

EN.553.633 Homework #5

Problem 2.21) Let the random variable  $X$  have pdf

$$f(x) = \begin{cases} \frac{1}{4} & , 0 < x < 1 \\ x - \frac{3}{4} & , 1 \leq x \leq 2 \end{cases}$$

Generate a random variable from  $f(x)$ , using

(a) the inverse-transform method.

Solution: To generate a random variable from  $f(x)$ , we first need the cdf of  $X$  and subsequently need the inverse cdf of  $X$ . Then with  $f(x)$  defined as above, we have:

For  $0 < x < 1$ :

$$F(x) = \int_0^x \frac{1}{4} dt = \frac{1}{4} t \Big|_0^x = \frac{1}{4} x$$

For  $1 \leq x \leq 2$ :

$$\begin{aligned} F(x) &= \int_0^x f(t) dt = \int_0^1 \frac{1}{4} dt + \int_1^x t - \frac{3}{4} dt \\ &= \frac{t}{4} \Big|_0^1 + \frac{t^2}{2} - \frac{3}{4} t \Big|_1^x = \frac{1}{4} + \left( \frac{x^2}{2} - \frac{3}{4} x \right) - \left( \frac{1}{2} - \frac{3}{4} \right) \\ &= \frac{x^2}{2} - \frac{3}{4} x + \frac{1}{2} \end{aligned}$$

For  $x > 2$ :  $F(x) = 1$  , For  $x < 0$ :  $F(x) = 0$

Then, the cdf of  $X$  is

$$F(x) = \begin{cases} 0 & , x < 0 \\ \frac{1}{4} x & , 0 < x < 1 \\ \frac{1}{2} x^2 - \frac{3}{4} x + \frac{1}{2} & , 1 \leq x \leq 2 \\ 1 & , x > 2 \end{cases} \Rightarrow$$

The inverse of  $F(x)$  is then,

For  $x \in (0, 1)$ :  $u = F(x) = \frac{x}{4} \Rightarrow 4u = x = F^{-1}(u)$

For  $x \in [1, 2]$ :  $u = F(x) = \frac{1}{2}x^2 - \frac{3}{4}x + \frac{1}{2}$

$$= \frac{1}{2}\left(x^2 - \frac{3}{2}x + 1\right) = \frac{1}{2}\left(x^2 - \frac{3}{2}x + \left(\frac{3}{4}\right)^2 + 1 - \left(\frac{3}{4}\right)^2\right)$$

(Completing the Square)  $= \frac{1}{2}\left(x - \frac{3}{4}\right)^2 + \frac{7}{32}$

$$\Rightarrow u = F(x) = \frac{1}{2}\left(x - \frac{3}{4}\right)^2 + \frac{7}{32}$$

$$\Rightarrow 2u - \frac{7}{16} = \left(x - \frac{3}{4}\right)^2 \Rightarrow \sqrt{2u - \frac{7}{16}} = x - \frac{3}{4}$$

(since  $x \in [1, 2] \Rightarrow x - \frac{3}{4} > 0$   
 $\Rightarrow$  we only take positive square root)

$$\Rightarrow x = \sqrt{2u - \frac{7}{16}} + \frac{3}{4} = F^{-1}(u)$$

which can be simplified:  $x = \frac{1}{4}(4\sqrt{2u - \frac{7}{16}} + 3)$   
 $= \frac{1}{4}(\sqrt{32u - 7} + 3)$

and  $x \in (0, 1) \Rightarrow 0 < u < \frac{1}{4}$  and  $x \in [1, 2] \Rightarrow \frac{1}{4} \leq u \leq 1$

$$\Rightarrow F^{-1}(u) = \begin{cases} 4u & , 0 < u < \frac{1}{4} \\ \frac{1}{4}(\sqrt{32u - 7} + 3) & , \frac{1}{4} \leq u \leq 1 \end{cases}$$

Now we have the inverse cdf of  $X$  to input into our algorithm to generate r.v. from  $f(x)$ . The method/algorithm then goes

1) We first generate a  $U$  from  $U(0, 1)$ .

2) If  $U \in (0, \frac{1}{4})$ , let  $X = 4U$  and Return  $X$

3) If  $U \in [\frac{1}{4}, 1]$ , let  $X = \frac{1}{4}(\sqrt{32U - 7} + 3)$  and Return  $X$ .

part (b) on next page.  $\rightarrow$

Problem A: Consider a Markov chain with state space  $\{1, 2, 3\}$  and with a stationary probability distribution of  $\pi = [\frac{3}{11}, \frac{3}{11}, \frac{5}{11}]$ .

(a) Determine the values of the entries in the transition matrix  $P$  under the constraint that the values in the upper left  $2 \times 2$  block are equal to each other (i.e.  $p_{11} = p_{12} = p_{21} = p_{22}$ ) and that the sum of the entries in the first column is  $13/15$ .

Solution: Suppose that for some Markov chain  $X = \{X_t : t \in \mathbb{N}\}$  with state space  $\mathcal{S} = \{1, 2, 3\}$  and transition matrix  $P$ , that the stationary probability distribution,  $\pi$ , is such that  $\pi = [\frac{3}{11}, \frac{3}{11}, \frac{5}{11}]$ . Then by Markov Convergence Theorem, we have  $\pi = \pi P$

$$(*) \Rightarrow \left[ \frac{3}{11}, \frac{3}{11}, \frac{5}{11} \right] \begin{bmatrix} p_{11} & p_{12} & p_{13} \\ p_{21} & p_{22} & p_{23} \\ p_{31} & p_{32} & p_{33} \end{bmatrix} = \left[ \frac{3}{11}, \frac{3}{11}, \frac{5}{11} \right]$$

For convenience and accuracy of notation, relabel the entries of  $P$  as  $a, b, c, \dots, i$ :

$$P = \begin{bmatrix} a & b & c \\ d & e & f \\ g & h & i \end{bmatrix}$$

We are given the constraints  $a = b = d = e$  and  $a + d + g = 13/15$   
 $\Rightarrow 2a + g = 13/15$ , by the first constraint.

Then we have,

$$\left[ \frac{3}{11}, \frac{3}{11}, \frac{5}{11} \right] \begin{bmatrix} a & a & c \\ a & a & f \\ g & h & i \end{bmatrix} = \left[ \frac{6}{11}a + \frac{5}{11}g, \frac{6}{11}a + \frac{5}{11}h, \frac{3}{11}c + \frac{3}{11}f + \frac{5}{11}i \right]$$

$$= \left[ \frac{3}{11}, \frac{3}{11}, \frac{5}{11} \right]$$

We also know, by properties of transition matrices, that the sum of the components in each row is 1

$$\begin{aligned} \text{i.e. } 2a + c &= 1 \\ 2a + f &= 1 \\ g + f + i &= 1 \end{aligned}$$

$\Rightarrow$   
continued...



We then have the system of equations:

$$\begin{array}{lcl}
 2a + c & = & 1 \\
 2a + f & = & 1 \\
 g + h + i & = & 1
 \end{array}
 \left. \vphantom{\begin{array}{lcl}} \right\} \text{By transition matrix property}$$

$$\begin{array}{lcl}
 \frac{6}{11}a + \frac{5}{11}g & = & \frac{3}{11} \\
 \frac{6}{11}a + \frac{5}{11}h & = & \frac{3}{11} \\
 \frac{3}{11}c + \frac{3}{11}f + \frac{5}{11}i & = & \frac{5}{11}
 \end{array}
 \left. \vphantom{\begin{array}{lcl}} \right\} \text{By Markov Convergence Theorem}$$

$$2a + g = \frac{13}{15} \left. \vphantom{\begin{array}{lcl}} \right\} \text{Constraint}$$

- \* The first 2 equations imply that  $c = f$
  - \* The 4<sup>th</sup> and 5<sup>th</sup> equations imply that  $g = h$
- We then have:

$$\begin{array}{lcl}
 \text{(Eq. 1)} & 2a + c = 1 & \\
 \text{(Eq. 2)} & 2g + i = 1 & \\
 \text{(Eq. 3)} & \frac{6}{11}a + \frac{5}{11}g = \frac{3}{11} \Rightarrow & \frac{5}{11} \times (\text{Eq. 5}) - (\text{Eq. 3}): \\
 \text{(Eq. 4)} & \frac{6}{11}c + \frac{5}{11}i = \frac{5}{11} & \frac{5}{11}(2a + g = \frac{13}{15}) \\
 \text{(Eq. 5)} & 2a + g = \frac{13}{15} & - \left( \frac{6}{11}a + \frac{5}{11}g = \frac{3}{11} \right) \\
 & & = \frac{4}{11}a = \frac{4}{33} \Rightarrow \boxed{a = \frac{1}{3}}
 \end{array}$$

\* Plugging  $a$  into (Eq. 1):  $c = 1 - 2(\frac{1}{3}) = \frac{1}{3} \Rightarrow \boxed{c = \frac{1}{3}}$

$\Rightarrow$  From Eq. 4:  $\frac{6}{11}(\frac{1}{3}) + \frac{5}{11}i = \frac{5}{11} \Rightarrow i = \frac{11}{5}(\frac{5}{11} - \frac{2}{11})$   
 $= \frac{3}{5} \Rightarrow \boxed{i = \frac{3}{5}}$

and lastly from (Eq. 2)  $g = \frac{1}{2}(1 - \frac{3}{5}) = \frac{1}{5} \Rightarrow \boxed{g = \frac{1}{5}}$

$\Rightarrow P = \begin{bmatrix} \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ \frac{1}{5} & \frac{1}{5} & \frac{3}{5} \end{bmatrix}$

To write algorithm a bit clearer (more legible)

- 1.) Generate  $X \sim U[0, 2]$  (i.e. from  $g(x) = 1/2$  on  $x \in [0, 2]$ )
  - 2.) Generate  $Y \sim U[0, 5/2]$  (i.e. from  $U[0, Cg(x)]$ )
  - 3.) { If  $Y \leq f(x)$   
Then  $\text{Accept} \leftarrow \text{True}$ ;  
Return  $X$ ;  
If not, return to (1) }
- end.

$$\Rightarrow P = \begin{bmatrix} 1/3 & 1/3 & 1/3 \\ 1/3 & 1/3 & 1/3 \\ 1/5 & 1/5 & 3/5 \end{bmatrix}$$

(b) Using the  $P$  in part (a), carry out a simulation of 5000 (see code)

The code generated 0.2973 for  $X_2 = 1$   
 0.2588 for  $X_2 = 2$   
 and 0.4071 for  $X_2 = 3$ .

(c) Code generated 0.2728 for  $X_{20} = 1$   
 0.2789 for  $X_{20} = 2$   
 0.4583 for  $X_{20} = 3$

We can see that now that the # of steps has increased to 20, the values approach the probabilities

in the lim. dist.  $\pi = [3/11, 3/11, 5/11]$

Problem B.) (a) If  $X \sim N(\mu, \sigma^2)$ , derive  $E[e^X]$ .

Solution: Since  $X \sim N(\mu, \sigma^2)$ , then

$$E[e^X] = \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^x e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx =$$

$$= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^{-\frac{1}{2\sigma^2}[(x-\mu)^2 - 2\sigma^2 x]} dx$$

\* Focusing only on the exponent of  $e^{-\frac{1}{2\sigma^2}[(x-\mu)^2 - 2\sigma^2 x]}$ :

$$-\frac{1}{2\sigma^2}[(x-\mu)^2 - 2\sigma^2 x] = -\frac{1}{2\sigma^2}[x^2 - 2x\mu + \mu^2 - 2\sigma^2 x] =$$

$$= -\frac{1}{2\sigma^2}[x^2 - 2(\mu + \sigma^2)x + \mu^2] =$$

$$= -\frac{1}{2\sigma^2}[x^2 - 2(\mu + \sigma^2)x + \mu^2 + (2\mu\sigma^2 + \sigma^4) - (2\mu\sigma^2 + \sigma^4)]$$

$$= -\frac{1}{2\sigma^2}[x^2 - 2(\mu + \sigma^2)x + (\mu^2 + 2\mu\sigma^2 + \sigma^4)] + \frac{1}{2\sigma^2}(2\mu\sigma^2 + \sigma^4)$$

$$= -\frac{1}{2\sigma^2}[x^2 - 2(\mu + \sigma^2)x + (\mu + \sigma^2)^2] + \mu + \frac{\sigma^2}{2}$$

$$= -\frac{1}{2\sigma^2}[x - (\mu + \sigma^2)]^2 + \mu + \frac{\sigma^2}{2}$$

Now, putting this back into the exponent above:

$$\frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^x e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx = \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^{-\frac{1}{2\sigma^2}[x - (\mu + \sigma^2)]^2 + \mu + \frac{\sigma^2}{2}} dx$$

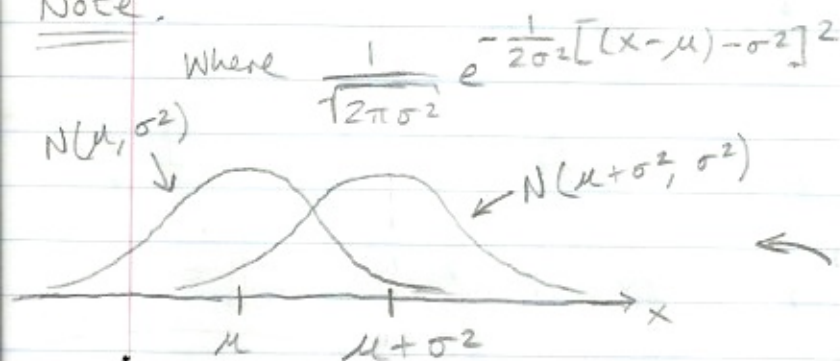
$\Rightarrow$   
cont'd...



$$= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^{-\frac{1}{2\sigma^2}[x-(\mu+\sigma^2)]^2} e^{\mu+\frac{1}{2}\sigma^2} dx$$

$$(*) = e^{\mu+\frac{1}{2}\sigma^2} \left( \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^{-\frac{1}{2\sigma^2}[(x-\mu)-\sigma^2]^2} dx \right) (*)$$

Note:



is also a probability density with mean  $\mu + \sigma^2$  and variance  $\sigma^2$ . Indeed, this is the  $\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}$  translated in the positive direction by  $\sigma^2$  units with respect to the domain.

Since the integral is taken over all of  $\mathbb{R}$ , the integral is still 1.

Thus,

$$(*) = e^{\mu+\frac{1}{2}\sigma^2} (1) = e^{\mu+\frac{1}{2}\sigma^2}$$

$$\Rightarrow E[e^x] = e^{\mu+\frac{1}{2}\sigma^2}$$



b) If  $W(t)$  is the standard Wiener process and  $U(W(t)) = \exp(t + \frac{W(t)}{2})$ , derive  $E[U(W(t))]$ .

Solution: Since  $W(t)$  is the standard Wiener process then  $W(t) - W(s) \sim N(0, t-s)$  for  $0 \leq s \leq t$  and  $W(0) = 0$ .  
Then  $W(t) - W(0) \sim N(0, t) \Rightarrow \underline{W(t) \sim N(0, t)}$   
i.e.  $E[W(t)] = 0$  and  $\text{Var}[W(t)] = t$ .

$$\begin{aligned} \text{Now consider } E[U(W(t))] &= E[e^{t + \frac{1}{2}W(t)}] \\ &= E[e^t e^{\frac{1}{2}W(t)}] = e^t E[e^{\frac{1}{2}W(t)}] \end{aligned}$$

let  $Y = \frac{W(t)}{2}$ ; then  $E[Y] = 0$ ,  $\text{Var}[Y] = \frac{1}{4}t$   
and clearly  $Y$  is a normal r.v.  $\Rightarrow Y \sim N(0, \frac{1}{4}t)$

In part (a) it was found that for a r.v.  $X \sim N(\mu, \sigma^2)$   
 $E[e^X] = e^{\mu + \sigma^2/2}$ .

$$\begin{aligned} \text{Then, by part (a), } E[e^Y] &= e^{E[Y] + \text{Var}[Y]/2} \\ &= e^{0 + t/8} = e^{t/8} \end{aligned}$$

$$\begin{aligned} \text{Then, } E[U(W(t))] &= E[e^{t + \frac{1}{2}W(t)}] = \\ &= e^t E[e^Y] = e^t e^{t/8} = e^{t + t/8} \\ &= e^{\frac{9t}{8}} \end{aligned}$$

$$\Rightarrow E[\exp(t + \frac{W(t)}{2})] = e^{9t/8}$$

(b) the acceptance-rejection method, using the proposal density  $g(x) = \frac{1}{2}x$ ,  $0 \leq x \leq 2$

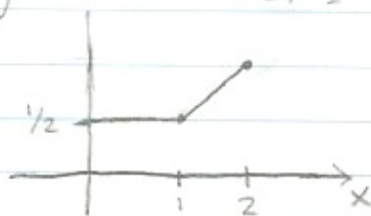
Solution: To use the A-R method to generate random variables from  $f(x)$ , we first need to find an appropriate constant,  $c$ , such that  $cg(x)$  majorizes  $f(x)$  for all  $x: 0 \leq x \leq 2$ .

That is, we need to find the minimum value of  $c$  such that  $\frac{f(x)}{g(x)} \leq c$ .

Since  $f(x)$  and  $g(x)$  are both continuous functions on  $[0, 2]$  then

$$c = \max_{x \in [0, 2]} \left( \frac{f(x)}{g(x)} \right) = \max_{x \in [0, 2]} \frac{f(x)}{1/2} = \max_{x \in [0, 2]} 2f(x)$$

Graph of  $2f(x)$ :



Since  $2f(x)$  is an increasing function (as opposed to strictly increasing) the max is at an end point. Namely,  $\arg \max_{x \in [0, 2]} 2f(x) = 2$

$$\text{Thus, } c = 2f(2) = 2\left(2 - \frac{3}{4}\right) = \frac{5}{2}$$

[Note that  $g(x) = 1/2$ ,  $x \in [0, 2] \Rightarrow g$  uniform dist on  $[0, 2]$ ]

The A-R Algorithm/Method is then:

- (1) Generate  $X$  from  $g(x) = 1/2$ ,  $x \in [0, 2]$  i.e.  $X \sim U[0, 2]$
- (2) Generate  $Y \sim U[0, 5/2]$
- (3) If  $Y \leq f(X)$  then  
Accept  $\leftarrow$  True; Return  $X$
- (4) Otherwise go back to (1).