

DIFFERENCE
IN
DIFFERENCES
THEORY
APPLICATION
APPENDIX

CAUSAL INFERENCE

DAY 2

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DIFFERENCE IN DIFFERENCES

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- Finding a valid instrument is hard!!

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- Finding a valid instrument is hard!!
- Let's look at another workhouse framework in economics:
Difference in differences (DiD)

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- This framework is much older than IV (mid 1800s)

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- Finding a valid instrument is hard!!
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- A group of units are assigned some treatment and then compared to a group of units that weren't
 - Therefore mostly used when you observe a large shock/policy change in time that only impacts certain units

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JOHN SNOW AND CHOLERA

- Three major waves of cholera in the early to mid 1800s in London
- John Snow believed cholera was spread through the Thames water supply which contradicted dominant theory about “dirty air” transmission
- Grand experiment: Lambeth moves its pipe between 1849 and 1854; Southwark and Vauxhall delay
- He can evaluate the effect in three ways (one of which is DiD)

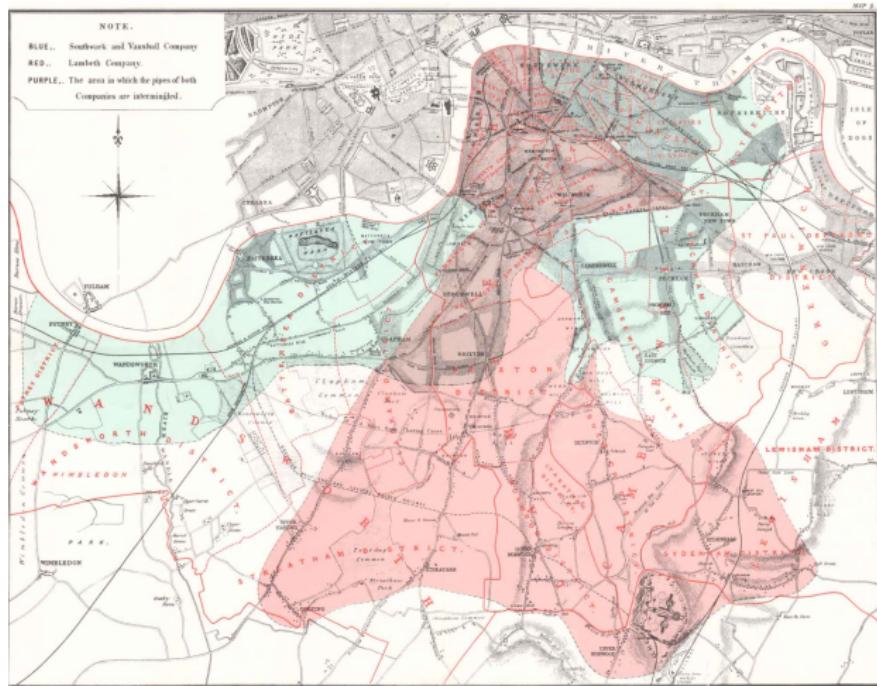
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FIGURE: Two water utility companies in London 1854



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FIGURE: Results

| Death-Rates per 1,000 of living Population in Two Epidemic Periods. | in 1854 as receiving their Water-supply— | |
|---|---|--|
| | from the LAMBETH Company. | from the SOUTHWARK and VAUXHALL Company. |
| CHOLERA | 1848-9.. | 12.5 |
| | 1853-4.. | 3.7 |
| DIARRHEA | 1848-9.. | 2.9 |
| | 1853-4.. | 2.1 |
| | | 11.8 |
| | | 13.0 |
| | | 2.7 |
| | | 3.3 |

1) SIMPLE CROSS-SECTIONAL DESIGN

TABLE: Lambeth and Southwark and Vauxhall, 1854

| Company | Cholera mortality |
|------------------------|-------------------|
| Lambeth | $Y = L + D$ |
| Southwark and Vauxhall | $Y = SV$ |

$$\widehat{\delta}_{cs} = D + (L - SV)$$

What is L and SV ?

1) SIMPLE CROSS-SECTIONAL DESIGN

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$$\widehat{\delta}_{cs} = D + (L - SV)$$

This is biased if $L \neq SV$ (selection bias). Give an example when we're pretty sure they are equal.

2) INTERRUPTED TIME SERIES DESIGN

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TABLE: Lambeth, 1849 and 1854

| Company | Time | Cholera mortality |
|----------------|-------------|--------------------------|
| Lambeth | 1849 | $Y = L$ |
| | 1854 | $Y = L + (T + D)$ |

$$\hat{\delta}_{its} = D + T$$

What is required for this estimator to be unbiased?

3) DIFFERENCE-IN-DIFFERENCES

TABLE: Lambeth and Southwark and Vauxhall, 1849 and 1854

| Companies | Time | Outcome | D_1 |
|------------------------|-------------|-------------------|-----------|
| Lambeth | Before | $Y = L$ | |
| | After | $Y = L + T_L + D$ | $T_L + D$ |
| | | | |
| Southwark and Vauxhall | Before | $Y = SV$ | |
| | After | $Y = SV + T_{SV}$ | T_{SV} |

$$\widehat{\delta}_{did} = D + (T_L - T_{SV})$$

TREATMENT EFFECT DEFINITIONS

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INDIVIDUAL TREATMENT EFFECT

The individual treatment effect, δ_i , equals $Y_i^1 - Y_i^0$

Individual causal effects cannot be calculated because one of the two needed potential outcomes will always be missing.

CONDITIONAL AVERAGE TREATMENT EFFECTS

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AVERAGE TREATMENT EFFECT ON THE TREATED (ATT)

The average treatment effect on the treatment group is equal to the average treatment effect conditional on being a treatment group member:

$$\begin{aligned} E[\delta|D = 1] &= E[Y^1 - Y^0|D = 1] \\ &= E[Y^1|D = 1] - E[Y^0|D = 1] \end{aligned}$$

We can estimate the ATT, but never be sure due to **missing potential outcomes** for the treated group

IDENTIFICATION WITHOUT RANDOMIZATION

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- We may be unable to randomize – not because we lack the imagination, but because we lack the permission
- If we cannot randomize, then how does DiD identify a treatment effect, and which treatment effect?
- DiD identifies the ATT, and since we are missing Y^0 for treated group, we will restrict counterfactual Y^0 in expectation
- One of the main advantages of DiD is the hope that we can identify the ATT *without* randomization

DiD EQUATION

DiD Equation: k and U notation for treated and untreated groups

$$\widehat{\delta}_{kU}^{2 \times 2} = \left(E[Y_k | Post] - E[Y_k | Pre] \right) - \left(E[Y_U | Post] - E[Y_U | Pre] \right)$$

k index people with Lambeth, U index people with Southwark and Vauxhall, $Post$ is after Lambeth moved pipe upstream, Pre before Lambeth moved its pipe (baseline), and $E[y]$ mean cholera mortality.

POTENTIAL OUTCOMES AND THE SWITCHING EQUATION

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$$\widehat{\delta}_{kU}^{2 \times 2} = \underbrace{\left(E[Y_k^1 | Post] - E[Y_k^0 | Pre] \right) - \left(E[Y_U^0 | Post] - E[Y_U^0 | Pre] \right)}_{\text{Switching equation}} + \underbrace{E[Y_k^0 | Post] - E[Y_k^0 | Post]}_{\text{Adding zero}}$$

PARALLEL TRENDS BIAS

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$$\widehat{\delta}_{kU}^{2 \times 2} =$$
$$\underbrace{E[Y_k^1 | Post] - E[Y_k^0 | Post]}_{ATT}$$
$$+ \underbrace{\left[E[Y_k^0 | Post] - E[Y_k^0 | Pre] \right] - \left[E[Y_U^0 | Post] - E[Y_U^0 | Pre] \right]}_{\text{Non-parallel trends bias in } 2 \times 2 \text{ case}}$$

IDENTIFICATION THROUGH PARALLEL TRENDS

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PARALLEL TRENDS

Assume two groups, treated and comparison group, then we define parallel trends as:

$$E(\Delta Y_k^0) = E(\Delta Y_U^0)$$

In words: “The evolution of cholera mortality for Lambeth *had it kept its pipe downstream* is the same as the evolution of cholera mortality for Southwark and Vauxhall”.

It's in red so you know it's a nontrivial assumption. But why?
Can't we just check?

PARALLEL TRENDS IS HARD

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- **There is no guaranteed “science” in parallel trends**
- This makes DiD a “hard” design – *because* it doesn’t rely on randomization (it relies on parallel trends), there are no “slam dunks” and evidence tends to be multi-layered
- Before we move into regression, let’s go through a simple exercise to really pin down these core ideas with simple calculations

Simple Exercise

OLS SPECIFICATION

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- Simple DiD equation will identify ATT under parallel trends
- But so will a particular OLS specification (two groups and no covariates)
- OLS was historically preferred because
 - OLS estimates the ATT under parallel trends
 - Easy to calculate the standard errors
 - Easy to include multiple periods
- People liked it also because of differential timing, continuous treatments and covariates, but those are more complex so we address them later

MINIMUM WAGES

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- Card and Krueger (1994) have a famous study estimating causal effect (ATT) of minimum wages on employment
- Exploited a policy change in New Jersey between February and November in mid-1990s where minimum wage was increased, but neighbor PA did not
- Using DiD, they do not find a negative effect of the minimum wage on employment which is part of its legacy today, but I mainly present it to illustrate the history and the design principles

OLS SPECIFICATION OF THE DiD EQUATION

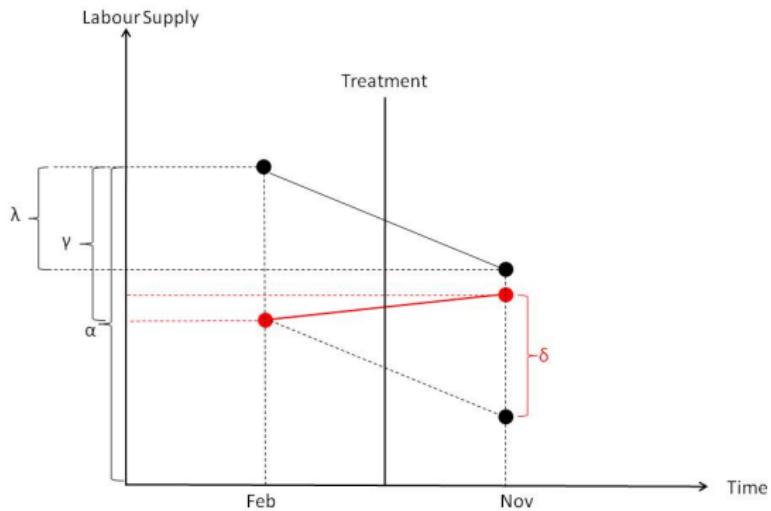
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- The correctly specified OLS regression is an interaction with time and group fixed effects:

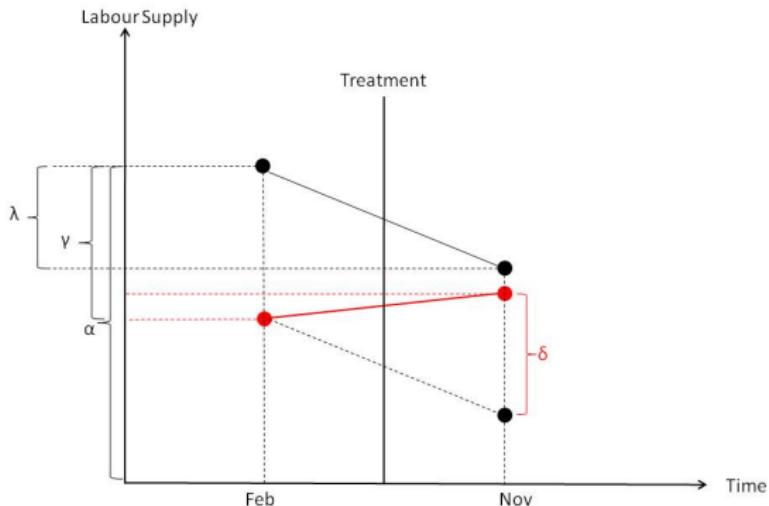
$$Y_{its} = \alpha + \gamma NJ_s + \lambda d_t + \delta(NJ \times d)_{st} + \varepsilon_{its}$$

- NJ is a dummy equal to 1 if the observation is from NJ
- d is a dummy equal to 1 if the observation is from November (the post period)
- This equation takes the following values
 - PA Pre: α
 - PA Post: $\alpha + \lambda$
 - NJ Pre: $\alpha + \gamma$
 - NJ Post: $\alpha + \gamma + \lambda + \delta$
- DiD equation: $(NJ \text{ Post} - NJ \text{ Pre}) - (PA \text{ Post} - PA \text{ Pre}) = \delta$

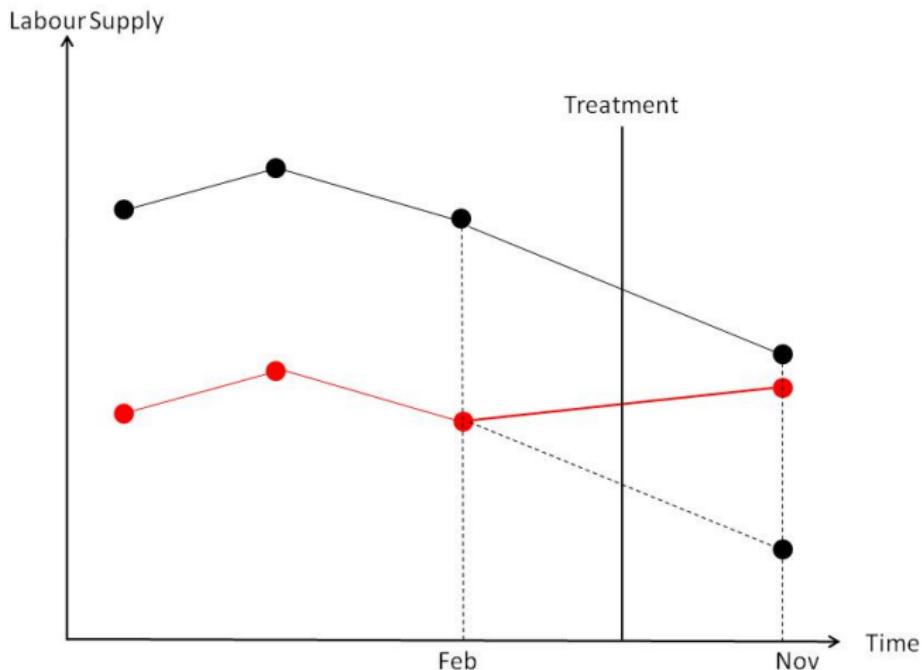
$$Y_{ist} = \alpha + \gamma NJ_s + \lambda d_t + \delta (NJ \times d)_{st} + \varepsilon_{ist}$$



$$Y_{ist} = \alpha + \gamma NJ_s + \lambda d_t + \delta(NJ \times d)_{st} + \varepsilon_{ist}$$



Notice how OLS is “imputing” $E[Y^0|D = 1, Post]$ for the treatment group in the post period? It is only “correct”, though, if parallel trends is a good approximation



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START OF APPENDIX SLIDES

SELECT VIRTUAL ENVIRONMENT

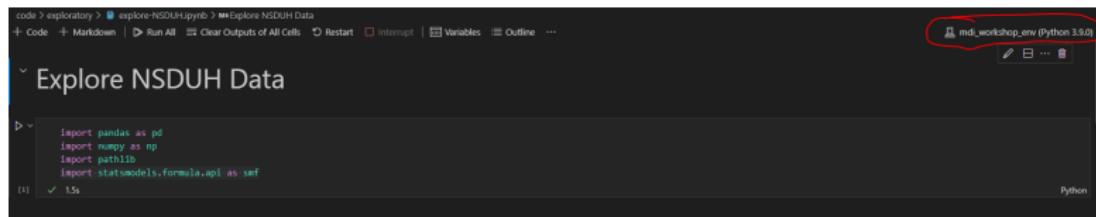
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The screenshot shows a Jupyter Notebook interface with the following details:

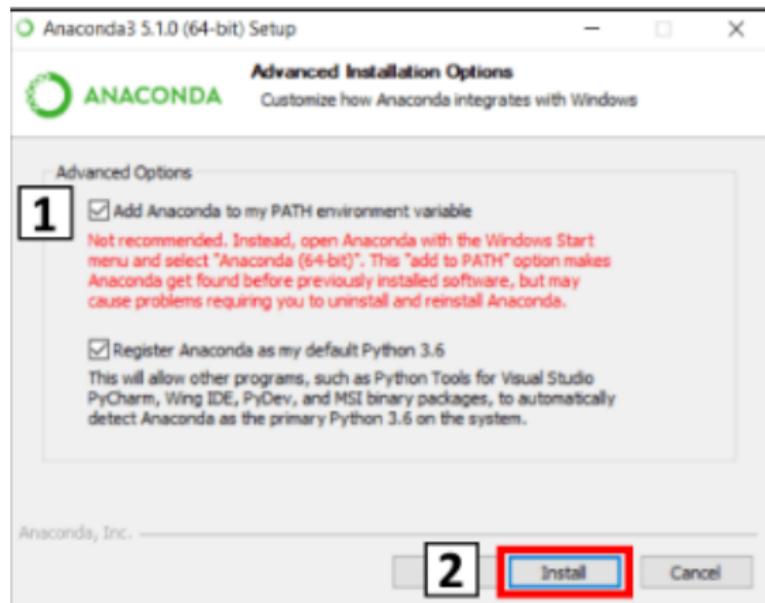
- Header: Code > exploratory > explore-NSDUH.ipynb > Explore NSDUH Data
- Toolbar: + Code, + Markdown, Run All, Clear Outputs of All Cells, Restart, Interrupt, Variables, Outline, ...
- Section: ▾ Explore NSDUH Data
- Code Cell (Cell 1):

```
import pandas as pd
import numpy as np
import pathlib
import statsmodels.formula.api as smf
```
- Cell Status: ✓ 1.5s
- Kernel: Python 3.9.0
- Virtual Environment: mdu_workshop_env (Python 3.9.0) (highlighted with a red box)

ADD CONDA TO PATH

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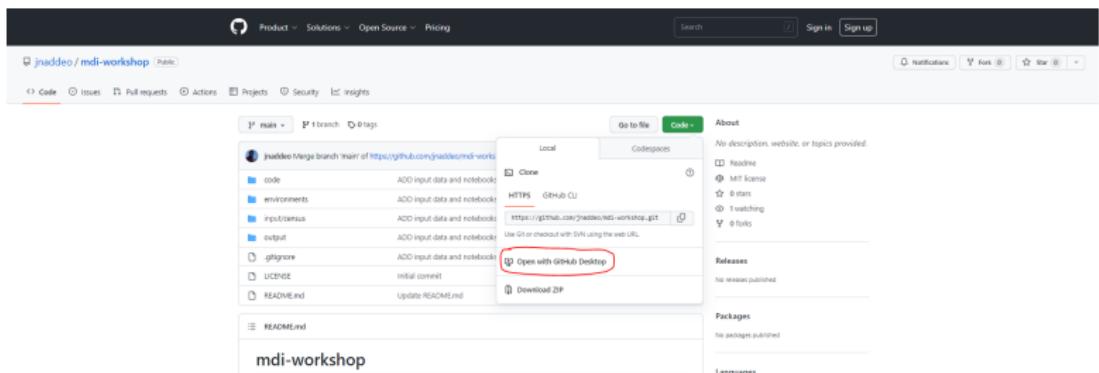
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OPEN REPO WITH DESKTOP

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RANDOM ASSIGNMENT BLACK

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figures/randomization_check_dismissal_black.png

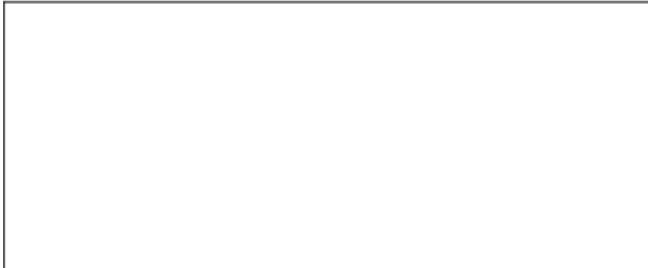
RANDOM ASSIGNMENT WHITE

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figures/randomization_check_dismissal_white.png

- ? provides joint test for null hypothesis that the exclusion and monotonicity assumptions hold, I fail to reject the null
- Also show assumption can be relaxed to average monotonicity \Rightarrow first stage should hold in all subsamples



figures/first_stage_subsamples.png