Total Variation Regularization Parameter Selection

Jodi Mead Department of Mathematics Boise State University



Outline

- TV Regularization Parameter Choices
- Residuals as Risk Estimators
- Degrees of Freedom Estimates
- $\bullet~\chi^2$ Test for Regularization Parameter Selection
- Imaging Examples

Inverse Problem

$$\mathbf{y} = \mathbf{A}\mathbf{x} + \boldsymbol{\epsilon}$$

$$\mathbf{y} \in \mathbb{R}^m$$
, $\mathbf{A} \in \mathbb{R}^{m imes n}$, $\mathbf{x} \in \mathbb{R}^n$, $\mathbb{E}(\epsilon) = \mathbf{0}$, $\mathrm{cov}(\epsilon) = \sigma^2 \mathbf{I}$

Total Variation minimization

$$\hat{\mathbf{x}}_{tv} \in \underset{\mathbf{x}}{\operatorname{arg\,min}} \ \frac{\lambda}{2} \|\mathbf{y} - \mathbf{A}\mathbf{x}\|_{2}^{2} + \|\nabla \mathbf{x}\|_{1}$$

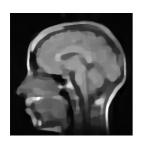
Three TV algorithms

- deconvtv (S. Chan 2011) Augmented Lagrangian method
- tvdeconv (P. Getreuer 2010) Split Bregman method
- FTVd (J. Yang, Y. Zhang, W. Yin 2008)) Augmented Lagrangian method

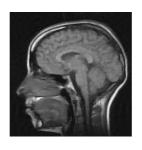
Focus of talk is on regularization parameter (λ) selection for any TV algorithm



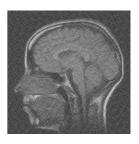
Choice of λ example







 $\lambda = 50000$



 $\lambda = 500000$

TV regularization parameter selection

Approaches

- **L-curve**: no guarantee that norm of the data residual vs.the solution will be L-shaped.
- TV function viewed or approximated with a quadratic functional:
 Discrepancy principle*, Unbiased Predictive Risk Estimator (UPRE)**,
 Generalized Cross Validation (GCV)***.



^{*}Wen et. al, 2012; **Lin et. al, 2012; ***Liao et. al, 2009

Residual properties

$$\mathbb{E}(\|\mathbf{y} - \mathbf{A}\mathbf{x}\|_2^2) = m\sigma^2$$

suggests solving nonlinear equation

$$f(\lambda) = \|\mathbf{y} - \mathbf{A}\mathbf{x}(\lambda)\|_2^2 - m\sigma^2 = 0$$

for λ . However

$$\hat{f}(\lambda) = \|\mathbf{y} - \mathbf{A}\hat{\mathbf{x}}(\lambda)\|_2^2 - m\sigma^2$$

is biased so choosing λ by solving $\hat{f}(\lambda)=0$ leads to oversmoothing (Discrepancy principle).

Effective Degrees of Freedom (EDF)

Tikhonov regularization

$$\hat{\mathbf{x}}_{ls} \in \operatorname*{arg\,min}_{\mathbf{x}} \|\mathbf{y} - \mathbf{A}\mathbf{x}\|_{2}^{2} + \lambda \|\mathbf{D}\mathbf{x}\|_{2}^{2}$$

gives ridge regression estimator

$$\hat{\mathbf{x}}_{ls} = (\mathbf{A}^T \mathbf{A} + \lambda \mathbf{D}^T \mathbf{D})^{-1} \mathbf{A}^T \mathbf{y}.$$

Predictors are

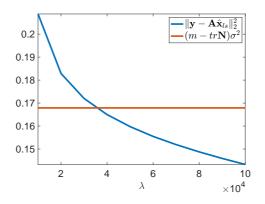
$$\mathbf{A}\hat{\mathbf{x}}_{ls} = \mathbf{N}(\lambda)\mathbf{y}$$

and*

$$\mathbb{E}(\|\mathbf{y} - \mathbf{A}\hat{\mathbf{x}}_{ls}\|_{2}^{2}) = (m - \operatorname{tr}\mathbf{N}(\lambda))\sigma^{2}$$

^{*}Hall et. al, 1987.

MRI example



Degrees of Freedom in Nonlinear Regression

Tikhonov regularization term defines smoothing matrix $A\hat{\mathbf{x}}_{ls} = \mathbf{N}\mathbf{y}$ while nonlinear smoothers (e.g. TV) have

$$\mathbf{A}\hat{\mathbf{x}}_{tv} = \delta(\mathbf{y}).$$

Degrees of freedom of δ are given by*

$$df(\mathbf{A}\hat{\mathbf{x}}_{tv}) = \sum_{i=1}^{m} \text{cov}((\hat{\mathbf{x}}_{tv})_i, \mathbf{y}_i) / \sigma^2$$

e. g. $df(\mathbf{A}\hat{\mathbf{x}}_{ls}) = tr(\mathbf{N})$.

^{*}Efron, 2004.

Degrees of Freedom for TV (generalized Lasso)*,**

$$\hat{\mathbf{x}}_{tv} \in \underset{\mathbf{x}}{\operatorname{arg\,min}} \frac{\lambda}{2} \|\mathbf{y} - \mathbf{A}\mathbf{x}\|_{2}^{2} + \|\mathbf{D}\mathbf{x}\|_{1}$$

No assumptions on A or D:

$$df(\mathbf{A}\hat{\mathbf{x}}_{tv}) = \mathbb{E}[dim(\mathbf{A}(null(\mathbf{D}_{-\mathcal{A}})))], \quad \mathcal{A} = \{i : (\mathbf{D}\hat{\mathbf{x}}_{tv})_i \neq 0\}$$

If \mathbf{A} is full rank, the number of non-zero predictors $\mathbf{D}\hat{\mathbf{x}}_{tv}$ is an unbaised estimate for $df(\mathbf{A}\hat{\mathbf{x}}_{tv})$.

^{*}Tibshirani 2012;**Dossal 2013.

TV Degrees of Freedom Example

$$nullity(\mathbf{D}_{-\mathcal{A}}) = n \to df(\mathbf{A}\hat{\mathbf{x}}_{tv}) = n$$

$$\hat{\lambda} \in \underset{\lambda}{\operatorname{arg\,min}} \|\mathbf{y} - \mathbf{A}\hat{\mathbf{x}}_{tv}\|_{2}^{2} + (n-m)\sigma^{2}$$

Similar λ choice as χ^2 test for Tikhonov

$$\frac{\|\mathbf{y} - \mathbf{A}\hat{\mathbf{x}}_{ls}\|_2^2}{\sigma^2} \sim \chi_{m-n}^2$$

Issue if $m \leq n$, instead use regularized residual*

$$\frac{\|\mathbf{y} - \mathbf{A}\hat{\mathbf{x}}_{ls}\|_2^2}{\sigma^2} + \lambda \|\mathbf{D}\hat{\mathbf{x}}_{ls}\|_2^2 \sim \chi_m^2$$

^{*}Mead et. al, 2008, 2009, 2013.

χ^2 Test for TV Functional

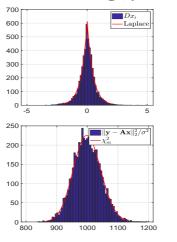
If $z_i \sim \mathsf{Laplace}(\theta, \beta)$ all independent, then

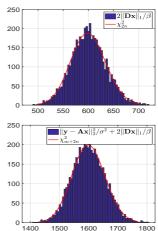
$$\sum_{i=1}^{n} \frac{2|z_i - \theta|}{\beta} \sim \chi_{2n}^2.$$

Since the TV functional is differentially Laplacian

$$\frac{\|\mathbf{y} - \mathbf{A}\mathbf{x}\|_2^2}{\sigma^2} + \frac{2\|\mathbf{D}\mathbf{x}\|_1}{\beta} \sim \chi_{m+2n}^2$$

Histograms illustrating χ^2 Test for TV Functional





χ^2 Test for TV Estimate

$$\frac{\|\mathbf{y} - \mathbf{A}\hat{\mathbf{x}}_{tv}\|_{2}^{2}}{\sigma^{2}} + \frac{2\|\mathbf{D}\hat{\mathbf{x}}_{tv}\|_{1}}{\beta} \sim \chi_{m-df(\mathbf{A}\hat{\mathbf{x}}_{tv})+df(\mathbf{D}\hat{\mathbf{x}}_{tv})}^{2}$$

Assuming **D** is full rank

$$df(\mathbf{D}\hat{\mathbf{x}}_{tv}) = df(\hat{\mathbf{x}}_{tv}) = \mathbb{E}[nullity(\mathbf{D}_{-\mathcal{A}})]$$

If \mathbf{A} is also full rank then $df(\mathbf{A}\hat{\mathbf{x}}_{tv}) = df(\mathbf{D}\hat{\mathbf{x}}_{tv})$ and we have

$$\hat{\lambda} \in \operatorname*{arg\,min}_{\lambda} \|\mathbf{y} - \mathbf{A}\hat{\mathbf{x}}_{tv}\|_{2}^{2} + \frac{2\|\mathbf{D}\hat{\mathbf{x}}_{tv}\|_{1}}{\lambda} - m\sigma^{2}$$

with $\lambda = \frac{\beta}{\sigma^2}$.



Numerical Tests - Evaluating Image Quality

MRI image filtered with a 15×15 uniform blur

Input noise:

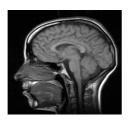
$$\mathbf{BSNR} = 20 \log_{10} \frac{\|\mathbf{y} - \mathbf{A}\mathbf{x}\|_2}{m\sigma}$$

Recovered image quality:

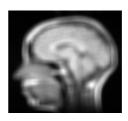
$$ISNR = 20 \log_{10} \frac{\|\mathbf{y} - \mathbf{x}\|_2}{\|\hat{\mathbf{x}} - \mathbf{x}\|_2}$$

Exact

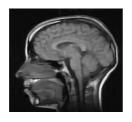
BSNR = 40

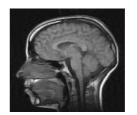


 $\chi^2 \ \mathrm{ISNR} = 8.22$

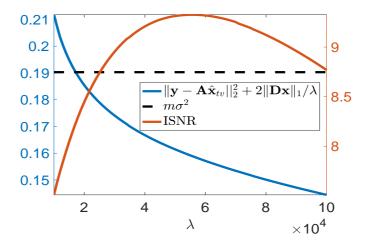


 $\mathsf{Maximum}\;\mathsf{ISNR} = 9.33$

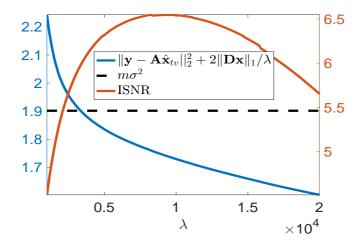




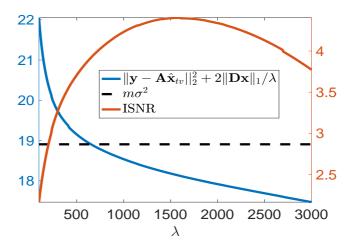
MRI BSNR = 40; χ^2 ISNR = 8.22; Max ISNR = 9.33



MRI BSNR = 30; χ^2 ISNR = 5.83; Max ISNR = 6.51



MRI BSNR = 20; χ^2 ISNR = 3.83; Max ISNR = 4.31



Summary and Conclusions

- We have developed a framework for automatic and efficient selection of TV regularization parameters. The approach extends results on residuals and risk estimators, in particular
 - The new measure of risk involves the regularized residual which follows a χ^2 distribution.
 - The degrees of freedom can be estimated from recent results on degrees of freedom for generalized Lasso.
- The proposed TV regularization parameter selection method* requires a data noise estimate and solves the TV problem multiple times during an optimization, rather than guess and checking.



^{*}Mead, in revision, J. Inv. Imag.