# Welcome to Week 16!

You made it!!

# A quicky review!

Emphasis on Final will be material listed here & in todays lecture.

# Hypothesis Testing Framework (Ch. 4-6)

The general outline of the process:

1. Set the hypotheses.

For a single proportion this will look like:

H₀: p = null value

H<sub>A</sub>: p < or > or ≠ null value

- 2. Check assumptions and conditions
- 3. Calculate a test statistic and a p-value
- 4. Make a decision, and interpret it in context
- If p-value < α, reject H<sub>0</sub>,
   there is sufficient evidence for [H<sub>A</sub>]
- If p-value > α, do not reject H₀,
   there is not sufficient for evidence for [H₄]

**English** 

Provides a rigorous way to determine the answer with a specific level of confidence.



**English** 

# Hypothesis Testing Framework (Ch. 4-6)

```
The general outline of the process:
                                                         (a) normal, large sample
                                                         (b) normal?, small sample
  1. Set the hypotheses.
       For a single proportion this will look like. (c) observations & theory
           H₀: p = null value
           H_A: p < or > or \neq null value
                                                             Test Statistics
2. Check assumptions and conditions
                                                            (a) Z-score -> P(Z)
3. Calculate a test statistic and a p-value
                                                            (b) T-Score -> P(T)
 4. Make a decision, and interpret it in context
                                                            (c) \chi^2 -> P(\chi^2)
     If p-value < \alpha, reject H_0,
                      there is sufficient evidence for [H<sub>A</sub>]
```

If p-value > α, do not reject H<sub>0</sub>,
 there is not sufficient for evidence for [H<sub>A</sub>]

# Anatomy of a test statistic

Only tricks are:

The general form of a test statistic is (1) picking what the point and null values are based on our hypotheses

point estimate – null value SE of point estimate

(2) what the form of the standard errors is based on what our underlying distribution looks like (normal, tdistribution,  $\chi^{2)}$ 

This construction is based on

- identifying the difference between a point estimate and an expected value if the null hypothesis was true, and
- standardizing that difference using the standard error of the point estimate.

These two ideas will help in the construction of an appropriate test statistic for count data.

# Decision errors (cont.)

There are two competing hypotheses: the null and the alternative. In a hypothesis test, we make a decision about which might be true, but our choice might be incorrect.

		Decision	
		fail to reject $H_0$	reject $H_0$
Truth	$H_0$ true	<b>✓</b>	Type 1 Error
	$H_A$ true	Type 2 Error	<b>✓</b>

A Type 1 Error is rejecting the null hypothesis when  $H_0$  is true. A Type 2 Error is failing to reject the null hypothesis when  $H_A$  is true.

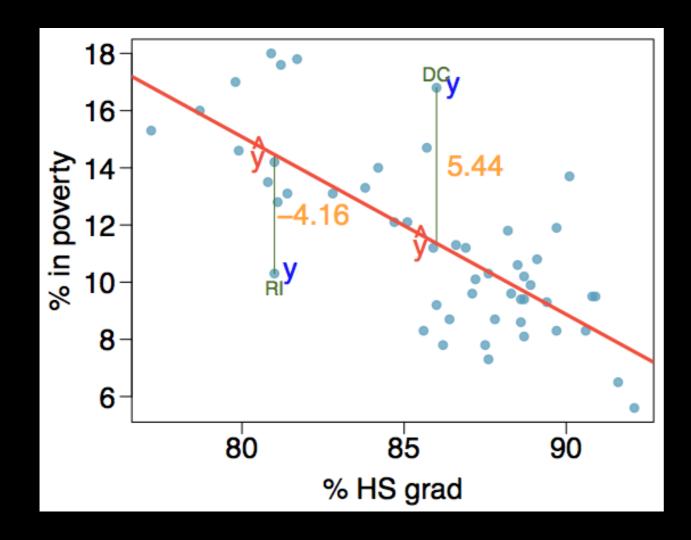
We (almost) never know if  $H_0$  or  $H_A$  is true, but we need to consider all possibilities.

## Residuals

Residuals are the leftovers from the model fit: Data = Fit + Residual

Aka a residual is the difference between the observed  $(y_i)$  and predicted  $\hat{y}_i$ .

$$e_i = y_i - \hat{y}_i$$



Here is a depiction of the residuals - how far each point is from our fitted line.

# p-values for Linear Regression

What's really going on here? Just the same calculations we've been doing the past few weeks!

p-value > 0.05 so we fail to reject H<sub>0</sub>

```
> summary(myLine)
Call:
lm(formula = BAC \sim Beers, data = BB)
Residuals:
     Min
                      Median
                10
-0.027118 -0.017350 0.001773 0.008623 0.041027
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.012701 0.012638 -1.005
                                           0 332
            0.017964 0.002402
Beers
                                  7.48% 2.97e-06
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' '1
Residual standard error: 0.02044 on 14 degrees of freedom
Multiple R-squared: 0.7998, Adjusted R-squared: 0.7855
F-statistic: 55.94 on 1 and 14 DF, p-value: 2.969e-06
```

H<sub>0</sub>: There is no relation between Beers and BAC - slope = 0 H<sub>A</sub>: There is a relationship between Beers and BAC - slope != 0

## Conditions to use MLR

- 1. Independence of observations of responses
- 2. Linearity of \*all\* variables linear relationship between response variable and each of the explanatory variables
- 3. Multicollinearity checked for does not mean we cannot use MLR, but we should be aware of how predictor/explanatory variables are related when quoting our results
- 4. Constant variance
- 5. Normality of Residuals
- 6. No influential points (outliers with strong leverage)

# Logistic Regression: A Morbid Example

Logistic regression is a GLM used to model a binary categorical variable using numerical and categorical predictors.

We assume a binomial distribution produced the outcome variable and we therefore want to model *p* the probability of success for a given set of predictors.

To finish specifying the Logistic model we just need to establish a reasonable link function that connects  $\eta$  to p. There are a variety of options but the most commonly used is the logit function.

$$logit(p) = log\left(\frac{p}{1-p}\right), \text{ for } 0 \le p \le 1$$

# Logistic Regression: A Morbid Example

Ok, so what does the totality of our model look like?

$$y_i \sim \mathsf{Binom}(p_i)$$
 
$$\eta = \beta_0 + \beta_1 x_1 + \dots + \beta_n x_n$$
 
$$\mathsf{logit}(p) = \eta$$

From which we back out the probability of survival based on parameters 1-n, for the *i*th observation:

$$p_{i} = \frac{\exp(\beta_{0} + \beta_{1}x_{1,i} + \dots + \beta_{n}x_{n,i})}{1 + \exp(\beta_{0} + \beta_{1}x_{1,i} + \dots + \beta_{n}x_{n,i})}$$

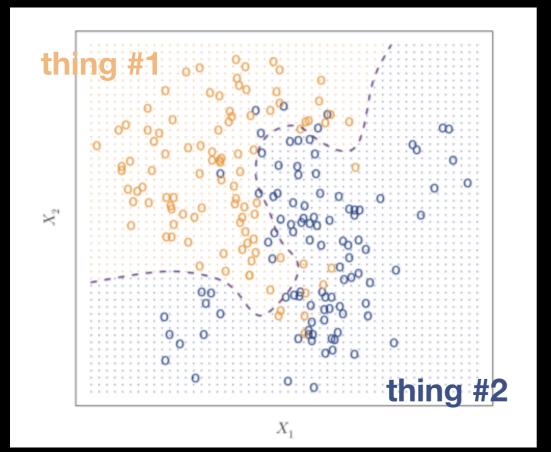
## So far...

$$\hat{y} = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + \dots + \beta_n x_n$$

So far we've been saying:
"I have a basic idea of what the functional form of y looks like (a line or a logit) - computer go find the parameters of that functional form.

This is nice because we have some hope of gaining intuition from our models.

# Now we classify...



"I want to know where thing #1 and thing #2 live in some 2D space - computer, go figure out the boundary between these two things and let me know"

This is nice because we don't have to assume some model beforehand.

# Bias-Variance Trade-Off (First Glance)

$$E\left(y_0 - \hat{f}(x_0)\right)^2 = \operatorname{Var}(\hat{f}(x_0)) + \left[\operatorname{Bias}(\hat{f}(x_0))\right]^2 + \operatorname{Var}(\epsilon)$$

mean square error if we kept estimating response variable y by our fitted function of our explanatory variables, f(x) with different sample datasets at point x<sub>0</sub>

Inherent error in our measurements

Inherent error (bias) in the fact that any model is only an approximation to reality

how much our function, f, changes if we use a different random sample (variance)

Fit model to data with a choice of parameters (e.g. degree of polynomial)





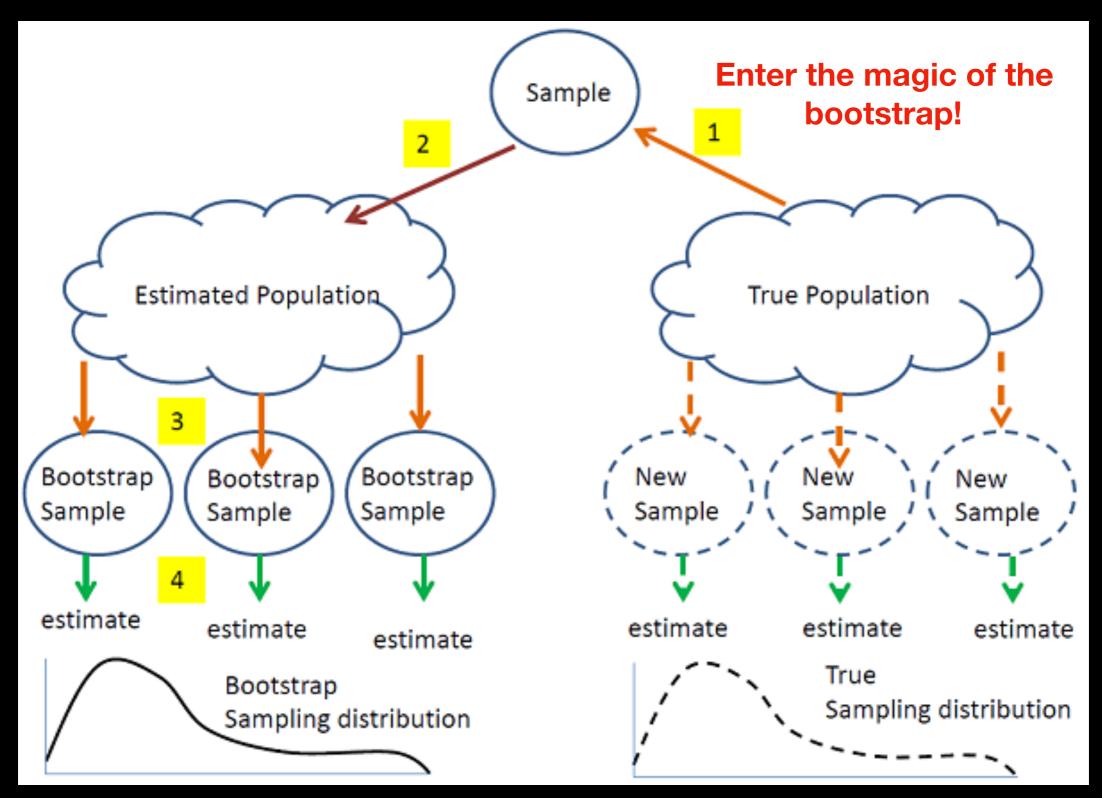
Calculate the mean square error (MSE) of subset data and model

$$MSE = \frac{1}{n} \sum_{i=1}^{n} (y_i - \hat{f}(x_i))^2$$

Repeat for a bunch of subsets of data

Repeat for different model parameters

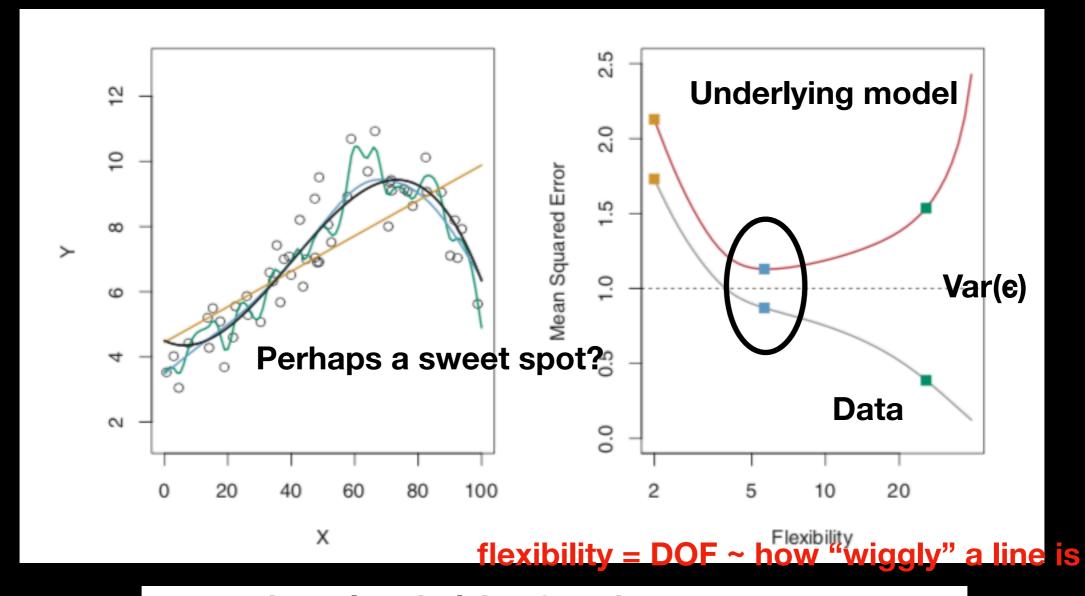
# **Bootstrapping**



Distribution of means, proportions, etc

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- Actual underlying function y
  o Simulated data with added error (€)
   Linear fit
   Low "flexibility" smooth spline
  - Low "flexibility" smooth spline
    High "flexibility" smooth spline

fits data well, but underlying model badly

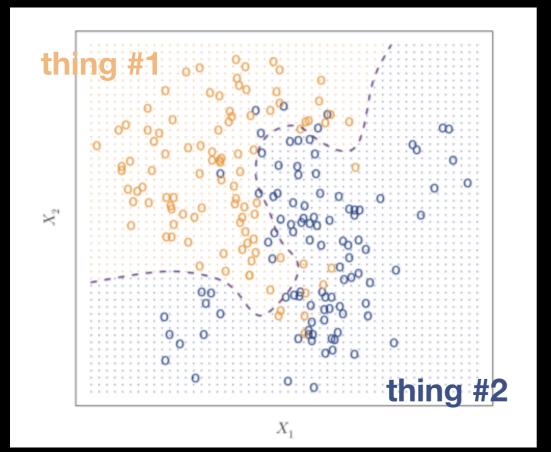
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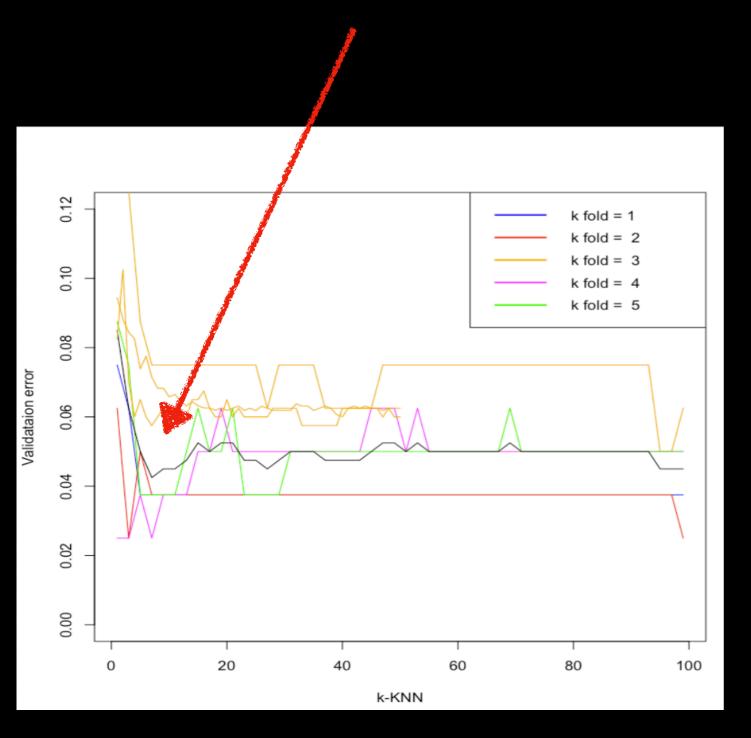
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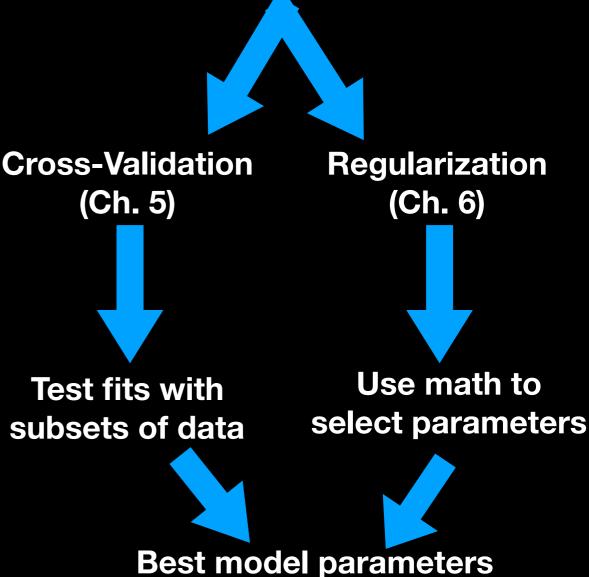


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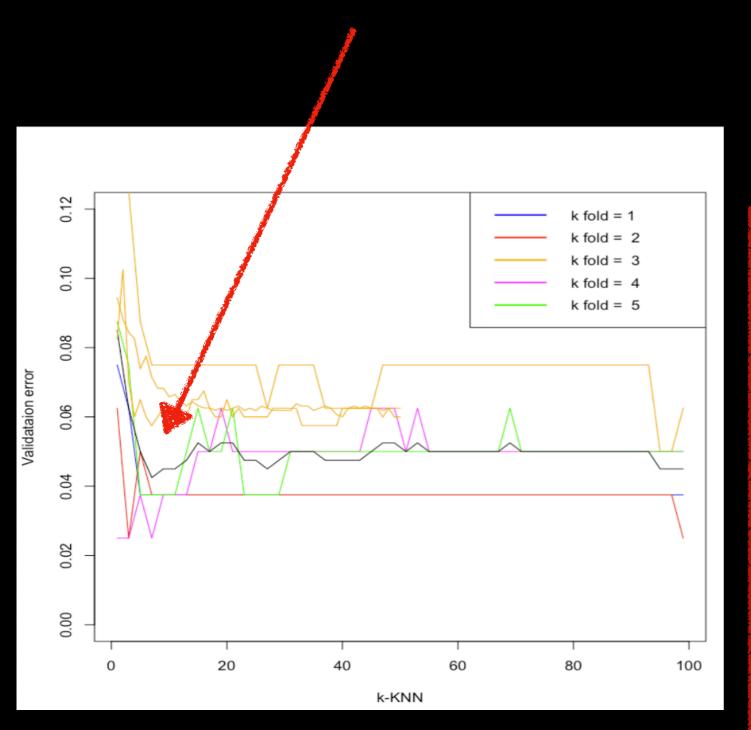
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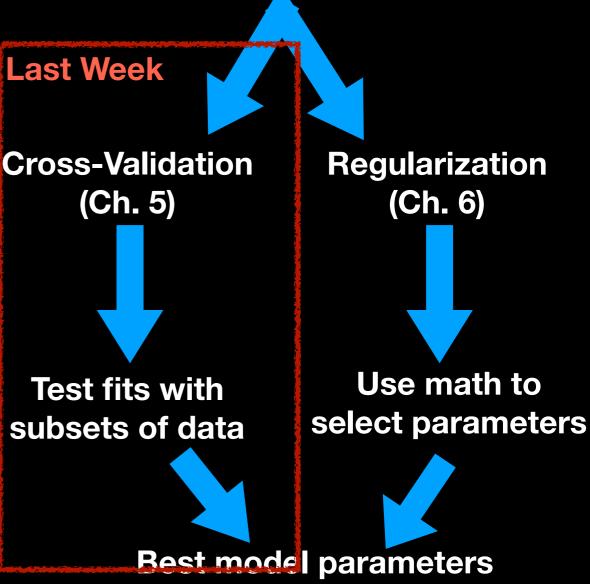
### **Best KNN k is about here**



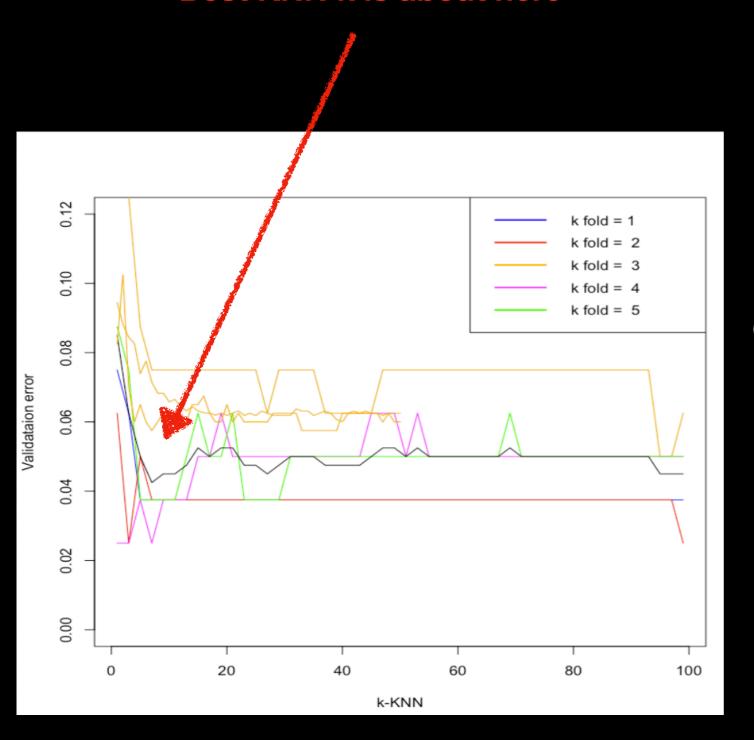


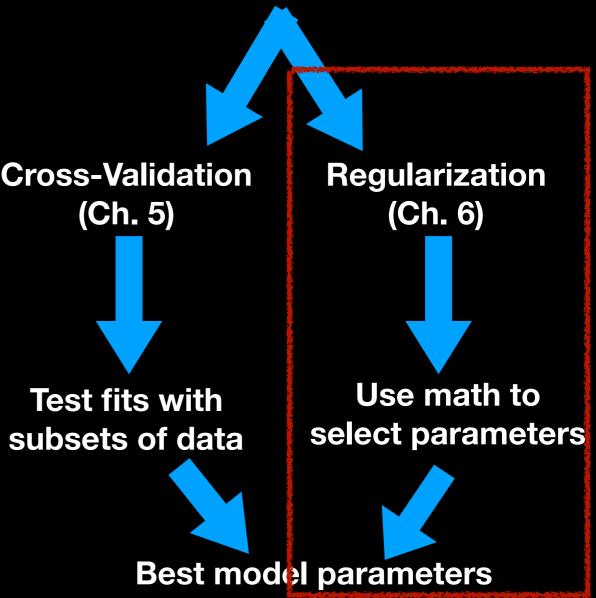
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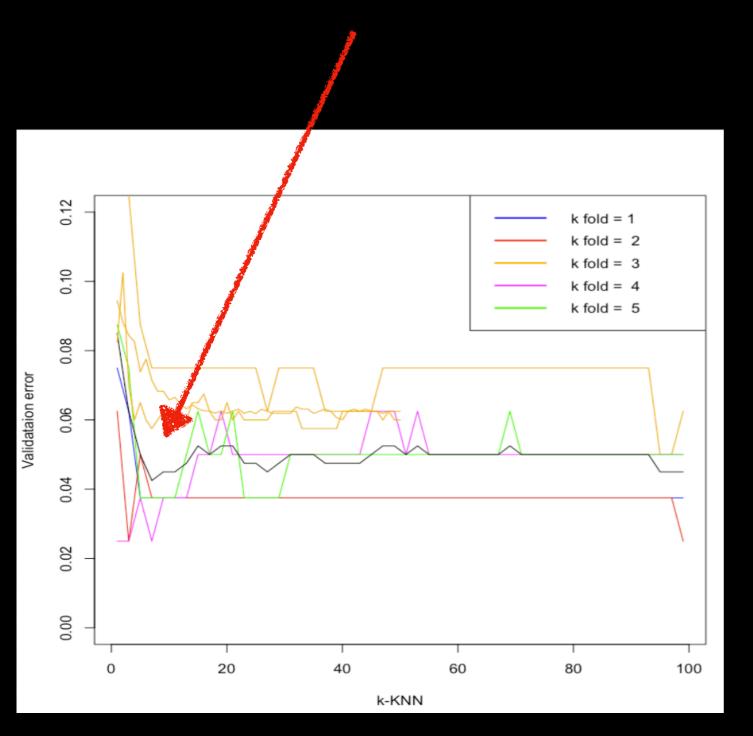


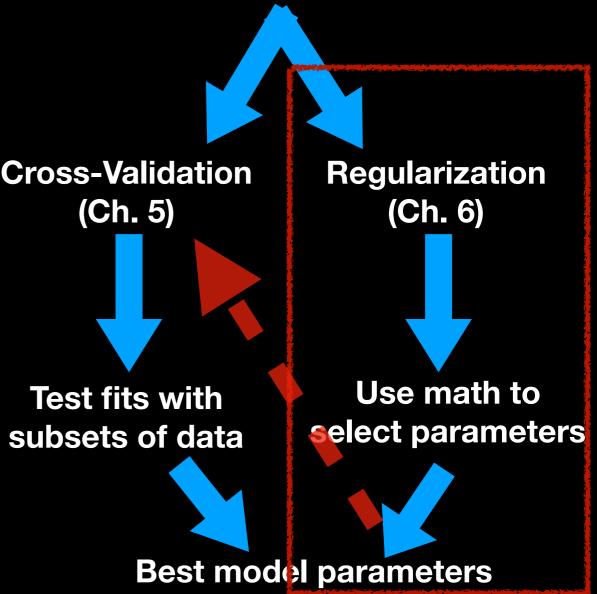
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Prediction Accuracy: especially when p > n, to control the variance.

Model Interpretability: By removing irrelevant features — that is, by setting the corresponding coefficient estimates to zero — we can obtain a model that is more easily interpreted. We will present some approaches for automatically performing feature selection.

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### Subset selection

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... we actually essentially already covered this!

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A few more details...

We identify a subset of the p predictors (of k possible) that we believe to be related to the response. We then fit a model using least squares on the reduced set of variables.

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  - (a) Fit all  $\binom{p}{k}$  models that contain exactly k predictors.
  - (b) Pick the best among these P models, and call it  $M_k$ . Here k best is defined as having the smallest RSS, or equivalently largest  $R^2$ .

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  (a) Fit all  $\binom{p}{k}$  mode is that containing this number can get big  $\sim 2^p$  p=2: 4 (including null) p=3: 8...
  - (b) Pick the best among these p=10: 1024, lk. Here k best is defined as having the smallest RSS, or equivalently largest R<sup>2</sup>.
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- Backward selection requires that the number of samples n is larger than the number of variables p (so that the full model can be fit). In contrast, forward stepwise can be used even when n < p, and so is the only viable subset method when p is very large.

How do we quantify how "good" each of our forward/backward selected models are?

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d predictors, n data points

estimate from  $SE^2/n = \sigma^2$  or RSE

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$$AIC = \frac{1}{n\hat{\sigma}^2} \left( RSS + 2d\hat{\sigma}^2 \right)$$

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BIC = 
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# Optional R notes.

Lets recall how we find a linear model:

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 Least squares minimization (i.e. minimize RSS)

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$$\sum_{i=1}^n \left(y_i-\beta_0-\sum_{j=1}^p \beta_j x_{ij}\right)^2 + \lambda \sum_{j=1}^p \beta_j^2 = \mathrm{RSS} + \lambda \sum_{j=1}^p \beta_j^2$$
 Ridge

$$\sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^p |\beta_j| = \text{RSS} + \lambda \sum_{j=1}^p |\beta_j|.$$
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where  $\lambda \ge 0$  is a tuning parameter, to be determined separately.

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 Ridge

Don't panic, its just a bit more math (and we'll get R to do it for us)

Lasso

$$\sum_{i=1}^{n} \left( y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{p} |\beta_j| = \text{RSS} + \lambda \sum_{j=1}^{p} |\beta_j|$$

where  $\lambda \ge 0$  is a tuning parameter, to be determined separately.

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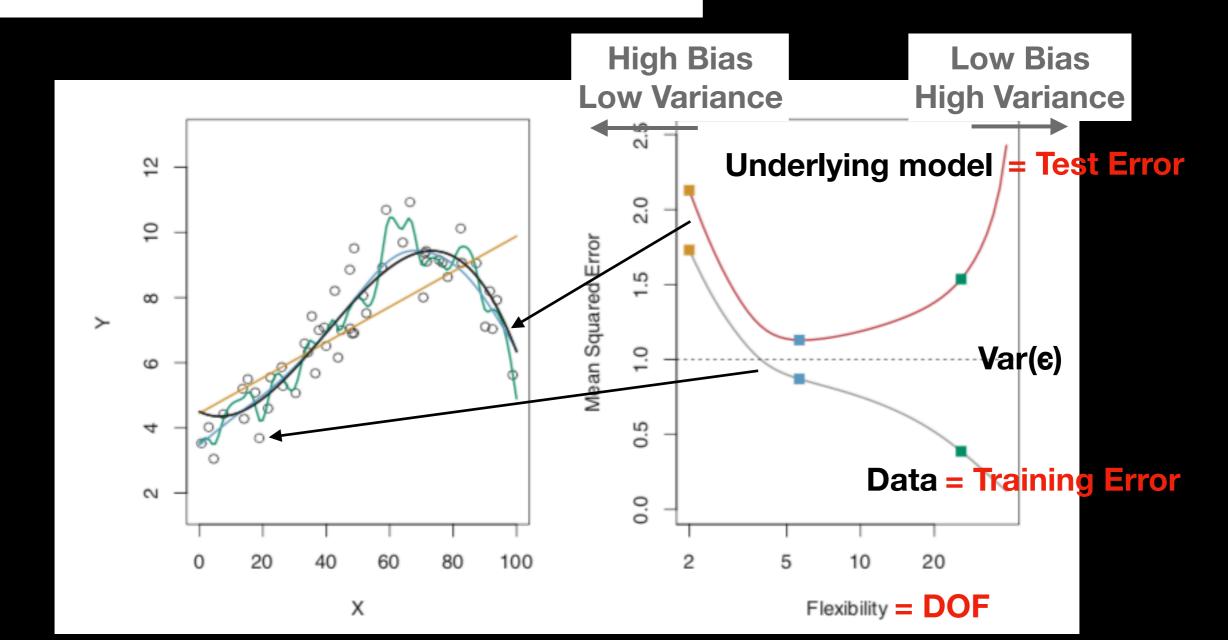
Why would we make our lives more complicated like this?

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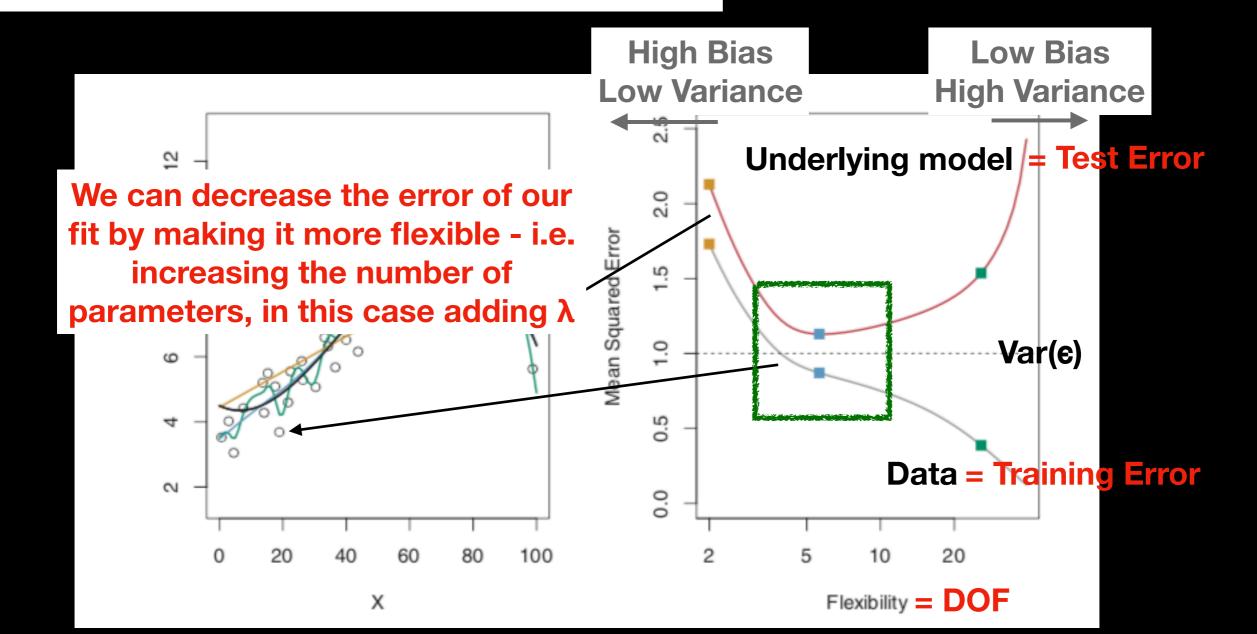


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"Shrinkage" of less important parameters to zero

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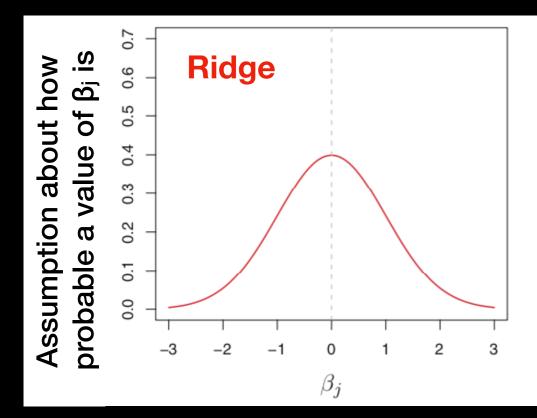
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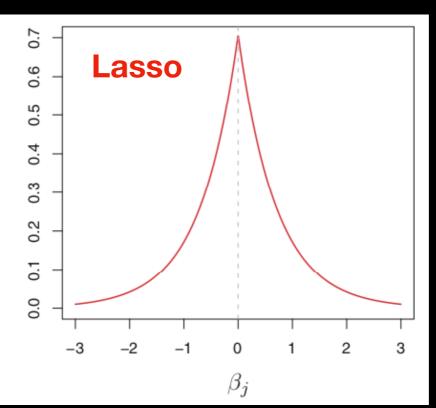
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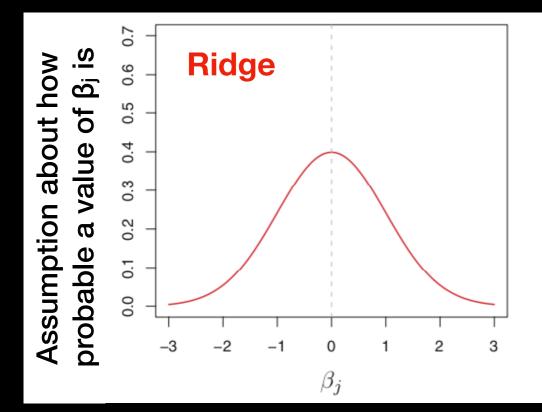
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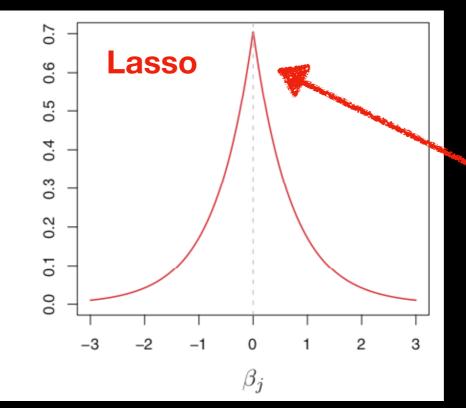
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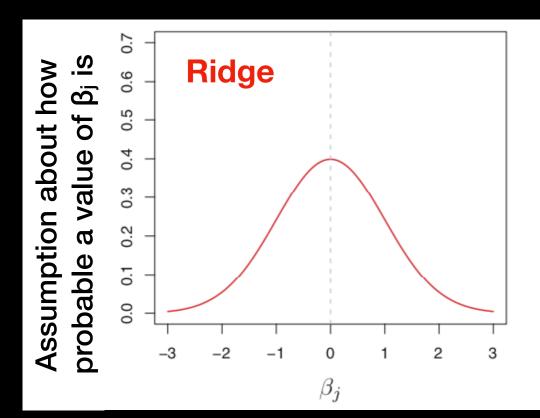
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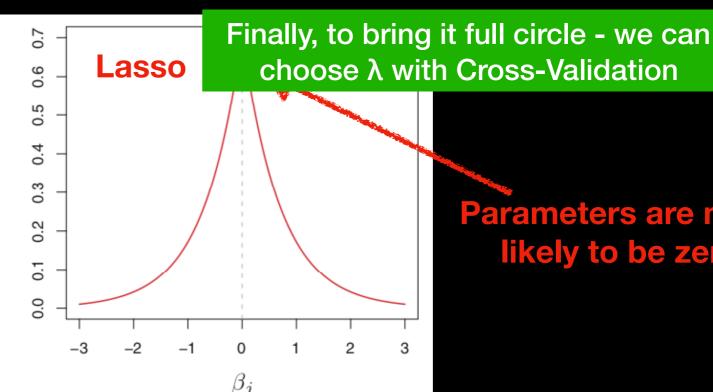
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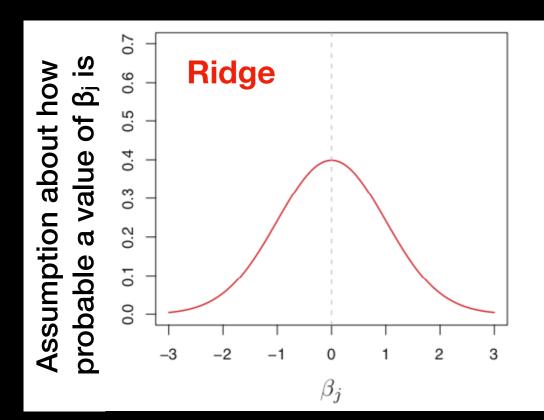
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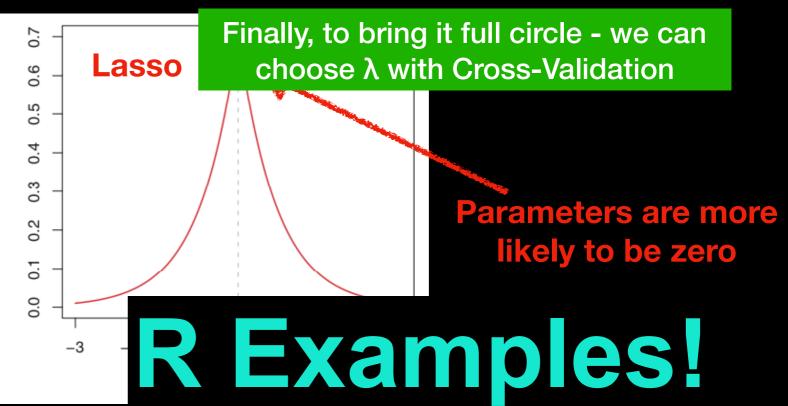
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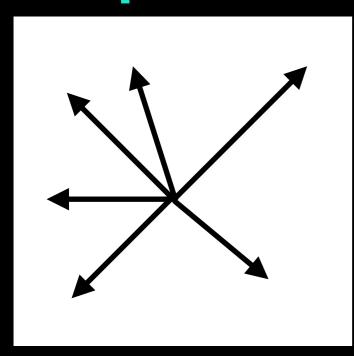




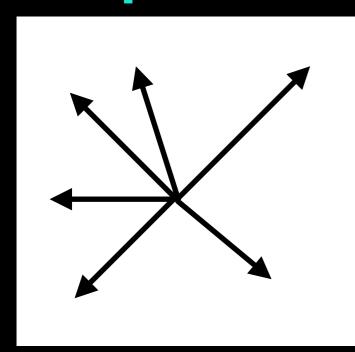
# **Unsupervised Learning: An intro to Principle Component Analysis**

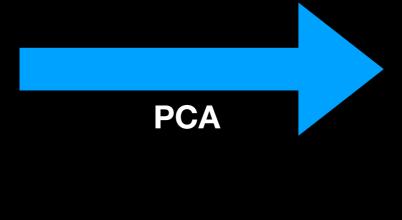
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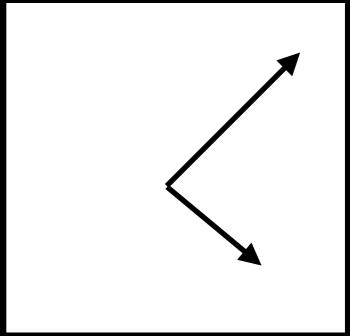
... what do we do when we don't know anything



How many vectors do I need to define a 2D space?

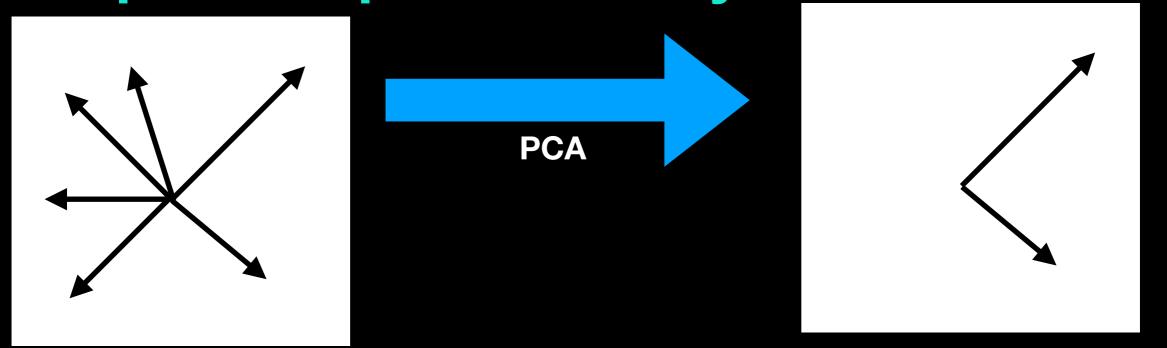






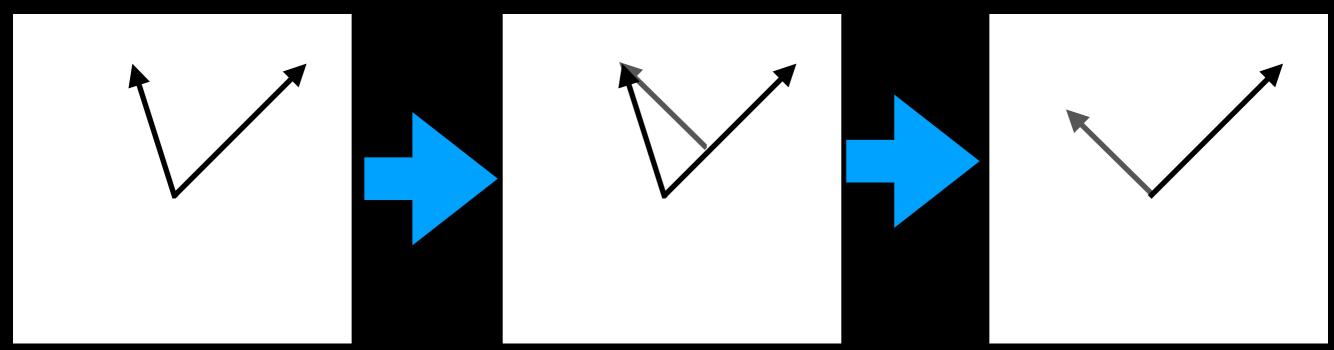
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Minimum number of vectors to define a space in a certain number of dimensions

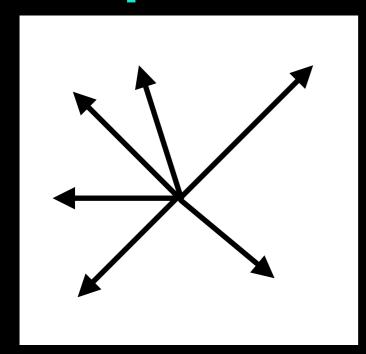


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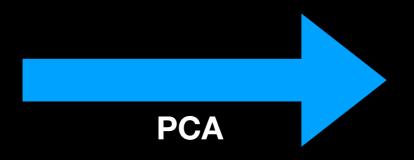
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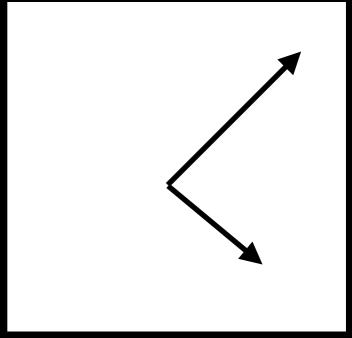
**Constructing orthogonal vectors** 



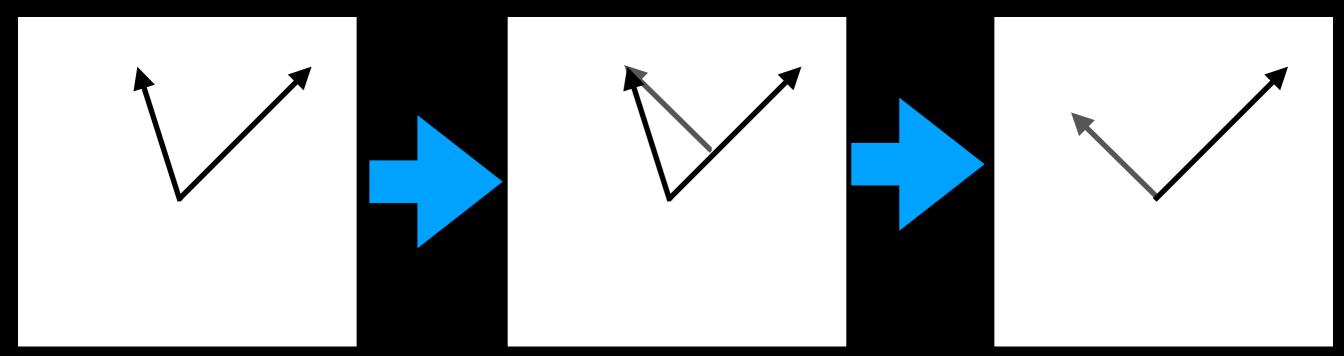
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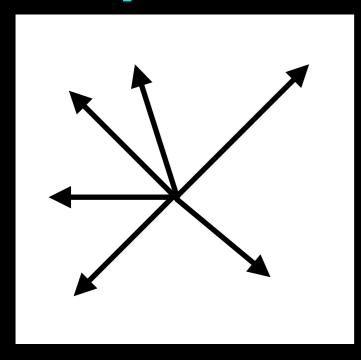


The "dimensions" of our space is dictated by the number of parameters we have

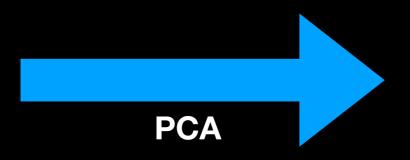


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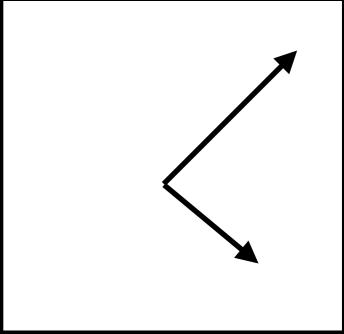




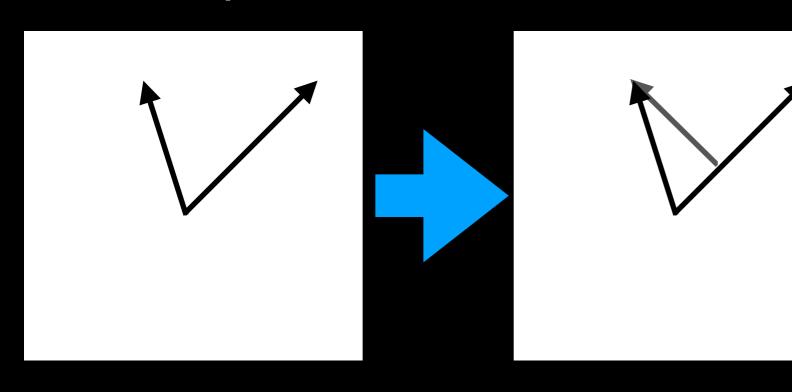
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Minimum number of vectors to define a space in a certain number of dimensions



Aside: these are also called "eigenvectors" and are used a lot in physics - for example to express states of atoms in quantum mechanics

**Constructing orthogonal vectors** 

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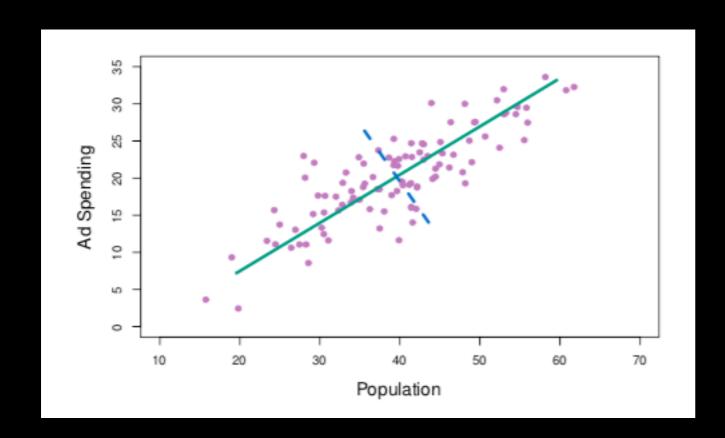
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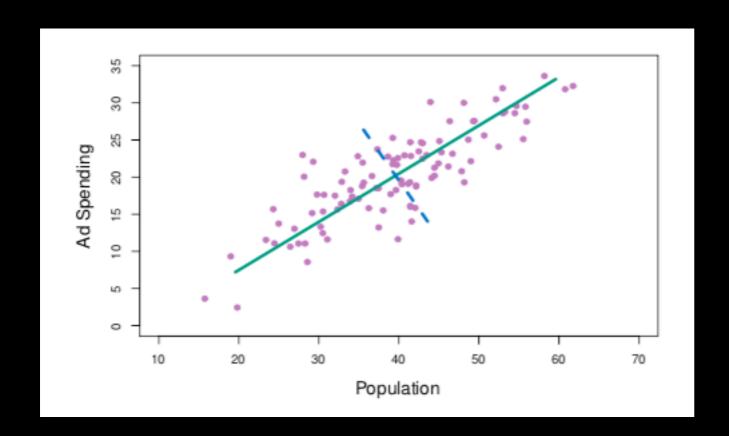
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Quick R example!

