

# Temporal Time Series Analysis

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University College London

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## 1 Course notes

## 2 Time series analysis

- Introduction to time series analysis
- Generation of time series
- Population measures used to describe time series
- Stationarity
- Sample measures used to describe time series
- Forecasting

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- On Spring 2023 I helped in the discussion sessions of the Neuroinformatics course by Prof. Ken Harris, for UCL undergraduate students in Neuroscience.
- Suggested to Klara Olofsdotter (SWC PhD program coordinator) and Sonja Hofer (SWC PhD program faculty coordinator) to ask SWC PhD students to take this course. They liked the idea.
- With Gatsby Unit PhD students and postdoctoral scholars, as well as researchers from elsewhere, we offered **Neuroinformatics 2024**, Ken taught the first five lectures and we taught the remaining ones.
- On 2025 we renamed the course *Statistical Neuroscience* and we invited lecturers and students from the Francis Crick Institute.

# A few of our motivations to run this course

- ① Learn by teaching.
- ② Gain more teaching experience.
- ③ Provide SWC PhD students with relevant neural data-analysis tools.
- ④ Contribute to better interactions between the SWC and the Gatsby Unit. Build a common language.

# Course structure

Refer to the course [repo](#).

**Lectures** : Monday 1-3pm, SWC lecture theatre.

**Practicals** : Friday 2-3:30pm, SWC lecture theatre.

**Office Hours** : Joaquín, Wednesday 4-5pm, or by an arranged appointment.

**Worksheets** : assigned on Mondays, due on the following Monday before 1pm. Worksheets by SWC PhD students will be graded. Solutions to all worksheets will be provided.

**Participation** : in-class participation, and off-class participation (e.g., by email), is greatly encouraged.

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# What is time series analysis?

- Time series analysis characterises **data that is correlated in time**.
- These correlations severely **restrict the applicability of conventional techniques** assuming data samples that are independent and identically distributed.
- These correlations allow to **forecast** future values of a time series based on present and past values.

# Relevance of time series analysis

**economics** daily stock market quotations, monthly unemployment figures.

**social scientists** birthrates, school enrolment.

**epidemiology** number of influenza cases observed over some time period.

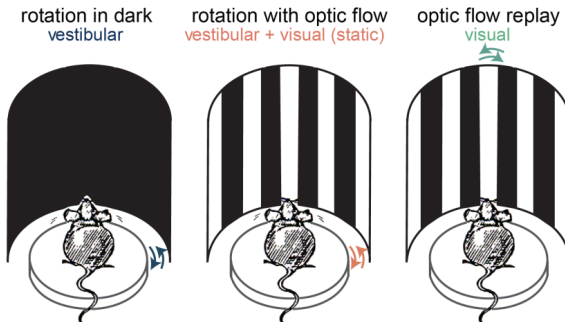
**medicine** blood pressure measurements traced over time.

# Examples of SWC time series analysis

- 1 aeon project: kinematic inference.



- 2 integration of visual/vestibular information, with Prof. Sepi Keshavarsi.



# Temporal vs spectral time series analysis

**temporal time series analysis** focuses on the analysis of lagged relationship (e.g., how does what happened today affect what will happen tomorrow?).

**spectral time series analysis** centres on the analysis of rhythms (e.g., can we observe rhythmic activity in local field potentials recorded from human brains?)

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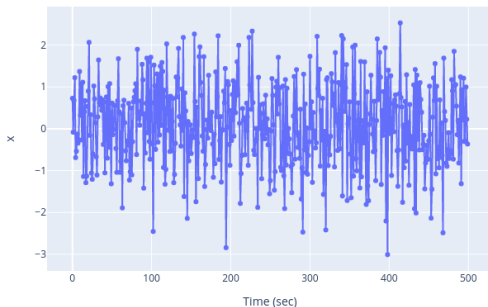
- Introduction to time series analysis
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# Generation of time series: white noise process

The first step to generate time series is to generate **white noise process**,  $\{w_t\}$  (i.e., independent Gaussian random variables with zero mean and fixed variance, **example**).

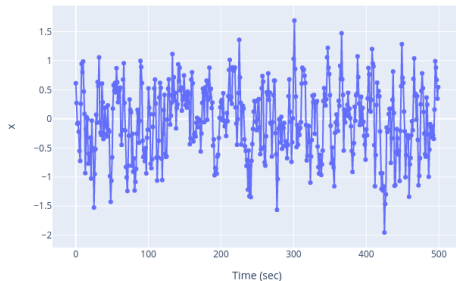
$$E\{w_t\} = 0$$
$$\text{Cov}\{w_t, w_s\} = \begin{cases} \sigma_w^2 & s = t \\ 0 & s \neq t \end{cases}$$



# Generation of time series: moving average model

In a white noise process  $w_t$ , for any pairs of time points,  $t_1$  and  $t_2$ , the random variables  $w_{t_1}$  and  $w_{t_2}$  are uncorrelated. The **moving average model** adds serial correlation to white noise (**example**).

$$\nu_t = \frac{1}{3}(w_{t-1} + w_t + w_{t+1}) \quad (1)$$





# Moving average model of order $q$

## Definition 1

An **moving average model** of order  $q$ , abbreviated as **MA**( $q$ ), is of the form

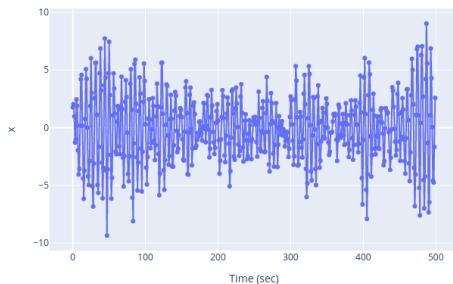
$$X_t = W_t + \theta_1 W_{t-1} + \theta_2 W_{t-2} + \dots + \theta_q W_{t-q}$$

where  $\{W_t\}$  is a white noise process and  $\theta_1, \dots, \theta_p$  are constants ( $\theta_p \neq 0$ ).

# Generation of time series: autoregressive model

Many neural time series, like local field potential recordings, exhibit oscillations of the type of sine waves. The **autoregressive model** generates oscillations (**example**). **Lecture on linear dynamical systems.**

$$x_t = x_{t-1} - 0.9x_{t-2} + w_t$$



## Definition 2

An **autoregressive model** of order  $p$ , abbreviated as **AR**( $p$ ), is of the form

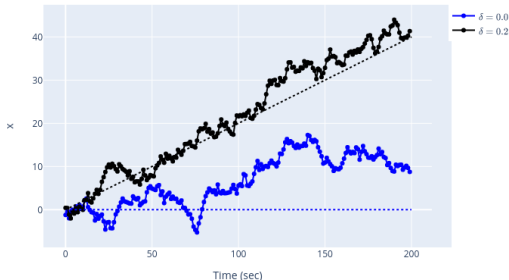
$$X_t = \phi_1 X_{t-1} + \phi_2 X_{t-2} + \dots + \phi_p X_{t-p} + w_t$$

where  $w_t$  is a white noise process,  $\phi_1, \dots, \phi_p$  are constants ( $\phi_p \neq 0$ ).

# Generation of time series: random walk with drift

The **random walk with noise** model is used to characterise trends in time series (**example**).

$$x_t = \delta + x_{t-1} + w_t \quad (2)$$



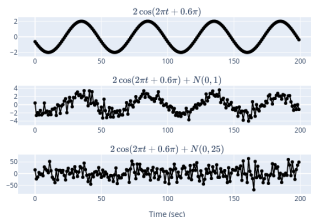
# Generation of time series: signal plus noise

Many realistic models of time series assume an underlying signal with a periodic variation contaminated by adding a random noise (**example**).

$$x_t = 2 \cos\left(2\pi \frac{t}{50} + 2\pi \frac{15}{50}\right) + w_t$$

$$A \cos(2\pi\omega t + \phi)$$

where  $A = 2, \omega = 1/50, \phi = 2\pi 15/50$ . **Lecture on spectral analysis of time series.**



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## Definition 3 (Mean function)

The mean function,  $\mu_t$ , is defined as  $\mu_t = E\{x_t\}$ .

## Example (Mean function of a moving average model)

Calculate the mean function of the moving average model in Eq. 5.

$$E\{\nu_t\} = \frac{1}{3}(E\{w_{t-1}\} + E\{w_t\} + E\{w_{t+1}\}) = \frac{1}{3}(0 + 0 + 0) = 0$$

## Example (Mean function of the autoregressive model of order 1)

Calculate the mean function of the autoregressive model of order 1, AR(1), in Eq. 3.

$$x_t = \phi x_{t-1} + w_t \quad (3)$$

An AR(1) model (Eq. 3) can be represented as a moving average of infinite order MA( $\infty$ ). See **MA( $\infty$ ) representation of AR(1) random process** in Appendix. Then

$$x_t = \sum_{i=0}^{\infty} \phi^i w_{t-i}$$

$$E\{x_t\} = \sum_{i=0}^{\infty} \phi^i E\{w_{t-i}\} = \sum_{i=0}^{\infty} \phi^i 0 = 0$$



## Example (Mean function of the random walk with drift model)

Calculate the mean function of the random noise with drift model, in Eq. 2.

The random noise with drift model in Eq. 2 can be represented as

$$x_t = t\delta + \sum_{i=0}^{\infty} w_{t-i}$$

$$E\{x_t\} = \delta t + \sum_{i=0}^{\infty} E\{w_{t-i}\} = \delta t + \sum_{i=0}^{\infty} 0 = \delta t$$

See the **figure** of samples of the random noise with drift random process.

# Autocovariance function

## Definition 4 (Autocovariance function)

The autocovariance function is defined as

$$\gamma(s, t) = \text{cov}(x_s, x_t) = E\{(x_s - \mu_s)(x_t - \mu_t)\}.$$

## Note

For  $s = t$  the autocovariance reduces to the variance, because

$$\gamma(t, t) = E\{(x_t - \mu_t)^2\} = \text{var}(x_t).$$

## Definition 5 (Autocorrelation function)

The autocorrelation function is defined as  $\rho(s, t) = \frac{\gamma(t, s)}{\sqrt{\gamma(t, t)\gamma(s, s)}}$ .

# Autocovariance function

## Example (Autocovariance function of moving average)

Calculate the autocovariance function of the moving average model in Eq. 5.

$$\gamma_\nu(s, t) = \text{cov}(\nu_s, \nu_t) = \text{cov}\left(\frac{1}{3}(w_{s-1} + w_s + w_{s+1}), \frac{1}{3}(w_{t-1} + w_t + w_{t+1})\right)$$

If  $s=t$ :

$$\begin{aligned}\gamma_\nu(t, t) &= \text{cov}(\nu_t, \nu_t) = \text{cov}\left(\frac{1}{3}(w_{t-1} + w_t + w_{t+1}), \frac{1}{3}(w_{t-1} + w_t + w_{t+1})\right) \\ &= \frac{1}{9} (\text{cov}(w_{t-1}, w_{t-1}) + \text{cov}(w_t, w_t) + \text{cov}(w_{t+1}, w_{t+1})) \\ &= \frac{1}{9} (\sigma_w^2 + \sigma_w^2 + \sigma_w^2) = \frac{3}{9} \sigma_w^2\end{aligned}$$

## Example (Autocovariance function of moving average)

If  $s=t+1$ :

$$\begin{aligned}\gamma_\nu(t+1, t) &= \text{cov}(\nu_{t+1}, \nu_t) \\ &= \text{cov}\left(\frac{1}{3}(w_t + w_{t+1} + w_{t+2}), \frac{1}{3}(w_{t-1} + w_t + w_{t+1})\right) \\ &= \frac{1}{9} (\text{cov}(w_t, w_t) + \text{cov}(w_{t+1}, w_{t+1})) \\ &= \frac{1}{9} (\sigma_w^2 + \sigma_w^2) = \frac{2}{9} \sigma_w^2\end{aligned}$$

## Example (Autocovariance function of moving average)

If  $s=t+2$ :

$$\begin{aligned}\gamma_\nu(t+2, t) &= \text{cov}(\nu_{t+2}, \nu_t) \\ &= \text{cov}\left(\frac{1}{3}(w_{t+1} + w_{t+2} + w_{t+3}), \frac{1}{3}(w_{t-1} + w_t + w_{t+1})\right) \\ &= \frac{1}{9} (\text{cov}(w_{t+1}, w_{t+1})) \\ &= \frac{1}{9} \sigma_w^2\end{aligned}$$

## Example (Autocovariance function of moving average)

$$\gamma_{\nu}(s, t) = \begin{cases} \frac{3}{9}\sigma_w^2 & \text{if } s = t, \\ \frac{2}{9}\sigma_w^2 & \text{if } |s - t| = 1, \\ \frac{1}{9}\sigma_w^2 & \text{if } |s - t| = 2, \\ 0 & \text{if } |s - t| > 2. \end{cases}$$

## Example (Autocovariance function of AR(1))

Calculate the autocovariance function of the autoregressive model of order 1 in Eq. 4.

An AR(1) model (Eq. 3) can be represented as a moving average of infinite order MA( $\infty$ ). See **MA( $\infty$ ) representation of AR(1) random process** in Appendix.

$$x_t = \sum_{i=0}^{\infty} \phi^i w_{t-i}$$

# Autocovariance function

## Example (Autocovariance function of AR(1))

$$\begin{aligned}\gamma(t-h, t) &= E\{(x_{t-h} - \mu_{t-h})(x_t - \mu_t)\} = E\{x_{t-h}x_t\} = E\left\{\left(\sum_{i=0}^{\infty} \phi^i w_{t-h-i}\right) \left(\sum_{j=0}^{\infty} \phi^j w_{t-j}\right)\right\} \\ &= E\left\{\sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \phi^i \phi^j w_{t-h-i} w_{t-j}\right\} = \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \phi^i \phi^j E\{w_{t-h-i} w_{t-j}\} \\ &= \sum_{i=0}^{\infty} \phi^i \phi^{i+h} E\{w_{t-h-i}^2\} = \phi^h \sigma_w^2 \sum_{i=0}^{\infty} \phi^{2i} = \phi^h \sigma_w^2 \frac{1}{1 - \phi^2}, \quad \text{if } |\phi| < 1\end{aligned}$$



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# Strictly stationary time series

## Definition 6 (Strict stationarity)

A **strictly stationary time series** is one for which the probabilistic behaviour of every collection of values

$$\{x_{t_1}, \dots, x_{t_n}\}$$

is identical to that of any shifted set

$$\{x_{t_1+h}, \dots, x_{t_n+h}\}$$

That is

$$P(x_{t_1} < c_1, \dots, x_{t_k} < c_k) = P(x_{t_1+h} < c_1, \dots, x_{t_k+h} < c_k)$$

for all  $k = 1, 2, \dots$ , all time points  $t_1, t_2, \dots, t_k$ , all numbers  $c_1, c_2, \dots, c_k$ , and all time shifts  $h = 0, \pm 1, \pm 2, \dots$

## Definition 7 (Weak or wide-sense stationarity)

A **weakly** or **wide-sense stationary time series** is a finite-variance process such that:

- i the mean function,  $\mu_t$ , is constant and does not depend on time  $t$ , and
- ii the autocovariance function,  $\gamma(s, t)$ , depends on  $s$  and  $t$  only through their difference  $|s - t|$ .

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# Sample mean, autocovariance and autocorrelation

## Definition 8 (Sample mean)

Let  $x_1, \dots, x_n$  be observations from a time series. The **sample mean** of  $x_1, \dots, x_n$  is

$$\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$

## Definition 9 (Sample autocovariance)

The **sample autocovariance function** is

$$\hat{\gamma}(h) = \frac{1}{n} \sum_{i=1}^{n-|h|} (x_{i+|h|} - \bar{x})(x_i - \bar{x}), \quad -n < h < n$$

## Definition 8 (Sample autocorrelation)

The **sample autocorrelation function** is

$$\hat{\rho}(h) = \frac{\hat{\gamma}(h)}{\hat{\gamma}(0)}$$

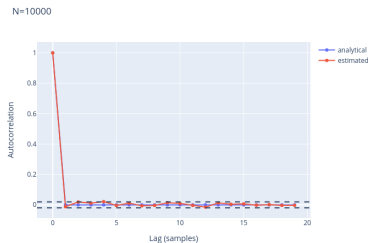
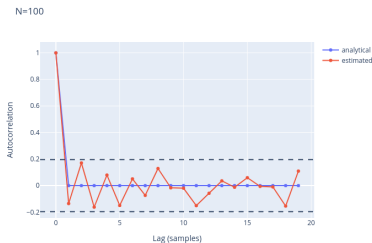
## Theorem 9 (Distribution of sample autocorrelation for white noise)

*For white noise, and a sample of size  $n$ , the sample autocorrelations,  $\hat{\gamma}(h)$ ,  $h > 0$ , are approximately independent and identically distributed  $N(0, 1/\sqrt{n})$ , for large  $n$  ([Brockwell and Davis, 1991](#)). Hence 95% of the sample autocorrelations should fall between the bound  $\pm 1.96/\sqrt{n}$*

# Analytical and estimated autocorrelation function for white noise

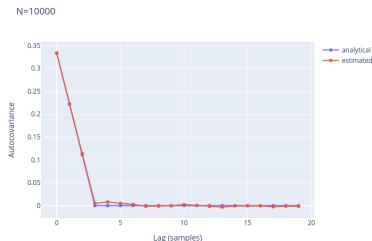
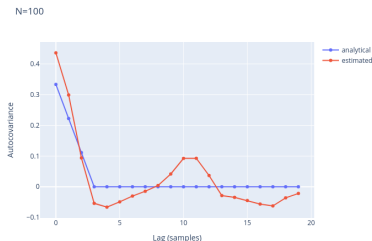
Simulate a **white noise** time series with  $N=100$  and  $N=100,000$  samples. For each  $N$ , plot the **analytical** and **estimated** autocorrelation function. Include the 95% confidence interval of the autocorrelation function.

**Solution.**



# Analytical and estimated autocovariance function for MA

Simulate the **previous** moving average time series with  $N=100$  and  $N=100,000$  samples. For each  $N$ , plot the **analytical** and **estimated** autocovariance function.

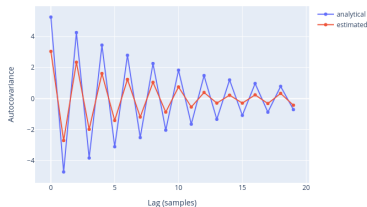




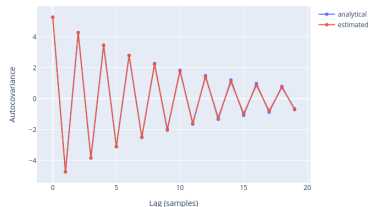
# Analytical and estimated autocovariance function for AR(1)

Simulate an AR(1) time series with  $N=100$  and  $N=100,000$  samples,  $\phi = 0.9$  and  $\sigma_w = 1.0$ . For each  $N$ , plot the **analytical** and **estimated** autocovariance function.

$N=100$



$N=10000$



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Forecasting is the problem of predicting the value of  $X_{n+h}$ ,  $h > 0$ , of a stationary time series, in term of the previous  $m$  values  $\{X_n, \dots, X_{n-(m-1)}\}$ . The mean of such predictor is

$$\text{mean}(\text{pred}(X_{n+h}|X_n, \dots, X_{n-(m-1)})) = \mu + \mathbf{a}_m^\top \begin{bmatrix} X_n - \mu \\ \dots \\ X_{n-(m-1)} - \mu \end{bmatrix}$$

and its variance is

$$\text{var}(\text{pred}(X_{n+h}|X_n, \dots, X_{n-(m-1)})) = \gamma(0) - \mathbf{a}_m^\top \gamma_m(h)$$

with

$$\Gamma_m \mathbf{a}_m = \gamma_m(h)$$

$$\Gamma_m = [\gamma(i-j)]_{i,j=1}^m = \begin{bmatrix} \gamma(0) & \gamma(1) & \gamma(2) & \gamma(3) & \dots & \gamma(m-1) \\ \gamma(1) & \gamma(0) & \gamma(1) & \gamma(2) & \dots & \gamma(m-2) \\ \gamma(2) & \gamma(1) & \gamma(0) & \gamma(1) & \dots & \gamma(m-3) \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ \gamma(m-1) & \gamma(m-2) & \gamma(m-3) & \gamma(m-4) & \dots & \gamma(0) \end{bmatrix}$$

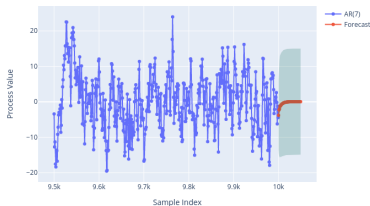
$$\mathbf{a}_m = [a_1, \dots, a_m]^\top$$

$$\gamma_m(h) = [\gamma(h), \gamma(h+1), \dots, \gamma(h+m-1)]^\top$$

# AR(p) forecasting example

## Example (Forecasting with an AR(p) model)

Simulate  $N=10,000$  samples from an  $AR(7)$  stochastic process with  $\phi = [5.0/6, -1.0/6, 0.5/6, -0.25/6, 0.5/6, -0.1/6, 0.05/6]$  and  $\sigma_w = 5.0$ . Use the last 500 samples to forecast 50 samples (i.e.,  $n = 10,000, m = 500, h = 1, \dots, 50$ ).



# Summary

- Brockwell and Davis (2002)
- Shumway and Stoffer (2016)
- Priestley (1981)

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# MA( $\infty$ ) representation of AR(1) random process

## Claim 1

Let  $|\phi| < 1$ , then

$$x_t = \phi x_{t-1} + w_t \quad \text{if and only if} \quad (4)$$

$$x_t = \sum_{i=0}^{\infty} \phi^i w_{t-i} \quad (5)$$

# MA( $\infty$ ) representation of AR(1) random process

## Proof.

We first show that  $x_t$ , as defined in Eq. 5, satisfies Eq. 4.

$$\begin{aligned}\phi x_{t-1} &= \phi \sum_{i=0}^{\infty} \phi^i w_{t-1-i} = \phi \sum_{j=1}^{\infty} \phi^{j-1} w_{t-j} = \sum_{j=1}^{\infty} \phi^j w_{t-j} \\ \phi x_{t-1} + w_t &= \sum_{j=0}^{\infty} \phi^j w_{t-j} = x_t\end{aligned}$$

# MA( $\infty$ ) representation of AR(1) random process

## Proof.

We now show that Eq. 5 is the unique solution to Eq. 3. Suppose  $y_t$  is stationary and satisfies Eq. 3, then

$$\begin{aligned}y_t &= \phi y_{t-1} + w_t \\&= \phi(\phi y_{t-2} + w_{t-1}) + w_t = \phi^2 y_{t-2} + \phi w_{t-1} + w_t \\&= \phi^{t+1} y_{t-(t+1)} + \phi^t w_t + \dots + \phi w_{t-1} + w_t \\&= \phi^{k+1} y_{t-(k+1)} + \sum_{i=0}^k \phi^k w_{t-i}\end{aligned}$$

$$E \left\{ \left( y_t - \sum_{i=0}^k \phi^i w_{t-i} \right)^2 \right\} = \phi^{2k+2} E \{ y_{t-(k+1)}^2 \} = \phi^{2k+2} \sigma^2$$

$$E \left\{ \left( y_t - \sum_{i=0}^{\infty} \phi^i w_{t-i} \right)^2 \right\} = \lim_{k \rightarrow \infty} \phi^{2k+2} \sigma^2 = 0$$

# MA( $\infty$ ) representation of AR(1) random process

Proof.

Thus  $y_t$  equals  $\sum_{i=0}^{\infty} \phi^i w_{t-i}$  in the mean-squared sense. □

Brockwell, P. J. and Davis, R. A. (1991). *Time series: Theory and methods*. Springer-Verlag, 2nd edition.

Brockwell, P. J. and Davis, R. A. (2002). *Introduction to time series and forecasting*. Springer.

Priestley, M. (1981). Spectral analysis and time series.

Shumway, R. H. and Stoffer, D. S. (2016). *Time series analysis and its applications*. Springer, 4 edition.