

Matching to Produce Causal Estimates in Non-experimental Settings

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Motivating Example

- Examine the effect of job training on unemployment
- Let z be treatment status
 - ▶ Job training ($z = 1$)
 - ▶ No job training ($z = 0$)
- Let r^{obs} be the observed outcome
 - ▶ Unemployed ($r^{obs} = 1$)
 - ▶ Employed ($r^{obs} = 0$).
- Let x be the vector of observed covariates

Potential Outcomes

An individual's possible outcomes given different treatments



A *causal effect* compares these two potential outcomes

Potential Outcomes

Only one potential outcome is ever observed

- "Missing data" problem
- Solution: look at causal effects on a population level
- Compare outcomes between *groups* of treated and non-treated units

The Estimand: the Odds Ratio

- Compares odds of unemployment in the job training group to that in the no job training group:

$$\gamma = \frac{\frac{\overline{r_1^{obs}}}{1 - \overline{r_1^{obs}}}}{\frac{\overline{r_0^{obs}}}{1 - \overline{r_0^{obs}}}}$$

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- 40 individuals: 20 treated, 20 non-treated
 - ▶ Among treated: 5 cases
 - ▶ Among non-treated: 10 cases

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$$\gamma = \frac{\frac{0.25}{0.75}}{\frac{0.5}{0.5}} = \frac{1}{3}$$

The Assumption: Covariate Balance

- Treated and non-treated groups must "look alike"
- Assumption not met \Rightarrow covariate imbalance
 - ▶ Covariate distributions differ
- Depends on treatment assignment mechanism

Study Designs: Randomized Experiment

- Random treatment assignment
- Covariate balance by design

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Non-treated (No job training)



Treated (Job training)

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- Treatment assignment is not random
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Study Designs: Retrospective Study

- Subjects sampled based on outcomes
- Oversample "case" individuals, then sample "non-case" individuals
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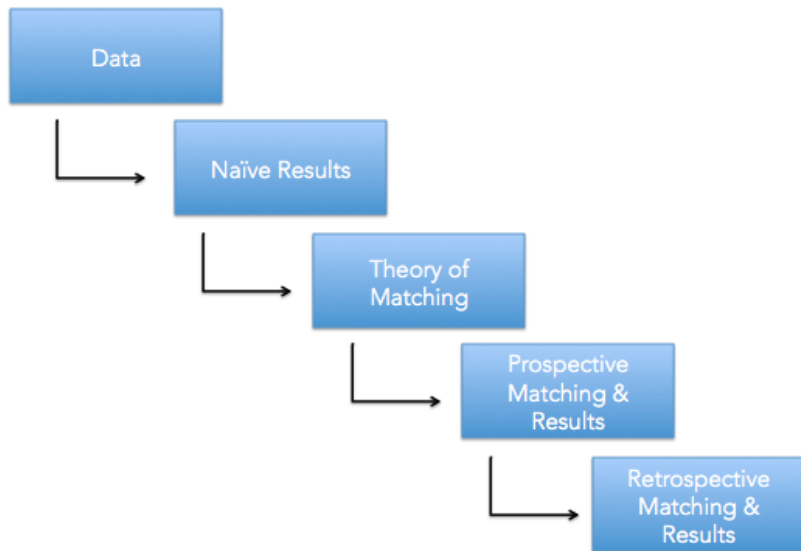


Non-case (Employed)



Case (Unemployed)

Outline



Our Data, Part I

The National Supported Work (NSW) Experiment

- Examined the effect of job training programs on employment
- 445 participants, 185 treated
- Outcome of interest: unemployment in 1978
- Covariates measured:
 - ▶ age
 - ▶ years of education
 - ▶ race (white, black, hispanic)
 - ▶ college degree
 - ▶ marital status
 - ▶ 1974 earnings
 - ▶ 1975 earnings

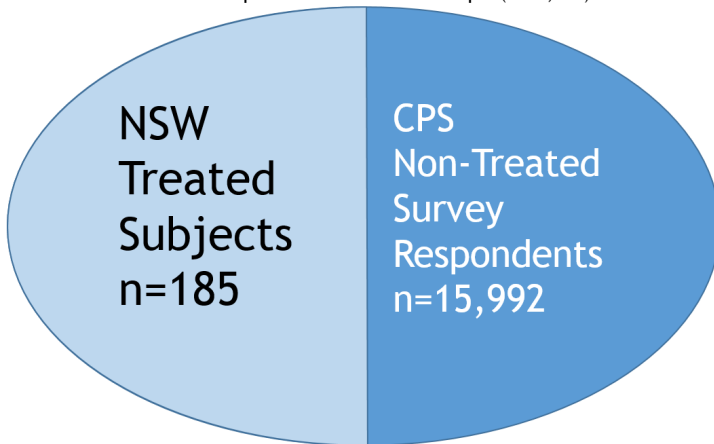
Our Data, Part II

Current Population Survey (CPS)

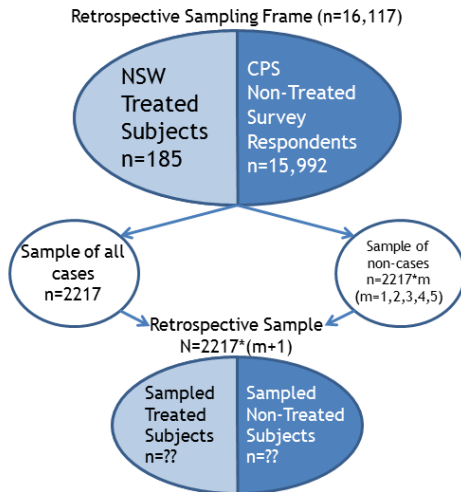
- Administered by the Census Bureau and the Bureau of Labor Statistics
- 15,992 respondents, all non-treated
- Recorded same covariates as NSW experiment

Simulating Prospective Study

Simulated Prospective Observational Sample (n=16,117)



Simulating Retrospective Study



Results from Randomized Experiment

Study	Naive Plug-in
Experimental	0.59

Table: Randomized experiment estimation of odds ratio γ for unemployment

$$\gamma = \frac{\frac{\overline{r_1^{obs}}}{1 - \overline{r_1^{obs}}}}{\frac{\overline{r_0^{obs}}}{1 - \overline{r_0^{obs}}}}$$

- Odds ratio of 1.00 indicates a treated individual is just as likely to be unemployed as an untreated individual
- This is a value we will hope to recover with alternative methods

Naive Plug-in Results

Study	Naive Plug-in
Experimental	0.59
Prospective	2.04
Retrospective 1:1	2.04
Retrospective 1:5	2.04

Table: Prospective Study Treated v. Non-Treated, unmatched estimation of odds ratio γ for unemployment indicator

Visualizing Covariate Imbalance

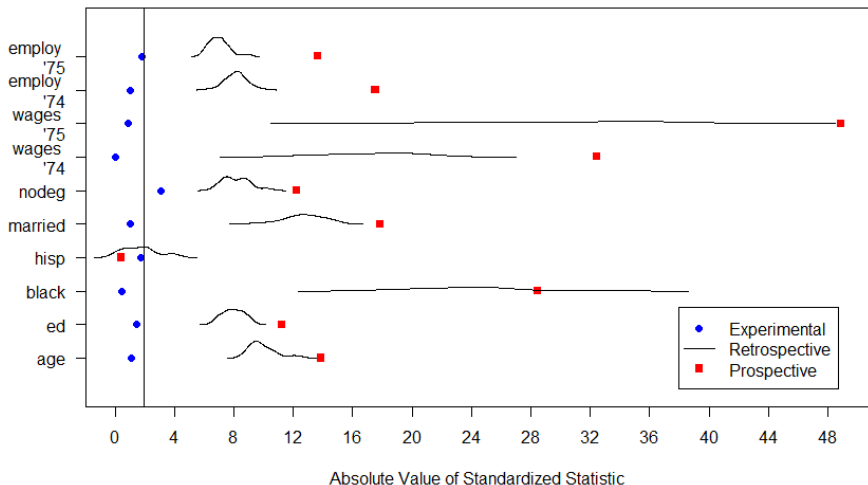


Figure: Love Plot depicting covariate imbalance

Logistic Regression

Let $p_i = P(r_i^{obs} = 1|x, z)$

$$\log \frac{p_i}{1 - p_i} = \beta_0 + \beta^{treat} z_i + \beta x_i, i = 1, \dots, n \quad (1)$$

- Controls for covariates
- Binary outcome variable \rightarrow Binary logistic Regression
- Produces Log odds ratio

Results of Logistic Regression

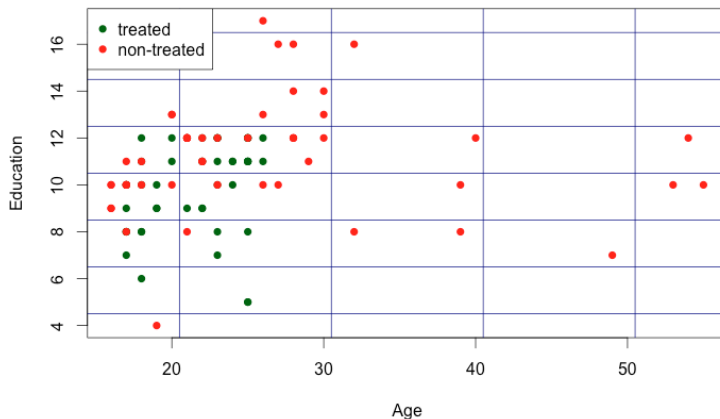
Study	Naive Plug-in	Logistic Regression
Experimental	0.59	0.58
Prospective	2.04	0.87
Retrospective 1:1	2.04	1.15
Retrospective 1:5	2.04	0.90

Table: Unmatched estimation of odds ratio γ for unemployment indicator

- Controlling for covariates means holding other covariates constant
- Covariates constant \neq Covariates equal

Matching

- Match treated unit to one or more non-treated units that have similar covariates
- "Recreate" randomized experiment



Difficulty in Matching

Multi-dimensional Matching Problem

In our example, we have 8 covariates to match on: Age, education, black, hispanic, marriage status, earnings in 1974, earnings in 1975

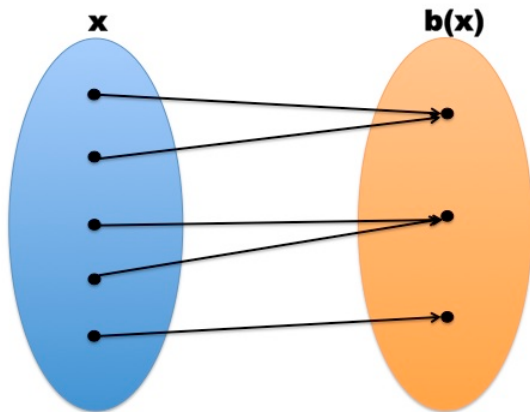
Solution

Balancing Score

Balancing Score

Definition

A balancing score $b(x)$ is a function of the observed covariates x such that $x \perp\!\!\!\perp z | b(x)$.



Propensity Score $e(x)$

Definition

Propensity score $e(x) = P(z = 1|x)$.

- The conditional probability of being assigned the treatment given the observed covariates.
- One-dimensional balancing score

Theorem 1

Theorem

Let $b(x)$ be a function of x . Then $b(x)$ is a balancing score if and only if propensity score $e(x)$ is a function of $b(x)$.

The proof uses:

- Definition of function
- Conditional Independence: $P(Z|X, Y) = P(Z|Y)$
- Law of Iterated Expectation: $E[E[Z|X]] = E[Z]$

Theorem 2

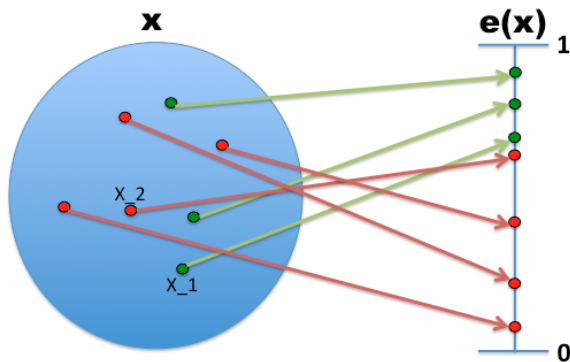
Theorem

Treatment assignment and the observed covariates are conditionally independent given the propensity score $e(x)$, that is,

$$x \perp\!\!\!\perp z | e(x),$$

- We can generate random assignment of treatment by conditioning on $e(x)$
- Matching on $e(x)$, we get treated and non-treated groups that look alike

Propensity Score Matching Example



Ex:

- ▶ x_1 = (25 years old, Hispanic, Not-married, Earned \$5,000 earnings in 1974,etc) $\rightarrow e(x_1) = .89$
- ▶ x_2 = (26 years old, White, Not-married, Earned \$4,900 in 1974,etc) $\rightarrow e(x_2) = .87$

Effects of Matching on a Balancing Score

- Covariate balance
- Obtain unbiased average treatment effect conditioned on $b(x)$

Prospective Matched Results

- Matched treated units to non-treated units using the propensity score

Study	Conditional Logistic Regression
Experimental	0.58
Prospective	0.67

Table: Estimation of odds ratio for unemployment

- “Conditional Logistic Regression” is a form of logistic regression that incorporates the matching information.

Covariate Imbalance After Matching

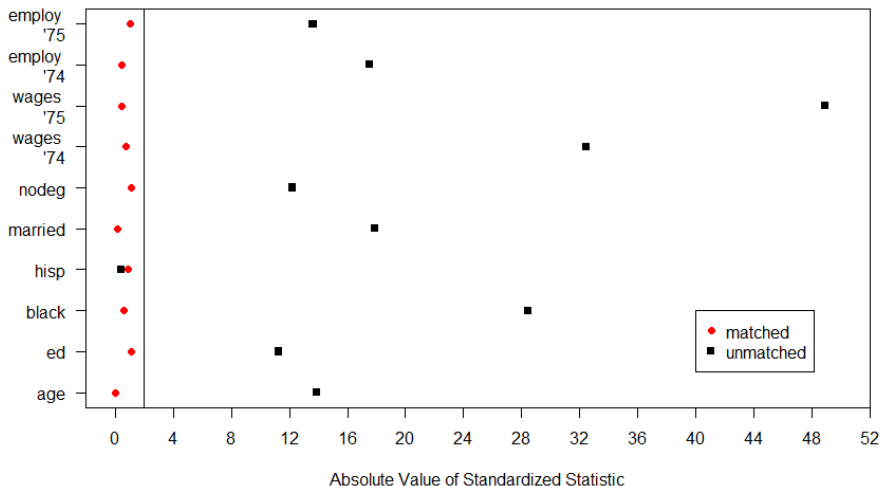
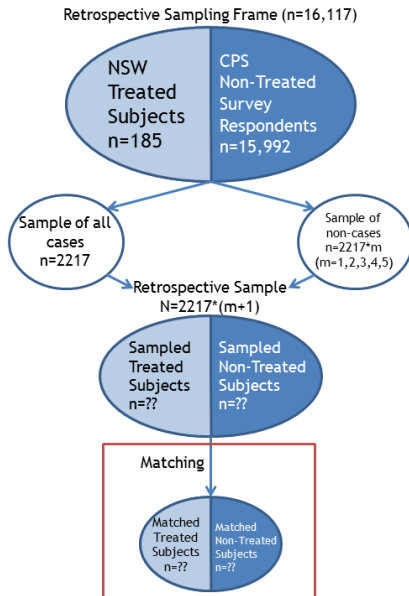


Figure: Love plot depicting covariate imbalance in matched prospective design using propensity score matching

Matching in the Retrospective Setting

- In prospective setting, matching and sampling are intertwined
- Retrospective challenge: we've already sampled subjects based on their case/non-case status
 - ▶ Observed relationship between treatment status and covariates conditioned on being sampled

Matching in the Retrospective Setting



Difficulties with Retrospective Matching

- Let s be a sampling indicator with

$$s = \begin{cases} 1 & \text{if retrospectively sampled} \\ 0 & \text{otherwise.} \end{cases}$$

- Regressing covariates x against treatment z in a retrospective sample estimates

$$P(z = 1|x, s = 1)$$

- ▶ The propensity score $P(z = 1|x) \neq P(z = 1|x, s = 1)$
- ▶ nor a function of it \Rightarrow not a balancing score

The Prospective Exposure Case (PEC) Score

PEC score

$$b_{pro}^{ty}(x) = P(z = t, r^{obs} = y|x) \text{ for } t, y \in \{0, 1\}$$

For a subject with covariates x , PEC score is three dimensional summary showing joint conditional probabilities of treatment and outcome:

	Non-Case ($y=0$)	Case ($y=1$)
Non-Treated ($t=0$)		$b_{pro}^{01}(x)$
Treated ($t=1$)	$b_{pro}^{10}(x)$	$b_{pro}^{11}(x)$

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Non-Treated (t=0)		$b_{pro}^{01}(x)$
Treated (t=1)	$b_{pro}^{10}(x)$	$b_{pro}^{11}(x)$

Theorem

$b_{pro}^{ty}(x)$ is a balancing score, i.e., $e(x) = f(b_{pro}^{ty}(x))$ for some function f .

Proof.

$$e(x) = P(z = 1|x) = b_{pro}^{10}(x) + b_{pro}^{11}(x)$$



The Retrospective Exposure Case (REC) Score

- Analogous to PEC score, but conditions on retrospective sampling.

REC Score

$$b_{ret}^{ty}(x) = P(z = t, r^{obs} = y | x, s = 1) \text{ for } t, y \in \{0, 1\}$$

	Non-Case (y=0)	Case (y=1)
Non-Treated (t=0)		$b_{ret}^{01}(x)$
Treated (t=1)	$b_{ret}^{10}(x)$	$b_{ret}^{11}(x)$

The Retrospective Exposure Case (REC) Score

Theorem

The REC score $b_{ret}^{ty}(x)$ is a balancing score, i.e., $e(x) = f(b_{ret}^{ty}(x))$ for some function f .

We prove this by writing the PEC score as a function of the REC score.

The Retrospective Exposure Case (REC) Score

Theorem

Let $p_1 = P(s = 1 | r^{obs} = 1)$, and $p_0 = P(s = 1 | r^{obs} = 0)$. Assume sampling is independent of treatment and covariates, given case status. Then the REC score $b_{ret}^{ty}(x)$ is a balancing score, i.e., $e(x) = f(b_{ret}^{ty}(x))$ for some function f .

Proof.

$$b_{ret}^{ty}(x) = \frac{p_y b_{pro}^{ty}(x)}{p_0 b_{pro}^{00}(x) + p_0 b_{pro}^{01}(x) + p_1 b_{pro}^{10}(x) + p_1 b_{pro}^{11}(x)}$$

Inverting this, we can show $b_{pro}^{ty}(x) \propto p_{1-y} b_{ret}^{ty}(x)$. Thus, $e(x)$ is a function of $b_{ret}^{ty}(x)$. □

Estimating the REC Score

To estimate the REC score from a retrospective sample we use *multinomial regression*

- similar to logistic regression, but for modeling a multinomial variable like REC score from covariates.
- predicts each dimension of REC Score for each subject

Matching on the RECS Score

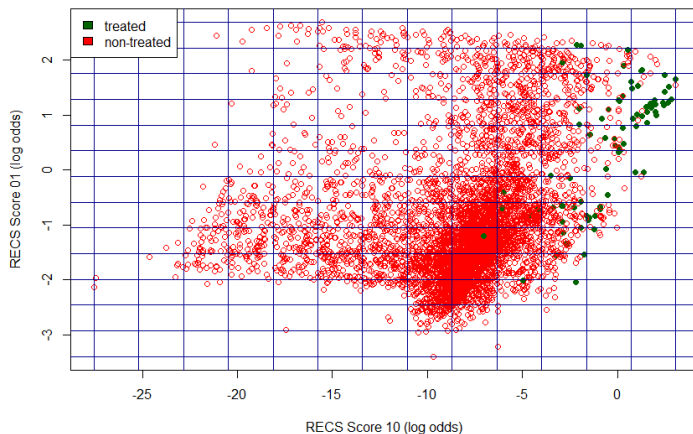


Figure: Plot showing two dimensions of RECS with strata

Covariate Imbalance After Matching

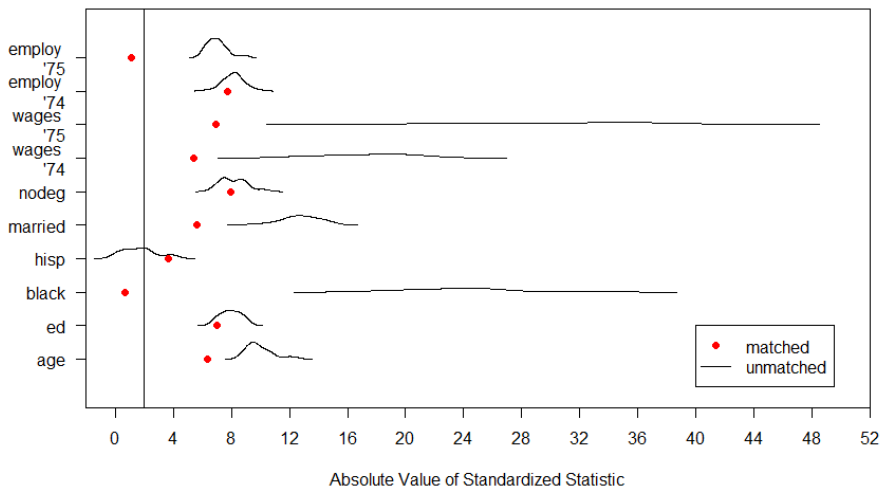


Figure: Love plot depicting covariate imbalance in matched retrospective dataset using RECS matching

Retrospective Matched Estimates

Study	Conditional Logistic Regression
Experimental	0.58
Retrospective 1:1	0.94
Retrospective 1:5	0.60

Table: Retrospective Study Treatment v. Control

- On average, 1:5 Retrospective analysis shows that receiving the job training decreases the odds of being unemployed by 40%.
 - ▶ Almost exactly recovers experimental value.

Primary Findings

- Matching is an effective way to fix covariate imbalance in prospective studies
- Results indicate REC score matching works
- Retrospectively matched sample still contains covariate imbalance

Limitations

- Matching

- ▶ Only observed variables are matched on
- ▶ Large sample needed to find good matches

Looking Ahead

- Put standard errors on estimates
- Implement REC score matching on another dataset
- Figure out why the RECS matching did so well
- Find a better balancing score to match on
- Lots of observational data and causal inference

Thank you!

- The inestimable Dave Watson
- Carleton Math & Stats faculty and mentors
- Our many friends and colleagues
- Donald Rubin & Paul Rosenbaum, inventors of $e(x)$