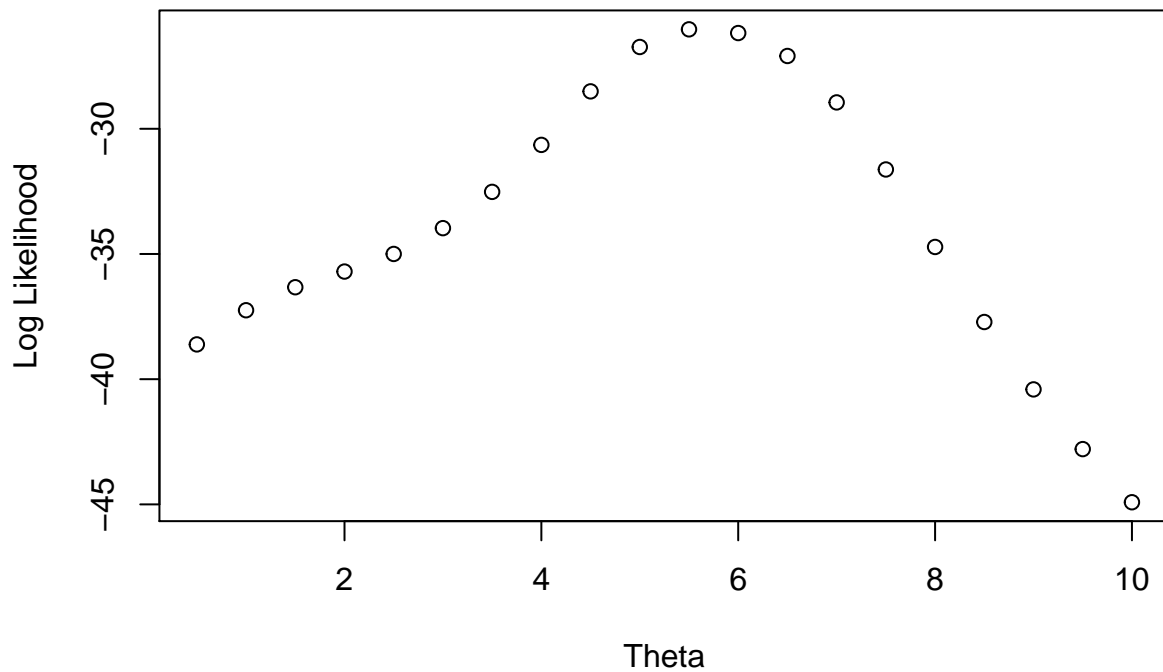
```
set.seed(909)

sampl<-rcauchy(10,location=5, scale=1) #scale is 1 for this problem

loglikCauchy <- function(x,theta, n=10){
  #x is a vector of values
  part <- sum(log(1+(theta - x)^2))
  -n*log(pi) - part
}
theta.set <- seq(0.5,10, by=.5)
y <- rep(0,20)

for(i in 1:20){
  y[i] <- loglikCauchy(sampl, i/2)
}
theta.set <- c(theta.set)
plot(theta.set, y, xlab= "Theta", ylab = "Log Likelihood", main= "Theta versus LogLikelihood")
```

## Theta versus LogLikelihood



3

```
#loglik function
loglikCauchy <- function(x){
  part <- sum(log(1+(x - sampl)^2))
  -10*log(pi) - part
}
#first derivative of loglik
l2.c <- function(t){
  -2*sum( (t-sampl)/ (1 + (t -sampl) ^2) )
}

#root via newton-taphson
uniroot(l2.c, c(-10,30))
```

```
## $root
## [1] 5.6357152455319985
##
## $f.root
## [1] -3.5072528556390203e-07
##
## $iter
## [1] 10
##
## $init.it
## [1] NA
```

```
##
## $estim.prec
## [1] 6.4550742709279518e-05
```

It appears that the MLE of theta is about 5.5.

4

To improve the process, we can attempt to use  $M = \text{loglikelihood's second derivative for the function } x_{t+1} = x_t - M_t^{-1}l'(x)$ .

```
# second derivative of loglik
l3 <- function(x){
  -2*sum( ( 1 -(x-sampl)^2) / ((1+(x-sampl)^2)^2))
}

#process
x=-10
for(i in 1:1000){
  x <- x + (l2.c(x))/(l3(x))
}
x
```

```
## [1] 5.2673645482712184
```

The results for this method do seem rather accurate, as it estimate the MLE to be theta.

5

The results of a univariate fixed-point iteration are shown below.

```
#Univariate Options
g1 <-function(t){
  1*l2.c(t) + t
}
g2 <-function(t){
  .64*l2.c(t) + t
}
g3 <-function(t){
  .25*l2.c(t) + t
}

#a=1
t=5
for(i in 1:100){
  t<-g1(t)
}
t
```

```
## [1] 5.3125345381118567
```

```
#a=.64
t=5
for(i in 1:100){
  t<-g2(t)
}
t
```

```
## [1] 3.8985495176085956
```

```
#a=.25
t=5
for(i in 1:100){
  t<-g3(t)
}
t
```

```
## [1] 5.6357151165410073
```

The iterative methods seem efficient at  $\alpha=1$ , and  $\alpha=.25$ , as both appear to converge to the true MLE for  $\theta$ . For  $\alpha=.64$ , it appears that the iteration is not convergent to the MLE.

For a fixed-point Newton-Raphson method, one can use  $M = \text{loglikelihood's second derivative of a good guess of the MLE (or the initial value)}$ , such as  $M = l''(5)$ . The results are given below.

```
x=-10
M = l3(5)
for(i in 1:100){
  x <- x - (l2.c(x))/M
}
x
```

```
## [1] 5.6357151165410064
```

This method provides an accurate estimate of the MLE, as seen in the earlier evaluations.

## 6

To use Fisher Scoring, we will use the form  $M = I_n = n/2$

```
#MLE via fisher scoring
x=-10
M= -(10/2)
for(i in 1:100){
  x <- x - (l2.c(x))/M
}
x
```

```
## [1] 5.6357151165410064
```

```
#MLE further refined via Newton-Raphson without hessian using M=l''(x)
for(i in 1:100){
  x <- x - (l2.c(x))/(l3(x))
}
x
```

```
## [1] 5.6357151165410073
```

The final estimate is 5.63571511654101

## 7

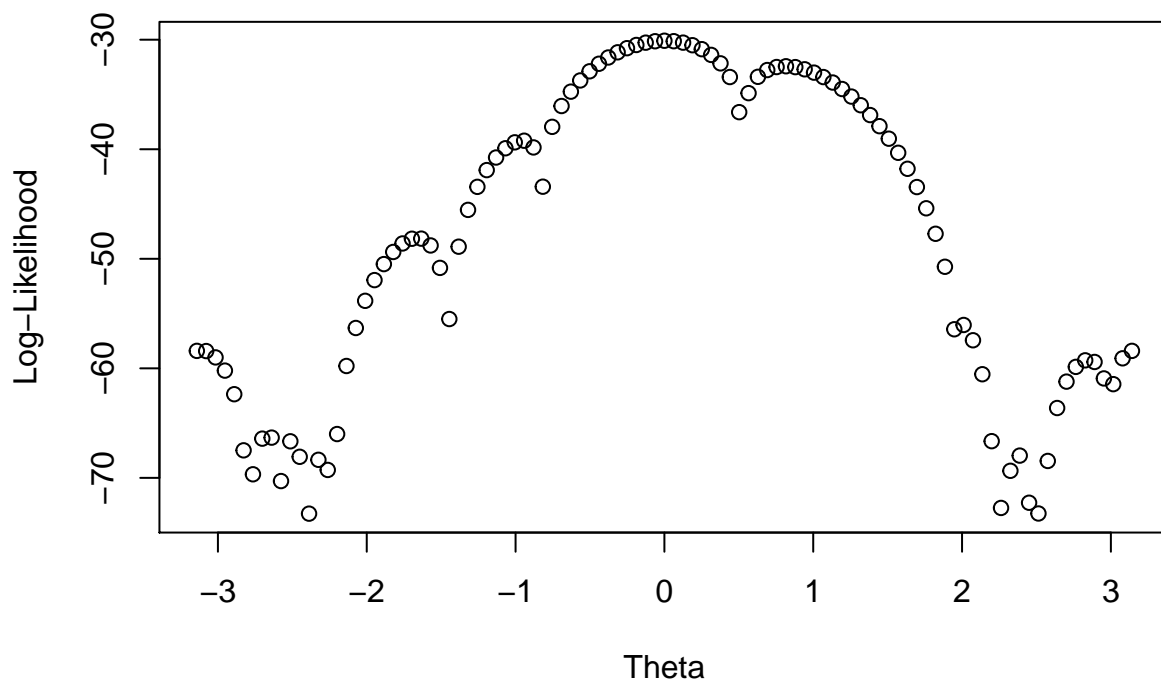
Many of the methods used above find similar results. For accuracy, the Fisher scoring followed by Newton-Raphson refining is likely the best, as it refined the estimate via two methods. This method takes the longest though, as it takes the evaluation of the information. The multivariate fixed-point iteration method appeared to be the easiest method to perform, as it only required a rough estimate of the MLE to provide a very accurate estimate (virtually the same as is seen in the Fisher scoring version).

It is also worth noting that the method used in part 4 of this question provided the result closest to 5 (the original parameter used).

```
x <- c(3.91, 4.85, 2.28, 4.06, 3.70, 4.04, 5.46, 3.53, 2.28, 1.96, 2.53, 3.88, 2.22, 3.47, 4.82, 2.46, 2.53)
n = length(x)
fn <- function(t){
  (1-cos(x-t))/(2*pi)
}
```

The likelihood function would be  $L(t) = \prod_{i=1}^n [1 - \cos(x-t)] / (2\pi)^n$ . Taking the log of this gives us  $l(t) = \sum (\log(1 - \cos(x-t))) - n \ln(2\pi)$

```
plot(set, lly, xlab = "Theta", ylab="Log-Likelihood", )
```



## 2

Note: t is the short hand for theta.

$E(x|t) = \text{integral}(x * (1 - \cos(x - t)) / (2\pi)) = -x((2\sin(x - t) - x) + 2\cos(x - t)) / 4\pi$  from  $2\pi$  to  $0 = \sin(t) + \pi = \text{Xbar}$ . From here we can further say that  $\sin(t) = \text{xbar} - \pi$ , and then  $\theta^{\text{hat}} = \arcsin(\text{xbar} - \pi)$ . Because we can calculate xbar from the sample, the MOM estimate of Theta is given as:

```
xbar <- mean(x)
t.mom = asin(xbar - pi)
t.mom
```

```
## [1] 0.095394067305843627
```

Note: this value does appear to be close to the max likelihood theta from the graph in part 1.

## 3

```
l11 <-function(t){
  sum(((sin(x-t)) / (1- cos(x-t))))
}

l12 <- function(t){
  sum( ( -(sin(x-t)^2) +(cos(x-t))^2 -(cos(x-t)))/((cos(x-t)-1)^2))
}
#newton-raphson with fixed point l''(tmom) by hand
tnew <- t.mom
for (i in 1:20){
  tnew <- tnew - ((l11(tnew))/ l12(tnew))
}
tnew
```

```
## [1] 0.51983029708841644
```

```
#using general optim
optim(t.mom,l1)$par
```

```
## [1] 0.52000001809650964
```

The MLE appears to be around .52, and varies depending on methods.

## 4

```
#at -2.7
optim(-2.7, l1)$par
```

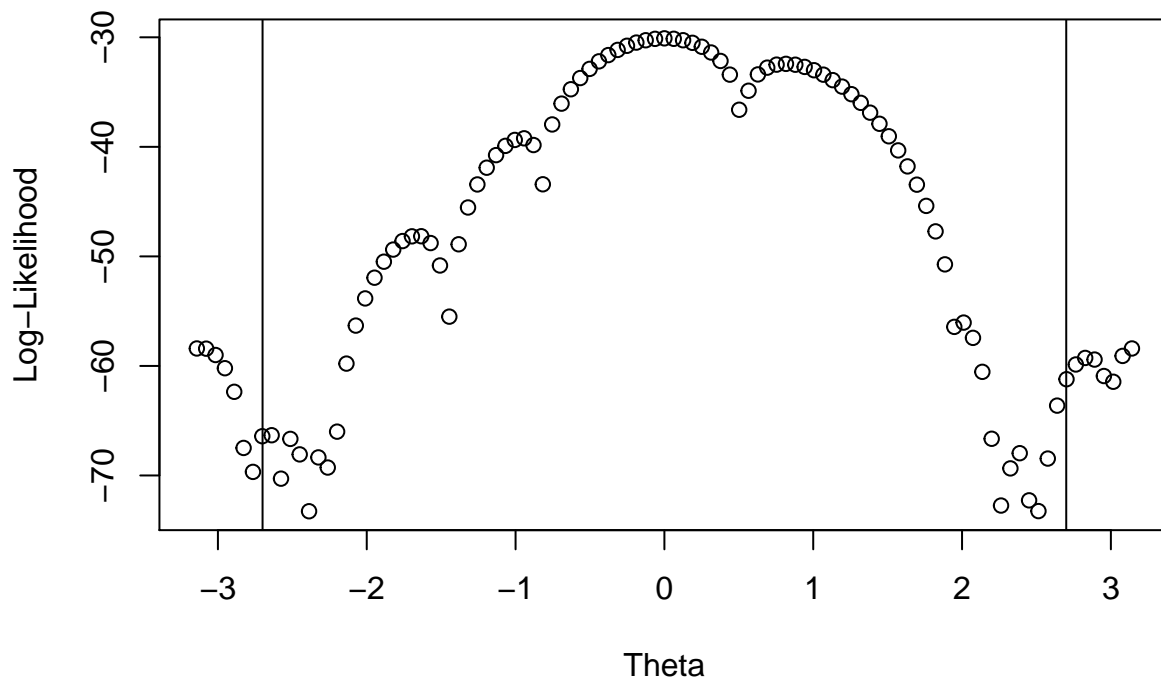
```
## [1] -2.3731852930784232
```

```
#at 2.7
optim(2.7,l1)$par
```

```
## [1] 2.9899999840557578
```

Both estimates are vastly different from the MLE estimate obtained in part 3. This is due to the existence of many local maximums. This can be seen on the graph, this time provided with verticals lines at both 2.7 and -2.7, which have local maximums closer to them then the ultimate maximum near  $t=0.1$ .

```
plot(set, lly, xlab = "Theta", ylab="Log-Likelihood", )
abline(v=2.7)
abline(v=-2.7)
```



5

The estimated zero for each given starting point is given in the following table as “output”. These outputs were then placed into a group corresponding with the output, and placed within a group with inputs that produced the same output. The levels (possible zeroes), are given in the output below, as well as the full print of each input/output and the corresponding group they received. Group is numbered with 1 corresponding to the lowest output (-2.814), and the next smallest being group 2, and so on.

```
range <- seq(-pi, pi, length.out= 200)
range <- c(range)

input <- range
output<- rep(0,200)
for(i in 1:200){
  temp <- input[i]
  output[i] <- optim(temp,ll)$par
}
dat <- cbind(input,output)
dat <- as.data.frame(dat)
options(digits=5)
dat$round <- round(dat$output,digits=3)
dat$group <- factor(dat$round)
levels(dat$group)
```

```
## [1] "-2.813" "-2.753" "-2.583" "-2.463" "-2.403" "-2.373" "-2.243" "-2.223"
## [9] "-1.463" "-1.433" "-0.823" "-0.821" "0.52" "2.22" "2.28" "2.46"
## [17] "2.53" "2.54" "2.99" "3.53" "3.7" "3.88" "3.91"
```

```

levels(dat$group) <- 1:23
fin<-dat[,c(1,2,4)]
fin

```

##	input	output	group
## 1	-3.141593	-2.75319	2
## 2	-3.110019	-2.75319	2
## 3	-3.078445	-2.81319	1
## 4	-3.046871	-2.58319	3
## 5	-3.015297	-2.37319	6
## 6	-2.983724	-2.40319	5
## 7	-2.952150	-2.40319	5
## 8	-2.920576	-2.37319	6
## 9	-2.889002	-2.81319	1
## 10	-2.857428	-2.37319	6
## 11	-2.825855	-2.75319	2
## 12	-2.794281	-2.75319	2
## 13	-2.762707	-2.75319	2
## 14	-2.731133	-2.58319	3
## 15	-2.699560	-2.37319	6
## 16	-2.667986	-2.40319	5
## 17	-2.636412	-2.37319	6
## 18	-2.604838	-2.37319	6
## 19	-2.573264	-2.58319	3
## 20	-2.541691	-2.40319	5
## 21	-2.510117	-2.40319	5
## 22	-2.478543	-2.24319	7
## 23	-2.446969	-2.40319	5
## 24	-2.415395	-2.40319	5
## 25	-2.383822	-2.40319	5
## 26	-2.352248	-2.58319	3
## 27	-2.320674	-2.37319	6
## 28	-2.289100	-2.40319	5
## 29	-2.257526	-2.37319	6
## 30	-2.225953	-2.24319	7
## 31	-2.194379	-2.40319	5
## 32	-2.162805	-2.40319	5
## 33	-2.131231	-2.22319	8
## 34	-2.099657	-2.37319	6
## 35	-2.068084	-2.40319	5
## 36	-2.036510	-2.24319	7
## 37	-2.004936	-2.40319	5
## 38	-1.973362	-2.37319	6
## 39	-1.941788	-2.40319	5
## 40	-1.910215	-2.40319	5
## 41	-1.878641	-2.24319	7
## 42	-1.847067	-2.40319	5
## 43	-1.815493	-2.40319	5
## 44	-1.783919	-2.40319	5
## 45	-1.752346	-1.46319	9
## 46	-1.720772	-1.46319	9
## 47	-1.689198	-1.43319	10
## 48	-1.657624	-1.43319	10
## 49	-1.626050	-1.46319	9



## 50	-1.594477	-1.43319	10
## 51	-1.562903	-1.46319	9
## 52	-1.531329	-1.43319	10
## 53	-1.499755	-1.46319	9
## 54	-1.468181	-1.43319	10
## 55	-1.436608	-1.43319	10
## 56	-1.405034	-1.46319	9
## 57	-1.373460	-1.46319	9
## 58	-1.341886	-1.46319	9
## 59	-1.310313	-1.46319	9
## 60	-1.278739	-1.43319	10
## 61	-1.247165	-1.43319	10
## 62	-1.215591	-1.46319	9
## 63	-1.184017	-1.43319	10
## 64	-1.152444	-1.46319	9
## 65	-1.120870	-2.24319	7
## 66	-1.089296	-1.46319	9
## 67	-1.057722	-1.46319	9
## 68	-1.026148	-1.46319	9
## 69	-0.994575	-0.82319	11
## 70	-0.963001	-0.82319	11
## 71	-0.931427	-0.82319	11
## 72	-0.899853	-0.82319	11
## 73	-0.868279	-0.82319	11
## 74	-0.836706	-0.82319	11
## 75	-0.805132	-0.82319	11
## 76	-0.773558	-0.82319	11
## 77	-0.741984	-0.82319	11
## 78	-0.710410	-0.82319	11
## 79	-0.678837	-0.82319	11
## 80	-0.647263	-2.58319	3
## 81	-0.615689	-0.82319	11
## 82	-0.584115	-0.82319	11
## 83	-0.552541	-2.40319	5
## 84	-0.520968	-0.82319	11
## 85	-0.489394	-0.82319	11
## 86	-0.457820	-1.46319	9
## 87	-0.426246	-0.82319	11
## 88	-0.394672	-0.82319	11
## 89	-0.363099	-1.43319	10
## 90	-0.331525	-0.82319	11
## 91	-0.299951	-2.40319	5
## 92	-0.268377	-0.82319	11
## 93	-0.236803	-2.37319	6
## 94	-0.205230	-0.82319	11
## 95	-0.173656	-2.22319	8
## 96	-0.142082	-0.82319	11
## 97	-0.110508	-2.40319	5
## 98	-0.078934	-0.82092	12
## 99	-0.047361	-2.46276	4
## 100	-0.015787	-0.82092	12
## 101	0.015787	0.52000	13
## 102	0.047361	0.52000	13
## 103	0.078934	0.52000	13

## 104	0.110508	0.52000	13
## 105	0.142082	0.52000	13
## 106	0.173656	0.52000	13
## 107	0.205230	0.52000	13
## 108	0.236803	0.52000	13
## 109	0.268377	0.52000	13
## 110	0.299951	0.52000	13
## 111	0.331525	0.52000	13
## 112	0.363099	0.52000	13
## 113	0.394672	0.52000	13
## 114	0.426246	0.52000	13
## 115	0.457820	0.52000	13
## 116	0.489394	0.52000	13
## 117	0.520968	0.52000	13
## 118	0.552541	0.52000	13
## 119	0.584115	0.52000	13
## 120	0.615689	0.52000	13
## 121	0.647263	0.52000	13
## 122	0.678837	0.52000	13
## 123	0.710410	0.52000	13
## 124	0.741984	0.52000	13
## 125	0.773558	0.52000	13
## 126	0.805132	2.99000	19
## 127	0.836706	3.91000	23
## 128	0.868279	2.54000	18
## 129	0.899853	2.28000	15
## 130	0.931427	2.54000	18
## 131	0.963001	2.28000	15
## 132	0.994575	2.28000	15
## 133	1.026148	2.46000	16
## 134	1.057722	2.28000	15
## 135	1.089296	2.28000	15
## 136	1.120870	2.53000	17
## 137	1.152444	2.28000	15
## 138	1.184017	2.53000	17
## 139	1.215591	2.53000	17
## 140	1.247165	2.54000	18
## 141	1.278739	2.28000	15
## 142	1.310313	2.53000	17
## 143	1.341886	2.28000	15
## 144	1.373460	2.54000	18
## 145	1.405034	2.53000	17
## 146	1.436608	2.28000	15
## 147	1.468181	2.53000	17
## 148	1.499755	2.54000	18
## 149	1.531329	2.22000	14
## 150	1.562903	2.46000	16
## 151	1.594477	2.53000	17
## 152	1.626050	2.28000	15
## 153	1.657624	2.54000	18
## 154	1.689198	2.54000	18
## 155	1.720772	2.28000	15
## 156	1.752346	2.28000	15
## 157	1.783919	2.46000	16

## 158	1.815493	2.53000	17
## 159	1.847067	2.28000	15
## 160	1.878641	2.28000	15
## 161	1.910215	2.28000	15
## 162	1.941788	2.53000	17
## 163	1.973362	2.54000	18
## 164	2.004936	2.28000	15
## 165	2.036510	2.54000	18
## 166	2.068084	2.28000	15
## 167	2.099657	2.54000	18
## 168	2.131231	2.46000	16
## 169	2.162805	2.54000	18
## 170	2.194379	2.28000	15
## 171	2.225953	2.54000	18
## 172	2.257526	2.28000	15
## 173	2.289100	2.28000	15
## 174	2.320674	2.53000	17
## 175	2.352248	2.53000	17
## 176	2.383822	2.46000	16
## 177	2.415395	2.28000	15
## 178	2.446969	2.53000	17
## 179	2.478543	2.54000	18
## 180	2.510117	2.54000	18
## 181	2.541691	2.54000	18
## 182	2.573264	2.54000	18
## 183	2.604838	2.54000	18
## 184	2.636412	2.46000	16
## 185	2.667986	2.53000	17
## 186	2.699560	2.99000	19
## 187	2.731133	2.99000	19
## 188	2.762707	2.99000	19
## 189	2.794281	2.54000	18
## 190	2.825855	2.54000	18
## 191	2.857428	2.28000	15
## 192	2.889002	2.53000	17
## 193	2.920576	2.53000	17
## 194	2.952150	2.53000	17
## 195	2.983724	2.99000	19
## 196	3.015297	2.99000	19
## 197	3.046871	3.88000	22
## 198	3.078445	3.70000	21
## 199	3.110019	3.91000	23
## 200	3.141593	3.53000	20