Problem 1: Suppose that X and Y are jointly continuous random variables with density

$$f(x,y) = \begin{cases} 24xy & : 0 \le x \le 1, \ 0 \le y \le 1, \ x+y \le 1, \\ 0 & : \text{otherwise.} \end{cases}$$

Compute the expectation E(XY).

Using the bivariate LotUS, we get:

$$E(XY) = \iint_{\mathbb{R}^2} xy f(x, y) \, dy dx = 24 \int_0^1 \int_0^{1-x} x^2 y^2 \, dy dx = \frac{2}{15}.$$

Problem 2: Suppose X and Y are jointly continuous random variables with the same density from Problem 1. Compute a formula for the conditional expectation $E(Y \mid X = x)$. Take care to precisely state the domain of this function.

We first get the marginal density for X:

$$f(x) = \int_{\mathbb{R}} f(x, y) \, dy = 24 \int_{0}^{1-x} xy \, dy = 12x(1-x)^{2},$$

where the second inequality holds for all $0 \le x \le 1$; otherwise, we have f(x) = 0. Note that the marginal f(x) is nonzero for 0 < x < 1. For these particular x-values, the conditional density is defined and is given by

$$f(y|x) = \frac{f(x,y)}{f(x)} = \begin{cases} \frac{2y}{(1-x)^2} & : 0 \le x \le 1, \ 0 \le y \le 1, \ x+y \le 1, \\ 0 & : \text{otherwise.} \end{cases}$$

Therefore, for all x with 0 < x < 1, we have

$$E(Y \mid X = x) = \int_{\mathbb{R}} y f(y|x) \, dy = \frac{2}{(1-x)^2} \int_0^{1-x} y^2 \, dy = \frac{2}{3} (1-x).$$

Problem 3: Let X and Y be two random variables on the probability space $S = \{a, b, c\}$. Suppose that the probability distribution P on S has mass function p(s) and that X and Y are defined according to the following table:

Compute the random variable $E(Y \mid X)$.

Our first goal is to get the joint mass function p(x,y). We compute:

$$\begin{array}{c|cccc} p(x,y) & y = 1 & y = 2 \\ \hline x = 1 & 0.3 & 0.2 \\ x = 2 & 0.5 & 0 \end{array}$$

By adding across the rows, we get the marginal:

$$\begin{array}{c|cc}
x & p(x) \\
\hline
1 & 0.5 \\
2 & 0.5
\end{array}$$

Then, from the formula p(y|x) = p(x,y)/p(x), we get

$$\begin{array}{c|cccc} p(y|x) & y = 1 & y = 2 \\ \hline x = 1 & 0.6 & 0.4 \\ x = 2 & 1 & 0 \end{array}$$

If $x \neq 1$ or 2, then the conditional mass p(y|x) is undefined. For x = 1 or 2, we have

$$E(Y \mid X = x) = \sum_{y=1}^{2} yp(y|x) = \begin{cases} 1 \cdot 0.6 + 2 \cdot 0.4 = 1.4 & : x = 1, \\ 1 \cdot 1 + 2 \cdot 0 = 1 & : x = 2. \end{cases}$$

Finally, recall that the random variable $E(Y \mid X)$ is defined as the composite $h(X) = h \circ X$, where $h(x) = E(Y \mid X = x)$ for x = 1, 2. Hence, we conclude the problem with:

Problem 4: Suppose that a point X = x is chosen uniformly in the interval (0,1). After x has been chosen, suppose that a second point Y = y is chosen uniformly in the interval [x,1]. Compute the expectation E(Y).

We will compute the expectation via the Law of Total Expectation:

$$E(Y) = E[E(Y \mid X)].$$

So, we begin by noting that $E(Y \mid X = x) = (x+1)/2 = x/2 + 1/2$, since the mean of a uniform distribution over an interval is the midpoint between its boundary points. Therefore $E(Y \mid X) = X/2 + 1/2$, and so

$$E(Y) = E(X/2 + 1/2) = E(X)/2 + 1/2 = 1/4 + 1/2 = 3/4.$$

Problem 5: The waiting time X in minutes between calls to a 911 center is exponentially distributed with mean $\mu = 2$ minutes. Compute the distribution of the transformed random variable Y = 60X that measures the waiting time in seconds.

Since the mean of $X \sim \mathcal{E}xp(\lambda)$ is $1/\lambda$, we see that $\lambda = 1/2$ and so

$$f_X(x) = \begin{cases} \frac{1}{2}e^{-x/2} & : x > 0, \\ 0 & : \text{otherwise.} \end{cases}$$

In the notation of the Density Transformation Theorem, the support of the density is $T=(0,\infty)$. If we define $r:T\to\mathbb{R}$ by r(x)=60x, then the range U of r is the open interval $(0,\infty)$. Note that the inverse function $s:U\to\mathbb{R}$ is given by s(y)=y/60, which is continuously differentiable. Therefore, we have everything that we need to apply the Density Transformation Theorem. We get:

$$f_Y(y) = \begin{cases} \frac{1}{120} e^{-y/120} & : y > 0, \\ 0 & : \text{ otherwise,} \end{cases}$$

and thus $Y \sim \mathcal{E}xp(1/120)$.

Problem 6: Suppose that X and Y are two random variables such that $Y = e^X$ and $X \sim \mathcal{N}(\mu, \sigma^2)$. Compute the density of Y.

The support of the density of X is all of \mathbb{R} ; so, in the notation of the Density Transformation Theorem we have $T = \mathbb{R}$. Define $r: T \to \mathbb{R}$ by setting $r(x) = e^x$, and note that the image U of r is the open interval $(0, \infty)$. The inverse function $s: U \to \mathbb{R}$ is given by $s(y) = \log y$, which is continuously differentiable. Therefore, we have everything that we need in order to use the Density Transformation Theorem. We compute:

$$f_Y(y) = \begin{cases} \frac{1}{y\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2\sigma^2} \left(\log y - \mu\right)^2\right] &: y > 0, \\ 0 &: \text{otherwise.} \end{cases}$$

Hence, Y is a lognormal random variable.

Problem 7: Suppose that $\mathbf{X} = (X_1, X_2)$ is a two-dimensional continuous random vector with density

$$f(x_1, x_2) = \begin{cases} 4x_1x_2 &: 0 < x_1 < 1, \ 0 < x_2 < 1, \\ 0 &: \text{otherwise.} \end{cases}$$

Letting T be the support of the density, define $r: T \to \mathbb{R}^2$ by setting

$$r(x_1, x_2) = \left(\frac{x_1}{x_2}, x_1 x_2\right)$$

for $(x_1, x_2) \in \mathbb{R}^2$. Compute the density of the random vector $\mathbf{Y} = r(\mathbf{X})$.

As I will explain during class, the image U of r is the open region in the (y_1, y_2) -plane bounded by the curves

$$y_1 = y_2, \quad y_2 = 0, \quad y_2 = \frac{1}{y_2}.$$

To compute an inverse $s:U\to\mathbb{R}^2$ of r, notice that the equations

$$y_1 = \frac{x_1}{x_2}$$
 and $y_2 = x_1 x_2$

can be solved uniquely for x_1 and x_2 yielding the solutions

$$x_1 = \sqrt{y_1 y_2}$$
 and $x_2 = \sqrt{\frac{y_2}{y_1}}$.

Thus we may define s by setting

$$s(y_1, y_2) = \left(\sqrt{y_1 y_2}, \sqrt{\frac{y_2}{y_1}}\right)$$

for all $(y_1, y_2) \in U$. To verify that s is continuously differentiable, we compute its Jacobian matrix:

$$\frac{\partial(s_1, s_2)}{\partial(y_1, y_2)}(y_1, y_2) = \begin{bmatrix} \frac{1}{2}\sqrt{\frac{y_2}{y_1}} & \frac{1}{2}\sqrt{\frac{y_1}{y_2}} \\ -\frac{1}{2}\sqrt{\frac{y_2}{y_1^3}} & \frac{1}{2}\sqrt{\frac{1}{y_1y_2}} \end{bmatrix}.$$

The four partial derivatives are all continuous on U, and thus s is continuously differentiable. We therefore have everything that we need to apply the Density Transformation Theorem. First, we compute the determinant of the Jacobian matrix:

$$\det \frac{\partial(s_1, s_2)}{\partial(y_1, y_2)}(y_1, y_2) = \frac{1}{2y_1}.$$

Then, we have

$$f\left(\sqrt{y_1y_2}, \sqrt{\frac{y_2}{y_1}}\right) \cdot \frac{1}{2y_1} = 4\sqrt{y_1y_2} \cdot \sqrt{\frac{y_2}{y_1}} \cdot \frac{1}{2y_1} = \frac{2y_2}{y_1}$$

for $(y_1, y_2) \in U$, and thus

$$f(y_1, y_2) = \begin{cases} \frac{2y_2}{y_1} & : (y_1, y_2) \in U, \\ 0 & : \text{ otherwise.} \end{cases}$$

Problem 8: Suppose that X is a continuous random variable with uniform distribution on [a, b]. Compute its moment generating function $\psi(t)$, and then find all moments $E(X^k)$, for $k \ge 1$.

We have

$$f(x) = \frac{1}{b-a}$$

for all $x \in [a, b]$ and f(x) = 0 otherwise. Thus, by the LotUS, we have

$$\psi(t) = E(e^{tX}) = \frac{1}{b-a} \int_a^b e^{tx} dx = \frac{e^{tb} - e^{ta}}{t(b-a)}.$$

If we expand the exponentials, we get

$$\psi(t) = \frac{1}{t(b-a)} \left[\sum_{k=0}^{\infty} \frac{t^k b^k}{k!} - \sum_{k=0}^{\infty} \frac{t^k a^k}{k!} \right]$$
$$= \frac{1}{b-a} \sum_{k=0}^{\infty} \frac{b^{k+1} - a^{k+1}}{(k+1)!} t^k$$

for all $t \in \mathbb{R}$. Thus,

$$E(X^k) = \psi^{(k)}(0) = \frac{b^{k+1} - a^{k+1}}{(k+1)(b-a)} = \frac{a^k + a^{k-1}b + \dots + ab^{k-1} + b^k}{k+1}.$$

Problem 9: Use moment generating functions to confirm that the mean and variance of a random variable $X \sim \mathcal{N}(\mu, \sigma^2)$ are indeed μ and σ^2 .

Letting $\psi(t)$ be the moment generating function of X, we have

$$\psi'(t) = \psi(t)(\mu + \sigma^2 t),$$

and so

$$E(X) = \psi'(0) = \mu.$$

To compute the variance, we first compute

$$\psi''(t) = \psi(t)(\mu + \sigma^2 t)^2 + \psi(t)\sigma^2.$$

Then

$$E(X^2) = \psi''(0) = \mu^2 + \sigma^2,$$

and so

$$V(X) = E(X^{2}) - E(X)^{2} = \mu^{2} + \sigma^{2} - \mu^{2} = \sigma^{2}.$$

Problem 10: Suppose that X and Y are random variables with the joint density function

$$f(x,y) = \begin{cases} 2xy + 0.5 & : 0 \le x \le 1, \ 0 \le y \le 1, \\ 0 & : \text{otherwise.} \end{cases}$$

Compute the covariance of X and Y.

Using the Shortcut Formula for Covariance, we compute:

$$\sigma_{XY} = E(XY) - E(X)E(Y).$$

But first, let's grab the expectations of X and Y. To do this, we integrate out y to get the density of x:

$$f(x) = \int_{\mathbb{R}} f(x, y) \, dy = \int_{0}^{1} (2xy + 0.5) \, dy = x + 0.5$$

for $0 \le x \le 1$, and f(x) = 0 otherwise. Then, we compute:

$$E(X) = \int_{\mathbb{R}} x f(x) \, dx = \int_{0}^{1} (x^{2} + 0.5x) \, dx = \frac{7}{12}.$$

Now, if you look at the joint density function, you'll notice that it is symmetric in x and y. This means that E(Y) = 7/12, as well. Finally, we compute the covariance from the shortcut formula:

$$\sigma_{XY} = E(XY) - \frac{7^2}{12^2}$$

$$= \iint_{\mathbb{R}^2} xy f(x, y) \, dy dx - \frac{7^2}{12^2}$$

$$= \int_0^1 \int_0^1 (2x^2 y^2 + 0.5xy) \, dy dx - \frac{7^2}{12^2}$$

$$= \frac{25}{72} - \frac{7^2}{12^2}$$

$$= \frac{1}{144}$$

$$\approx 0.007.$$

Problem 11: Compute the correlation ρ_{XY} of the random variables in the previous problem.

First, we compute the variances:

$$\sigma_X^2 = \int_{\mathbb{R}} x^2 f(x) \, dx = \int_0^1 x^2 (x + 0.5) \, dx = \frac{5}{12}.$$

By symmetry, we also have $\sigma_Y^2 = 5/12$. Thus,

$$\rho_{XY} = \frac{\sigma_{XY}}{\sigma_X \sigma_Y} = \frac{1/144}{(5/12)^2} = \frac{1}{60} \approx 0.017.$$