Jason Seawright

j-seawright@northwestern.edu

Jan. 16, 2024

# Regression in Multi-Method Research

 We are interested in the effects of a dichotomous treatment (i.e., independent variable).

- We are interested in the effects of a dichotomous treatment (i.e., independent variable).
- This variable can be written as  $D_i = (t, c)$ .

For a given case, i, we either observe
 D<sub>i</sub> = t or D<sub>i</sub> = c. If D<sub>i</sub> = t, let us denote
 the value of the dependent variable as y<sub>i,t</sub>.
 If D<sub>i</sub> = c, let us denote the value of the
 dependent variable as y<sub>i,c</sub>

- The causal effect of *D* on *y* is:
  - $\bullet$   $y_{i,t} y_{i,c}$

### The Average Treatment Effect

• Sometimes, we are interested in developing an estimate of the effect of *D* on *y* in some population, from which we have a random sample (or even the whole population) split randomly into treatment and control cases.

#### The Average Treatment Effect

- Sometimes, we are interested in developing an estimate of the effect of *D* on *y* in some population, from which we have a random sample (or even the whole population) split randomly into treatment and control cases.
- Here, interest focuses on the "average treatment effect":
  - $\bullet \ E(y_{i,t}) E(y_{i,c})$



## **Experiments and Causal Inference**

• Under random assignment, the set of cases where  $D_i = t$  produces a random sample from the population of  $y_t$ . Likewise, the set of cases where  $D_i = c$  produces a random sample from the population of  $y_c$ . Thus:

•  $E(y_{i,t}|D_i = t) = E(y_{i,t})$ 

### **Experiments and Causal Inference**

• Under random assignment, the set of cases where  $D_i = t$  produces a random sample from the population of  $y_t$ . Likewise, the set of cases where  $D_i = c$  produces a random sample from the population of  $y_c$ . Thus:

- $\bullet \ E(y_{i,t}|D_i=t)=E(y_{i,t})$
- $\bullet \ E(y_{i,c}|D_i=c)=E(y_{i,c})$

### Experiments and Causal Inference

• Under random assignment, the set of cases where  $D_i = t$  produces a random sample from the population of  $y_t$ . Likewise, the set of cases where  $D_i = c$  produces a random sample from the population of  $y_c$ . Thus:

- $E(y_{i,t}|D_i = t) = E(y_{i,t})$
- $\bullet \ E(y_{i,c}|D_i=c)=E(y_{i,c})$
- $E(y_{i,t}) E(y_{i,c}) = E(y_{i,t}|D_i = t) E(y_{i,c}|D_i = c)$



Recall that  $D_i$  as an indicator of treatment assignment in an experiment. Let's change the coding, such that  $D_i$  equals 1 for treatment cases and 0 for control cases. Then we can write:

$$Y_i = D_i Y_{i,t} + (1 - D_i) Y_{i,c}$$
 (1)

Recall that  $D_i$  as an indicator of treatment assignment in an experiment. Let's change the coding, such that  $D_i$  equals 1 for treatment cases and 0 for control cases. Then we can write:

$$Y_i = D_i Y_{i,t} + (1 - D_i) Y_{i,c}$$
 (1)

$$Y_i = Y_{i,c} + D_i(Y_{i,t} - Y_{i,c})$$
 (2)

$$Y_{i} = \bar{Y_{i,c}} + D_{i}(\bar{Y_{i,t}} - \bar{Y_{i,c}}) + [Y_{i,c} - \bar{Y_{i,c}} + D_{i}(Y_{i,t} - \bar{Y_{i,t}} - Y_{i,c} + \bar{Y_{i,c}})]$$
(3)

Suppose we use OLS to estimate:

$$Y_i = \beta_0 + \beta_1 D_i + \epsilon_i$$

Will this work?

$$\hat{\beta}_1 = \frac{cov(D, Y)}{var(D)} = \frac{\sum (D - \bar{D})(Y - \bar{Y})}{\sum (D - \bar{D})^2}$$

$$\hat{eta}_1 = rac{cov(D, Y)}{var(D)} = rac{\sum (D - \bar{D})(Y - \bar{Y})}{\sum (D - \bar{D})^2}$$

Let  $\pi$  equal the proportion of cases assigned to the treatment group.

$$\hat{eta}_1 = rac{cov(D, Y)}{var(D)} = rac{\sum (D - \bar{D})(Y - \bar{Y})}{\sum (D - \bar{D})^2}$$

Let  $\pi$  equal the proportion of cases assigned to the treatment group.

$$\hat{eta}_1 = rac{\sum (D-\pi)(Y-ar{Y})}{N\pi(1-\pi)}$$

|ロ ト 4 団 ト 4 豆 ト 4 豆 ト 9 Q C

$$\hat{eta}_1 = rac{\sum_{D_i=1} (1-\pi)(Y-ar{Y}) - \sum_{D_i=0} \pi(Y-ar{Y})}{N\pi(1-\pi)}$$

$$\hat{\beta}_1 = \frac{\sum_{D_i=1} (1-\pi)(Y-\bar{Y}) - \sum_{D_i=0} \pi(Y-\bar{Y})}{N\pi(1-\pi)}$$

$$\hat{\beta}_{1} = \frac{\sum_{D_{i}=1}(Y_{i} - E(Y_{i}))}{N\pi} - \frac{\sum_{D_{i}=0}(Y_{i} - E(Y_{i}))}{N(1-\pi)}$$



$$\hat{eta}_1 = rac{\sum_{D_i=1} Y_i}{N\pi} - rac{\sum_{D_i=0} Y_i}{N(1-\pi)}$$

$$\hat{\beta}_1 = \frac{\sum_{D_i=1} Y_i}{N\pi} - \frac{\sum_{D_i=0} Y_i}{N(1-\pi)}$$

 $N\pi$  is just the number of cases where  $D_i=1$ , and  $N(1-\pi)$  is the number of cases where  $D_i=0$ . So the last expression simplifies to:

$$\hat{\beta}_1 = \bar{Y}_{i,D_i=1} - \bar{Y}_{i,D_i=0}$$

Multivariate regression for experiments is *not* guaranteed to be unbiased in the way that bivariate regression is. Multivariate regression can be (even very badly) biased if:

Multivariate regression for experiments is *not* guaranteed to be unbiased in the way that bivariate regression is. Multivariate regression can be (even very badly) biased if:

 Some control variable W is included in the model that is in any way caused by the treatment, or

Multivariate regression for experiments is *not* guaranteed to be unbiased in the way that bivariate regression is. Multivariate regression can be (even very badly) biased if:

 The causal effect of interest is highly heterogeneous across categories of the control variables.

$$Y_{i} = E(Y_{i,c}) + D_{i}\{E(Y_{i,t}) - E(Y_{i,c})\} + [Y_{i,c} - E(Y_{i,c})] + D_{i}([Y_{i,t} - E(Y_{i,t})] - [Y_{i,c} - E(Y_{i,c})])$$

$$= \mu_{0} + D_{i}(\mu_{1} - \mu_{0}) + \{\nu_{0} + D_{i}(\nu_{1} - \nu_{0})\}$$

• A collection of control variables  $\mathbb{X}$  will allow regression to produce an unbiased estimate of  $(\mu_1 - \mu_0)$  when:

- A collection of control variables  $\mathbb{X}$  will allow regression to produce an unbiased estimate of  $(\mu_1 \mu_0)$  when:
  - **1** *D* is uncorrelated with  $\{\nu_0 + D_i(\nu_1 \nu_0)\}$  within each group defined by  $\mathbb{X}$ , and

- A collection of control variables  $\mathbb{X}$  will allow regression to produce an unbiased estimate of  $(\mu_1 \mu_0)$  when:
  - ① D is uncorrelated with  $\{\nu_0 + D_i(\nu_1 \nu_0)\}$  within each group defined by  $\mathbb{X}$ , and
  - 2 the residual causal effect is not correlated with  $\mathbb{X}$ , and

- A collection of control variables  $\mathbb{X}$  will allow regression to produce an unbiased estimate of  $(\mu_1 \mu_0)$  when:
  - ① D is uncorrelated with  $\{\nu_0 + D_i(\nu_1 \nu_0)\}$  within each group defined by  $\mathbb{X}$ , and
  - 2 the residual causal effect is not correlated with  $\mathbb{X}$ , and
  - $oldsymbol{3}$  a fully flexible parameterization of  $\mathbb X$  is used.

$$Y_i = \mu_0 + D_i(\mu_1 - \mu_0) + \{\nu_0 + D_i(\nu_1 - \nu_0)\}$$

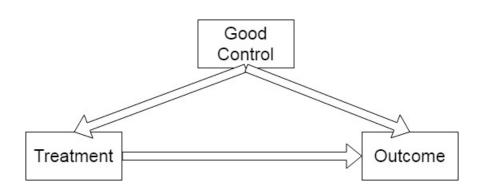
$$Y_i = \mu_0 + D_i(\mu_1 - \mu_0) + X_i\beta + \{\nu_0^* + D_i(\nu_1^* - \nu_0^*)\}$$

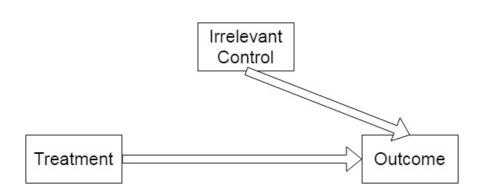
These conditions imply that:

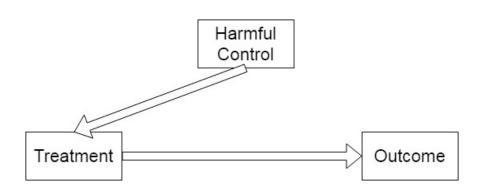
- These conditions imply that:
  - 1 No element of X is on any causal path from D to y, and

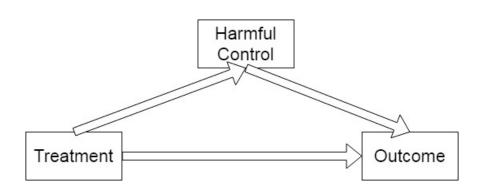
- These conditions imply that:
  - 1 No element of X is on any causal path from D to y, and
  - 2 no element of X is caused by D or any of its unmeasured causes and some other unmeasured cause of y, and

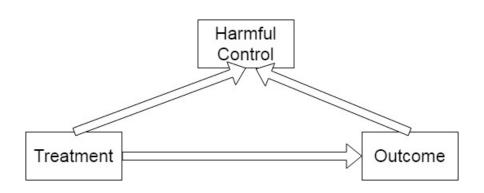
- These conditions imply that:
  - 1 No element of X is on any causal path from D to y, and
  - 2 no element of X is caused by D or any of its unmeasured causes and some other unmeasured cause of y, and
  - 3 all causes of D that are also causes of y have some element of X somewhere on the causal path from the unmeasured initial cause to either D or y.











#### Measurement Error



## The Legacy of Political Violence across Generations •• •

Noam Lupu Vanderbilt University
Leonid Peisakhin New York University-Abu Dhabi

Table A8. Effects of first-generation victimization on third-generation attitudes and behaviors

	Victimization				
Dependent variable	effect	SE	Obs.	Families	R <sup>2</sup>
Political identities					
In-group attachment	0.058*	0.022	957	298	0.015
Victimhood	$0.080^{*}$	0.023	801	286	0.028
Threat perception	0.970*	0.047	951	298	0.009
Radicalism					
Support for radical Islam	0.009	0.058	645	267	0.000
Religiosity	0.049	0.031	988	298	0.005
Crimean Tatar issues					
Support for CT leaders	0.178*	0.050	724	285	0.032
Celebrate CT holiday	0.038*	0.019	947	297	0.009
Attitudes toward Russia					
Support for Chechen rebels	0.101*	0.021	639	270	0.046
Support for annexation	-0.074*	0.026	778	285	0.025
Pro-Russia vote choice	-0.089*	0.044	964	296	0.008
Political engagement					
Turnout	0.094*	0.042	974	298	0.010
Willingness to participate	0.120*	0.044	727	287	0.015
[Past participation]	0.080*	0.040	996	298	0.007

<sup>\*</sup> p<0.05

Notes: Linear regression estimates with standard errors clustered by family. Constant terms not shown.

**Table A9**. Effects of first-generation victimization on third-generation identities, attitudes, and behaviors, with pre-deportation controls

100	Victimization				
Dependent variable	effect	SE	Obs.	Families	R <sup>2</sup>
Political identities					
In-group attachment	0.052*	0.026	699	214	0.058
Victimhood	0.074*	0.026	605	210	0.120
Threat perception	0.163*	0.053	687	214	0.033
Radicalism					
Support for radical Islam	-0.002	0.066	467	191	0.059
Religiosity	0.033	0.037	715	214	0.033
Crimean Tatar issues					
Support for CT leaders	0.187*	0.061	528	207	0.057
Celebrate CT holiday	0.039	0.024	686	214	0.064
Attitudes toward Russia					
Support for Chechen rebels	0.101*	0.024	469	195	0.087
Support for annexation	-0.064*	0.028	569	205	0.107
Pro-Russia vote choice	-0.087	0.055	707	213	0.051
Political engagement					
Turnout	0.096	0.050	706	214	0.054
Willingness to participate	0.096	0.050	534	207	0.031

<sup>\*</sup> p<0.05

Notes: Linear regression estimates with standard errors clustered by family. Constant terms and pre-deportation control variables (pre-Soviet household wealth, dekulakized, Soviet opposition, pre-Soviet religiosity, pre-deportation region, and deportation republic) not shown.

**Table A10**. Effects of first-generation victimization on third-generation identities, attitudes, and behaviors, with third-generation demographic controls

	Victimization				
Dependent variable	effect	SE	Obs.	Families	R <sup>2</sup>
Political identities					
In-group attachment	$0.054^*$	0.022	940	298	0.030
Victimhood	0.072*	0.024	788	285	0.045
Threat perception	0.095*	0.047	934	298	0.017
Radicalism					
Support for radical Islam	0.016	0.057	635	267	0.017
Religiosity	0.054 0.031 970	298	0.028		
Crimean Tatar issues					
Support for CT leaders	0.164*	0.050	710	285	0.045
Celebrate CT holiday	0.037	0.019	927	297	0.018
Attitudes toward Russia					
Support for Chechen rebels	0.102*	0.021	628	269	0.058
Support for annexation	-0.071*	0.026	765	285	0.038
Pro-Russia vote choice	-0.091*	0.043	944	296	0.024
Political engagement					
Turnout	$0.096^*$	0.041	956	298	0.028
Willingness to participate	0.126*	0.042	709	286	0.065

p<0.05

Notes: Linear regression estimates with standard errors clustered by family. Constant terms and demographic control variables (wealth index, education, age, gender, and marital status) not shown

Table A6. Endogeneity tests

Variable	Linear regression	Ordered probit		
Pre-Soviet household wealth	0.077	0.079		
rie-soviet ilouseiloid weartii	(0.068)	(0.072)		
Dekulakized	0.040	0.064		
Dekulakized	(0.152)	(0.158)		
Carriet annualtian	0.132)	0.116		
Soviet opposition	(0.092)			
D. C. i.e. Histories	0.092)	(0.097)		
Pre-Soviet religiosity		0.159		
	(0.119)	(0.124)		
Pre-deportation region				
Southwest (reference group)				
	122			
Southeast	-0.030	-0.039		
	(0.235)	(0.248)		
Northwest	0.137	0.172		
	(0.248)	(0.238)		
Northeast	0.046	0.038		
	(0.457)	(0.510)		
Deportation republic				
Uzbekistan (reference group)	V22			
Other Central Asia	-0.269	-0.280		
	(0.312)	(0.336)		
Russia	-0.425	-0.479		
	(0.252)	(0.285)		
Observations	212	212		
R <sup>2</sup> /Pseudo-R <sup>2</sup>	0.041	0.015		

Notes: Estimates with robust standard errors in parentheses. Constant term not shown.



## Heterogeneity

 If the causal effect is not constant across all cases, regression will not give a consistent estimate of the average treatment effect.

## Heterogeneity

- If the causal effect is not constant across all cases, regression will not give a consistent estimate of the average treatment effect.
- Instead, it estimates a covariance-adjusted weighted average of cases' treatment effects.

#### Aronow and Samii 2016

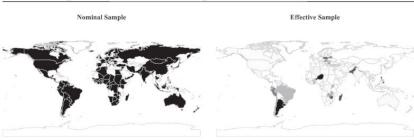
$$\hat{\beta} \stackrel{p}{\rightarrow} \frac{\mathrm{E}\left[w_i \tau_i\right]}{\mathrm{E}\left[w_i\right]}, \text{ where } w_i = (D_i - \mathrm{E}\left[D_i | X_i\right])^2$$

#### Aronow and Samii 2016

$$E[w_i|X_i] = Var[D_i|X_i].$$

#### Aronow and Samii 2016

FIGURE 1 Example of nominal and effective samples from Jensen (2003)



Note: On the left, the shading shows countries in the nominal sample for Jensen (2003) estimate of the effects of regime type on FDI. On the right, darker shading indicates that a country contributes more to the effective sample, based on the panel specification used in estimation.

## Weights for Correct Causal Inference

It might sometimes be possible to get sample average treatment effects by weighting cases by  $1/w_i$ . But this may break down if some cases have weights equal to, or very close to, zero.

## Weights for Correct Causal Inference

As long as all weights are strictly greater than zero, then the population causal inference can be derived from a regression where all control variables are centered around their means and where all possible interaction terms between control variables and the treatment are included.

36 / 49

## Is Democracy Good for the Economy?

## Democracy and Growth: A Case Study

- Which dependent variable should we use?
  - For this literature, there isn't much ambiguity: economic growth is the outcome of interest.

## Democracy and Growth: A Case Study

- 2. How should the dependent variable be measured, and perhaps transformed?
  - Once again, there is little controversy here: annual measures or period-averages.

# Democracy and Growth: A Case Study

- 2. How should the dependent variable be measured, and perhaps transformed?
  - Once again, there is little controversy here: annual measures or period-averages.
  - And yet...
    - Should we focus on short-term or long-term growth?
    - Should we use a transformation to limit the influence of extreme years, or not?

- 3. Which independent variables should we use?
  - Scholars have not achieved any degree of consensus regarding the proper set of control variables for a democracy-and-growth regression.

#### 3. Which independent variables should we use?

The section of the se	Barro (1997)	Feng (1997)	Durham (1999)	Gasiorowski (2000)	Przeworski et al. (2000)	Kurzman et al. (2002)	Baum and Lake (2003)
Capital stock			Tarabian Tarabian	JOHN VIII -	Х		1 -11 - 111)
Education/literacy	X	X	X	X	Λ	X	Nr. 18
GDP	X	X	X	X		X	X
Government consumption	X	har zi	X	in A		X	X
Inflation	X	X		X			
Investment		X	X	X		v	77
Labor supply			5.110.70	X	v	Λ	X
Life expectancy	X			24	Λ	v	X
Money supply growth				X		Λ	A
Peaceful unrest				X			
Population growth	X		X	A		v	
Population size			1			Λ	V 32
Regime change		X					A
Region dummies	X		X				
Rule of law	X						
Terms of trade	X						
Гrade		X	X				
Violent unrest			17 18 1	X		X	
Wage growth				X		Λ	

- 3. Which independent variables should we use?
  - Consider education.

- 3. Which independent variables should we use?
  - Consider education.
    - Education, strangely enough, has a negative relationship with economic growth.

- 3. Which independent variables should we use?
  - Consider education.
    - Education, strangely enough, has a negative relationship with economic growth.
    - Education has a positive relationship with democracy.

- 3. Which independent variables should we use?
  - Consider education.
    - In a growth regression controlling only for lagged GDP, democracy has a significant coefficient of -0.00151.

- 3. Which independent variables should we use?
  - Consider education.
    - In a growth regression controlling only for lagged GDP, democracy has a significant coefficient of -0.00151.
    - In a similar regression also controlling for primary school attainment, democracy has a significant coefficient of 0.04371.

- 3. Which independent variables should we use?
  - Consider education.
    - Is education a cause or a consequence of democracy?

- 4. What functional form should the equation take?
  - In practice, the chosen functional form is almost always linear, but sometimes quadratic.
  - Interaction terms are rarely considered.
  - In particular, it is generally assumed that democracy has a single, universal causal effect on economic growth.

**IPSA 2024** 

5. Should observations be treated as independent? If not, then how?

- 5. Should observations be treated as independent? If not, then how?
  - Many studies assume total independence.
     Some assume independence conditional on constant country and/or period effects.

- 5. Should observations be treated as independent? If not, then how?
  - Many studies assume total independence.
     Some assume independence conditional on constant country and/or period effects.
  - Models exist that could consider diffusion of economic growth patterns along geographic or trade-network lines, for example.

- 6. What kind of interpretations can the resulting coefficients reasonably be given?
  - Strictly speaking, the coefficient estimates represent the slope, in one direction, of the best-fitting hyperplane through the observed data.

- 6. What kind of interpretations can the resulting coefficients reasonably be given?
  - Analysts virtually always interpret the coefficients as if they provided a description of causal counterfactuals applicable to each country in the data set.

- 6. What kind of interpretations can the resulting coefficients reasonably be given?
  - Less ambitious substantive interpretations are probably appropriate, but it is hard to know exactly what those interpretations should be. More theoretical development and more attention to mechanisms might help us guess what the various available coefficient estimates might mean.