## THE CONVERGENCE RATE OF THE CONDITIONAL LOGIT ESTIMATOR

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We discuss the convergence rate of the conditional logit estimator and give some supporting calculations.

The Chamberlain (1992, 2010) two-period logit model has n outcome variables  $y_1, \ldots, y_n$ , and each outcome  $y_i \equiv (y_{i1}, y_{i2})'$  is generated independently according to

$$\Pr(y_{i1} = 1) = \frac{1}{1 + e^{-\alpha_i}}, \quad \Pr(y_{i2} = 1) = \frac{1}{1 + e^{-(\alpha_i + \beta)}},$$

for unknown parameters  $\alpha_1, \ldots, \alpha_n$  and  $\beta$ .

Let  $\Delta y_i \equiv y_{i2} - y_{i1}$ . Because

$$\Pr(\Delta y_i = 1 | \Delta y_i \neq 0) = \frac{1}{1 + e^{-\beta}}, \qquad \Pr(\Delta y_i = -1 | \Delta y_i \neq 0) = \frac{e^{-\beta}}{1 + e^{-\beta}}$$

do not depend on  $\alpha_i$  the conditional log-likelihood function

$$\ell(\beta) \equiv \sum_{i: \Delta y_i \neq 0} \left( \frac{1 + \Delta y_i}{2} \right) \log \left( \frac{1}{1 + e^{-\beta}} \right) + \left( \frac{1 - \Delta y_i}{2} \right) \log \left( \frac{e^{-\beta}}{1 + e^{-\beta}} \right)$$

separates estimation of the  $\alpha_i$  from inference on  $\beta$ . The conditional-logit estimator of  $\beta$  (Rasch 1960) is

$$b_n = \arg\max_{\beta} \ell(\beta).$$

In practice,  $b_n$  is computed via a standard logit programme, applied to the subsample of informative units, the index set  $\{i : \Delta y_i \neq 0\}$ , commonly referred to as movers.

Let  $n^* \equiv \sum_{i=1}^n 1\{\Delta y_i \neq 0\}$ , the number of informative observations. Note that  $n^*$  is random. Furthermore, its distribution clearly depends on the  $\alpha_i$ ; we have

$$E(n^*) = \sum_{i=1}^{n} \Pr(\Delta y_i \neq 0) = \sum_{i=1}^{n} \frac{e^{-\alpha_i} + e^{-(\alpha_i + \beta)}}{(1 + e^{-\alpha_i})(1 + e^{-(\alpha_i + \beta)})}.$$

Consistency of  $b_n$  requires that  $E(n^*) \to \infty$  as  $n \to \infty$ . This limits the speed at which  $\Pr(\Delta y_i \neq 0)$  is allowed to shrink to zero as i grows. In particular, we require that

$$n p_n \to \infty$$

as  $n \to \infty$ , where we let  $p_n = E(n^*/n)$ . Thus, the expected fraction of movers is allowed to shrink with the sample size n, but at a rate no faster than  $n^{-1}$ .

The Fisher information on  $\beta$  is

$$I \equiv \sum_{i=1}^{n} \frac{e^{-\beta}}{(1 + e^{-\beta})^2} \Pr(\Delta y_i \neq 0).$$

In large samples,  $b_n \sim \mathcal{N}(\beta, I^{-1})$ . The rate at which information accrues is  $\sqrt{np_n} = \sqrt{E(n^*)}$  and may be very slow.

Note that the difference  $n^* - np_n$  converges to zero in probability as  $n \to \infty$ . Furthermore, the inverse of the conditional information

$$I_* \equiv \sum_{i=1}^n \frac{e^{-\beta}}{(1+e^{-\beta})^2} \, \mathbf{1} \{ \Delta y_i \neq 0 \} = \sum_{i: \Delta y_i \neq 0} \frac{e^{-\beta}}{(1+e^{-\beta})^2}$$

is a valid large-sample variance for  $b_n$ . That is,  $I_* - I$  converges to zero in probability as  $n \to \infty$ . The quantity  $I_*$  is delivered by any standard logit optimization programme when applied to the subsample of movers. Thus, in practice, we base inference on the approximation

$$b_n \stackrel{a}{\sim} \mathcal{N}(\beta, I_*^{-1}),$$

which becomes more precise at the rate  $(np_n)^{-1/2}$ .

The convergence rate can be interpreted as a function of the growth rate of the  $\alpha_i$ . Moreover, because

$$\Pr(\Delta y_i \neq 0) \approx e^{-|\alpha_i|} \text{ as } |\alpha_i| \to \infty,$$

we have

$$E(n^*) \simeq \sum_{i=1}^n e^{-|\alpha_i|}.$$

If  $\alpha_i$  is finite for all i then  $0 < \Pr(\Delta y_i \neq 0) < 1$ . Consequently,  $E(n^*)$  grows like n and the convergence rate of  $b_n$  is  $n^{-1/2}$ , the parametric rate. More generally, if the  $\alpha_i$  are drawn from a distribution whose tails are sufficiently thin to ensure that  $p_n$  converges to a positive constant as  $n \to \infty$  the parametric rate remains attainable. The normal distribution would be one example. On the other hand, if the  $\alpha_i$  are allowed to become unbounded the convergence rate will decrease. For example, if  $\alpha_i \asymp \log(i)$ , then  $E(n^*) \asymp \sum_{i=1}^n i^{-1}$ , which is the nth harmonic number. For large n, the nth harmonic number behaves like  $\log(n)$ . Therefore, in this case,  $\operatorname{var}(b_n)$  shrinks like  $(\log(n))^{-1}$ , which is extremely slow. When  $\alpha_i \asymp c \log(i)$  for a constant c,  $E(n^*)$  converges to the Euler-Riemann zeta function at c, which is a finite constant for any c > 1. In such a case,  $np_n \to c$  as  $n \to \infty$  and  $b_n$  is not consistent.

## REFERENCES

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