

Exercises: inference in the linear Gaussian model

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1 Inferring location of a static submarine from its sonar measurements

A static submarine is located in a 2D planar surface deep in the sea. A priori, we model its unknown location with a 2D random variable \mathbf{z} with

$$p(\mathbf{z}) = \mathcal{N}(\mathbf{z}|\mu_z, \Sigma_z) \quad (1)$$

We obtain noisy measurements of the location of the submarine with a sonar. We represent a 2D sonar measurement with random variable \mathbf{y}_n , with

$$p(\mathbf{y}_n|\mathbf{z}) = \mathcal{N}(\mathbf{y}_n|\mathbf{z}, \Sigma_y) \quad (2)$$

- (a) Sample 100 a priori locations of the submarine (i.e., $\mathbf{z}_1, \dots, \mathbf{z}_{100}$) using a mean $\mu_z = [2, 3]^\top$, a standard deviation along the horizontal direction $\sigma_{zx} = 1.0$, a standard deviation along the vertical direction $\sigma_{zy} = 2.0$, and a correlation coefficient between the vertical and horizontal directions $\rho_z = 0.7$.

Plot μ_z , the 95% confidence ellipse for \mathbf{z} , the samples of the a-priori submarine location and verify that approximately 95% of these samples lie inside the 95% confidence ellipse.

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You may want to complete the script `doExSubmarine.a.py` to address this item.

- (b) Select the submarine location \mathbf{z}_1 generated in the previous item. Sample $N = 5$ sonar measurements, assuming the submarine is at location \mathbf{z}_1 (i.e., sample from $p(\mathbf{y}|\mathbf{z}_1)$ to obtain $\mathbf{y}_1, \dots, \mathbf{y}_N$). Use a standard deviation of 1.0 for the measurement noise along the horizontal and vertical directions, and assume that this noise is uncorrelated along these directions.

Plot \mathbf{z}_1 , the 95% confidence ellipse for the sonar measurements given that the submarine located at \mathbf{z}_1 , the sonar measurements samples, and verify that approximately 95% of these samples lie inside the 95% confidence band.

You may want to complete the script `doExSubmarine.b.py` to address this item.

- (c) derive a mathematical expression for the posterior of the submarine location, given sonar measurements; i.e., $p(\mathbf{z}|\mathbf{y}_1, \dots, \mathbf{y}_N)$.

Hints:

- The posterior of the submarine location given sonar measurements is proportional to the joint distribution of the submaring location and sonar measurements; i.e., $p(\mathbf{z}|\mathbf{y}_1, \dots, \mathbf{y}_N) = \frac{p(\mathbf{y}_1, \dots, \mathbf{y}_N, \mathbf{z})}{p(\mathbf{y}_1, \dots, \mathbf{y}_N)} = K p(\mathbf{y}_1, \dots, \mathbf{y}_N, \mathbf{z})$, where K is a value that does not depend on \mathbf{z} . Thus, to obtain the posterior we can just keep the terms of the joint that depend on \mathbf{z} and normalize the resulting expression to integrate to one.
- The joint is the product of the likelihood and the prior; i.e., $p(\mathbf{y}_1, \dots, \mathbf{y}_N, \mathbf{z}) = p(\mathbf{y}_1, \dots, \mathbf{y}_N|\mathbf{z})p(\mathbf{z})$. Thus, to keep the terms of the joint that depend on \mathbf{z} , we should just keep the term of the likelihood that depend on \mathbf{z} and combine the result with the prior.
- As shown in Claim. ??, the terms of the likelihood that depend on \mathbf{z} are proportional to a Gaussian distribution with mean \mathbf{z} and covariance $\frac{1}{N}\Sigma$; i.e., $p(\mathbf{y}_1, \dots, \mathbf{y}_N|\mathbf{z}) = K \mathcal{N}(\bar{\mathbf{y}}|\mathbf{z}, \frac{1}{N}\Sigma_y)$, where K is a value that does not depend on \mathbf{z} .
- From the previous arguments, to obtain the posterior of \mathbf{z} we can multiply $\mathcal{N}(\bar{\mathbf{y}}|\mathbf{z}, \frac{1}{N}\Sigma_y)$ with the prior $p(\mathbf{z}) = \mathcal{N}(\mathbf{z}|\mu_z, \Sigma_z)$ and

normalize the result. To do this we can use the expression for the posterior of the linear Gaussian model described in class.

- (d) plot the sample mean of the measurements, its 95% confidence ellipse, the mean of the posterior, its 95% confidence ellipse, and check if the population mean of the measurements, \mathbf{z}_1 , lies within this ellipse.

You may want to complete the script `doExSubmarine.d.py` to address this item.

- (e) repeat (b) and (d) with $N \in \{3, 10, 50, 100, 1000\}$ measurements, and show the plots generated in (d). How do the posterior and sample mean estimates change as N increases?
- (f) write expressions of the posterior mean and covariances to show that:
- as the number of measurements increases, the relative contribution of the prior to estimates of the posterior mean and covariance decreases,
 - in the limit when the number of measurements approaches infinity, the posterior covariance approaches zero and the posterior mean approaches the measurements sample mean. That is, for an infinite number of measurements, the posterior estimate becomes deterministic and the contribution of the prior to this estimate disappears.

Can you see these points in the previous simulations?

Claim 1. If $P(\mathbf{y}_i|\mathbf{z}) = \mathcal{N}(\mathbf{y}_i|\mathbf{z}, \Sigma)$, $i = 1, \dots, N$,
and $P(\mathbf{y}_1, \dots, \mathbf{y}_N|\mathbf{z}) = \prod_{i=1}^N P(\mathbf{y}_i|\mathbf{z})$,
then $P(\mathbf{y}_1, \dots, \mathbf{y}_N|\mathbf{z}) = K\mathcal{N}(\bar{\mathbf{y}}_N|\mathbf{z}, \frac{1}{N}\Sigma)$
where K is a value unrelated to \mathbf{z} .

Proof. By induction: $P_n = P(\mathbf{y}_1, \dots, \mathbf{y}_n|\mathbf{z}) = K\mathcal{N}(\bar{\mathbf{y}}_n|\mathbf{z}, \frac{1}{n}\Sigma)$
 P_1 :

$$P(\mathbf{y}_1|\mathbf{z}) = \mathcal{N}(\mathbf{y}_1|\mathbf{z}, \Sigma) = \mathcal{N}(\bar{\mathbf{y}}_1|\mathbf{z}, \frac{1}{1}\Sigma)$$

$$P_n \rightarrow P_{n+1}:$$

$$\begin{aligned} P(\mathbf{y}_1, \dots, \mathbf{y}_n, \mathbf{y}_{n+1}|\mathbf{z}) &= \prod_{i=1}^{n+1} P(\mathbf{y}_i|\mathbf{z}) \\ &= P(\mathbf{y}_1, \dots, \mathbf{y}_n|\mathbf{z})P(\mathbf{y}_{n+1}|\mathbf{z}) \\ &= \mathcal{N}(\bar{\mathbf{y}}_n|\mathbf{z}, \frac{1}{n}\Sigma)\mathcal{N}(\mathbf{y}_{n+1}|\mathbf{z}, \Sigma) \end{aligned}$$

then

$$\begin{aligned} \log P(\mathbf{y}_1, \dots, \mathbf{y}_n, \mathbf{y}_{n+1}|\mathbf{z}) &= K - \frac{1}{2}(\bar{\mathbf{y}}_n - \mathbf{z})^\top n\Sigma^{-1}(\bar{\mathbf{y}}_n - \mathbf{z}) - \\ &\quad \frac{1}{2}(\mathbf{y}_{n+1} - \mathbf{z})^\top \Sigma^{-1}(\mathbf{y}_{n+1} - \mathbf{z}) \\ &= K_1 - \frac{1}{2}(\mathbf{z}^\top(n+1)\Sigma^{-1}\mathbf{z} - 2\mathbf{z}^\top n\Sigma^{-1}\bar{\mathbf{y}}_n - 2\mathbf{z}^\top \Sigma^{-1}\mathbf{y}_{n+1}) \\ &= K_1 - \frac{1}{2}\left(\mathbf{z}^\top(n+1)\Sigma^{-1}\mathbf{z} - 2\mathbf{z}^\top \Sigma^{-1}\sum_{i=1}^n \mathbf{y}_i - 2\mathbf{z}^\top \Sigma^{-1}\mathbf{y}_{n+1}\right) \\ &= K_1 - \frac{1}{2}\left(\mathbf{z}^\top(n+1)\Sigma^{-1}\mathbf{z} - 2\mathbf{z}^\top \Sigma^{-1}\sum_{i=1}^{n+1} \mathbf{y}_i\right) \\ &= K_1 - \frac{1}{2}(\mathbf{z}^\top(n+1)\Sigma^{-1}\mathbf{z} - 2\mathbf{z}^\top(n+1)\Sigma^{-1}\bar{\mathbf{y}}_{n+1}) \end{aligned}$$

Therefore

$$P(\mathbf{y}_1, \dots, \mathbf{y}_n, \mathbf{y}_{n+1} | \mathbf{z}) = K_2 \mathcal{N} \left(\mathbf{z} \left| \bar{\mathbf{y}}_{n+1}, \frac{1}{n+1} \Sigma \right. \right) = K_2 \mathcal{N} \left(\bar{\mathbf{y}}_{n+1} \left| \mathbf{z}, \frac{1}{n+1} \Sigma \right. \right)$$

□