

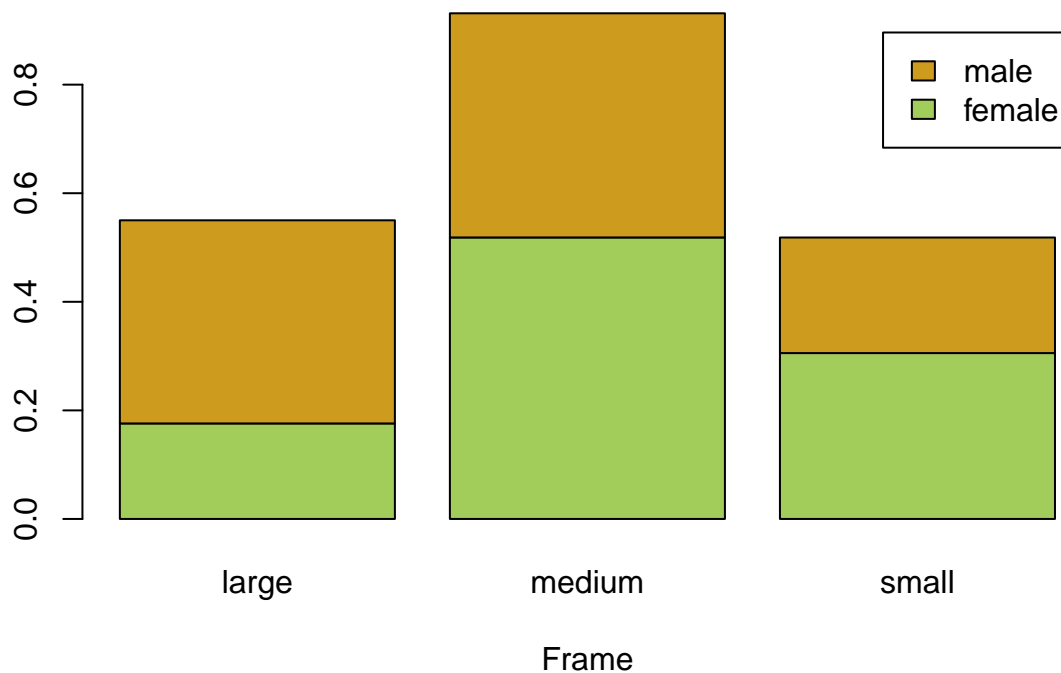
5. a)

```
diabetes <- read.csv("diabetes_fall2022.csv")
table1 <- table(diabetes$gender, diabetes$frame)
addmargins(table1)
```

```
##
##      large medium small Sum
## female    38    112    66 216
## male      58     64    33 155
## Sum       96    176    99 371
```

5. b)

```
barplot(prop.table(table1,margin=1), xlab="Frame", legend.text = rownames(table1),
        col=c("darkolivegreen3", "goldenrod3"))
```



5. c)

$H_0$  : frame and gender are independent vs  $H_1$  : frame and gender are associated  $\alpha = 0.05$

```
chisq.test(table1, correct=F)
```

```
##
```

```
## Pearson's Chi-squared test
##
## data:  table1
## X-squared = 18.734, df = 2, p-value = 8.548e-05
```

```
qchisq(1-0.05, df = (3-1)*(2-1))
```

```
## [1] 5.991465
```

We get a critical value of 5.9915 and a p-value of 0.00008548. Since  $5.9915 < 18.734 = X^2$  and  $0.00008548 < 0.05 = \alpha$ , we reject the null hypothesis that frame and gender are independent and conclude that there is an association between frame and gender. **6. a)**

$H_0 : \sigma_m^2 = \sigma_f^2$  vs  $H_1 : \sigma_m^2 \neq \sigma_f^2$   $\alpha = 0.01$

```
s_m <- var(diabetes[diabetes$gender=="male",]$stab.glu)
n_m <- length(diabetes[diabetes$gender=="male",]$stab.glu)
s_f <- var(diabetes[diabetes$gender=="female",]$stab.glu)
n_f <- length(diabetes[diabetes$gender=="female",]$stab.glu)

f_obs <- s_m/s_f
f_obs
```

```
## [1] 2.114595
```

```
2*(1-pf(f_obs, n_m-1, n_f-1))
```

```
## [1] 4.123336e-07
```

Since  $p\text{-value} = 4.1233 \times 10^{-7} < 0.01 = \alpha$ , we reject the null hypothesis that the variances of stabilized glucose for gender are the same.

**6. b)**

$H_0 : \mu_m - \mu_f \leq 0$  vs  $H_1 : \mu_m - \mu_f > 0$

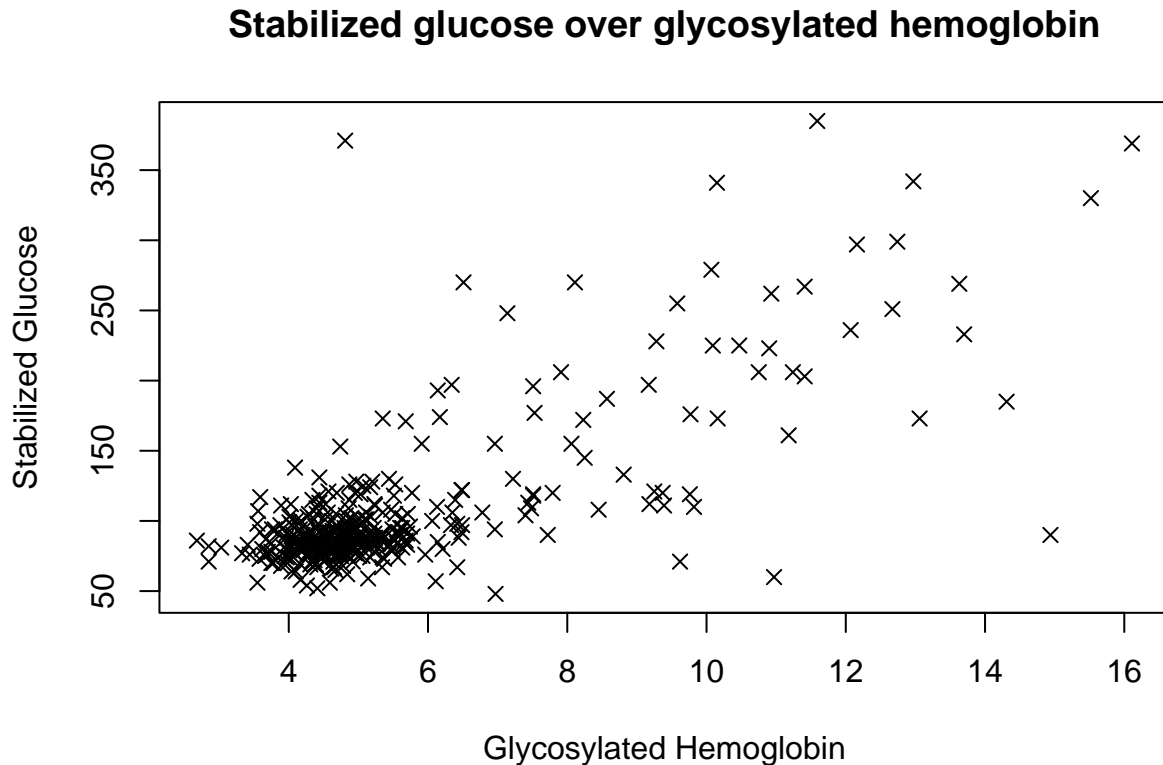
$\alpha = 0.05$

```
t.test(diabetes[diabetes$gender=="male",]$stab.glu, diabetes[diabetes$gender=="female",]$stab.glu,
       alternative="greater", conf.level = 0.95)
```

```
##
## Welch Two Sample t-test
##
## data:  diabetes[diabetes$gender == "male",]$stab.glu and diabetes[diabetes$gender == "female",]$stab.glu
## t = 1.5167, df = 255.2, p-value = 0.06529
## alternative hypothesis: true difference in means is greater than 0
## 95 percent confidence interval:
##  -0.8043766      Inf
## sample estimates:
## mean of x mean of y
## 112.4194 103.3241
```

Since  $p\text{-value} = 0.0653 > 0.05 = \alpha$ , we fail to reject the null hypothesis that the mean stabilized glucose of males is less than or equal to the stabilized glucose of females. **7. a)**

```
plot(diabetes$glyhb, diabetes$stab.glu, pch=4, main="Stabilized glucose over glycosylated hemoglobin",
     ylab = "Stabilized Glucose", xlab = "Glycosylated Hemoglobin")
```



```
cor(diabetes$glyhb, diabetes$stab.glu)
```

```
## [1] 0.7408235
```

There appears to be a moderately strong linear relationship between the two variables based on the correlation of 0.7408 and the scatterplot, indicating linear regression to be appropriate.

**7. b)**  $H_0 : \rho = 0$  vs  $H_1 : \rho \neq 0$   $\alpha = 0.05$

```
cor.test(diabetes$glyhb, diabetes$stab.glu)
```

```
##
## Pearson's product-moment correlation
##
## data: diabetes$glyhb and diabetes$stab.glu
## t = 21.186, df = 369, p-value < 2.2e-16
## alternative hypothesis: true correlation is not equal to 0
## 95 percent confidence interval:
```

```
## 0.6911384 0.7835390
## sample estimates:
##      cor
## 0.7408235
```

```
qt(1-(0.05/2), df = length(diabetes$glyhb)-2)
```

```
## [1] 1.966414
```

We find the p-value to be less than  $2.2 \times 10^{-16}$  and the test statistic to be 21.186. Since  $2.2 \times 10^{-16} < 0.05 = \alpha$ , we reject the null hypothesis that stabilized glucose and glycosylated hemoglobin are independent. **7. c)**

```
lm1 <- lm(diabetes$stab.glu~diabetes$glyhb)
lm1
```

```
##
## Call:
## lm(formula = diabetes$stab.glu ~ diabetes$glyhb)
##
## Coefficients:
##      (Intercept)  diabetes$glyhb
##           6.945           17.930
```

The regression equation is  $y = 17.930x + 6.945$

**7. d)**  $\alpha = 0.01$

```
confint(lm1, level=0.99)
```

```
##              0.5 %    99.5 %
## (Intercept)  -6.229659 20.11944
## diabetes$glyhb 15.739136 20.12179
```

We are 99% confident that the slope for the regression equation falls between 15.7391 and 20.1218.

**7. e)**

```
summary(lm1)
```

```
##
## Call:
## lm(formula = diabetes$stab.glu ~ diabetes$glyhb)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -184.826  -15.391   -3.784   10.282  277.810
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    6.9449     5.0882   1.365    0.173
## diabetes$glyhb 17.9305     0.8463  21.186 <2e-16 ***
```

```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 36.2 on 369 degrees of freedom
## Multiple R-squared:  0.5488, Adjusted R-squared:  0.5476
## F-statistic: 448.9 on 1 and 369 DF,  p-value: < 2.2e-16
```

54.88% of the variability in the glycosylated hemoglobin is explained by its linear relationship with stabilized glucose.

7. f)

```
lm2 <- lm(diabetes$stab.glu~diabetes$glyhb+diabetes$bmi+diabetes$frame)
lm2
```

```
##
## Call:
## lm(formula = diabetes$stab.glu ~ diabetes$glyhb + diabetes$bmi +
##     diabetes$frame)
##
## Coefficients:
##             (Intercept)          diabetes$glyhb          diabetes$bmi
##                9.3176             17.7060             0.1715
## diabetes$framemedium    diabetes$framesmall
##                -9.2878             -6.2408
```

$y_{large} = 17.7060x_{glyhb} * 0.1715x_{bmi} + 9.3176$   $y_{medium} = 17.7060x_{glyhb} * 0.1715x_{bmi} + 0.0298$   $y_{small} = 17.7060x_{glyhb} * 0.1715x_{bmi} + 3.0768$  The slope for all the equations is the same but each different frame has a different intercept.

7. g)

```
summary(lm2)
```

```
##
## Call:
## lm(formula = diabetes$stab.glu ~ diabetes$glyhb + diabetes$bmi +
##     diabetes$frame)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -178.827  -15.397   -4.319   11.285   278.504
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)      9.3176    11.7298   0.794  0.4275
## diabetes$glyhb    17.7060     0.8583  20.630 <2e-16 ***
## diabetes$bmi       0.1715     0.3137   0.547  0.5849
## diabetes$framemedium -9.2878     4.7065  -1.973  0.0492 *
## diabetes$framesmall -6.2408     5.7661  -1.082  0.2798
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
```

```
## Residual standard error: 36.1 on 366 degrees of freedom
## Multiple R-squared:  0.5548, Adjusted R-squared:  0.55
## F-statistic: 114 on 4 and 366 DF, p-value: < 2.2e-16
```

Multiple R-squared is 0.5548 and the adjusted R-squared is 0.55, since the adjusted R-squared is lower, that means that we added a variable that is not impactful on the output.

7. *h)*  $H_0$  : The bmi variable does not predict stabilized glucose vs  $H_1$  : The bmi variable significantly predicts stabilized glucose  $H_0$  : The frame variable does not predict stabilized glucose vs  $H_1$  : The frame variable significantly predicts stabilized glucose  $\alpha = 0.05$

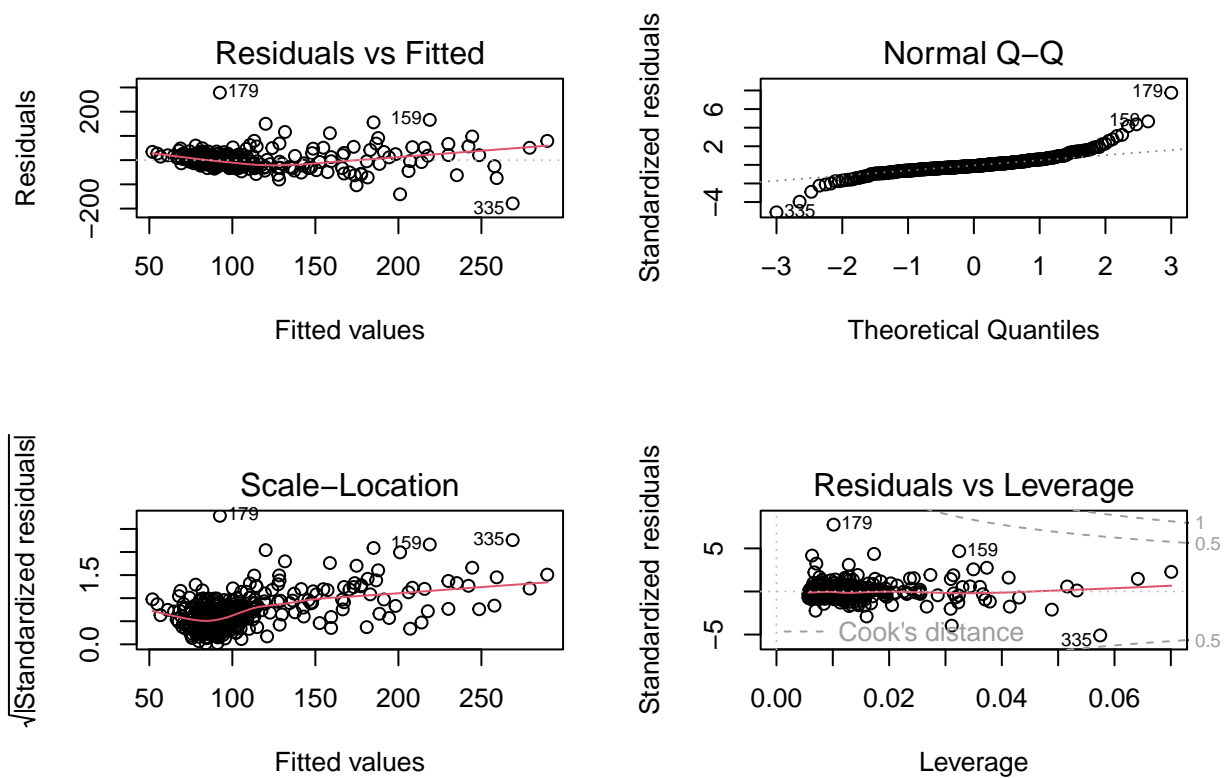
```
lm3 <- lm(diabetes$stab.glu~diabetes$bmi+diabetes$frame)
anova(lm3)
```

```
## Analysis of Variance Table
##
## Response: diabetes$stab.glu
##          Df Sum Sq Mean Sq F value    Pr(>F)
## diabetes$bmi      1   19271    19271   6.8555 0.009203 **
## diabetes$frame     2   20566    10283   3.6580 0.026727 *
## Residuals       367 1031652     2811
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

We find the test statistic for bmi to be 6.8555 and the test statistic for frame to be 3.6580. Since the bmi p-value = 0.009203 < 0.05 =  $\alpha$ , we reject the null hypothesis that the bmi variable does not predict stabilized glucose. Since the frame p-value = 0.026726 < 0.05 =  $\alpha$ , we reject the null hypothesis that the frame variable does not predict stabilized glucose.

7. *i)*

```
par(mfrow=c(2,2))
plot(lm2)
```



In the residuals vs fitted graph, we see increasing variability which indicates that the relationship is potentially non-linear and our regression line is not appropriate for modeling the data. For our homoscedasticity, we do not see a constant variance across all values of  $x$ . The QQ plot shows that the data skews right which means that our values do not have a normal distribution.