

SS3.633 HW #6

Problem A:

(a) $E[X(t)]$: $X(t) = e^{(\lambda - \frac{\mu^2}{2})t + \mu W(t)}$

Then since $W(t) \sim N(0, t)$ (standard Wiener process)

for $Z \sim N(0, 1)$, $W(t) = \sqrt{t} Z$.

We now proceed,

$$\begin{aligned} E[X(t)] &= E[\exp[(\lambda - \frac{\mu^2}{2})t + \mu W(t)]] = \\ &= E[e^{(\lambda - \frac{\mu^2}{2})t} e^{\mu W(t)}] = e^{(\lambda - \frac{\mu^2}{2})t} E[e^{\mu W(t)}] \\ &= e^{(\lambda - \frac{\mu^2}{2})t} E[e^{\mu \sqrt{t} Z}] \quad (\text{from above}) \\ &= e^{(\lambda - \frac{\mu^2}{2})t} \int_{-\infty}^{\infty} e^{\mu \sqrt{t} z} \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz \quad (\text{Since } Z \sim N(0, 1)) \\ &= \frac{e^{(\lambda - \frac{\mu^2}{2})t}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{\frac{-(z^2 - 2\mu\sqrt{t}z)}{2}} dz \\ &= \frac{e^{(\lambda - \frac{\mu^2}{2})t}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{\frac{-(z - \mu\sqrt{t})^2}{2} + \frac{\mu^2 t}{2}} dz \\ &= \frac{e^{(\lambda - \frac{\mu^2}{2})t}}{\sqrt{2\pi}} e^{\mu^2 t/2} \int_{-\infty}^{\infty} e^{\frac{-(z - \mu\sqrt{t})^2}{2}} dz \end{aligned}$$

\Rightarrow

$$= e^{\lambda t} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{(z - \mu\sqrt{t})^2}{2}} dz = e^{\lambda t} \cdot (1)$$

$$= e^{\lambda t} \quad \text{Thus, } E[X(t)] = e^{\lambda t} //$$

(b) $E\left[\lambda \int_0^t X(s) ds\right]$

$$E\left[\lambda \int_0^t X(s) ds\right] = \lambda \int_0^t E[X(s)] ds \quad \text{by linearity of expectation}$$

(and by part (a) we showed that $E[X(t)] = e^{\lambda t}$)

$$= \lambda \int_0^t e^{\lambda s} ds = e^{\lambda s} \Big|_0^t = e^{\lambda t} - e^0 = e^{\lambda t} - 1$$

$$\Rightarrow E\left[\lambda \int_0^t X(s) ds\right] = e^{\lambda t} - 1 //$$

(c) $E\left[\mu \int_0^t X(s) dW(s)\right]$

From $dX(t) = \lambda X(t)dt + \mu X(t)dW(t)$ and $X(0) = 1$
Then, using the Ito integral equation (from slide 8)

$$X(t) = X(0) + \int_0^t \lambda X(s) ds + \int_0^t \mu X(s) dW(s)$$

We can use the values from (a) and (b) and solve.



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$$\Rightarrow E[X(t)] = E[X(0)] + E\left[\lambda \int_0^t X(s) ds\right] + E\left[\mu \int_0^t X(s) dW(s)\right]$$

$$\Rightarrow e^{\lambda t} = 1 + (e^{ts} - 1) + E\left[\mu \int_0^t X(s) dW(s)\right]$$

$$\Rightarrow E\left[\mu \int_0^t X(s) dW(s)\right] = e^{\lambda t} - e^{\lambda t} = 0$$


$$\Rightarrow E\left[\mu \int_0^t X(s) dW(s)\right] = 0$$

Problem B

```
In[202]:= sigma = 10;  
mu = 20;  
theta = 1;  
t = 0.01;  
x0 = 0;  
n = 50;  
proc = OrnsteinUhlenbeckProcess[mu, sigma, theta, x0];
```

```
In[238]:= x = ItoProcess[proc];
```

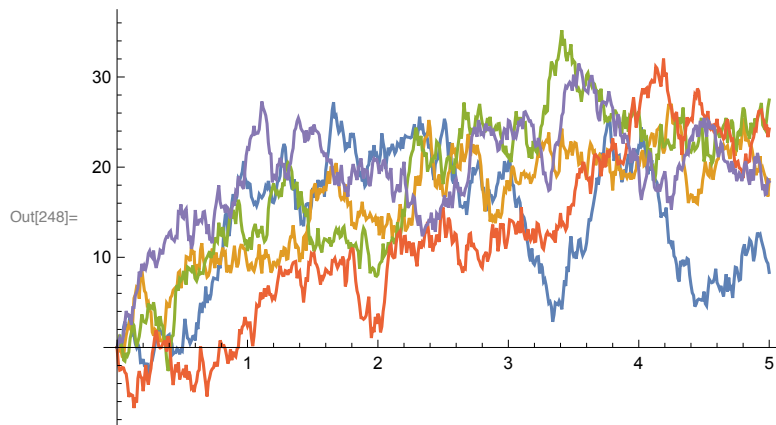
```
In[246]:= data = RandomFunction[x, {0, 5, t}, n]
```

```
Out[246]= TemporalData[ Time: 0. to 5.  
Data points: 25050 Paths: 50 ]
```

```
In[247]:= alldata = Table[data, n];
```

Part (a)

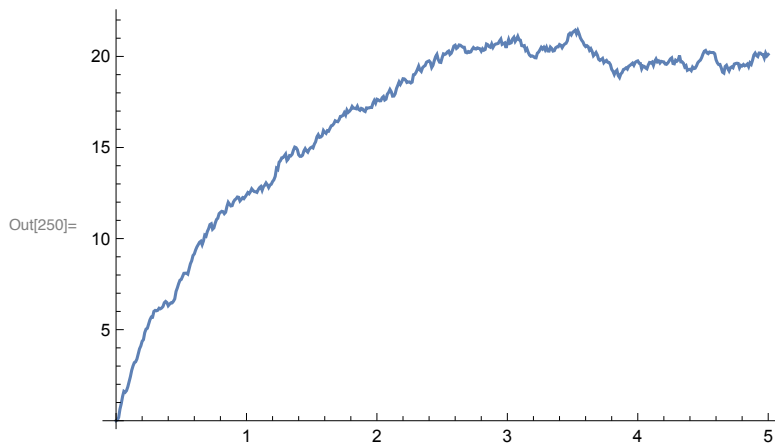
```
In[248]:= ListLinePlot[data["Path", {1, 2, 3, 4, 5}], PlotRange → All]
```



Part (b)

```
In[249]:= meanpath = Mean[data["Path", {1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12, 13, 14, 15, 16,  
17, 18, 19, 20, 21, 22, 23, 24, 25, 26, 27, 28, 29, 30, 31, 32, 33,  
34, 35, 36, 37, 38, 39, 40, 41, 42, 43, 44, 45, 46, 47, 48, 49, 50}]]];
```

```
In[250]:= ListLinePlot[meanpath, PlotRange -> All]
```



From Slide 10 “Example 2 of Ito SDE: Ornstein-Uhlenbeck Process”, $E[X(t)] \rightarrow \mu = 20$. It can be seen in the plot of the mean path above that the sample mean of the 50 simulated paths converge to 20 as t gets larger.

Part (c)

```
In[258]:= meanx2 = Mean[data[2]];
```

```
In[259]:= stdevx2 = StandardDeviation[data[2]];
```

Hypothesis Test: $H_0: \mu = 20$, $H_1: \mu \neq 20$

```
In[260]:= tstat2 = ((meanx2 - mu) / (stdevx2)) * Sqrt[n]
```

```
Out[260]:= -2.59505
```

Thus, our t-statistics is -2.59505

```
In[271]:= pvalue2 = CDF[StudentTDistribution[n-1], tstat2] +
           (1 - CDF[StudentTDistribution[n-1], -tstat2])
```

```
Out[271]:= 0.0124431
```

Thus, our P-value is 0.0124431. At the 0.05 significance level, since our p-value $< 0.05/2 = 0.025$, then we reject the null hypothesis. That is, there is not enough evidence to claim that the mean of $X(2)$ approaches the true mean.

```
meanx5 = Mean[data[5]];
```

```
stdevx5 = StandardDeviation[data[5]];
```

Hypothesis Test: $H_0: \mu = 20$, $H_1: \mu \neq 20$

```
In[263]:= tstat5 = ((meanx5 - mu) / (stdevx5)) * Sqrt[n]
```

```
Out[263]:= 0.101619
```

Thus, our t-statistics is 0.101619

```
In[272]:= pvalue5 = CDF[StudentTDistribution[n - 1], -tstat5] +  
            (1 - CDF[StudentTDistribution[n - 1], tstat5])
```

```
Out[272]= 0.919473
```

For $X(5)$ we get a much larger P-Value = 0.919473. Then, at any significance level (say 0.05) since $0.919 < 0.05/2 = 0.025$, we conclude that there is not enough evidence to reject the null hypothesis that the 50 simulations at $X(5)$ approaches the true mean.